

# Quantifying Albedo Susceptibility Biases in Shallow Clouds

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**Abstract.** The evaluation of radiative forcing associated with aerosol-cloud interactions remains a significant source of uncertainty in future climate projections. The problem is confounded by the fact that aerosol particles influence clouds locally, and that averaging to larger spatial and/or temporal scales carries biases that depend on the heterogeneity and spatial correlation of the interacting fields and the non-linearity of the responses. Mimicking commonly applied satellite data analyses for calculation of albedo susceptibility  $S_o$ , we quantify  $S_o$  aggregation biases using an ensemble of 127 large eddy simulations of marine stratocumulus. We explore the cloud field properties that control this spatial aggregation bias, and quantify the bias for a large range of shallow stratocumulus cloud conditions manifesting a variety of morphologies and range of cloud fractions. We show that  $S_o$  spatial aggregation biases can be on the order of 100s of percent, depending on methodology. Key uncertainties emanate from the typically applied adiabatic drop concentration  $N_d$  retrieval, the correlation between aerosol and cloud fields, and the extent to which averaging reduces the variance in cloud albedo  $A_c$  and  $N_d$ .  $S_o$  biases are more often positive than negative, and are highly correlated to biases in the liquid water path adjustment. Temporal aggregation biases are shown to offset spatial aggregation biases. Both spatial and temporal biases have significant implications for observationally based assessments of aerosol indirect effects and our inferences of underlying aerosol-cloud-radiation effects.

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## 15 1 Introduction

Shallow liquid clouds are a poorly quantified component of the climate system and one of the greatest sources of uncertainty for climate projections (e.g. Bony and Dufresne, 2005; Bony et al., 2017). The problem is multifaceted and encompasses fundamental understanding of how these clouds are affected by the thermodynamic structure of the atmosphere, how they might change in a warmer world, how they are influenced by the atmospheric aerosol, and how all of these components are represented in climate models. The difficulty in quantifying the radiative effects of shallow clouds emanates, to a large extent, from the large range of spatiotemporal scales involved: aerosol-cloud interaction processes need to be understood and resolved at the scale of centimeters (e.g. Hoffmann et al., 2019) while cloud fields and their organization are driven by larger scale circulations at scales of 100s to 1000s of kms (e.g. Norris and Klein, 2000). Importantly, aerosol-cloud interactions acting at

scales on the order of 100 m need to be resolved; (a) because they can lead to fundamental changes in the radiative state of  
25 a cloud system by changing the cloud albedo, cloud fraction, and spatial distribution of condensate (e.g. Sharon et al., 2006; Stevens et al., 2005; Wang and Feingold, 2009); and (b) because non-linearities in aerosol-cloud-radiation interactions mean that the methodology of averaging small-scale properties to larger-scales might generate biases in the radiative response.

This paper focuses on the ramifications of small-scale processes for cloud albedo susceptibility and cloud liquid water path  
15 adjustments (changes in liquid water path in response to changes in aerosol concentration) within the context of how they are treated in satellite-based analyses. Two key aspects are addressed: the first relates to spatial averaging from the level of the satellite pixel (order 1 km) to commonly used aggregation scales on the order of 10s – 100s of kms; the second relates to temporal aggregation from the individual scene snapshot up to a timeframe on the order of months. Large scale analyses often aggregate data both spatially and temporally into data sets that might cover e.g.  $4^\circ \times 4^\circ$  and five years (Chen et al., 2014), or a range of scales from kms - 10s kms for cloud microphysics, liquid water path, and radiation,  $1^\circ \times 1^\circ$  for aerosol, and three  
35 month seasonal responses (Lebsack et al., 2008). While it is impractical not to aggregate to some degree – e.g., to smooth noisy retrievals or extract signals from a noisy background – the implications for quantifying aerosol-cloud interactions are still not well understood.

In the following, we will use the terms ‘aggregation’ and ‘averaging’ synonymously; ‘aggregation’ tends to be used when speaking more broadly about including data from a larger range of spatial and temporal scales, whereas ‘averaging’ is used in  
40 more of a mathematical sense.

We quantify aerosol-cloud interactions using the albedo susceptibility metric (Platnick and Twomey, 1994) defined here as

$$S_o = \frac{d\ln A_c}{d\ln N_d} = \frac{(1 - A_c)}{3} \left[ 1 + \frac{5}{2} \frac{d\ln \mathcal{L}}{d\ln N_d} \right], \quad (1)$$

where  $A_c$  is the cloud albedo,  $N_d$  is the drop concentration, and  $\mathcal{L}$  is the liquid water path. The expression comprises the cloud brightening or ‘Twomey’ component  $(1 - A_c)/3$  and the adjustment term  $\mathcal{L}_o = d\ln \mathcal{L}/d\ln N_d$ , and assumes no changes in drop  
45 distribution width (e.g. Feingold and Siebert, 2009). The factor of 5/2 suggests a potentially strong but uncertain contribution from  $\mathcal{L}$  adjustments; even the sign of this term varies from positive for precipitating clouds to negative for non-precipitating clouds (Christensen and Stephens, 2011; Glassmeier et al., 2021).

Using satellite-based observation systems, e.g., the MODerate Imaging Spectroradiometer (MODIS; Salomonson et al., 1998), one can derive a drop concentration  $N_{d,a}$ , based on adiabatic assumptions, from retrieved visible cloud optical depth  $\tau$  and cloud-top drop effective radius  $r_e$ :

$$N_{d,a} = \frac{\sqrt{5}}{2\pi k} \left( \frac{f_{ad} c_w(T, P) \tau}{Q_{ext} \rho_w r_e^5} \right)^{1/2}, \quad (2)$$

where  $c_w(T, P)$  is related to the condensation rate and is a known function of cloud-base temperature  $T$  and pressure  $P$ ,  $f_{ad}$  is the adiabatic fraction (assumed in this paper to be 0.8),  $Q_{ext}$  is the extinction coefficient ( $\approx 2$  in the visible part of the spectrum),  $\rho_w$  is the density of liquid water, and  $k$  is a factor that is inversely proportional to the width of the drop size distribution (assumed to be 0.8). When  $f_{ad} = 1$ ,  $N_{d,a}$  is the adiabatic drop concentration.

Liquid water path  $\mathcal{L}$  is derived from MODIS data as

$$\mathcal{L} = \frac{5}{9} f_{ad} r_e \tau. \quad (3)$$

Further details of these derivations can be found in Brenguier et al. (2000) and Grosvenor et al. (2018).  $A_c$  can be derived from  $\tau$  using a simple two-stream approximation for a plane-parallel cloud (Bohren, 1987)

$$60 \quad A_c = \frac{\tau}{\gamma + \tau} \quad (4)$$

where  $\gamma$  depends on the degree of forward scattering. Equation(4) also assumes an overhead sun, no absorption, and a dark underlying surface. We do not consider 3-dimensional radiative transfer.

As an example of how averaging of data can affect the quantification of derived variables, we note that Eq. (4) is a concave function, which following Jensen's inequality, means that for an inhomogeneous cloud field,  $f(\bar{\tau}) > \overline{f(\tau)}$ . Thus calculating  $\tau$  based on large length scale-averaged cloud properties, and then calculating  $A_c = f(\bar{\tau})$  using Eq.(4), will generate a high bias in  $A_c$  that is inherently a function of the inhomogeneity of the cloud field. Because this well-known albedo bias (e.g. Cahalan et al., 1994; Oreopoulos and Davies, 1998) is not the topic of this paper, we will assume in all calculations that  $A_c$  is measured directly by an instrument like Clouds and the Earth's Radiant Energy System (CERES) at the desired measurement length-scale, and therefore does not suffer from an averaging bias. Instead, we explore similar biases that affect quantification of  $S_o$ . For example, Eq. (2) is a highly non-linear function of  $\tau$  and  $r_e$  so that whether one elects to calculate  $N_{d,a}$  before or after averaging of component variables  $\tau$  and  $r_e$  will potentially have a strong effect on  $S_o$ .

