



1 Interpretable machine learning prediction of fire emissions and 2 comparison with FireMIP process-based models

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8 **Abstract.** Annual burned areas in the United States have increased twofold during the past decades. With more
9 large fires resulting in more emissions of fine particulate matter, an accurate prediction of fire emissions is critical
10 for quantifying the impacts of fires on air quality, human health, and climate. This study aims to construct a machine
11 learning (ML) model with game-theory interpretation to predict monthly fire emissions over the contiguous US
12 and to understand the controlling factors of fire emissions. By comparing the predicted fire PM_{2.5} emissions from
13 the interpretable ML model with the Global Fire Emissions Database (GFED) observations and predictions from
14 process-based models in the Fire Modeling Intercomparison Project (FireMIP), the ML model is also used to
15 diagnose the process-based models to inform future development. Results show promising performance for the ML
16 model, Community Land Model (CLM), and Joint UK Land Environment Simulator-Interactive Fire And Emission
17 Algorithm For Natural Environments (JULES-INFERN0) in reproducing the spatial distributions, seasonality, and
18 interannual variability of fire emissions over CONUS. Regional analysis shows that only the ML model and CLM
19 simulate the realistic interannual variability of fire emissions for most of the subregions ($r > 0.95$ for ML and
20 $r = 0.14 \sim 0.70$ for CLM), except for Mediterranean California, where all the models perform poorly ($r = 0.74$ for ML
21 and $r < 0.30$ for the FireMIP models). Regarding seasonality, most models capture the peak emission in July over
22 western US. However, all models except for the ML model fail to reproduce the bimodal peaks in July and October
23 over Mediterranean California, which may be explained by the coarse spatial resolutions of the processed-based
24 models or atmospheric forcing data or limitations in model parameterizations for capturing the effects of Santa Ana
25 winds on fire activity. Furthermore, most models struggle to capture the spring peak in emissions in southeastern
26 US, probably due to underrepresentation of human effects and the influences of winter dryness on fires in the
27 models. As for extreme events, both the ML model and CLM successfully reproduce the frequency map of extreme
28 emission occurrence but overestimate the number of months with extremely large fire emissions. Comparing the
29 fire PM_{2.5} emissions from the interpretable ML model with process-based fire models highlights their strengths and
30 uncertainties for regional analysis and prediction and provides useful insights on future directions for model
31 improvements.

32 1. Introduction

33 Large fires have increased across the United States over the past two decades, especially in the western
34 US. While the total area burned in 2020 increased by 51% compared to the 10-year average for 2010-2019, the
35 total number of fires in 2020 is smaller than the 10-year average. This indicates the contribution of larger and more
36 powerful fires to the growing burned areas (NIFC, 2020). Large fires can directly lead to property damages and
37 pose a threat to human lives (Thomas et al., 2017). Meanwhile, fine particulate matter (PM_{2.5}, particles with an



38 aerodynamic diameter smaller than and equal to $2.5 \mu\text{m}$) emitted from fires not only have negative impacts on
39 human health but also affect climate and ecosystems (Johnston et al., 2012; Ward et al., 2012; Rap et al., 2013;
40 Kaulfus et al., 2017; Liu et al., 2018; Wang et al., 2018; Stowell et al., 2019). Driven by stronger fire heating and
41 with higher injection height, aerosols emitted from large fires can be transported to broader area and stay in
42 atmosphere longer. Given the increasing trend of fire emissions, fire smokes may become the predominant source
43 of $\text{PM}_{2.5}$ in the US in the future (Yue et al., 2013; Liu et al., 2016; Ford et al., 2018). Thus, an accurate prediction
44 of fire emissions is imperative for investigating the impacts of historical and future fires on air quality, human
45 health, and climate.

46 One of the widely used methods for predicting fire emission is process-based fire parameterization. These
47 process-based models generally employ universal functions depicting non-linear relationships between fires and
48 the input variables and apply the same functions to all grid cells in a model (Pechony and Shindell, 2009; Thonicke
49 et al., 2010). In addition, the parameters of the process-based model are usually determined by empirical or
50 statistical functions, assuming that the same parameters apply to all the regions or regions with limited fire
51 observations (Crevoisier et al., 2007; Parisien et al., 2016). Process-based models are usually included in the
52 dynamic global vegetation models (DGVMs) to simulate fire dynamics, vegetation dynamics, and biogeochemistry
53 driven by atmospheric forcing and socio-economic data (Li et al., 2013; Knorr et al., 2016). Fire emissions,
54 including trace gases and aerosols, are calculated from the simulated fire carbon emissions and the emission factors,
55 with the former computed as the product of the burned area, fuel load, and combustion completeness. The process-
56 based models in DGVM coupled with other components of Earth system models can be used to assess the impacts
57 of environmental factors on fires and the feedback between fire emissions, land processes, and climate (Kloster et
58 al., 2010). In 2014, the Fire Model Intercomparison Project (FireMIP) was initiated to compare nine DGVMs that
59 include fire modules to better understand the performances of the global fire models (Rabin et al., 2017). The
60 FireMIP enables comprehensive evaluation and comparison across various process-based models and provides a
61 dataset of long-term fire simulations for regional and global analysis (Li et al., 2019; Hantson et al., 2020).

62 Besides process-based fire models, data-driven statistical models are also commonly used to estimate fire
63 activities using relationships between fires and predictor variables. Multiple linear regression (MLR) is a popular
64 simple statistical method used for fire modeling (Spracklen et al., 2009; Morton et al., 2013; Urbieto et al., 2015;
65 Williams et al., 2019). MLR can achieve a good performance, but it fails to capture the non-linear relationships
66 between fires and predictors, and it is sensitive to the collinearity and combinations of predictors (Littell et al.,
67 2009). Unlike MLR, machine learning (ML) is a novel tool for advancing fire modeling, given its strengths in
68 resolving the complex relationships between the target and predictor variables. Different ML approaches have been
69 used to estimate fire occurrence, burned areas, or emissions at various time scales and spatial scales (Cortez and
70 Morais, 2007; Aldersley et al., 2011; Dillon et al., 2011; Birch et al., 2015; Kane et al., 2015; Coffield et al., 2019;
71 Wang and Wang, 2020). Even though ML models generally achieve higher accuracy than simple statistical models,
72 their decision processes are often inscrutable, and hence lack interpretability. The development of explainable ML
73 represents major advances for scientific applications beyond predictions (Gunning, 2017; Barredo Arrieta et al.,
74 2020). For example, Wang et al. (2021) used the Extreme Gradient Boosting (XGBoost) algorithm and Shapley
75 Additive explanation (SHAP) to predict wildfire burned area and revealed the relationships between burned areas
76 and predictor variables. As process-based and data-driven models have their own advantages and weaknesses, as
77 listed in Table 1, comparing these models and assessing their uncertainties in historical simulations and future
78 projections are important. Yue et al. (2013) applied an MLR and a parameterization method to estimate burned
79 areas in ecoregions of the western US and found that both models explained $\sim 50\%$ of the variance in the observed
80 burned areas. Although they compared the burned areas estimated by the two methods and quantified their
81 uncertainties in fire projections, both methods are only driven by meteorology while the effects of fuels and human
82 activities are not considered.



83 The FireMIP dataset provides long-term simulations of multiple DGVMs with fire modules, allowing
84 comparisons between process-based and data-driven models, with all models considering all the potential factors
85 influencing fires, including climate, weather, vegetation, and human activities. This study aims to develop an ML
86 model with game theory interpretation for fire emission prediction and to understand controls of fire emissions.
87 The interpretable ML model is then used to reveal the important factors controlling fire emissions and diagnosis
88 the process-based FireMIP models. The ML model predicts the monthly $\text{PM}_{2.5}$ emissions from fires during 2000-
89 2020 at a spatial resolution of $0.25^\circ \times 0.25^\circ$ over the contiguous US (CONUS). It uses the XGBoost algorithm and
90 incorporates various predictors, including local and large-scale meteorology, land surface characteristics, and
91 socioeconomic variables, which are common input variables also used by the FireMIP models while some are
92 specifically related to fire activities in CONUS. The ML model and FireMIP models are optimized using different
93 data or predictors at various scales, which enables us to use the ML to diagnose the performance of FireMIP models
94 over CONUS through the comparisons of their performances and variable importance from the ML model. We
95 evaluate and compare the predicted fire emissions from the ML and FireMIP models against the GFED fire
96 emission product, focusing on spatial distributions, seasonality, and interannual variability over selected regions in
97 CONUS. Additionally, the ML model and the SHAP importance are used to identify the important drivers of fire
98 emissions in different regions and compare them with the corresponding parameterizations in the process-based
99 models. Lastly, we compare the process-based and ML model performances in simulating extremely large fire
100 emissions, including the spatial distributions of frequency and two case studies.

101 2. Data

102 2.1 Fire-induced $\text{PM}_{2.5}$ emission data

103 Monthly fire $\text{PM}_{2.5}$ emission data is obtained from the Global Fire Emissions Database (GFED). GFED
104 version 4 provides monthly burned area at 0.25° spatial resolution from 1997 to present, based on a combination of
105 the MODIS burned area product with active fire data from the Tropical Rainfall Measuring Mission (TRMM)
106 Visible and Infrared Scanner (VIRS) and Along-Track Scanning Radiometer (ATSR) family of sensors (Giglio et
107 al., 2013). The GFED fire $\text{PM}_{2.5}$ emissions are estimated by combining the burned area boosted by small fire burned
108 area (Randerson et al., 2012) and the emission factor data with a revised version of the Carnegie-Ames-Stanford
109 Approach (CASA) biogeochemical model that estimates fuel loads and combustion completeness for each monthly
110 time step (van der Werf et al., 2017). We use the GFED fire $\text{PM}_{2.5}$ emission as the target variable in the machine
111 learning model development and for model evaluation.

112 To reduce spatial heterogeneity and help model learning, we apply the inverse distance weighting (IDW)
113 (Bartier and Keller, 1996; Shepard, 1968) to interpolate the monthly gridded fire $\text{PM}_{2.5}$ emission at $0.25^\circ \times 0.25^\circ$.
114 The IDW method determines the value at a grid cell as the weighted average of the surrounding values within a
115 search distance, with the weights proportional to the inverse of the distance raised to the power value p . Here we
116 choose a value of 1 for p and a search distance of 35 km for IDW processing. Note that the total fire emitted $\text{PM}_{2.5}$
117 within a search distance after IDW processing is constrained to be the same as the original data. In this study, we
118 only include grids with more than eight months of fire emissions larger than zero (in a total of 250 months),
119 encompassing 90% of the total fire emissions and ensuring sufficient data for the XGBoost model training. The
120 interpolated fire emission is normalized based on its 21-year mean and standard deviation for each grid to reduce
121 the skewness and improve data symmetry.



122 2.2 Predictor variables

123 We develop an empirical model at $0.25^\circ \times 0.25^\circ$ grid resolution driven by various predictor variables at a
124 monthly scale from January 2000 to October 2020. Given the datasets have different spatial resolutions, all the
125 predictor variables are resampled to the spatial resolution of $0.25^\circ \times 0.25^\circ$ by linear interpolation. The predictor
126 variables used in the model along with their original spatial and temporal resolutions are included in Table 2. Most
127 variables were also used in Wang et al. (2021) for developing an ML model of fire burned area over the contiguous
128 U.S.