McComiskey and Feingold (2012) analyzed large eddy simulation (LES) output of three cloud fields characterized by different degrees of inhomogeneity and showed that an aerosol-cloud interaction metric,  $d\ln\tau/d\ln N_d$ , increased as a result of averaging – more so for heterogeneous fields than for homogeneous fields. To put the topic on firmer footing we first establish a theoretical framework for assessing the biases. We then use a large number of LESs (127) as proxy data, which we then use to simulate satellite retrievals. We apply typical methodologies used in satellite retrievals, as well as variants, to assess the provenance of the  $S_o$  bias. Both spatial and temporal aggregation scales are considered, with a focus on the former. Because of the limited domain size available to LES, we concern ourselves with the effects of spatial averaging from scales on the order of 1 km to 10 km. The multiple snapshots of different scenarios available from the LES output provide the basis for temporal aggregation. We note that Grandey and Stier (2010) also looked at the effect of aggregation scale on quantification of aerosol indirect effects (the radiative forcing of aerosol-cloud interactions). Their analysis addressed much larger scales ( $1^\circ$ ,  $4^\circ$ , and  $60^\circ$ ) in climate models. Here we focus on a model framework that explicitly resolves aerosol-cloud interactions at the cloud scale.

## 2 Theory

### 85 2.1 Spatial aggregation of variables derived from non-linear functions

The two fundamental geophysical variables associated with aerosol-cloud interactions are  $N$  (a generic concentration; drop or aerosol number concentration, which are well-correlated) and  $\mathcal{L}$ . In the case of homogeneous aerosol and cloud fields, averaging of data to different scales has no effect on derived quantities such as  $N_d$ ,  $\mathcal{L}$ ,  $A_c$ , and  $S_o$ , and the order of calculation of these fields is of no consequence. In reality, however, cloud fields exhibit different degrees of inhomogeneity: Condensation 90 of cloud water responds to local updrafts, and to some extent, availability of cloud condensation nuclei. Drop concentration depends on aerosol concentration – typically a less variable field than cloud water – as well as local supersaturation driven by updrafts. Under these conditions, the quantification of aerosol-cloud interaction metrics like  $S_o$  and the influence of averaging could be far more important.

The theoretical framework for addressing this question is well-known from similar examples in the atmospheric sciences, 95 notably biases in rain formation processes that result from large-scale averaging of the cloud water and drop concentration terms in expressions for autoconversion and accretion (e.g. Lebsock et al., 2013; Zhang et al., 2019). In the interest of completeness, we repeat the key equations here. Assuming a lognormal PDF of quantity  $x$ :

$$P(x) = \frac{1}{\sqrt{2\pi} x \ln \sigma_{g,x}} \exp[-(x - x_g)/(2 \ln^2 \sigma_{g,x})], \quad (5)$$

where the lognormal parameters are geometric mean (or median)  $x_g$  and geometric standard deviation  $\sigma_{g,x}$ . Quantity  $x$  represents 100  $A_c$  or  $\mathcal{L}$ , and  $N$ . Using well-known integral properties of the lognormal function (e.g. Feingold and Levin, 1986) it is easy to show that the bias in a moment  $x^{\beta_x}$  associated with averaging can be written as:

$$B_x = \frac{\overline{x^{\beta_x}}}{\bar{x}^{\beta_x}} = (D_x^2 + 1)^{\frac{\beta_x^2 - \beta_x}{2}}, \quad (6)$$

with the relative dispersion  $D_x$ , the ratio of the standard deviation to the mean, defined as

$$D_x = ((\bar{x}^2 - \bar{x})^2)^{1/2} / \bar{x}. \quad (7)$$

105 If the two interacting fields, e.g.  $\mathcal{L}$  and  $N$ , are assumed to follow a bivariate lognormal distribution, the bias associated with the covariance between  $\mathcal{L}$  and  $N$  is

$$B_{cov} = \exp(r(\mathcal{L}, N) \cdot \beta_{\mathcal{L}} \cdot \beta_N \cdot \sigma_{g,\mathcal{L}} \cdot \sigma_{g,N}), \quad (8)$$

where  $r(\mathcal{L}, N)$  is the spatial correlation between  $\mathcal{L}$  and  $N$ . (Note that we elect to present the theory in terms of  $\mathcal{L}$  and  $N$  rather than  $A_c$  and  $N$  because  $A_c$  includes compounded dependence on  $N$  via  $\tau$ . Since  $\mathcal{L}$  and  $A_c$  are highly correlated, this choice 110 does not affect the general framework for discussion.)

The overall bias associated with averaging for two covarying fields is then given by

$$B = B_{\mathcal{L}} \cdot B_N \cdot B_{cov}. \quad (9)$$

The equations allow theoretical calculation of the  $S_o$  biases associated with two covarying fields  $\mathcal{L}$  and  $N$ , each characterized by its own heterogeneity  $D_{\mathcal{L}}$  and  $D_N$ , respectively.

115 To apply this analysis to biases in  $S_o$  associated with interacting  $\mathcal{L}$  and  $N$  fields, we make a number of assumptions. In the absence of an analytical form for the  $\mathcal{L}$  adjustment, we assume it is negligible. We also assume that  $A_c$  depends linearly on  $\tau$ , which holds over a large range of  $\tau$ .  $S_o$  is then approximated as being proportional to  $(1 - \tau)$  (Eq. 1). Note that because the % bias in  $(1 - \tau)$ , is equivalent to the % bias in  $\tau$ , we can apply

$$\tau = c \mathcal{L}^{5/6} N^{1/3} \quad (10)$$

120 (e.g. Brenguier et al., 2000), where  $c$  is a constant that does not require specification for the purpose of the current analysis. Then, with  $\beta_{\mathcal{L}} = 5/6$  and  $\beta_N = 1/3$ , Eq. 9 can be calculated for given values of  $D_{\mathcal{L}}$ ,  $D_N$ , and  $r(\mathcal{L}, N)$ . These are shown in Fig. 1 in  $D_{\mathcal{L}}$ ;  $D_N$  space, for three values of  $r(\mathcal{L}, N)$ . We note that large positive and negative biases can result from averaging; for negative  $r(\mathcal{L}, N)$ , biases are positive, and on the order of 20 – 60 % whereas for positive values of  $r(\mathcal{L}, N)$ , biases are negative, and on the order of -20 – -70%. When correlation between the fields is zero, biases are 10 – 20 %. The central role of  $r(\mathcal{L}, N)$  is further illuminated by plotting the bias in  $S_o$  as a function of  $r(\mathcal{L}, N)$  for specified  $D_{\mathcal{L}}$  and  $D_N$  combinations (Fig. 2).

125 Note that the assumption of a bivariate lognormal distribution is common when dealing with geophysical fields. Another cautionary note is that given the various assumptions applied to generate the results in Figs. (1) and (2), they should be considered illustrative – i.e., they are primarily intended to highlight key variables that control aggregation bias. When we embark on our analysis of LES output, we will assume that  $\tau$  and  $r_e$  are known exactly, and quantitative comparison with Figs. 130 (1) and (2) should be avoided.

## 2.2 Effects of spatial and temporal averaging on variance, and correlation between fields

The second framework for assessment of biases derives from the basic definition of the linear regression fit:

$$\hat{b} = r_{x,y} \frac{\sigma_y}{\sigma_x} \quad (11)$$

135 where  $\hat{b}$  is the regression slope and  $\sigma_x$  is the standard deviation of field  $x$ . In our case,  $x$  = aerosol or drop concentration (we will use  $N_d$ ) and  $y$  = cloud variable ( $A_c$  or  $\mathcal{L}$ ). As shown by McComiskey and Feingold (2012), averaging increases  $r(x, y)$  but decreases  $\sigma_x$  and  $\sigma_y$  to varying degrees. Of interest is therefore, the extent to which averaging changes  $r(x, y)$  and the ratio  $\sigma_y/\sigma_x$ .

## 3 Large Eddy Simulation as a Data Source

140 We calculate  $S_o$  biases using output from 127 large eddy simulations of marine stratocumulus under a range of conditions from fairly homogeneous overcast to broken open cellular structures.

Simulations are generated by the System for Atmospheric Modeling (SAM; Khairoutdinov and Randall, 2003). Input conditions are derived from ERA-5 reanalysis in the stratocumulus regime off the coast of California. The model set-up is similar to

Feingold et al. (2016) and Glassmeier et al. (2021), with initial meteorological and aerosol conditions sampled using the Latin hypercube method. Simulations are nocturnal and of 12 h duration. The first 5 h of output from each simulation is discarded  
 145 to avoid cloud scenes that are not fully developed. The domain size is  $48 \text{ km} \times 48 \text{ km} \times 2.5 \text{ km}$  with a 500 m damping layer  
 below the model top at 2.5 km. The profile is extended up to 35 km (5 hPa) for radiation calculations. Grid spacings are  $dx = dy = 200 \text{ m}$ ,  $dz = 10 \text{ m}$ . A two-moment bulk microphysical method that calculates all warm cloud processes (including  
 supersaturation and activation) is applied (Feingold et al., 1998). The simulations differ from Glassmeier et al. (2021) in one  
 150 key respect, namely surface fluxes are calculated interactively. Resultant cloud fields exhibit varying degrees of heterogeneity,  
 including closed-, open-, and transitions from closed-to-open cellular convection.

The LES output provides microphysical fields of drop number concentration and liquid water content as three-dimensional  
 prognostic variables, under the assumption of a bimodal lognormal distribution of cloud drops and rain drops with fixed  
 distribution width ( $\sigma_g = 1.2$  for both). However to mimic satellite retrievals, we work with the derived model variables  $\tau$  and  
 155 cloud-top  $r_e$  to calculate  $N_d$ , and  $\mathcal{L}$  based on Equations (2) and (3). Both cloud water and rain water contribute to  $\tau$  and  
 $r_e$ . Cloud-top is calculated based on a liquid water mixing ratio threshold ( $0.01 \text{ g kg}^{-1}$ ). Because these clouds are strongly  
 capped, the first grid point exceeding this value (when working from above, and downward towards the cloud-top) almost  
 always exceeds the threshold significantly, providing a maximum, or near-maximum  $r_e$  in each model column.  $A_c$  is calculated  
 based on the modeled value of  $\tau$  (Eq. 4) and then averaged, and therefore does not suffer from averaging bias.  $S_o$  is calculated  
 160 directly for each scene based on the definition ( $S_o = d\ln A_c / d\ln N_d$ ) using least squares regression to the natural logarithms of  
 $A_c$  and  $N_d$ .

### 3.1 Spatial Aggregation

#### 3.1.1 Level 2 Analysis

The standard averaging method follows the MODIS level 2 (L2) methodology in the sense that calculations are performed at  
 high resolution prior to averaging. In the current work, Eqns. (2) and (3) are calculated based on the native 200 m model mesh  
 165 ( $n = 1$ ) and then averaged to  $n \times n$  tiles. Results will be shown for  $n = 30$  vs.  $n = 4$ , i.e.,  $6 \text{ km} \times 6 \text{ km}$  boxes vs.  $800 \text{ m} \times$   
 $800 \text{ m}$  boxes. These choices result from a desire for a reasonable number of regression points ( $n = 30$ ) and for some small  
 amount of smoothing ( $n = 4$ ) to reduce the noise in  $N_d$  retrievals (Eq. 2). The choice of  $n = 4$  also brings us close to the typical  
 1 km length-scale used for analysis of L2 MODIS data. (We did not use  $n = 5$ , i.e., 1 km, because the number of points in  
 the domain is even, and our desire is to maximize our coverage and simplify analysis.) Note that L2 methodology removes  
 170 the biases associated with non-linear functions applied to averaged data discussed in section 2.1 (Jensen's inequality) since  
 $N_d$  is calculated using Eq. (2) at the 200 m level and then averaged up. The same is true for  $A_c$ , which as previously noted is  
 calculated based on Eq. (4).

Biases are defined as

$$(\bar{X} - X)/X, \quad (12)$$

175 where  $X$  represents  $S_o$  calculated at  $n = 4$  and  $\bar{X}$  represents  $S_o$  at  $n = 30$ . Correlations  $r(\mathcal{L}, N_d)$  refer to calculations for  $n = 4$  unless otherwise stated.  $N_d$  in the L2 analysis represents cloudy-column averaged drop concentration, i.e.,  $N_d = N_{d,a}$  based on Eq.(2).

The  $N_d$  retrieval (Eq. 2) is very sensitive to thin clouds particularly when  $r_e$ , but also  $\tau$  are small. As is common in MODIS data analyses, calculations are only applied to thicker clouds, in our case, to  $r_e > 3 \mu\text{m}$  and  $\tau > 3$ .

### 180 3.1.2 Variants

Two variants of the calculations will be presented.

1. Mimic MODIS level 3 (L3) analysis. Here satellite-based retrievals are based on aggregated data. In other words Eqns. (2) and (3) for  $N_d$  and  $\mathcal{L}$ , respectively, are applied to data averaged to  $n = 4$  and  $n = 30$ . By doing so the averaging biases associated with Jensen's inequality are introduced. The biases are expected to derive from a mix of influences: 185 low for  $N_d$  (a convex function in  $r_e$  dominates a concave function in  $\tau$ ; Eq. 2), and negligible for  $\mathcal{L}$  (Eq. 3). The effect of smoothing associated with level 3 aggregation is thus likely to be highly dependent on cloud field heterogeneity.
2. Mimic MODIS L2 analysis but eliminate uncertainties in the (sub)adiabatic retrieval of  $N_d$  by assuming a 'perfect'  $N_d$ , which is taken directly from the LES. Because the LES  $N_d$  is a 3-dimensional variable, we use in-cloud, column average values. This methodology will be referred to as L2<sub>N</sub>. Note that  $\mathcal{L}$  is still calculated as per Eq. (3). The goal here is to 190 understand the extent to which the retrieval of  $N_d$  drives the regression biases.

### 3.2 Temporal aggregation

To address temporal aggregation we consider the same individual LES snapshots described above but calculate  $S_o$  in two ways: (i)  $S_o$  regression fits to individual cloud scenes are simply averaged up over all cloud scenes; (ii) LES output is temporally aggregated to a large data set, from which  $S_o$  is calculated via regression. The first approach preserves the individual cloud 195 scene susceptibilities while the second aggregates many different cloud fields before performing the regression. The biases are calculated based on Eq. (12), with the overbar indicating the second approach (ii). The methodology is followed (separately) for both  $n = 4$  and  $n = 30$  spatial averaging, and for L2, L3, and L2<sub>N</sub>.

## 4 Results

### 4.1 Spatial Aggregation

#### 200 4.1.1 Effect of Spatial Aggregation on $S_o$ and correlation

Figure 3 presents results for the three approaches (L2, L3, and L2<sub>N</sub>). We start with the figures showing  $\bar{S}_o$  vs.  $S_o$  to avoid ambiguity in the sign of the bias associated with Eq. (12). (According to Eq. 12, conditions under which  $\bar{S}_o$  is more negative than  $S_o$  also manifest as positive biases.) The solid line is the 1:1 line.

Points are colored by cloud fraction  $f_c$  (defined at the native 200 m grid spacing) since it also serves as a good proxy for cloud field heterogeneity. (The correlation between  $f_c$  and  $D(\mathcal{L})$  is 0.86.) For L2, we note that both  $\bar{S}_o$  and  $S_o$  are almost always positive and that low  $f_c$  states do not suffer from worse biases than high  $f_c$  states. On the contrary, high  $f_c$  is often associated with the largest biases. The reasons for this will become apparent in the subsequent analysis.

Responses for the L3 analysis are distinctly different in a number of ways: first, both  $\bar{S}_o$  and  $S_o$  are almost always negative, and low  $f_c$  cloud scenes often have lower biases than high  $f_c$  scenes. This is because L3 averaging has a stronger smoothing effect on broken cloud fields, and therefore somewhat unexpectedly *reduces* the averaging bias for broken cloud fields compared to solid cloud fields. Nevertheless, the non-physical shift in the sign of  $S_o$  associated with L3 methodology should act as a cautionary note. We present an explanation for the reversal of sign in Appendix A.

L2<sub>N</sub> analysis yields strongly positive  $S_o$  and a clearer dependence of the bias on  $f_c$ . For broken cloud scenes  $S_o$  is sometimes negative but biases tend to be scattered and relatively small. These low  $f_c \approx 0.3$  states are dominated by cumulus cells with stronger updrafts that result in coherent  $N_d$ . Since  $S_o$  and  $\bar{S}_o$  are only calculated in cloudy regions (above the  $r_e$  and  $\tau$  thresholds), this coherence in  $N_d$  results in a small bias. Thus the reasons for small susceptibility bias at low  $f_c$  differ for L3 and L2<sub>N</sub>. With increasing  $f_c$ , biases tighten around the 1:1 line but start to deviate for  $f_c > 0.85$ , and exhibit increasingly large values.