129 **Local meteorology:** Same as the local meteorological predictors used in Wang et al. (2021), we include monthly
130 data of mean surface temperature, relative humidity (RH) at 2 m, daily precipitation, zonal (U) and meridional (V)
131 components of wind at 10 m from the North American Regional Reanalysis (NARR) (Mesinger et al., 2006) and
132 1000-hour dead fuel moisture (FM1000), Energy Release Component (ERC), and vapor pressure deficit (VPD)
133 from the gridMET dataset (Abatzoglou and Kolden, 2013; Coffield et al., 2019). Drought is a natural phenomenon
134 that influences fires through ignition efficiency, fuel availability, and fuel moisture. Thus, we include the monthly
135 Standardized Precipitation Evapotranspiration Index (SPEI), a multiscale drought index based on climatic data
136 (Vicente-Serrano et al., 2010). Given that lightning is one of the major ignition sources of fires and makes up
137 approximately 75% of burned areas in western US (Pyne, 1984; Stephens, 2005), in this study, we add the cloud-
138 to-ground (CG) lightning flash density from Severe Weather Data Inventory (SWDI) based on the National
139 Lightning Detection Network (NLDN) (Cummins and Murphy, 2009; NOAA, 2006). The daily number of CG
140 lightning flashes is summarized in 0.1° tiles and we aggregate the daily data to monthly scale.

141 **Large-scale meteorological patterns:** Large-scale meteorological patterns at a synoptic scale have been found to
142 link to large fire events (Crimmins, 2006; Trouet et al., 2009; Zhong et al., 2020; Dong et al., 2021). Furthermore,
143 it has been shown that including predictors of large-scale meteorological patterns conducive to wildfires
144 significantly improves the prediction of burned areas over CONUS (Wang et al., 2021). Thus, we follow the
145 methods developed by Wang et al. (2021) using the singular value decomposition (SVD) method to construct
146 predictors representing the synoptic patterns driving fire emission variability. Note that the only difference between
147 Wang et al. (2021) and this study is that they used wildfire burned area data and we use fire emissions to construct
148 the SVDs. Three regions where large fires periodically occur are selected for constructing SVDs: Northern
149 California, southern Rocky Mountains, and southeastern US, as defined in Wang et al. (2021). For each region, we
150 calculate the daily mean fire $PM_{2.5}$ emissions over the region and compute the day-to-day correlations between the
151 regional mean fire $PM_{2.5}$ emissions and the five gridded daily meteorological variables (surface temperature, 2-
152 meter RH, U-wind and V-wind at 850 hPa, and geopotential height at 500 hPa) for all $1^\circ \times 1^\circ$ grid cells within the
153 large-scale domain, giving a correlation map for each meteorological variable. The correlation maps are then used
154 to derive the SVD modes representing the large-scale meteorological patterns related to fires. Finally, we compute
155 the monthly standard deviation of the daily SVD time series for the first two SVD modes, representing the month-
156 to-month variations of synoptic fluctuations and atmospheric instability. The detailed methods and discussions
157 about the SVDs are provided in Wang et al. (2021). Overall, the identified SVDs for the three regions are similar
158 to the SVDs in Wang et al. (2021) calculated using wildfire burned areas (Figs. S1-3).

159 **Land-surface properties:** We use the same set of variables in the burned area model that represent the effects of
160 fuel and land surface states on fire emissions, including evapotranspiration (ET), surface soil moisture, land types,
161 and topography (Wang et al., 2021). Monthly mean ET, vegetation fraction, and surface soil moisture are obtained
162 from the North American Land Data Assimilation System (NLDAS-2) (Xia et al., 2012). Land cover data of the
163 LAI classification scheme is obtained from the Terra and Aqua combined MODIS Land Cover Climate Modeling



164 Grid (CMG) Version 6 data (Friedl, 2015). Since the land cover data is at yearly intervals from 2001 to 2020, we
165 use the land cover data of 2001 for 2000. Topography data of slope and elevation is obtained from Amatulli et al.
166 (2018).

167 Besides the above-mentioned variables that were also used in Wang et al. (2021), in this study, we consider
168 the effect of fuel load on fire emissions, since fuel load is critical to fire emissions through its controls on fuel
169 consumption and burned areas (Parks et al., 2012; Liu and Wimberly, 2015). As there are limited observations of
170 fuel load, we use LAI to approximate the canopy bulk density and vegetation fraction to represent the existing
171 amount of vegetation. LAI is taken from MODerate resolution Imaging Spectroradiometer (MODIS) instruments
172 (Myneni et al., 2015) and vegetation fraction is obtained from the NLDAS-2. We also include fuel load simulated
173 by Community Land Model (CLM). Monthly fuel load data from 2000 to 2015 is obtained from a simulation by
174 CLM version 5 with biogeochemistry and prognostic crop, driven by atmospheric forcing from GSWP3v1
175 (Lawrence et al., 2019). The fuel load after 2015 is taken from a simulation under the SSP3 (shared socioeconomic
176 pathways) scenario. Additionally, we include normalized fuel load as a predictor to capture the effects of temporal
177 variation of fuel load, as the influence of fuel load on fire emissions is mainly attributed to its spatial variation
178 rather than the temporal variation (Lasslop and Kloster, 2015).
179

180 **Socioeconomic variables:** We use population density and gross domestic product (GDP) per capita to represent
181 human effects on wildfires. The population density data is obtained from the Gridded Population of the World data
182 collection (GPW V4) for the years 2000, 2010, 2015, and 2020, with a spatial resolution of 30 arc-second (CIESIN-
183 Columbia University, 2017). The populations in other years are linearly interpolated between the abovementioned
184 four years. The GDP per capita is taken from a gridded global dataset for 2000-2015 with a spatial resolution of 5
185 arc minutes (Kummu et al., 2018). For the GDP after 2015, we use the data of 2015.

186 3. Description of fire emission models

187 3.1 Process-based fire emission models

188 The Fire Model Intercomparison Project (FireMIP) includes a set of common fire modeling experiments
189 from nine DGVMs driven by the same forcing data, allowing a better understanding of global fire models (Rabin
190 et al., 2017). The FireMIP dataset provides global gridded burned area fraction and fire emissions, including carbon
191 and 33 species of trace gases and aerosols over 1700-2012. Nine DGVMs with different fire modules are included
192 in FireMIP, including Community Land Model version 4.5 (CLM4.5) with the CLM5 fire module, Canadian
193 Terrestrial Ecosystem Model (CTEM), Jena Scheme for Biosphere-Atmosphere Coupling in Hamburg with Spread
194 and InTensity fire model (JSBACH-SPITFIRE; hereafter referred to as JSBACH), Joint UK Land Environment
195 Simulator with Interactive Fire And Emission Algorithm For Natural Environments (JULES-INFERNO; hereafter
196 referred to as JULES), Lund-Potsdam-Jena General Ecosystem Simulator with Global FIRE Model (LPJ-GUESS-
197 GlobFIRM; hereafter referred to as LPJ-Glob), LPJ-GUESS with SIMple FIRE model and Blaze-Induced Land-
198 Atmosphere Flux Estimator (LPJ-GUESS-SIMFIRE-BLAZE; hereafter referred to as LPJ-SIM), LPJ-GUESS with
199 SPITFIRE model (LPJ-GUESS-SPITFIRE; hereafter referred to as LPJ-SPI), MC2, and Organizing Carbon
200 Hydrology In Dynamic Ecosystems with SPITFIRE model (ORCHIDEE-SPITFIRE; hereafter referred to as
201 ORCHIDEE) (Rabin et al., 2017).

202 The nine DGVMs in FireMIP are driven by the CRU-NCEP v5.3.2 atmospheric forcing data with a spatial
203 resolution of 0.5° and a 6-hourly temporal resolution (Wei et al., 2014; Rabin et al., 2017). Other forcing data,



204 including annual global atmospheric CO₂ concentration, land use and land cover, and population density from 1700
205 to 2012 is taken from various data sources (Klein Goldewijk et al., 2010; Hurtt et al., 2011; Le Quéré et al., 2014).
206 Monthly cloud-to-ground lightning frequency with a resolution of 0.5° × 0.5° over 1901-2012 is calculated based
207 on the observed relationship between present-day lightning and convective available potential energy (CAPE)
208 anomalies (Pfeiffer et al., 2013). Fire emissions in FireMIP are calculated considering the fire carbon emissions
209 and vegetation characteristics based on the plant functional type (PFT) from the FireMIP historical transient control
210 run (SF1). SF1 breaks the simulation period into three phases: the spin-up phase in 1700, the transient phase in
211 1701-1900, and the transient phase in 1901-2012 (see the detailed descriptions and model settings in Rabin et al.,
212 2017, Li et al., 2019, and Hantsan et al., 2020). In the 1901-2012 transient phase, the models are driven by time-
213 varying atmospheric forcing, CO₂ concentration, LULCC, population density, and lightning data. Note that the
214 MC2 and CTEM runs start from 1901 and 1861, while the rest of the models start from 1700. As the spatial
215 resolutions of the FireMIP models are different, the regridded model outputs with 1° × 1° resolution obtained from
216 Li et al. (2019) are used to compare with the GFED data and the ML model.

217 3.2 ML-based approach: An eXtreme Gradient Boosting (XGBoost) model

218 The eXtreme Gradient Boosting (XGBoost) is a decision-tree-based ensemble machine learning method
219 using the gradient boosting approach (Chen and Guestrin, 2016). The XGBoost model builds multiple decision
220 trees that are added subsequently and learn the errors of the previous tree to reduce the loss and obtain the best
221 prediction. Unlike the gradient boosting machine (GBM) that also uses the gradient boosting approach, XGBoost
222 utilizes a more regularized model formalization to prevent over-fitting and improve the computational efficiency.
223 The formula for the prediction at step t and grid location i can be defined as follows:

$$224 \hat{y}_i^t = \sum_{k=1}^t f_k(x_i) = \hat{y}_i^{(t-1)} + f_t(x_i)$$

225 where $f_t(x_i)$ is the tree model at step t , \hat{y}_i^t and $\hat{y}_i^{(t-1)}$ are the predictions at steps t and $t-1$, and x_i are the predictor
226 variables. The parameters of the model $f_t(x_i)$ are selected by optimizing the objective function that measures how
227 well the model fit the training data:

$$228 \text{Obj}^t = \sum_{i=1}^n L^t + \Omega^t$$

229 which is composed of the loss function L^t and the regularizing term Ω^t in each step. L_t is defined as $l(y_i, \hat{y}_i^{t-1} +$
230 $f_t(x_i))$ and Ω^t is defined as $\gamma T + \frac{1}{2} \lambda \|\omega\|^2$, where γ is the regularization term which penalizes the number of
231 leaves in the tree T and λ is the regularization term which penalizes ω , the weights of different leaves.

232 We use grid search to choose the set of suitable hyperparameters and achieve the best ML model
233 performance. Grid search is a tuning technique for computing the optimal values of hyperparameters considering
234 a range of numbers with a given increment. The parameter set that yields the best 5-fold cross-validation score is
235 selected as the final set of hyper-parameters. The considered hyper-parameters, their search domains, and the final
236 values are denoted in Table S1.

237 The 10-fold cross-validation (CV) technique is applied to evaluate the model and avoid overfitting. First,
238 we randomly divide the fire emission dataset (2000-2020 over CONUS) into ten equal-sized splits. Then, we train
239 the model with nine splits of the data and use the trained model to predict fire emissions for the remaining one split.



240 This process is repeated ten times for each split. Finally, the predictions are evaluated by grids and regions using
241 root mean square error (RMSE), correlation coefficient (R), and the index of agreement (IoA). The IoA represents
242 the ratio of the mean square error and the potential error, and the value closer to 1 indicates better agreement.