To quantify the biases, these same analyses are shown as % biases in  $S_o$  (Eq. 12) as a function of the correlation between  $\mathcal{L}$  and  $N_d$  ( $r(\mathcal{L}, N_d)$ ; Fig. 4), the calculations of which are consistent with the derivation of variables averaged to  $n = 4$ ; e.g., L2 and L3 calculations apply  $r(\mathcal{L}, N_d)$  based on Eqns. 2 and 3, and L2<sub>N</sub> calculates  $r(\mathcal{L}, N_d)$  using the true  $N_d$  and Eq. 3.

L2 biases are almost always positive and can reach values of many 100s of % (Fig. 4). As expected from Section 2.1,  $r(\mathcal{L}, N_d)$  has a strong influence over the  $S_o$  bias, particularly for L2, with the bias increasing noticeably with decreasing  $r(\mathcal{L}, N_d)$ , in a manner qualitatively similar to Fig. 2. The high values and high variability in the bias as one approaches  $r(\mathcal{L}, N_d) \approx 0$  are to some extent a consequence of an uncertain regression fit when the correlation between the  $\mathcal{L}$  (or the closely related  $A_c$ ) and  $N_d$  fields is poor.

Of note is that L3 analysis methodology (aggregate first, then derive) changes the sign of  $r(\mathcal{L}, N_d)$  to negative values, as it did the sign of  $S_o$  (Fig. 3). L3 biases are more widely dispersed and show no clear trend with  $r(\mathcal{L}, N_d)$  or  $f_c$ . L2<sub>N</sub> biases tend to be capped at about 100% and values of  $r(\mathcal{L}, N_d)$  are noticeably more positive than those in L2. These results reinforce two points: (i) that L3 analysis generates non-physical results (negative  $S_o$  and a change in the sign of  $r(\mathcal{L}, N_d)$ ); and (ii) that the use of the (sub)adiabatic assumption (Eq. 2) as a proxy for  $N_d$  incurs a significant increase in the  $S_o$  bias relative to L2<sub>N</sub>, most noticeably at low  $r(\mathcal{L}, N_d)$  where Eq. 2 results in a significantly reduced (but for the most part, positive) correlation.

#### 4.1.2 Effect of Spatial Aggregation on Regression

It is useful to turn to the underlying definitions of regression analysis (section 2.2) to explore more deeply the influence of averaging on the  $S_o$  bias. According to Eq. (11),  $S_o$  biases are related to the ratio of  $\hat{b}$  (6 km) to  $\hat{b}$  (800 m):

$$\hat{b}/\hat{b} = \frac{\bar{r}_{x,y}}{r_{x,y}} \left( \frac{\bar{\sigma}_y}{\bar{\sigma}_x} \right) / \left( \frac{\sigma_y}{\sigma_x} \right), \quad (13)$$

with  $\bar{b}$  denoting 6 km averaging and  $\hat{b}$  denoting 800 m averaging. We therefore consider (i) the ratio  $\bar{\sigma}_{A_c}/\bar{\sigma}_{N_d}$  to  $\sigma_{A_c}/\sigma_{N_d}$  (henceforth the ‘ $\sigma$ -ratio’; Fig. 5), and (ii) the ratio  $\bar{r}(A_c, N_d)/r(A_c, N_d)$  (henceforth the ‘ $r$ -ratio’; Fig. 6), which by Eq. (11) are both determinants of the bias. Note that for the  $r$ -ratio, we use  $r(A_c, N_d)$  to adhere more rigorously to the regression analysis definitions. The interpretation of these ratios is non-trivial. They express the extent to which  $\sigma_{A_c}/\sigma_{N_d}$  and  $r(A_c, N_d)$  are modified by averaging. In the case of  $\sigma_{A_c}/\sigma_{N_d}$  this amounts to interpreting a ‘ratio of ratios’. Later we will delve into this in more detail.

240 L2 results show a clear dependence of the  $S_o$  bias on the  $\sigma$ -ratio, and especially large biases when the  $\sigma$ -ratio is high (Fig. 5a). These also happen to be points exhibiting high  $f_c$  (cf. Fig 4a). Also apparent is that the separation of positive and negative 245 biases is demarcated at a  $\sigma$ -ratio of 1. The results suggest a strong correlation between the  $\sigma$ -ratio and  $f_c$ . At high  $r(\mathcal{L}, N_d)$ , the  $S_o$  bias increases systematically with increasing  $\sigma$ -ratio (Fig. 5a) but with decreasing  $r(\mathcal{L}, N_d)$  the strong, and orthogonal influence of the  $r$ -ratio becomes more important (Fig. 6a). Clearly evident in Fig. 6a is the anticipated unstable calculation of the  $r$ -ratio in the vicinity of  $r(\mathcal{L}, N_d) = 0$ .

250 The L3 methodology exhibits an even clearer dependence of the  $S_o$  bias on the  $\sigma$ -ratio, except for cloud fields with  $r(\mathcal{L}, N_d) \approx -0.2$  (Fig. 5b) where the high bias is clearly related to both the  $\sigma$ - and  $r$ -ratios (Fig. 6b). The aforementioned vertically oriented green colored points have a  $\sigma$ -ratio of about 1 and  $r$ -ratio of 2.5, i.e., the bias is driven by the  $r$ -ratio.

255 For the L2<sub>N</sub> methodology, here too the  $\sigma$ -ratio (Fig. 5c) provides a clearer indication of the magnitude of the bias compared to either the  $r$ -ratio (Fig. 6c) or  $f_c$  (Fig. 4).

260 Based on Eq. (13), the influence of the  $\sigma$ - and  $r$ -ratios on  $S_o$  bias is expected. What is more revealing is the influence of averaging on the *components* of these ratios. To this end, Fig. 7 examines the effect of averaging on the reduction in  $\sigma_{A_c}$  and  $\sigma_{N_d}$ . In other words, we ask: to what extent does averaging smooth the  $A_c$  field relative to the smoothing in the  $N_d$  field? Accordingly, axes in Fig. 7 are calculated as normalized reductions in the variables.

265 L2 analysis shows that at low  $f_c$  the normalized reductions in  $\sigma(N_d)$  tend to be smaller than those in  $\sigma(A_c)$  but that the reverse tends to be true for  $f_c > 0.75$  (Fig. 7a), i.e. at high  $f_c$ , averaging smooths the  $N_d$  field more than it smooths the  $A_c$  field. We will show below that these high  $f_c$  scenes, although inherently more homogeneous, often manifest significant inhomogeneity in  $N_d$  as a result of the  $N_d$  retrieval. (See further discussion in section 4.1.4; Examples.)

270 The clear exception to the trend of averaging smoothing  $N_d$  more than  $A_c$  with increasing  $f_c$  is the group of low  $f_c$  points that exists in Fig. 4a at  $r(\mathcal{L}, N_d) < 0.2$ . These anomalous points appear below the 1:1 line in Fig. 7a and show up as lower  $r$ -ratio points (reddish points) in a sea of higher  $r$ -ratio points (brown colors) in Fig. 8a. Another distinct feature is the group of vertically oriented high  $f_c$  (Fig. 7a) and low  $r$ -ratio (Fig. 8a) points for which smoothing of  $N_d$  increasingly exceeds smoothing in  $A_c$  as one moves below the 1:1 line. Because these points are characterized by small values of the  $r$ -ratio, there appears to be an offsetting of  $\sigma$ -ratio and  $r$ -ratio effects.

275 A closer look shows that these points manifest as a negative  $S_o$  bias (Fig. 9a); in other words, the reduction in the  $r$ -ratio dominates the increase in the  $\sigma$ -ratio. Although these negative bias points tend to be more rare, they can be identified in Fig. 4a at high  $f_c$  and low  $r(\mathcal{L}, N_d) (< 0.3)$ .

Analysis of L3 shows some similarities and some differences from L2. First, there is a much more significant scatter in points, particularly at lower  $f_c$  (Fig. 7b); second, as in L2, averaging-related smoothing of  $N_d$  tends to exceed smoothing in  $A_c$  at higher  $f_c$  (Fig. 7b); third, and different from L2, values of the  $r$ -ratio of  $\approx 1$  (Fig. 8b, green colors) are associated with higher  $S_o$  biases (Fig. 9b), which must derive from the  $\sigma$ -ratio.