243 3.3 Shapley additive explanations (SHAP)

244 We utilize the SHAP to identify the relative importance of the predictor variables. SHAP is a novel
245 approach to resolve and explain variable importance based on game theory (Lundberg and Lee, 2017). Within the
246 scope of game theory, the goal is a prediction for a single observation. Each predictor variable is referred to as a
247 “player” in this game and contributes to the goal (“payout”). For each predictor, the SHAP variable importance
248 measures the marginal contribution considering all possible combinations of the predictor variables. The marginal
249 contribution is calculated by comparing the differences between the model fit $f_x(S \cup \{i\})$ including the predictor i
250 and another model fit $f_x(S)$ without predictor i . When there is more than one predictor i , the marginal contribution
251 also depends on the interactions with other predictors. Thus, the calculation repeats considering the whole set of
252 the predictors. The final contribution ϕ_i of predictor i is the weighted average of all marginal contributions:

$$253 \quad \phi_i = \sum_{S \subseteq F \setminus \{i\}} \frac{|S|!(F - |S| - 1)!}{F!} [f_x(S \cup \{i\}) - f_x(S)]$$

254 where F is the total number of features, S is the subset of predictors from all predictors except for predictor i ,
255 $\frac{|S|!(F - |S| - 1)!}{F!}$ is the weighting factor counting the number of permutations of the subset S . $f_x(S)$ is the expected
256 output given the predictors subset S . $[f_x(S \cup \{i\}) - f_x(S)]$ is the difference made by predictor i .

257 Compared to the commonly used feature importance, such as gain, or split count, SHAP is more consistent
258 and faithful to the model (Lundberg et al., 2019). More importantly, SHAP provides local importance that measures
259 the variable importance for each sample, while most of the feature importance metrics only have global importance
260 that measures variable contributions limited to the entire dataset. The global importance by SHAP is the average
261 of the absolute SHAP values for each predictor, providing an overall picture of the predominant variables
262 controlling fire emissions in CONUS. The local importance will be used to identify the important predictors for
263 large fire events in the ML model and diagnose the deficiency of the process-based models.

264 4. Results

265 4.1 XGBoost model performance and variable importance

266 Table 3 shows the whole CONUS and regional model performance, including RMSE, IoA, and correlation.
267 The model performs well at grid level over CONUS, with an RMSE of 0.16 g/m² and an IoA of 0.84. Figure 1a
268 shows the map of correlation between the observed and predicted monthly fire emission time series for each grid
269 over CONUS. Overall, the results indicate the ML model can reproduce the interannual variability of fire emissions
270 at 0.25° resolution over CONUS, with a mean correlation of 0.58 and more than 70% of the grids having
271 correlations larger than 0.4. To better assess model performance in different regions, Table 3 summarizes the model
272 performance for several selected regions: (1) western forest area, (2) Mediterranean California, (3) southwestern
273 US, and (4) southeastern US (color boxes in Fig. 1a). The regions where fires frequently occur are selected by the
274 similarity of ecoregions, vegetation types, and fire regimes. Figs. 1b-e show the time series of observed and
275 predicted fire PM_{2.5} emissions averaged over several regions. Generally, the ML model reproduces the interannual



276 variability of fire emissions for the selected regions ($r=0.84-0.98$). Among these regions, Mediterranean California
277 has the smallest correlation coefficient and largest RMSE compared to other regions, which can be explained by
278 the fact that fires in this region interact with multiple factors, including human activity, complex terrain, and Santa
279 Ana winds (Syphard et al., 2008; Yue et al., 2014). The interactions between fires and these factors pose
280 uncertainties and challenges in fire prediction over this region. It is also worth noting that the ML model captures
281 the large fire events in September 2020 in Oregon and California but underestimates the peak values by $\sim 30\%$
282 (Figs. 1b and 1c).

283 To improve understanding of the ML prediction, we utilize the SHAP method to quantify the contributions
284 of each predictor variable to the prediction and identify the key contributing factors of fire $PM_{2.5}$ emission. Fig. 2
285 shows the 20 most important variables for the model ranked by the absolute mean SHAP values. Among the top
286 10 variables, seven of them are local meteorological variables, indicating local meteorology is the predominant
287 control of fire emissions, as these variables control fire activity directly (Liu and Wimberly, 2015; Abatzoglou et
288 al., 2016; Wang et al., 2021). Besides local meteorology, the predictors of large-scale meteorology (SVD1_SElag2
289 and SVD2_SElag2) are identified as the eighth and tenth important variables, showing that meteorology is not only
290 important at local scale but also at synoptic scale (Trouet et al., 2009; Pollina et al., 2013; Dong et al., 2021).
291 Finally, in addition to meteorology, fuel load is identified as the fifth important variable in the model, as fuel load
292 affects emission through controlling burned area and fuel consumption (Seiler and Crutzen, 1980). Considering the
293 important variables in different regions, the selected regions in western US (western forest area, Mediterranean
294 California, and southwestern US) generally share the common top 10 variables (Fig. S4). Over western US,
295 predictors controlling fuel dryness and fuel amount, including RH, fuel moisture (FM1000), ERC, vegetation
296 fraction, and fuel load, contribute more to fire emissions. On the other hand, large-scale meteorological patterns
297 (SVDs_SElag2) are more important for fire emissions in southeastern US.

298 4.2 General comparison between GFED, ML, and FireMIP models

299 This section compares the performance of the ML and FireMIP models benchmarked against observations
300 from GFED, and the evaluations are based on spatial distributions, seasonality, and interannual variability of fire
301 $PM_{2.5}$ emissions. Since the spatial resolutions of the GFED data, ML models, and FireMIP models are different,
302 they are all regridded to $1^\circ \times 1^\circ$ using bilinear interpolation. Note that the simulation period of FireMIP models
303 ends in 2012, so we use the overlapping period of 2000-2012 for comparison and exclude the MC2 model because
304 its simulation ends in 2008.

305 4.2.1 Spatial distributions of fire $PM_{2.5}$ emissions and sensitivities to RH and temperature

306 Fig. 3 compares the observed and simulated spatial distributions of long-term mean monthly fire $PM_{2.5}$
307 emissions averaged over 2000-2012. Among the models, the ML model, CLM, and JULES have better performance
308 in reproducing the spatial distributions of fire emissions over CONUS, with a correlation coefficient of 0.83, 0.52,
309 0.40, respectively. The ML model shows the best agreement with GFED, though it overestimates fire emissions
310 over Northern California. Both CLM and JULES simulate more $PM_{2.5}$ emissions over southeastern US, and JULES
311 overestimates fire emissions in Northern California. Some other models, such as CTEM, JSBACH, and LPJ-SIP,
312 tend to overestimate fire emissions over central US (e.g., Great Plains and Texas). LPJ-SIP captures the hotspots
313 of fire emissions over western US and southeastern US, but it simulates much more $PM_{2.5}$ emissions over the Rocky
314 Mountain and northeastern US. In terms of the total amount of $PM_{2.5}$ emissions, all models except ORCHIDEE-
315 SPITFIRE overestimate $PM_{2.5}$ emissions (8.33-79.49 Tg), compared to the GFED estimate of 4.98 Tg during 2000-
316 2012 over CONUS (Table 4).



317 The overestimations in some models may be explained by the sensitivities of fire emissions to individual
318 meteorological variables. Fig. 4 shows the slopes for the dependence of annual mean fire $\text{PM}_{2.5}$ emissions on annual
319 mean RH from the CRUNCEP atmospheric forcing data for GFED and the ten models based on linear regression.
320 Since the ML model uses NARR meteorology as predictors, we also include sensitivities of the fire emissions
321 predicted by the ML model to the NARR meteorology (Fig. 4b). Almost all models capture the negative dependence
322 of $\text{PM}_{2.5}$ emissions on RH over western US ($r=-0.06\sim 0.84$), but the sensitivities in the models are much stronger
323 (steeper negative slope) than in GFED. For temperature, positive sensitivity is shown over western US in GFED
324 (Fig. 5), with the most significant slope in northern California. The sensitivities to temperature in models agree
325 with the observed sensitivities ($r=-0.06\sim 0.64$), but some models show much stronger sensitivities over western,
326 central, and southeastern US. Generally, the spatial distributions of the long-term mean fire emissions shown in
327 Fig. 3 match well with the spatial distributions of sensitivities to RH or temperature, suggesting an important role
328 of the sensitivities in the model biases of predicting fire emissions. However, the correspondence of large fire
329 emissions to the sensitivities to RH or temperature shows regional differences. For instance, in western US, the
330 stronger sensitivities to both RH and temperature correspond to the overestimations in this region for most models,
331 including the ML model, CLM, CTEM, JSBACH, JULES, LPJ-SIM, and LPJ-SPI (Figs. 4 and 5). On the other
332 hand, over central US, larger $\text{PM}_{2.5}$ emissions simulated by CTEM and JSBACH only correspond to stronger
333 sensitivity to temperature (Fig. 5). Similar to central US, in southeastern US, the overestimations in CLM and
334 JULES only correlate with stronger sensitivity to temperature (Fig. 5). Regional differences in the correspondences
335 between the predicted fire emissions and their sensitivity to meteorology can be explained by several factors. For
336 western US, the overestimations of fire emissions correspond to both stronger sensitivities to RH and temperature,
337 given that fire activities are sensitive to fuel aridity that is controlled by temperature and fuel moisture (Abatzoglou
338 and Williams, 2016; Holden et al., 2018). As for southeastern US, fuels in this region typically burn at higher RH
339 and the interannual RH variation (standard deviation) is smaller (Balch et al., 2017; Brey et al., 2018). With higher
340 RH values and less variation in RH, the fire emissions in southeastern US show weaker sensitivity to RH than to
341 temperature in observation (Table S2). The above analysis shows that the overestimation of fire emissions in the
342 models may be attributed to the stronger sensitivities to meteorology. However, fire activities are controlled by
343 meteorology and other factors such as vegetation and human, so the analysis of fire emission sensitivity to
344 meteorology only provides a potential explanation to the overestimation of fire emissions in the models (Forkel et
345 al., 2019).

346 4.2.2 Seasonality and interannual variability over CONUS

347 In addition to evaluating spatial distributions, it is also important to compare the models' ability to
348 reproduce the temporal variability of fire emissions. As the models may systematically over-or underestimate fire
349 emissions, we normalize the emissions by the mean and standard deviation and focus only on its temporal
350 variability. Fig. 6a shows the seasonality of normalized fire $\text{PM}_{2.5}$ emission over CONUS. Most models capture the
351 seasonality of fire emission successfully ($r>0.85$), except LPJ-SIM which simulates peak emission in August-
352 October ($r=0.65$). Among the models, the ML model has the highest correlation coefficient between prediction and
353 observation from GFED ($r=0.98$) and successfully reproduces the peak in August. The seasonal peaks simulated
354 by the FireMIP models are broader and flatter than the peak in GFED, with an early peak in June-July continuing
355 to September (Fig. 6a).

356 In terms of interannual variability (Fig. 6b), the ML model, CLM, and JULES perform better than other
357 models, with larger correlation coefficient between simulated and observed fire $\text{PM}_{2.5}$ emissions ($r=0.87, 0.71,$ and
358 0.55 for ML, CLM, and JULES, respectively; Table 4). Other models have relatively poor performance in capturing
359 the interannual variability. The interannual variability of fire emissions shows several peaks in 2002, 2007, and
360 2012 (black line in Fig. 6b), when western US contributes 76% of the total emissions to the peaks in these years.



361 Almost all models except ORCHIDEE capture the peak in 2012. However, most models miss the peaks in 2002
362 and 2007. Among all models, LPJ-Glob model simulates the peaks in the two years, while ML, JULES, and CLM
363 only capture the largest emission in 2007 (Fig. 6b).

364 4.2.3 Seasonality and interannual variability by regions

365 As the temporal variability of fire activities varies by region, we compare the performance between GFED
366 and the ML and FireMIP models by the regions defined in Sec. 4.1. Fig. 7 shows the seasonality and interannual
367 variability of normalized fire $PM_{2.5}$ emission over western forest area, Mediterranean California, southwestern US,
368 and the southeastern US. All models generally capture the seasonality of the western forest area peaking in summer,
369 with correlation coefficients larger than 0.8 (Table 4). Even though the FireMIP models generally reproduce the
370 peaks in summer, the predicted peaks are broad and flat, indicating a relatively longer fire season starting in June
371 and ending in September (Fig. 7a). When looking at the interannual variability, we find that the ML model has the
372 best performance with a correlation coefficient of 0.93, and it successfully captures the largest fire emission in
373 2007. CLM, JULES, and LGJ-Glob perform better than the rest of the models ($r=0.70, 0.60,$ and 0.51 for CLM,
374 JULES, and LPJ-Glob, respectively; Table 4), but all of them still miss the peaks in 2007 and overestimate fire
375 emissions in 2001 and 2003 (Fig. 7b). The emission peak in 2007 is mainly attributed to the large fires in Idaho,
376 which were associated with synoptic weather patterns characterized by positive geopotential height and temperature
377 anomalies over the Pacific Coast and western US (Zhong et al., 2020). Consistent with prior findings, SHAP
378 importance shows that in the ML model SVD predictors (SVD_NCA and SVD_RM in July and August 2007 Fig.
379 8a) are the dominant factors of fire emissions in 2007 (contribute 27% and 28% for July and August 2007,
380 respectively), which are characterized by high pressure, low RH, and northeasterly winds over western US (Figs.
381 S1 and S2). Thus, the underestimation of peak emission in 2007 may be explained by the fact that the influences
382 of large-scale meteorology on fire activity are not fully considered in the FireMIP models, which are point models
383 driven only by local atmospheric forcing.

384 In Mediterranean California, the seasonality of fire emissions shows a bimodal pattern, peaking in August
385 and October. The peak in October is mainly due to the extremely large fires associated with Santa Ana winds in
386 2003 and 2007 (Keeley et al., 2009; Yue et al., 2014). The ML model simulates a flatter peak from July to October,
387 while all the FireMIP models except ORCHIDEE capture the first emission peak in summer but fail to simulate the
388 large fire emission in October (Fig. 7c). The underestimation associated with the Santa Ana winds is also shown in
389 the interannual time series in Fig. 7d. Several models, including LPJ-Glob, CTEM, LPJ-SPI, and JULES, capture
390 the peak in 2007 but only the ML model predicts both peaks in 2003 and 2007 even though the peak in 2003 is
391 underestimated. According to the SHAP importance from the ML model, the peak emissions in October 2003 and
392 October 2007 are mainly contributed by the SVD predictors and ERC (SVD2_NCA and SVD1_RM together
393 contribute 20% to the fire emissions for October 2003 and SVDs_SElag2 and SVD2_RM together contribute 31%
394 to the fire emissions for October 2007) and ERC (15% and 18% for October 2003 and 2007, respectively) (Fig.
395 8b). The results indicate that the ML model captures the effect of synoptic weather patterns on fire activity by
396 including the SVD predictors. Even though the wind speed is included in simulating fire spread in the FireMIP
397 models, the spatial resolutions of the models and/or the atmospheric forcing data may not be fine enough to resolve
398 the strengthened offshore winds through the complex terrain, and subsequently, they may not well capture the
399 effects of Santa Ana winds on fires. Besides the above-mentioned shortfall, all the models have problems
400 reproducing the interannual variability of the fire emissions over Mediterranean California, with very low
401 correlations ($r<0.25$) for the FireMIP models and a relatively low correlation ($r=0.72$) for the ML model (Table 4;
402 Fig. 7d). The poor performance for this region may be due to the complex relations between fires and multiple
403 factors, including meteorology, complex terrain, fuel, and human, which may not be fully represented in the models
404 (Mann et al., 2016; Radeloff et al., 2018).



405 Both the ML model and LPJ-SIM successfully reproduce the seasonality of fire emission in southwestern
406 US peaking in June ($r=0.99$ and 0.94 for ML and LPJ-SIM, respectively), while other models simulate relatively
407 smooth seasonality (Fig. 7e and Table 4). The ML model, LPJ-SIM, and ORCHIDEE have better performance for
408 the interannual variability, with correlation coefficients of 0.95 , 0.40 , and 0.45 , respectively (Table 4). However,
409 most FireMIP models show larger variability in fire emissions than the GFED, and they all fail to capture the
410 extremely large fire emission in 2011 (Fig. 7f). The peak fire emission in 2011 over southwestern US was caused
411 by extremely low atmospheric moisture along with moderately high temperature, leading to record-breaking VPD
412 and wildfire activities (Williams et al., 2015). To explain why the FireMIP models fail to capture the peak of 2011,
413 we compare the VPD calculated from CRUNCEP data and the VPD data from gridMET used in the ML model. As
414 Fig. S5 shows, CRUNCEP shows smaller positive anomalies of VPD over southwestern US in 2011 summer, while
415 gridMET data demonstrates a significantly larger VPD anomaly. The biases in CRUNCEP data may partially
416 explain the underestimations in all FireMIP models.

417 For southeastern US, the seasonal cycle of fire $PM_{2.5}$ emissions displays a bimodal pattern, peaking in
418 spring (March-April) and fall (September and October) (Fig. 7g). Most models fail to reproduce the bimodal fire
419 emissions, but the ML model, LPJ-SIM, and LPJ-SPI can capture the bimodal pattern. Although LPJ-SIM and LPJ-
420 SPI predict the bimodal peaks, the first peak simulated by LPJ-SIM shows a one-month delay, and the second peak
421 simulated by LPJ-SIM and LPJ-SPI is one month early and one month late, respectively (Fig. 7g). In addition, the
422 ML model, CLM, and JSBACH reproduce the interannual variability of fire $PM_{2.5}$ emissions relatively well ($r=0.96$,
423 0.57 , and 0.72 for the ML model, CLM, and JSBACH, respectively) (Table 4 and Fig. 7h). Interestingly, CLM and
424 JSBACH can capture several peaks in 2007, 2010, and 2011 but they do not simulate seasonality correctly, which
425 may be explained by the underestimation in spring compensated by the overestimations in summer related to
426 abnormal dryness or drought.

427 4.3 Performance in modeling extreme events

428 Fire activity in the US is becoming more hazardous, particularly over western US, due to more frequent
429 hotter and drier conditions as climate continues to warm (Williams et al., 2019). Thus, it is necessary to assess
430 whether the ML model and process-based models can capture the extreme events in terms of their magnitude,
431 frequency, timing, and location, which is essential to future projection and adaptation. As CLM performs relatively
432 well among the FireMIP models, we select CLM for comparison with the ML model at $1^\circ \times 1^\circ$ resolution, focusing
433 on the spatial patterns of extreme event frequency and two case studies with extremely large fire emissions.

434 4.3.1 Frequency of extreme event occurrence

435 Fig. 9 shows the frequency maps of months with large fire emissions during 2000-2012 for GFED, the ML
436 model, and CLM. Large fire emission is defined as monthly fire $PM_{2.5}$ emissions greater than the 95th percentile of
437 fire $PM_{2.5}$ emission considering all the grids over CONUS in 2000-2012. GFED shows hot spots with a higher
438 frequency over northern California, the Pacific Northwest, and southeastern US, with total counts ranging from 15
439 to 105 (Fig. 9a). The ML model captures the spatial patterns ($r=0.74$), but it overestimates the number of months
440 by a factor of two to three compared to GFED, especially over western US (Fig. 9b). The spatial patterns of large
441 fire emission occurrence simulated by CLM are generally consistent with the observed distribution by GFED
442 ($r=0.35$). However, it overestimates the frequency, particularly over Idaho and northeastern US, and simulates more
443 significant numbers of months with extreme events over large spatial extents, may be due to its coarse spatial
444 resolution (Fig. 9c).



445 4.3.2 Case studies

446 To evaluate how well the models simulate the large fire emissions, we compare model performance for the
447 two recent cases reported to be the largest fire events during 2000-2012, including the fires in southern US in 2011
448 and western US in 2012. During 2011, a severe drought leading to large wildfires was observed over southern US,
449 including Arizona, New Mexico, and Texas (NOAA, 2012; Wang et al., 2015). Fig. 10 shows the maps of annual
450 mean fire $PM_{2.5}$ emissions over southern US from GFED, the ML model, and CLM. GFED shows the largest fire
451 emissions close to the border of Arizona and New Mexico in conjunction with other small hotspots over New
452 Mexico, Texas, and Louisiana (Fig. 10a). The ML model overall reproduces the spatial distributions of the fire
453 emissions ($r=0.96$) and captures the largest fire emission in Arizona and New Mexico in 2011 (Fig. 10b). However,
454 CLM does not capture the hotspots observed in GFED over Arizona and New Mexico but simulates larger fire
455 emissions in Louisiana instead (Fig. 10c). In terms of the time series, the ML model reproduces the temporal
456 variability of fire emissions and successfully captures the peak of total fire $PM_{2.5}$ emissions in June 2011 ($r=0.98$;
457 Figs. 10d and 10e). Although CLM simulates the peak in June, it overestimates fire emissions in the following
458 months by a factor of 4 ($r=0.52$; Figs. 10d and 10e).

459 In 2012, western US experienced several major wildfires (NOAA, 2013). The warm and dry conditions led
460 to large wildfires in California, Oregon, New Mexico, and Colorado (Fig. 11). Both the ML model and CLM
461 capture the hotspots with large fire emissions (Fig. 11b and 11c) and have correlation coefficients of 0.56 and 0.37,
462 respectively. However, the ML model tends to overestimate fire emissions, especially in areas surrounding the
463 grids with extremely large fire emissions (Fig. 11b). CLM misses some large fire emissions in Colorado and New
464 Mexico and underestimates the larger fire emissions in several hotspots (Fig. 11a), which may be explained by its
465 coarse resolution. The time series of normalized fire $PM_{2.5}$ emissions in 2012 show one peak in August. The ML
466 model captures the peak and presents a high correlation between the simulated and observed normalized and total
467 $PM_{2.5}$ fire emissions ($r=0.98$). CLM captures the peak in August but overestimates emissions in September and
468 October ($r=0.84$; Figs. 11d and 11e).

469 5. Discussion and conclusions

470 This study provides the first assessment to evaluate the performance of data-driven and process-based
471 models in predicting fire $PM_{2.5}$ emissions over CONUS. We first demonstrate that the developed ML model
472 performs well in predicting monthly fire $PM_{2.5}$ emissions nationwide at grid cells of $0.25^\circ \times 0.25^\circ$ resolution from
473 2000 to 2020, with an RMSE of 0.16 g/m^2 and IoA of 0.84. The ML model outperforms prior statistical models
474 predicting fire activities at similar spatial and temporal scales (Carvalho et al., 2008; Bedia et al., 2014).
475 Considering the performance at a regional scale, the ML model reproduces the interannual variability of fire
476 emissions for the selected regions, with correlation coefficients ranging from 0.84 to 0.98. Therefore, the ML model
477 has a promising performance in predicting fire emission over CONUS at a relatively fine spatial resolution.
478 Compared to the wildfire burned area model in Wang et al. (2021), the fire emission model in this study shows
479 slight degradations in capturing the interannual variability of fire emission at grid level (e.g., percentage of grids
480 with correlations larger than 0.4). This may be explained by the fact that the fire emission model may not effectively
481 resolve the relationships between fires and predictors when more grids with less fire occurrence are included (i.e.,
482 more zeros or unburned grids) without reliable information about ignition. As a side note, both burned area and
483 emission ML models have relatively poor performance over Mediterranean California, indicating the challenges in
484 modeling fire activities in this region where the terrain and land use are complex. The SHAP variable importance
485 shows that meteorology at both local and synoptic scale as well as fuel loads are important variables controlling
486 fire emissions over CONUS. Regional analysis of predictors indicates that fuel dryness such as fuel moisture and



487 energy release component (ERC) and fuel load are important for predicting fire emissions in western US, while
488 large-scale meteorological patterns (SVDs_SElag2) contribute more to fire emissions in southeastern US.