275 Finally,  $L_{2N}$  reveals a somewhat richer palette of responses. First the commonalities: (i) the general trend for smoothing of  $N_d$  to exceed smoothing of  $A_c$  with increasing  $f_c$  and increasing  $r$ -ratio is relatively robust (cf. Fig. 7c and Figs 7ab); (ii) when the  $r$ -ratio  $\approx 1$ , and smoothing of the fields is similar (Figs. 8a and 8c), the  $S_o$  bias is capped at about 100% (Figs. 9a and 9c). *In fact it is clear from Fig. 9 that  $S_o$  biases exceeding 100% always occur when averaging-related smoothing has a stronger effect on  $N_d$  than on  $A_c$* , although the magnitude and even sign of these biases vary significantly depending on the 280 methodology. The richness (lack of monotonicity) in responses under these conditions depends on the relative strength of the  $r$ -ratio, and the extent to which it amplifies or counteracts the  $\sigma$ -ratio.

We note one interesting difference between L2 and  $L_{2N}$ : for the latter, low  $f_c$  points reside in conditions under which averaging-related smoothing is dominated by more *as well as less* significant smoothing of  $N_d$  vs  $A_c$ . (Fig. 7c). In very rare 285 cases the low  $f_c$  points above, but close to the 1:1 line in Fig. 9c cause negative  $S_o$  biases (cf. Fig. 6c). Finally, some very high positive  $S_o$  biases (up to 500 %) do exist for  $L_{2N}$ . These can be traced to conditions when the  $r$ -ratio is large (Fig. 8c) and averaging smooths  $N_d$  more than  $A_c$ . In this case the effects of averaging on the  $r$ -ratio and the  $\sigma$ -ratio work in unison to amplify the bias.

#### 4.1.3 $S_o$ bias vs. $\mathcal{L}$ adjustment bias

The topic of  $\mathcal{L}$  adjustments is of great interest given that the term may both enhance or offset the overall albedo susceptibility 290 (Eq. 1) (e.g. Glassmeier et al., 2021). For example, a value of  $\mathcal{L}_o = d\ln\mathcal{L}/d\ln N_d < -0.4$  will change the sign of  $S_o$  from positive to negative. Numerous recent articles, based on models and observations, point to  $\mathcal{L}_o$  being positive in precipitating conditions, following the familiar Albrecht (1989) ‘cloud lifetime’ hypothesis which posits that aerosol perturbations will suppress collision-coalescence and decrease precipitation, and therefore  $\mathcal{L}$  losses, while it is negative in the non-precipitating regime, as a result of enhanced evaporation-entrainment feedbacks (Wang et al., 2003; Ackerman et al., 2004; Xue et al., 2008; 295 Christensen and Stephens, 2011; Gryspeert et al., 2019). Fig. 10 shows the relationship between spatial averaging-related  $S_o$  and  $\mathcal{L}_o$  biases, with points colored by  $f_c$ . For clarity we show a subsample of 58 of the total 127 simulations to avoid points clustering and obscuring points below. These 58 samples represent Latin hypercube sampling of the full data set and do not exhibit any bias relative to the full set.

First, we note a strong positive correlation between the two biases, with  $\mathcal{L}_o$  biases larger than  $S_o$  biases in L2 and  $L_{2N}$  (Fig. 300 10a, c). The reverse is true for L3 (Fig. 10b). For L2 and L3, two distinct branches appear; the first is a tight relationship for  $f_c > 0.7$ , while the second is somewhat less well-defined and associated with lower  $f_c$ . There is a saturation in the ratio for  $f_c$  on the order of 0.3 in L2 (Fig. 10a), while L3 shows a distinctly stronger  $S_o$  bias at low  $f_c$  compared to the approximately linear relationship for high  $f_c$  (Fig. 10b). Detailed analysis of this flip in the relative slope for high  $f_c$  vs. low  $f_c$  for L2 and L3 can be traced to relative differences in the degree of aggregation-related smoothing between  $A_c$  and  $\mathcal{L}$  (see Appendix B).

305 The separation of these branches is much less distinct for L2<sub>N</sub> (Fig. 10c) but in general  $\mathcal{L}_o$  biases are larger than  $S_o$  biases, although to a lesser extent than in L2. The absence of a clear separation in the branches is a result of the use of the correct  $N_d$  such that cloud heterogeneities affect the biases to a similar degree. While the strong positive correlation between  $S_o$  and  $\mathcal{L}_o$  biases is not surprising given the expected tight relationship between  $S_o$  and  $\mathcal{L}_o$ , it is clear that using the methodologies applied here, satellite-based analyses of the  $\mathcal{L}_o$  bias have the potential to be at least as severe as those associated with  $S_o$  biases.

310 **4.1.4 Examples**

While our goal has been to provide a broad assessment of susceptibility biases in terms of cloud field properties, some examples are helpful to illustrate the issues. We focus on L2 and L2<sub>N</sub> to isolate the effect of the  $N_d$  retrieval for individual cases, chosen randomly based on their visual physical characteristics, but supported by other cases. Figure 11 presents an LES-generated stratocumulus scene characterized by high  $f_c$  (= 0.98) and a very realistic closed-cellular structure (Fig. 11a,c).

315 Drop concentration fields for the (sub)adiabatically-derived  $N_d$  (Eq. 2) and the ‘true’ (LES)  $N_d$  show the problem very clearly. The retrieved  $N_d$  shows a great deal more fine-scale structure than the true  $N_d$ , and although in the mean the retrieved  $N_d$  is not highly unrealistic (retrieved  $N_d = 167 \text{ cm}^{-3}$ ; true  $N_d = 227 \text{ cm}^{-3}$  — an error of -26 %), the two fields are strongly negatively correlated, with a fit slope of  $N_d$  (retrieved) =  $378 - 0.9 \times N_d$  (true). To quantify this further, we consider the correlations between  $\mathcal{L}$  and  $N_d$ : Applying the true  $N_d$ ,  $r(\mathcal{L}, N_d) = 0.78$  whereas using the retrieved  $N_d$ ,  $r(\mathcal{L}, N_d) = -0.24$ . This shift from 320 strong positive  $r(\mathcal{L}, N_d)$  to negative  $r(\mathcal{L}, N_d)$  has implications for the  $S_o$  bias (e.g. Fig. 2): The L2-derived  $S_o$  bias is -2017 % whereas the L2<sub>N</sub> bias is +80 %. Here the use of true  $N_d$  reduces the  $S_o$  error very significantly. This example serves to explain why high  $S_o$  biases can exist in high  $f_c$  scenes (e.g. Fig. 4). We emphasize that the significant inhomogeneity in  $N_d$  is a result of the  $N_d$  retrieval, rather than an inherent property of the cloud field, and that it is this increase in inhomogeneity and reduction in  $r(\mathcal{L}, N_d)$  that drives up the  $S_o$  bias.

325 The second example (Fig. 12) is a low  $f_c$  case (0.23) exhibiting classic open-cellular structure. Here the mean retrieved  $N_d = 21 \text{ cm}^{-3}$  and the true  $N_d = 12 \text{ cm}^{-3}$  — an error of +75 %. The best fit linear regression between the two yields  $N_d$  (retrieved) =  $20 + 0.075 \times N_d$  (true). Examining the true  $N_d$ , we find  $r(\mathcal{L}, N_d) = 0.13$  whereas using the retrieved  $N_d$ ,  $r(\mathcal{L}, N_d) = -0.05$ . The  $S_o$  biases are 285 % and 803 % for L2 and L2<sub>N</sub> respectively, i.e., the true  $N_d$  *degrades* the  $S_o$  bias. We have identified two 330 contributing factors to this unexpected result: (i) in the case of L2<sub>N</sub>, the proximity of  $r(\mathcal{L}, N_d)$  to zero generates an unstable  $r$ -ratio and unstable  $S_o$  bias (Eq. 13); (ii) more generally, unexpected results can occur when the base  $S_o$  ( $n = 4$ ) is small, which by Eq. (12), will generate unstable bias calculations. For the current case,  $S_o$  is small because the open cell walls are already very bright, i.e.,  $(1 - A_c)$  in Eq. (1) is small.

335 Note that the results in Fig. 12 may seem counterintuitive at first glance since from a satellite remote sensing perspective, a negative bias in retrieved  $N_d$  is generally expected in broken cloud scenes due to a positive bias in  $r_e$  and a negative bias in  $\tau$  (e.g., Grosvenor et al., 2018). However, the aforementioned bias is a satellite-based measurement bias. In our case, we assume that  $r_e$  and  $\tau$  are correct (taken directly from the model). We take the opportunity to emphasize that this work focuses on the effect of averaging cloud fields characterized by different morphologies, and not with the broken cloud  $r_e$  and  $\tau$  biases. The latter would need to be considered with a MODIS instrument simulator for a more complete assessment.

While examination of case studies proves useful, we argue that the broader statistical view is essential to understanding the  
340 error landscape that might be encountered in highly variable natural cloud scenes. In this regard, Figs. 3 – 6, supported by Figs.  
7 – 9, are essential.

## 4.2 Temporal Aggregation

Although the focus of the work has been on spatial aggregation, we now briefly consider the effects of temporal aggregation,  
which by Eq. (11) will also affect  $S_o$ . Large scale analyses often aggregate data over multiyear timescales (e.g. Chen et al.,  
345 2014), or three month seasonal timescales (Lebsack et al., 2008), extending the range of conditions and changing the variance  
and correlation between the fields. The result is that the regression fit to a longterm, temporally aggregated data set will be  
different from the short timeframe fits, averaged up to include the same data.

Table 1 summarizes the results for the three methodologies (L2, L3, and L2<sub>N</sub>), and for spatial averaging of  $n = 4$  and  $n = 30$ . To represent the temporally unaggregated approach, first a regression fit to  $\ln A_c$  vs  $\ln N_d$  is performed to each scene  
350 (the entire domain) to generate a value of  $S_o$ ; then the individual  $S_o$  values are averaged (represented by  $\Sigma S_o$ ). This is done  
for all scenes that meet the criteria discussed in Section 3. The temporally aggregated approach ( $\bar{S}_o$ ) aggregates all scenes  
into one large data-set before performing the regression. Because of the large number of individual  $A_c, N_d$  pairs required to  
calculate  $\bar{S}_o$ , calculations are limited to 58 of the 127 simulations, as in Section 4.1.3. The biases are calculated based on Eq.  
355 (12). Of immediate note is that the analysis shows that temporal aggregation results in a reduced  $S_o$  on the order of 70 – 110 %,  
depending on the methodology applied. For L2 and L2<sub>N</sub> the average of local-in-time cloud albedo susceptibility is *larger* than  
that calculated by temporally aggregating many scenes. This bias is of *opposite sign* to the typical biases associated with spatial  
aggregation (e.g. Fig. 4) and not significantly different in magnitude. This offsetting of errors should be seen as a cautionary  
flag when making choices of how to aggregate data, rather than as a fortuitous occurrence, since as we have shown above, the  
biases are highly dependent on cloud field properties.

360 In the case of L3, the temporal aggregation of many different scenes results in an increase in albedo susceptibility and a  
change in sign from negative to weakly positive. Here too, the spatial and temporal aggregation has opposite effects (cf. Fig.  
3b). (Note that the % differences for L3 in Table 1 are somewhat misleading: because of the change in sign, increases in  $S_o$   
due to temporal averaging still show up as negative differences because of the normalization by the negative  $\Sigma S_o$ .)

Finally, given the close relationship between the  $S_o$  and  $\mathcal{L}_o$  spatial aggregation biases, we surmise that temporal averaging  
365 will have similar effects on  $\mathcal{L}_o$  as on  $S_o$ .

## 5 Discussion and Summary

Satellite-based measurements are our best means of assessing aerosol-cloud-radiation interactions at the global scale and  
providing model constraints for one of the most uncertain forcings of the atmospheric system, namely aerosol indirect effects.  
Space-based data draw heavily on polar-orbiting satellites carrying passive instruments from which we infer aerosol properties  
370 (e.g., aerosol optical depth), cloud properties (cloud optical depth, cloud-top drop effective radius, liquid water path), and

radiative fluxes. Typical approaches to quantification of aerosol-cloud-radiation interactions average these inferred properties spatially and temporally to suppress the ‘noisy’ retrievals inherent to these measurements. The goal of this paper has been to address the ramifications of spatial and temporal aggregation for standard metrics of indirect effects in the form of cloud albedo susceptibility  $S_o$ , which is in turn strongly dependent on the liquid water path adjustment ( $\mathcal{L}_o$ ) (Eq. 1). The question 375 is central to our ability to quantify aerosol indirect effects, and raises fundamental questions of non-linearities in the aerosol-cloud interaction system, and natural co-variability of the interacting fields. Early work recognized the effects of aggregation on cloud albedo ( $A_c$ ) – the so-called “albedo bias” (e.g. Cahalan et al., 1994). The current study has assumed no aggregation error in  $A_c$  (as in the case of a direct CERES-based retrieval) but has focused instead on derivatives of the form of Eq. (1) 380 that include highly non-linear satellite-based retrievals of drop concentration ( $N_d$ ). In addition, we have derived liquid water path  $\mathcal{L}$  from the product of spectrometer-derived cloud optical depth and drop effective radius (Eq. 3), as is typically done for MODIS-based retrievals, and have not addressed the question of how this derivation might affect the  $S_o$  bias.

To provide context for the problem we start with a theoretical framework based on spatial distributions of the interacting aerosol and cloud fields, which helps identify key variables that control the  $S_o$  bias (the variance in the fields and the correlation 385 between the fields). Then, picking up on earlier work (McComiskey and Feingold, 2012), which analyzed three stratocumulus cloud scenes with varying degrees of cloud field variance, we extend the analysis to 127 simulations exhibiting a wide variety of stratocumulus cloud scenes. The LES provides all essential microphysical properties from which standard retrievals can be performed. The key cloud field properties are cloud-top drop effective radius  $r_e$ , cloud optical depth  $\tau$  (both of which are taken directly from the LES) and  $A_c$  (based on LES  $\tau$  and Eq. 4).  $N_d$  is derived from Eq. (2) and  $\mathcal{L}$ , from Eq. (3). Within 390 the framework of standard regression analysis, we quantify the effects of aggregation on the variance in the fields and the correlation between the fields, and the implications for  $S_o$  biases.

The assessment of spatial aggregation biases considers three methodologies. The first is standard level 2 (L2) satellite methodology, which retrieves cloud properties at high resolution (order 1 km) and averages them up to a scale on the order of 100 km. Given the limitations in our LES domain size (40 km), we instead compare  $S_o$  based on 800 m and 6 km scales. This forms the basis of our spatial scale averaging. Recognizing that some might choose to use the more compact aggregated data 395 as a start point, our second approach is level 3 (L3) methodology, which applies microphysical retrievals to spatially averaged data. Considering that a key and uncertain variable in the  $S_o$  calculation is the derived drop concentration ( $N_d$ ), a third set of calculations repeats the level 2 analysis but uses the ‘true’ (LES-calculated)  $N_d$ , which we refer to as the L2<sub>N</sub> methodology. All consider perfect cloud albedo retrievals based on LES cloud optical depth (Eq. 4), removing the well-known cloud albedo averaging bias from the discussion. Furthermore, derived variables  $r_e$  and  $\tau$ , which are central to the  $S_o$  bias analysis have been 400 assumed to be free of measurement error. Real-world satellite retrieval errors of these variables, especially in heterogeneous or broken cloud fields, could amplify or perhaps counter the biases identified here. Therefore, applying the conclusions directly to satellite studies should be done with caution.

Key results pertaining to spatial aggregation biases are

1.  $S_o$  biases are generally positive for all three approaches and, consistent with theory, tend to increase with decreasing  
405 correlation between the fields (Fig. 4). For L2, biases can reach 500 %, whereas they are capped at about 100 % for L2<sub>N</sub>. These biases will translate to biases in aerosol-cloud radiative forcing.

2. L3 methodology (falsely) creates negative correlations between the cloud fields, resulting in generally unpredictable  $S_o$  bias behavior (Fig. 4b). Moreover, it generates negative  $S_o$  values (see Appendix A), exacerbated by increasing degrees of averaging (Fig. 3b). The positive  $S_o$  biases for L3 are therefore misleading (Figs. 3b and 4b).

410 3. High cloud fraction  $f_c$  states are equally or even more prone to  $S_o$  bias than low  $f_c$  (high  $D(\mathcal{L})$ ) states. This is in part due to the nature of the  $N_d$  retrieval, which introduces heterogeneity into the field (see example Fig. 11), but L2<sub>N</sub> cases also tend to exhibit this trend. The literature tends to consider high  $f_c$  states as homogeneous, but this work shows that this is not necessarily the case since the  $N_d$  retrieval generates unrealistic small-scale heterogeneity in relatively homogeneous conditions.