489 We then compare the simulated fire $PM_{2.5}$ emissions from the ML model and FireMIP models against
490 GFED from 2000 to 2012 at the spatial resolution of $1^\circ \times 1^\circ$. The ML model, CLM, and JULES reproduce the
491 spatial distribution more reasonably than the rest of the FireMIP models ($r=0.83, 0.52,$ and 0.40 for the ML, CLM,
492 and JULES, respectively). Both CLM and JULES simulate more fire $PM_{2.5}$ emissions over southeastern US, which
493 can be explained by several reasons. First, it has been shown that the satellite-observed burned areas in southeastern
494 US are much smaller than the burned areas estimated from the ground-based fire records, which might have resulted
495 from the small prescribed and agricultural fires (Hu et al., 2016; Nowell et al., 2018). In addition, large differences
496 exist among different satellite estimated fire $PM_{2.5}$ emissions in southeastern US (Li et al., 2019). As a consequence,
497 these studies highlighted uncertainties about the GFED estimated burned area and emission over southeastern US.
498 Second, cropland fires are one of the predominant fire types in this region. Among the FireMIP models, CLM is
499 the only model that simulates cropland fires (Li et al., 2013). For JULES, even though it does not simulate cropland
500 fires, it treats croplands as natural grasslands. The emission factors of grasslands and croplands used in the FireMIP
501 models are larger than in GFED4s, thereby causing larger fire $PM_{2.5}$ emissions in southeastern US in CLM and
502 JULES (van der Werf et al., 2017; Li et al., 2019). Furthermore, Li et al. (2019) noted that CLM4.5 simulates
503 higher fuel loads in croplands than the CASA model used by GFED4s, leading to higher fire carbon emissions
504 estimated by CLM than by GFED. It is worth noting that the ML model incorporates fuel load simulated by CLM4.5
505 but it predicts fire emissions closer to GFED4s, indicating a smaller sensitivity of fire emission to fuel load in the
506 ML model. The overestimation of fire $PM_{2.5}$ emissions can also be explained by the sensitivity to meteorology. The
507 spatial distributions of the long-term mean fire emissions shown in Fig. 3 correlate with the spatial distributions of
508 sensitivities to RH and/or temperature, with regional differences. For western US, large fire emissions are
509 associated with stronger sensitivities to both RH and temperature in the ML and most FireMIP models. For central
510 and southeastern US, overestimation of fire $PM_{2.5}$ emissions only corresponds to stronger sensitivity to temperature
511 in some FireMIP models.

512 Besides comparing model performance aggregated over CONUS, we analyze the model performance for
513 several regions, including the western forest area, Mediterranean California, southwestern US, and southeastern
514 US. For the western forest area, the ML model performs well in capturing both seasonality and interannual
515 variability of fire $PM_{2.5}$ emissions, with correlation coefficients of 0.98 and 0.96, respectively. In contrast, the
516 FireMIP models generally reproduce the seasonality well but do not simulate the interannual variability well,
517 especially underestimating the peak in 2007, which related to large-scale meteorological patterns favorable for fires
518 in Pacific Northwest (Zhong et al., 2020). For Mediterranean California, all FireMIP models only capture the first
519 peak in August but fail to simulate the second peak in October, which is caused by large fires related to Santa Ana
520 winds in 2003 and 2007. Such lack of peak emission is also shown in the interannual variability, as all FireMIP
521 models show limited ability to simulate the peaks in these two years. By contrast, the ML model successfully
522 predicts the bimodal seasonality and the large fire emissions related to the Santa Ana winds in 2003 and 2007. The
523 underestimation of the peak in the FireMIP models may be attributed to the underrepresentation of the effects of
524 large-scale meteorology in the two regions, as the ML model and SHAP importance show that SVD predictors
525 have larger contributions to the fire emissions in both events. The results of the two regions in the western US
526 suggest that fire parameterization in the FireMIP models could be improved by including the effects of regional
527 and large-scale meteorology (e.g., Santa Ana winds) on fire activity (Yue et al., 2014). Modeling the effect of Santa
528 Ana winds on wildfires may be particularly challenging as the offshore Santa Ana winds exhibit variability related
529 to both synoptic scale pressure anomaly over the Great Basin and local thermodynamic forcing associated with
530 strong desert-ocean temperature gradient (Hughes and Hall, 2010).

531 As for southwestern US, the ML model and LPJ-SIM estimate the peak in June ($r=0.99$ and 0.94 for ML
532 and LPJ-SIM, respectively), which highly agrees with the GFED observation. Interestingly, most FireMIP models



533 fail to capture the extremely large fire emission in the 2011 summer mainly due to the low biases of VPD anomalies
534 in CRUNCEP (Tang et al., 2017). Unlike southwestern US, the seasonality of southeastern US has peaks in March-
535 April and September-October. The two peaks of fire emissions correspond to wildfires (Mar-Apr and Sep),
536 cropland fires (Feb-Mar and Aug-Oct), and prescribed fires (Feb-Apr and Oct) that include burnings for pest
537 controls and land cleaning (Knapp et al., 2009; Lin et al., 2014). Most models fail to reproduce the bimodal fire
538 emissions, but the ML model, LPJ-SIM, and LPJ-SPI can capture the bimodal pattern. Even though the seasonality
539 of fires over this region is not simulated accurately, the CLM and JSBACH well reproduce the interannual
540 variability of fire PM_{2.5} emissions and predict the peaks. The FireMIP models' shortfall in reproducing the bimodal
541 seasonality can be explained by two reasons. First, the relationships between human and fire spread implemented
542 in the process-based models may not be realistic compared to the observed relationships. Parisien et al. (2016)
543 demonstrated the large spatial variability of human impacts on burned areas in North America, which is not well
544 represented in the FireMIP models (Li et al., 2019). Second, drier conditions in winter would promote fires in
545 springtime (Wear and Greis, 2013; Wang et al., 2021), which may not be directly considered in the FireMIP models
546 but are incorporated as SVD predictors in the ML model. Overall, the representations of the effects of human and
547 large-scale meteorology on fires may explain why the models simulate the seasonality incorrectly in southeastern
548 US. In addition to the comparison of general model performance, we also compare the ability of the data-driven
549 and processed-based models in predicting extremely large fire emissions. Both ML and CLM models reproduce
550 the spatial pattern of extreme fire events and reasonably simulate the historical events of large fires in southwestern
551 and western US.

552 To summarize, we utilize the ML model with SHAP importance to diagnose the fire emissions simulated
553 by process-based models and attributed model biases to several factors. First, the sensitivities of fire emissions to
554 meteorology in the models are stronger than the observed, leading to overestimations. Second, the large-scale
555 meteorological patterns conducive to fires are not fully considered in the process-based models, which are
556 important contributors of large fire emissions in western US and southeastern US. Third, the spatial resolutions of
557 models and/or the atmospheric forcing they used may be too coarse to resolve the effects of regional weather
558 phenomenon such as Santa Ana winds. Fourth, biases in the atmospheric forcing data may result in biases of fire
559 emission predictions. Last but not least, human activities are a critical component shaping fire regimes but the
560 human effects on fire activities in the FireMIP models may not reflect the human-fire relationships in the real-
561 world. This is also an issue in the ML model as the human-related predictors in the ML model may be too simple
562 to represent the human influences. The underrepresentation of human effects in both types of models may cause
563 additional uncertainties in projecting future fire activities and their impacts on climate. By training the ML model
564 using the GFED emissions, the ML model is able to better explain fire emissions in the US, which makes it a useful
565 tool for diagnosing processes or relationships that may be missing or not well represented in the process-based
566 models to guide future development for improving their performance. Besides its use in diagnosing process-based
567 models, the interpretable ML model provides a different and novel approach to simulate fire emissions more
568 accurately and identify the important predictor variables. While the ML model generally has higher accuracy than
569 the FireMIP models, the feedbacks between fire emissions and climate are not included, which could potentially
570 affect the reliability of ML-based models in fire emission prediction under future climate change scenario (Zou et
571 al., 2020). Lastly, due to limited training data, the ML model cannot predict fires in regions with longer fire return
572 intervals, posing additional uncertainties in their use for making future projections.

573
574
575 *Code availability.* Model code is available upon request to the first author.

576
577 *Data availability.* The ML prediction and predictor dataset used in this study are publicly accessible online at
578 <https://zenodo.org/record/5076646#.YOZ14zZKjOQ>.



579
580 *Author contributions.* SW, YQ, RL conceived the research ideas. SW wrote the initial draft of the paper, performed
581 analyses, and model development. All authors contributed to the interpretation of the results and the preparation of
582 the manuscript.

583
584 *Acknowledgements.*
585 This research was performed at PNNL and funded under Assistance Agreement No. RD835871 by the U.S.
586 Environmental Protection Agency to Yale University through the SEARCH (Solutions for Energy, AiR, Climate,
587 and Health) Center. It has not been formally reviewed by EPA. The views expressed in this document are solely
588 those of the SEARCH Center and do not necessarily reflect those of the Agency. EPA does not endorse any products
589 or commercial services mentioned in this publication.

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839 **Table 1.** Advantages and limitations of different types of fire models

	Representative method	Advantages	Limitations
Data-driven model	Multiple Linear regression (MLR)	<ol style="list-style-type: none"> 1. Computationally efficient 2. Simple model 3. It is easy to interpret 	<ol style="list-style-type: none"> 1. It cannot capture the non-linear relationships between fires and predictors 2. It assumes that the predictor variables are independent 3. It is sensitive to outliers
	Machine learning method (e.g., neural network, decision tree etc.)	<ol style="list-style-type: none"> 1. Computationally cheap 2. The performance improves when more training data are included 3. It can easily handle multi-dimensional data 	<ol style="list-style-type: none"> 1. It requires a lot of training data 2. It is relatively hard to interpret 3. The interactions between fires and vegetation/atmosphere cannot be updated to the model
Process-based model	Dynamic global vegetation model (DGVM)	<ol style="list-style-type: none"> 1. Physics-driven 2. The simulations can include feedbacks between fires and climate or vegetation 	<ol style="list-style-type: none"> 1. Computationally expensive 2. The same parameterization may not be applied to all regions 3. It only parameterizes the known processes or phenomena

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843 **Table 2.** Predictor variables used in the ML model

Variables	Abbreviation	Categories	Temporal resolution	Spatial resolution	Data Source	References
Monthly mean surface temperature	temp	Local meteorology	monthly	32 km	North American Reanalysis (NARR)	Mesinger et al. (2006)
Monthly mean relative humidity	RH	Local meteorology	monthly	32 km	North American Reanalysis (NARR)	Mesinger et al. (2006)
Monthly mean of daily precipitation	precip	Local meteorology	monthly	32 km	North American Reanalysis (NARR)	Mesinger et al. (2006)
Monthly mean zonal component of wind speed	U	Local meteorology	monthly	32 km	North American Reanalysis (NARR)	Mesinger et al. (2006)
Monthly mean meridional component of wind speed	V	Local meteorology	monthly	32 km	North American Reanalysis (NARR)	Mesinger et al. (2006)
Monthly Standardized Precipitation Evapotranspiration Index	SPEI	Local meteorology	monthly	0.5°×0.5°	SPEI	Vicente-Serrano et al. (2010)
Monthly mean 1000-hour dead fuel moisture	FM1000	Local meteorology	daily	4 km	gridMET	Abatzoglou (2013)
Monthly mean energy release component	ERC	Local meteorology	daily	4 km	gridMET	Abatzoglou (2013)
Monthly mean vapor pressure deficit	VPD	Local meteorology	daily	4 km	gridMET	Abatzoglou (2013)
Monthly lightning flashes density	lightning	Local meteorology	daily	0.1°×0.1°	SWDI/NLDN	NOAA (2006); Cummins and Murphy (2009)
Monthly standard deviation of daily SVDs for northern California	SVD1_NCA and SVD2_NCA	Large-scale meteorological patterns	monthly	Regional	North American Reanalysis (NARR)	Wang et al. (2021)