415 4. Using regression theory (Eq. 11) we can interpret biases in  $S_o$  based on the extent to which averaging changes (i) the ratio of the correlation  $r_{x,y}$  ( $r$ -ratio) and (ii) the ratio of  $\sigma_y/\sigma_x$  ( $\sigma$ -ratio) at the 6 km vs. the 800 m scales. Fig. 5 shows that the  $S_o$  bias is strongly dependent on the  $\sigma$ -ratio, particularly for L3 and L2<sub>N</sub>, and for L2 at high  $r(\mathcal{L}, N_d)$ . The  $r$ -ratio has a weaker control over the  $S_o$  biases (Fig. 6).

420 5. Since the  $\sigma$ -ratio is a ratio-of-ratios, we further expand our analysis of this term to assess the extent to which averaging changes  $\sigma_{A_c}$  relative to  $\sigma_{N_d}$ . We find that while averaging can reduce  $\sigma_{A_c}$  as much as  $\sigma_{N_d}$ , there is a tendency for larger reductions in  $\sigma_{N_d}$  (Fig. 7), particularly for high  $f_c$  and low  $r$ -ratio L2 cases (Fig. 8) and for low  $f_c$  L2<sub>N</sub> cases (Fig. 7).

425 6.  $S_o$  biases exceeding 100% always occur when averaging-related smoothing has a stronger effect on  $N_d$  than on  $A_c$  (Fig. 9) although the magnitude and even sign of these biases vary significantly depending on the methodology. The lack of monotonicity in responses under these conditions depends on the relative strength of the  $r$ -ratio, and the extent to which it amplifies or counteracts the  $\sigma$ -ratio. This in turn, depends on the cloud fields themselves in ways that are not always easy to clarify.

7. As anticipated by Eq. (1) the  $S_o$  bias is a strong function of the  $\mathcal{L}_o$  bias (Fig. 10). In the case of L2, the  $\mathcal{L}_o$  bias is noticeably smaller than the  $S_o$  bias at low  $f_c$ , while the reverse is true for L3 (see Appendix B).

Regarding temporal aggregation, we note that if aerosol-cloud interaction metrics such as  $S_o$  are based on aggregation of  
430 cloud scenes over an extended period of time, bias can be expected as a result of extending the range of conditions beyond the natural local fluctuations inherent to covarying meteorological and aerosol conditions. To assess the effects of this temporal aggregation, we consider the difference between the average of  $S_o$  from individual cloud scenes, and the  $S_o$  that is derived from a best-fit to the data from all those scenes, with the former reflecting the average of local-in-time  $S_o$  and the latter reflecting the effect of temporal aggregation. Calculations are performed at both 800 m and 6 km scales (Table 1). Of note is that temporal averaging, as calculated in this manner, *reduces* (in % terms)  $S_o$  at both averaging scales, and for L2, L3, and L2<sub>N</sub> analysis.

The negative bias is on the order of 70 – 80 % for L2 and L2<sub>N</sub>. As noted in section 4.1.1, L3 spatial aggregation methodology generates *negative*  $S_o$ , and temporal aggregation has the beneficial effect of creating small positive, and therefore more realistic  $S_o$  (relative to temporally unaggregated) values. (In percentage change terms, however, the bias is negative.)

We emphasize that these offsetting effects of spatial and temporal  $S_o$  biases are very situationally-dependent and require  
440 further investigation. As in the case of the regression analysis (Eq. 11) performed for spatial aggregation, the pertinent question becomes the extent to which temporal aggregation will affect the  $\sigma$ - and  $r$ -ratios. Further work will need to place this offsetting effect on firmer footing.

Finally, it is of great importance that similar assessments of  $S_o$  biases be considered in real-world satellite-based data. This will allow the community to assess the degree of coherence between the aforementioned studies and the model-based results  
445 presented here, and the implications for quantification of aerosol-cloud radiative forcing.

**Table 1.** Temporal averaging values and % differences.  $\Sigma S_o$  refers to a simple average of the  $S_o$  for individual cloudy scenes;  $\bar{S}_o$  refers to the  $S_o$  regression fit to the temporally aggregated data from all cloudy scenes; the % difference is defined as  $(\bar{S}_o - \Sigma S_o)/\Sigma S_o$ .  $n = 4$  and  $n = 30$  refer to the 800 m and 6 km spatial averaging scales, respectively.

| Temporal Averaging Values and % Differences |              |             |     |              |             |      |                 |             |     |
|---|--------------|-------------|-----|--------------|-------------|------|-----------------|-------------|-----|
|   | L2           |             |     | L3           |             |      | L2 <sub>N</sub> |             |     |
|   | $\Sigma S_o$ | $\bar{S}_o$ | %   | $\Sigma S_o$ | $\bar{S}_o$ | %    | $\Sigma S_o$    | $\bar{S}_o$ | %   |
| $n = 4$                                     | 0.564        | 0.259       | -74 | -0.652       | 0.044       | -107 | 1.012           | 0.301       | -70 |
| $n = 30$                                    | 0.966        | 0.298       | -69 | -0.959       | 0.073       | -108 | 1.373           | 0.296       | -78 |

*Data availability.* TEXT

Model output is available on request.

## Appendix A

450 Sorting or aggregation of data can generate unexpected responses. A common example of this is the relationship between ‘success at basketball’ and ‘height’ (Fig. A1). Although the expected positive correlation between these two variables emerges if one sorts the data by age group, the reverse occurs if one aggregates all the data together (‘first aggregate, then fit’), simply because older players, while taller, do not typically perform as well as younger players. This apparent contradiction was noted by Simpson (1951) and is known as Simpson’s Paradox. It is receiving more attention during the COVID-19 pandemic as  
455 it explains apparent contradictions between vaccination rates and hospitalizations when data are not sorted by confounding factors.

Although hard to prove, the commonality for the current data is that L2 methodology takes the approach of fitting the un-aggregated data, whereas L3 methodology first averages the data and then performs a fit. In our case the data are not as clearly separated as in the simple example in Fig. A1 but L2 methodology has a much higher chance of retaining the  
460 underlying physical relationship between variables than L3 methodology. Indeed, L2 produces positive  $S_o$  whereas L3 produces a counterintuitive negative  $S_o$ .

## Appendix B

We address the flip in the relative slopes of low and high  $f_c$  points for L2 and L3 in Fig. 10. Based on Fig. 7 and Eq. (13), our  
465 intuition is that it is likely a function of the change in averaging-smoothing in  $A_c$  vs. averaging-smoothing in  $\mathcal{L}$ . To test this, we repeat Fig. 7 but now also look at smoothing in  $\mathcal{L}$  and  $N_d$ .

What is apparent is that for L3, there is much more smoothing in  $A_c$  than in  $\mathcal{L}$  at low  $f_c$  (the L3 points lie closer to the 1:1 line in Fig. B1b vs. Fig. 7b for low  $f_c$ ). For L2, more low  $f_c$  points tend to lie below the 1:1 line in Fig. B1a vs. Fig. 7a, but because of this migration of points from above to below the line, it is more difficult to interpret. For L2N the smoothing in  $\mathcal{L}$  470 and  $A_c$ , look similar, in line with our intuition (cf. Fig. 10 where one sees less of a bias in the relationship for low and high  $f_c$  points).

To dig deeper, we look more closely at the  $\sigma$ -ratio (Eq. 13) for (i)  $A_c, N_d$  ( $\sigma$ -ratio ( $A_c, N_d$ )) and (ii)  $\mathcal{L}, N_d$  ( $\sigma$ -ratio ( $\mathcal{L}, N_d$ )) for the low  $f_c$  points. We then fit a linear relation between points for  $0.3 < f_c < 0.43$  (Fig. B2) and  $f_c > 0.85$  (not shown) and obtain the following: For  $0.3 < f_c < 0.43$ , L2:  $y = 0.06 + 0.78x$ , L3:  $y = -0.25 + 1.03x$ , L2N:  $y = -0.22 + 1.02x$  (with  $x =$   
475 the  $\sigma$  ratio for  $\mathcal{L}$  and  $N_d$ , and  $y$ , the  $\sigma$ -ratio for  $A_c$  and  $N_d$ ). For  $f_c > 0.85$ , L2:  $y = 0.1 + 0.90x$ , L3:  $y = 0.04 + 0.94x$ , L2N:  $y = 0.23 + 0.81x$ .