Monthly standard deviation of daily SVDs for southern Rocky Mountain	SVD1_SRM and SVD2_SRM	Large-scale meteorological patterns	monthly	Regional	North American Reanalysis (NARR)	Wang et al. (2021)
Monthly standard deviation of daily SVDs for southeastern US (with 2-month lag)	SVD1_SElag2 and SVD2_SElag2	Large-scale meteorological patterns	monthly	Regional	North American Reanalysis (NARR)	Wang et al. (2021)
Monthly mean evapotranspiration	ET	Land-surface properties	monthly	0.25°×0.25°	North American Land Data Assimilation System (NLDAS-2)	Xia et al. (2012)
Monthly mean surface soil moisture	soilm	Land-surface properties	monthly	0.25°×0.25°	Global Land Data Assimilation System (GLDAS-2)	Xia et al. (2012)
Monthly mean vegetation fraction	Veg_frac	Land-surface properties	monthly	0.25°×0.25°	Global Land Data Assimilation System (GLDAS-2)	Xia et al. (2012)
Monthly mean Leaf Area Index	LAI	Land-surface properties	8 days	500 m	MODerate resolution Imaging Spectroradiometer (MODIS); LAI classification scheme	Myneni et al. (2015)
Monthly fuel load/normalized fuel load	fuel_load/fuel_load_nor	Land-surface properties	monthly	0.9°×1.25°	Community Land Model (CLM)	Lawrence et al. (2019)
Land cover percentage	p_	Land-surface properties	Yearly	0.05°×0.05°	MODerate resolution Imaging Spectroradiometer (MODIS); LAI classification scheme	Friedl (2015)
Median Topography (slope and elevation)	Slope and elevation	Land-surface properties	Not change by time	100 km		Amatulli et al. (2018)
Gross domestic product	GDP	Socioeconomic and coordinate variables	Yearly	5 arc		Kummu et al. (2018)
Population density	Pop	Socioeconomic and coordinate variables	Yearly	30 arc	Gridded Population of the World data collection (GPW v4)	CIESIN-Columbia University (2017)



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847 **Table 3.** The ML model performance for different regions: western forest area, Mediterranean California, southwestern US,
 848 and southeastern US

	Western forest area	Mediterranean California	Southwestern US	Southeastern US	Whole US
Grid scale (individual grid)					
RMSE (km ²)	0.29	0.32	0.10	0.02	0.16
Correlation (r)	0.79	0.51	0.76	0.84	0.75
IoA	0.86	0.60	0.85	0.90	0.84
Percentage of grids with correlation > 0.4 (%)	68	47	52	80	74
Regional scale (summation over the region)					
RMSE (km ²)	37.80	13.94	2.76	3.37	49.98
Correlation (r)	0.98	0.81	0.94	0.97	0.97
IoA	0.98	0.81	0.95	0.98	0.97

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851 **Table 4.** The model performance for the ML model and FireMIP models

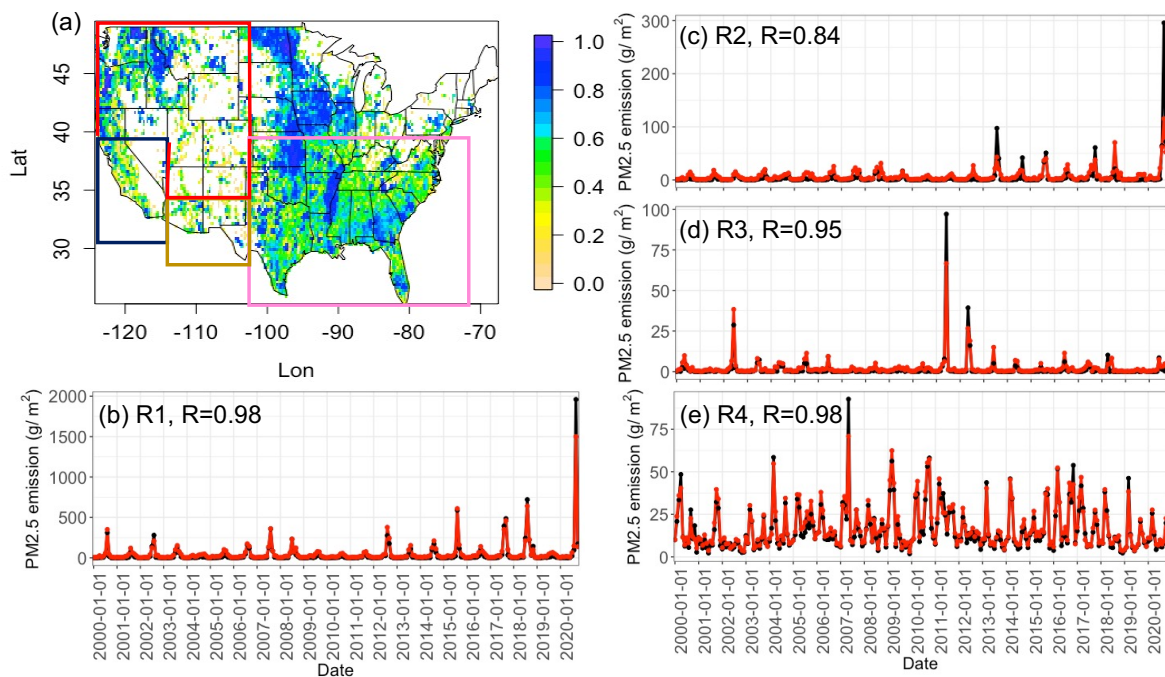
	ML model	CLM	CTEM	JSBACH	LPJ-SPI	LPJ-Glob	LPJ-SIM	ORCHIDEE	JULES	
Total amounts of fire PM_{2.5} emissions (Tg=10¹² g) (GFED: 4.89 Tg)										
Total fire PM _{2.5} emissions over 2000-2012 (Tg)	8.33	16.54	41.50	19.92	16.23	79.49	35.38	2.43	33.43	
Correlation of interannual/seasonal variability for the CONUS										
Correlation (interannual/seasonal)	0.87/0.98	0.71/0.92	0.28/0.87	0.15/0.89	0.15/0.92	0.02/-	0.23/0.65	0.03/0.91	0.55/0.93	
Correlation of interannual/seasonal variability for the selected regions										
Western forest area	0.93/0.98	0.70/0.93	0.33/0.88	0.21/0.88	0.38/0.79	0.51/-	0.46/0.98	0.05/0.94	0.60/0.92	
Mediterranean California	0.72/0.97	-0.01/0.87	0.05/0.94	-0.30/0.89	-0.07/0.90	-0.14/-	-0.19/0.83	0.25/0.13	-0.21/0.87	



Southwestern US	0.95/0.99	0.14/0.85	-0.26/0.62	-0.28/0.45	0.34/0.42	0.30/-	0.40/0.94	0.45/0.72	-0.07/0.69
Southeastern US	0.96/0.99	0.57/-0.27	-0.16/0.09	0.72/-0.14	0.08/0.35	0.39/-	0.18/0.68	0.16/0.13	0.36/0.01

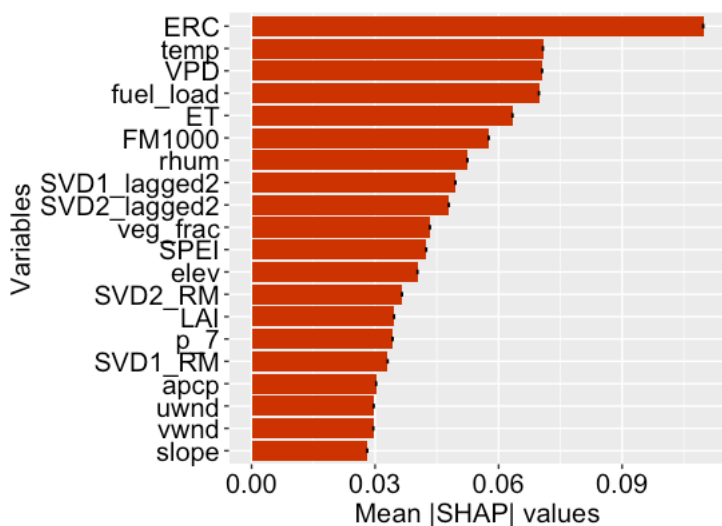
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855 **Fig. 1.** (a) The map of temporal correlation between observed and predicted PM_{2.5} fire emission for each grid. Time series of
 856 observed (black) and predicted (red) PM_{2.5} fire emission average across (b) western forest area (red box in 1a), (c)
 857 Mediterranean California (blue box), (d) southwestern US (dusty box), (e) southeastern US (pink box).

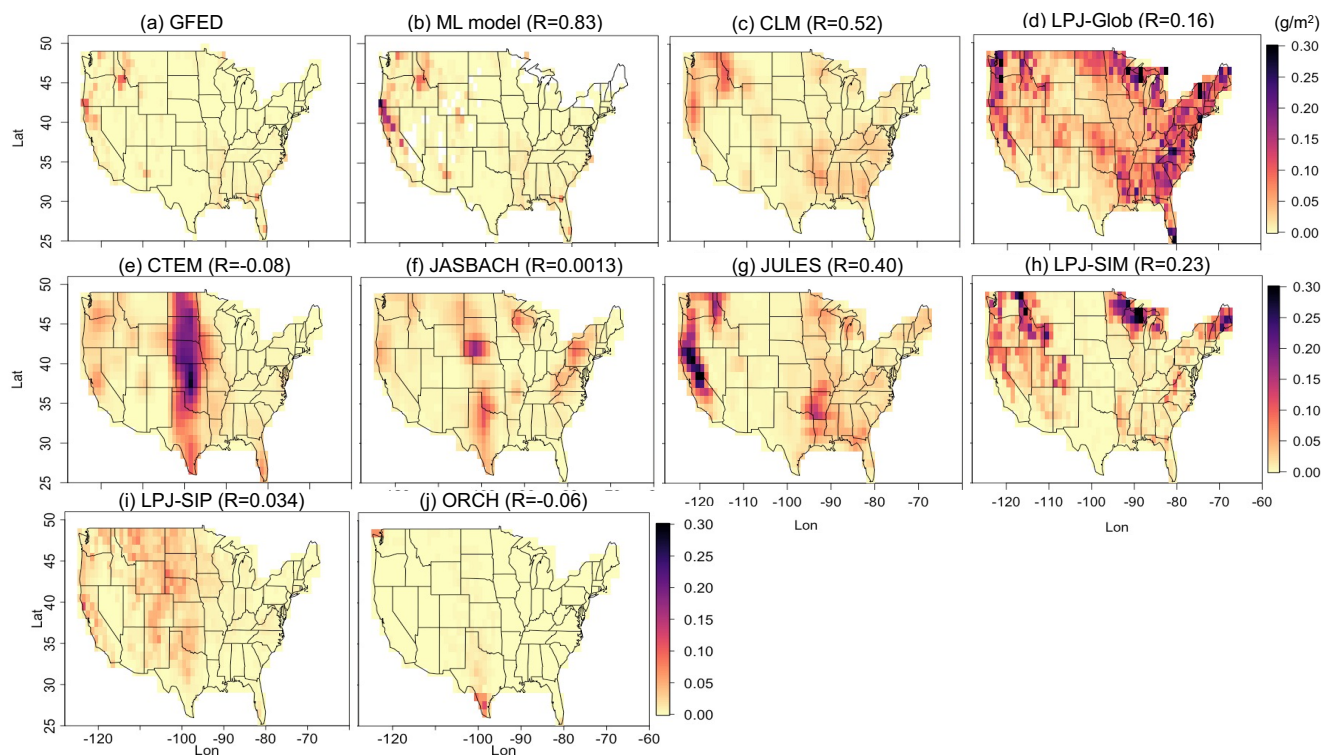


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859 **Fig. 2.** Top 20 variables for the model based on the mean absolute SHAP value with the 95% confidence intervals.

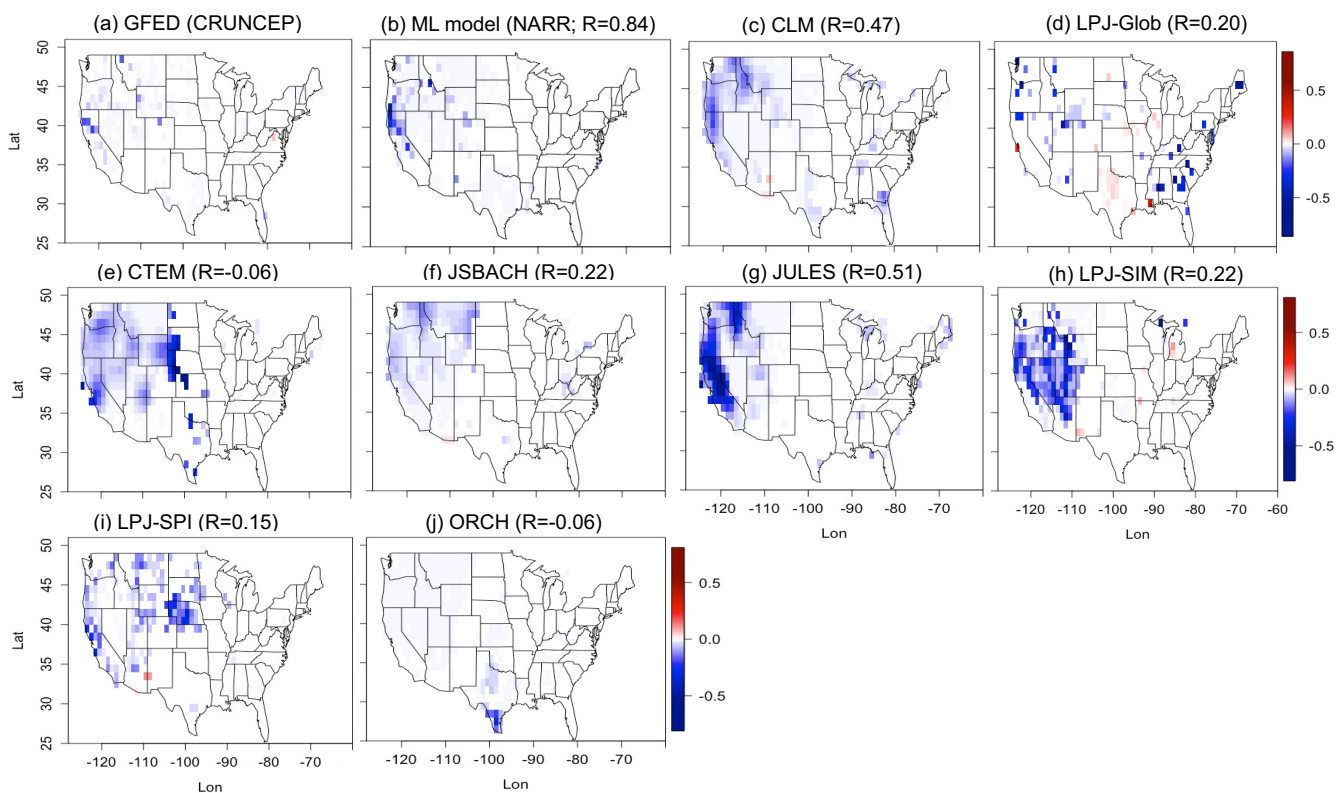
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862 **Fig. 3.** Spatial distributions of the monthly mean PM_{2.5} fire emission (g/m²/month) over 2000-2012.

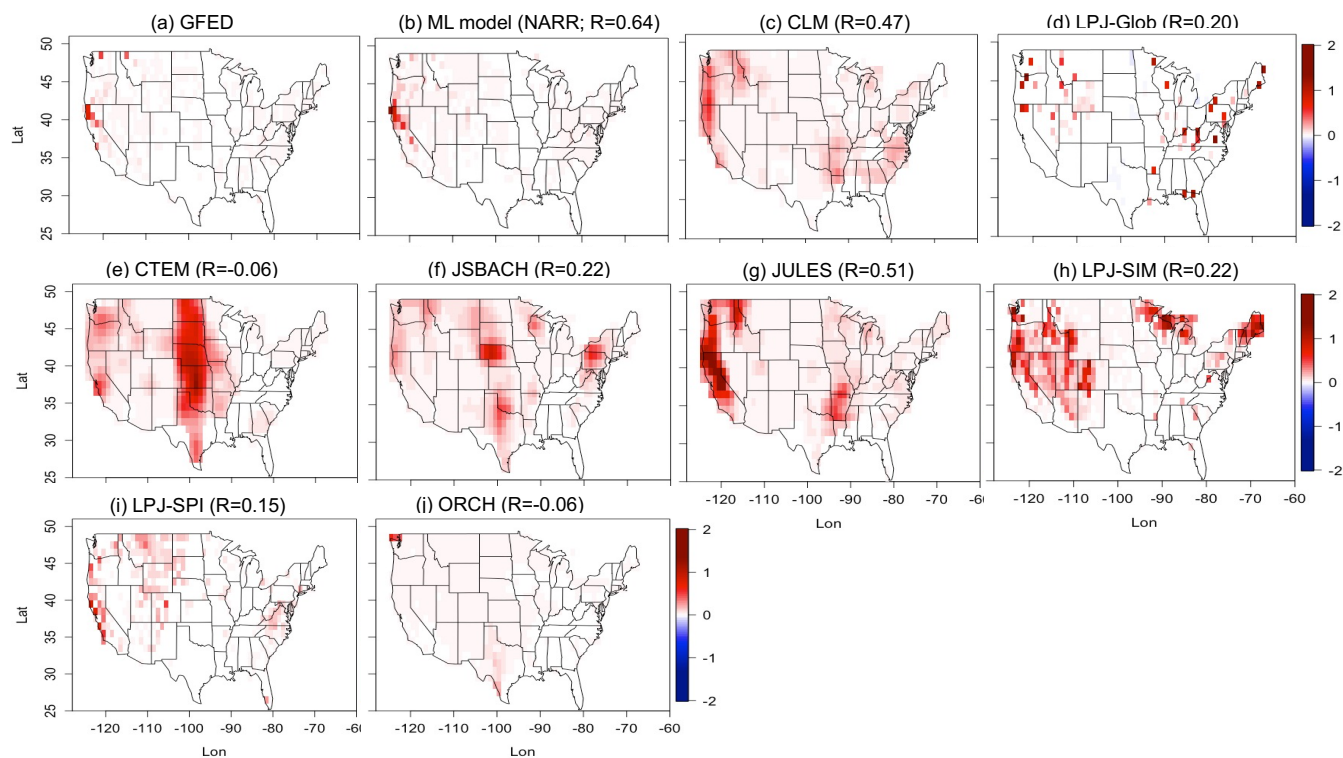
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865 **Fig. 4.** Spatial distributions of the linear regression slope for the dependence of annual mean $PM_{2.5}$ fire emissions on annual
866 mean RH. Only the grids with slopes that are statistically significant ($p < 0.05$) are shown.

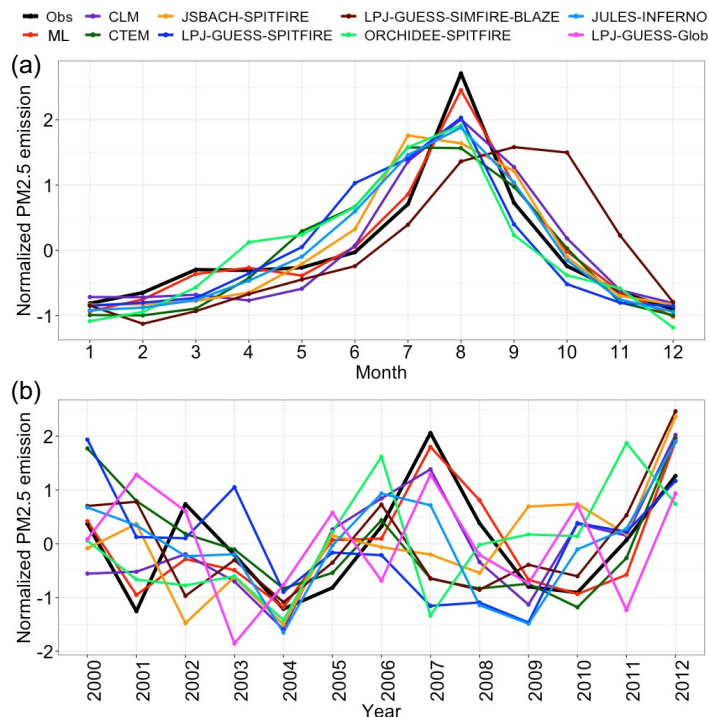
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869 **Fig. 5.** Spatial distributions of the linear regression slope for the dependence of annual mean $PM_{2.5}$ fire emissions on annual
870 mean temperature. Only the grids with slopes that are statistically significant ($p < 0.05$) are shown.