The high  $f_c$  slopes for L2 and L3 are similar (0.90 vs. 0.94, respectively) but there is a more significant difference in the low  $f_c$  slopes (0.78 vs. 1.03 for L2 and L3 respectively), and in the direction consistent with the differences between L2 and L3 (Fig. 10ab). Note the steepening of the low  $f_c$  points in L3 and the negative intercept for L3 at low  $f_c$  (Fig. B2), which is  
480 consistent with Fig. 10b. This points strongly to the differences in degree of aggregation-related smoothing between  $A_c$  and  $\mathcal{L}$  in these low  $f_c$  scenes being responsible for the flip in the relative sign of  $S_o$  vs.  $\mathcal{L}$  adjustment slopes between Fig. 10a (L2) and Fig. 10b (L3).

We note that these differences in smoothing derive from the derivations of  $A_c$  (Eq. 4) and  $\mathcal{L}$  (Eq. 3), with  $A_c$  a (non-linear) function of  $\tau$  only, and  $\mathcal{L}$  a function of the product of  $\tau$  and  $r_e$ . Anticipating how the aggregation biases play out is not intuitive,  
485 and requires, in our experience, analyses of the kind shown here.

*Author contributions.* The study was conceived by GF. GF and TG designed the study. GF performed the analysis. Simulations were done by TY. All authors were engaged in reviewing results.

*Competing interests.* GF is an Associate Editor of ACP. The authors declare no other competing interests.

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## Figure Captions

Figure 1: Theoretical calculations of the albedo susceptibility bias (in %) for a range of relative dispersions in  $\mathcal{L}$  and  $N$  ( $D(\mathcal{L})$  and  $D(N)$ , respectively), for three different correlations ( $r$ ) between  $\mathcal{L}$  and  $N$ . Note the values of large positive and negative biases.

Figure 2: As in Fig. 4 but for the  $S_o$  bias as a function of the spatial correlation between  $\mathcal{L}$  and  $N$ , for a range of relative dispersions in  $\mathcal{L}$  ( $D(\mathcal{L})$ ) and fixed  $D(N)$ . Of note is that for the conditions shown the biases are large and range from about -150 % (for positive  $r$  and large  $D(\mathcal{L})$ ) to + 70 % (negative  $r$ , large  $D(\mathcal{L})$ ).

Figure 3:  $S_o$  aggregated to a 6 km scale ( $\bar{S}_o$ ) vs.  $S_o$  aggregated to an 800 m scale ( $S_o$ ) for (a) L2, (b) L3 and (c) L2<sub>N</sub> methodology as described in the text. Solid line is the 1:1 line. Dashed lines are drawn at ordinate and abscissa values of zero. Note the general overestimate in  $S_o$  with increasing aggregation scale in (a) and (c) and the change of sign in  $S_o$  associated with L3 aggregation in (b). Of note is that high cloud fraction/low heterogeneity conditions are often associated with high biases.

Figure 4:  $S_o$  bias in % as defined in Eq. (12) as a function of the correlation between liquid water path and drop concentration  $r(\mathcal{L}, N_d)$  with points colored by cloud fraction  $f_c$  for (a) L2, (b) L3 and (c) L2<sub>N</sub> methodology. In (a) and (c) the  $S_o$  bias tends to increase with decreasing  $r(\mathcal{L}, N_d)$  and increasing cloud fraction  $f_c$ , with noted exceptions. Use of the true  $N_d$  in (c) restricts the  $S_o$  bias to about 100%.

Figure 5: As in Fig. 4 but with points colored by the ratio of  $\sigma_{A_c}/\sigma_{N_d}$  at 6 km to  $\sigma_{A_c}/\sigma_{N_d}$  at 800 m (the  $\sigma$ -ratio). Note the clear dependence of the  $S_o$  bias on the  $\sigma$ -ratio and that the bias  $\sim 0$  when the  $\sigma$ -ratio  $\sim 1$  (green colors). Exceptions to the latter occur for L2 at low  $r(\mathcal{L}, N_d)$ .

Figure 6: As in Fig. 4 but with points colored by the ratio of correlations  $r(\mathcal{L}, N_d)$  at 6 km to  $r(\mathcal{L}, N_d)$  at 800 m (the  $r$ -ratio). Comparison with Fig. 5 shows approximately orthogonal dependence of the  $\sigma$ - and  $r$ -ratios on the  $S_o$  bias. In (a) and (b), calculation of the  $r$ -ratio is mathematically unstable at  $r(\mathcal{L}, N_d) = 0$  as evidenced by saturating colors.

Figure 7: Effects of aggregation on normalized reduction in  $\sigma(A_c)$  vs. normalized reduction in  $\sigma(N_d)$  to shed light on results in Fig. 5. Points are colored by  $f_c$ . At high (low)  $f_c$  averaging tends to more significantly reduce (increase)  $\sigma(N_d)$  relative to  $\sigma(A_c)$ . Notable exceptions occur in (c) for some low  $f_c$  conditions. See text for further discussion.

Figure 8: As in Fig. 7 but with points colored by the  $r$ -ratio. See text for discussion.

Figure 9: As in Fig. 7 but with points colored by  $S_o$  bias. Biases exceeding 100 % always occur when averaging-related smoothing has a stronger effect on  $N_d$  than on  $A_c$ .

520 Figure 10:  $S_o$  bias vs. the  $\mathcal{L}$ -adjustment bias ( $d\ln\mathcal{L}/d\ln N_d$  in Eq. 1) for a limited set of model scenes (for clarity). Note a strong positive correlation between the two biases, as expected from Eq. (1). In (a) and (b), the relative magnitude of these biases depends on  $f_c$ , although with opposite trend. In (a) and (c), the  $\mathcal{L}$ -adjustment bias tends to be larger than the  $S_o$  bias.

525 Figure 11: High resolution ( $n = 1$ ) 2-D snapshot of (a) ‘true’  $N_d$ , (b) retrieved  $N_d$  (Eq. 2); (c)  $\mathcal{L}$ , and (d) the relationship between (b) and (a). Note the different scales between (a) and (b). Although the mean values of  $N_d$  in (a) and (b) differ by only -36 %, they exhibit negative correlation over the scene. The retrieved  $N_d$  introduces significant heterogeneity into the field. Use of the true  $N_d$  reduces the absolute value of the  $S_o$  error from -2017 % to +80 %.

Figure 12: As in Fig. 11 but for an open-cellular case. Note the different scales between (a) and (b). Mean values of  $N_d$  in (a) and (b) differ by -75 %. Use of the true  $N_d$  degrades the  $S_o$  bias (285 % for L2 vs. 803 % for L2<sub>N</sub>).

530 Figure A1: Demonstration of Simpson’s Paradox (Simpson, 1951) with the often-quoted relationship between ‘success at basketball’ (y-axis) and ‘height’ (x-axis). If data are separated by age, then the expected positive relationship emerges. If data are first aggregated – i.e., age is not taken into account – then a negative relationship emerges. The analogy to the current work is that aggregating blue points and red points separately and then fitting (as in L3 methodology) may explain negative  $S_o$  values.

535 Figure B1: Effects of aggregation on normalized reduction in  $\sigma(\mathcal{L})$  vs. normalized reduction in  $\sigma(N_d)$  to shed light on results in Fig. 10. Points are colored by  $f_c$ . In Fig. B1b, and at low  $f_c$ , aggregation tends to less significantly reduce  $\sigma(\mathcal{L})$  relative to  $\sigma(A_c)$  (Fig. 7b). (a) L2, (b) L3, (c) L2<sub>N</sub>. See Appendix B for further discussion.

Figure B2: Plots of  $\sigma$ -ratio (Eq. 13) for  $A_c$ ,  $N_d$  ( $\sigma$ -ratio ( $A_c$ ,  $N_d$ )) vs.  $\mathcal{L}$ ,  $N_d$  ( $\sigma$ -ratio ( $\mathcal{L}$ ,  $N_d$ )). The solid lines are best fit lines to the low  $f_c$  points. The dash-dotted line is the 1:1 line. Note that the log scale distorts the solid-line linear relationships. (a) L2, (b) L3, (c) L2<sub>N</sub>. See Appendix B for further discussion.