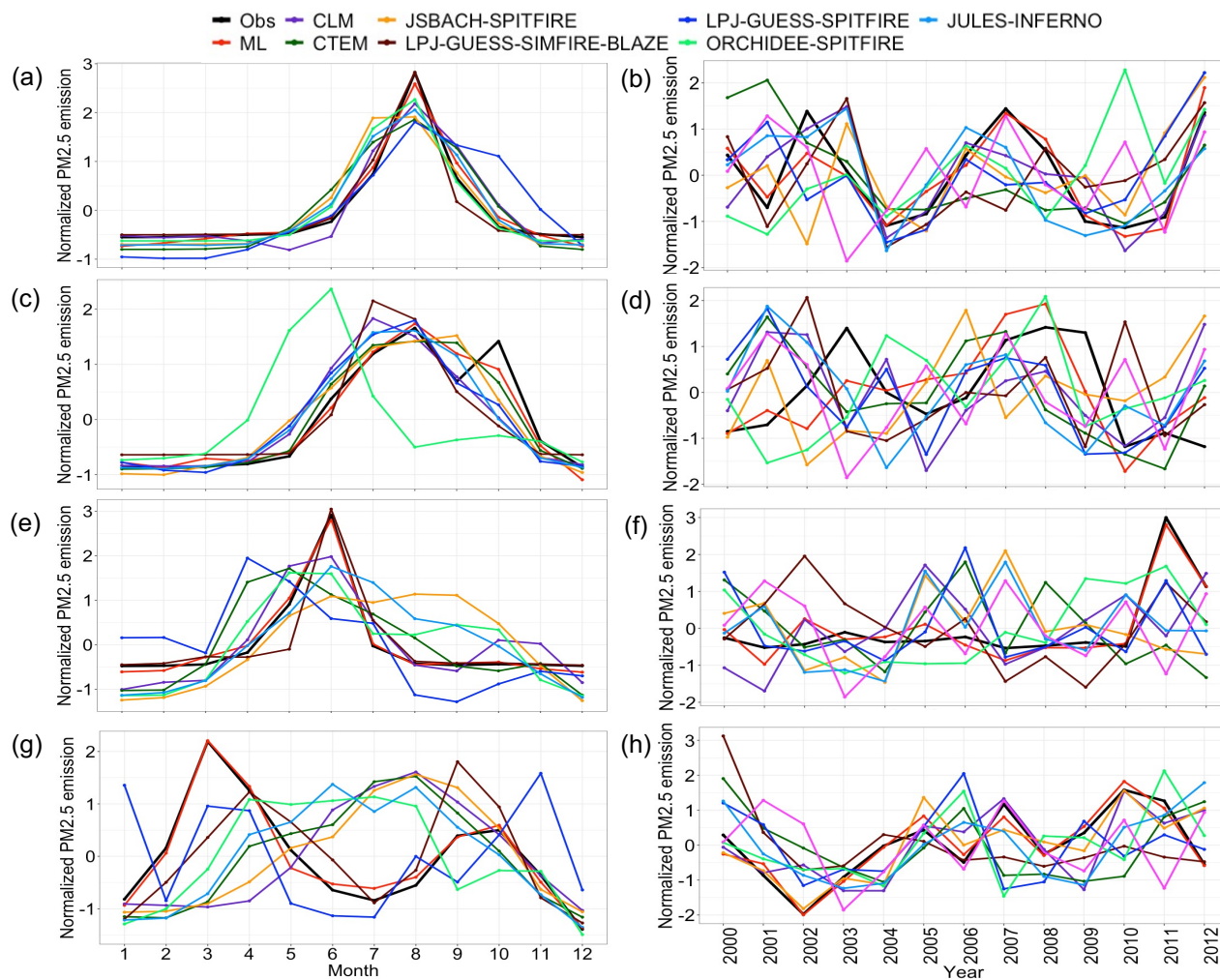
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873 **Fig. 6.** (a) Seasonality and (b) interannual variability of the normalized averaged $PM_{2.5}$ fire emission from the GFED (black
874 line), ML model (red line), and the FireMIP models (color lines). The $PM_{2.5}$ fire emissions are first averaged over CONUS
875 and normalized by the monthly (annual) mean and standard deviation for seasonality (interannual variability) plots.

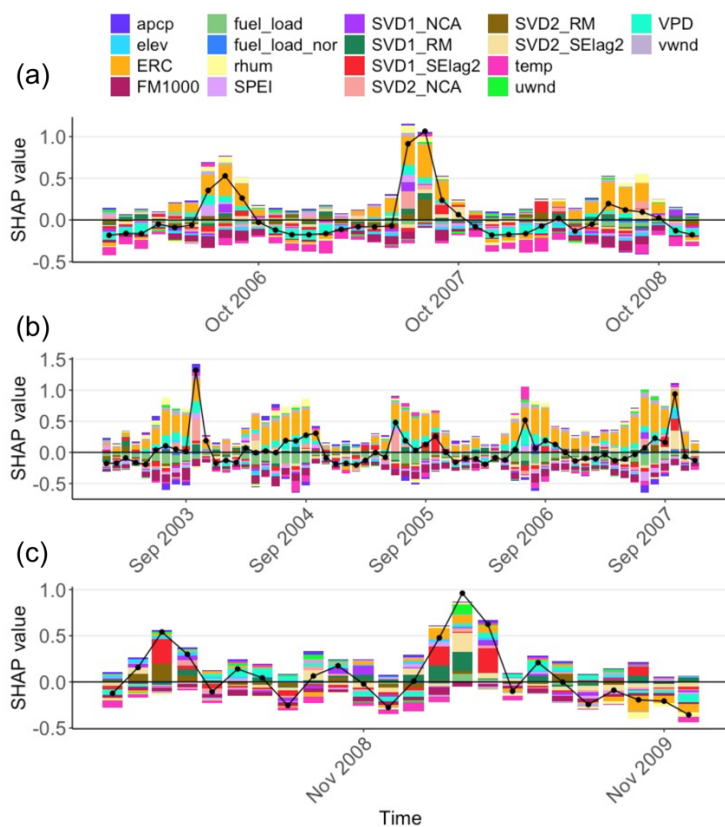
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878 **Fig. 7.** Seasonality and interannual variability of the PM_{2.5} fire emission from the GFED (black line), ML model (red line),
879 and the FireMIP models (color lines) for (a, b) western forest area, (c, d) Mediterranean California, (e, f) southwestern US,
880 and (g, h) southeastern US.

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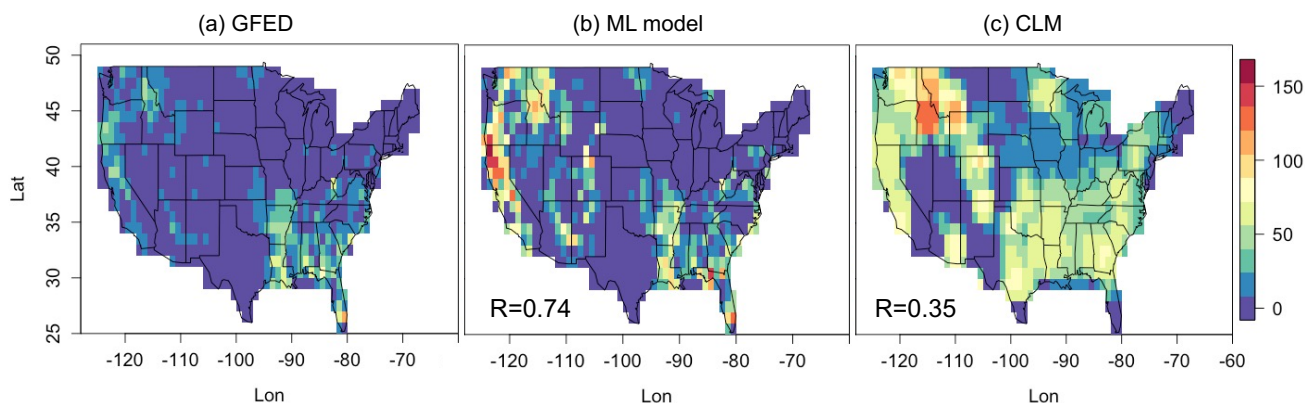


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883 **Fig. 8.** Time series of the average SHAP values (bar) and predicted normalized $PM_{2.5}$ fire emission (line) for (a) western forest
 884 area from 2006 to 2008, (b) Mediterranean California from 2003 to 2007, and (c) southeastern US from 2008 to 2009. The
 885 SHAP values indicate the contribution of the predictors to the prediction of normalized fire emission.

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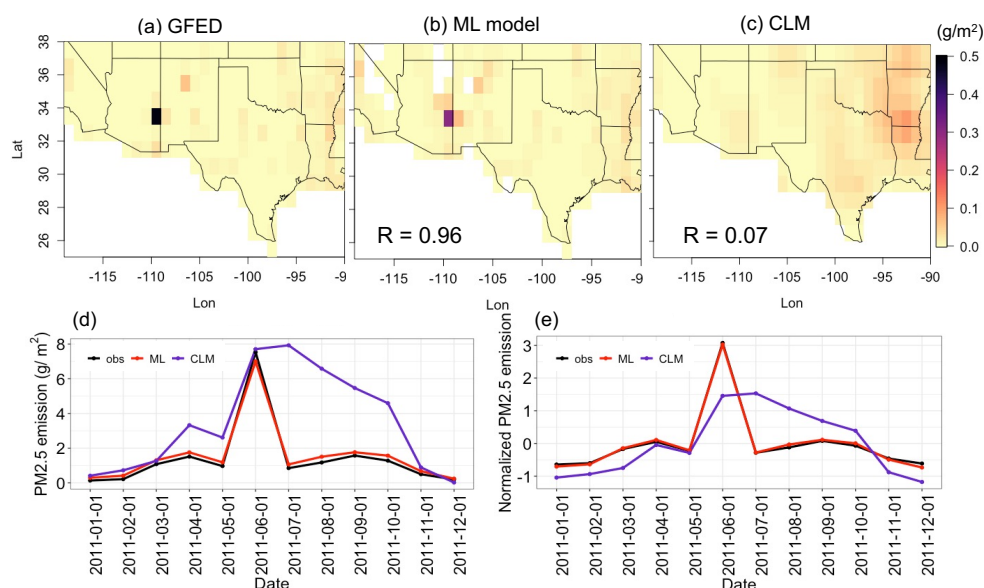


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889 **Fig. 9.** Spatial distributions of number of months with large fire emissions ($> 95^{\text{th}}$ percentiles of $PM_{2.5}$ fire emission over all
 890 the grids in 2000-2012) for (a) GFED, (b) ML model, and (c) CLM.



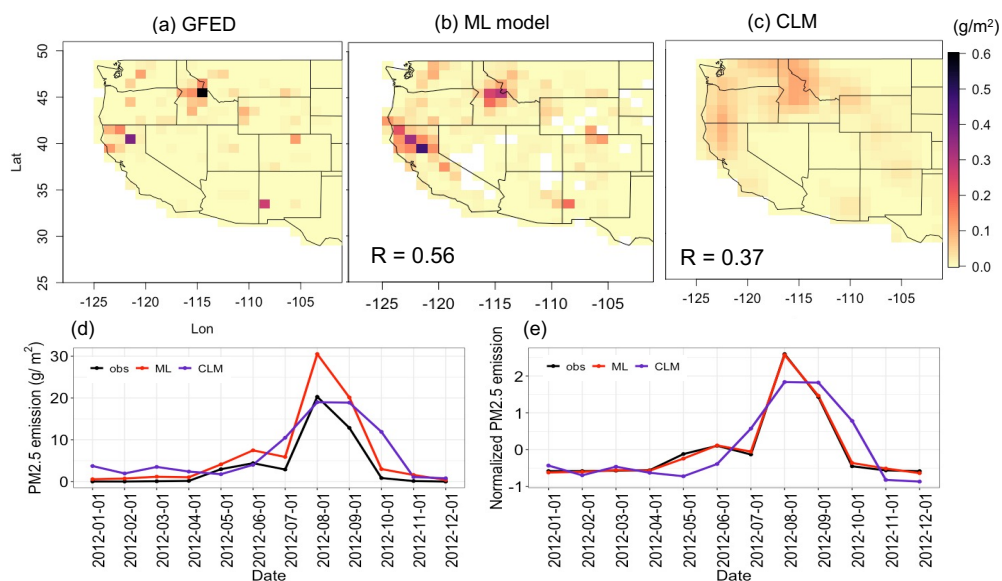
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893 **Fig. 10.** Top panel: Spatial distributions of the annual mean PM_{2.5} fire emission in 2011 for (a) GFED, (b) ML model, and (c)
894 CLM. Bottom panel: Time series of the (d) total PM_{2.5} fire emissions and (e) normalized PM_{2.5} fire emission over southern US
895 domain during 2011.

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898 **Fig. 11.** Top panel: Spatial distributions of the annual mean PM_{2.5} fire emission in 2012 for (a) GFED, (b) ML model, and (c)
899 CLM. Bottom panel: Time series of the (d) total PM_{2.5} fire emissions and (e) normalized PM_{2.5} fire emission over western US
900 domain during 2012.