

# Adding value to Extended-range Forecasts in Northern Europe by Statistical Post-processing Using Stratospheric Observations

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**Abstract.** The strength of the stratospheric polar vortex influences the surface weather in the Northern Hemisphere in winter; a weaker (stronger) than average stratospheric polar vortex is connected to negative (positive) Arctic Oscillation (AO) and colder (warmer) than average surface temperatures in Northern Europe within weeks or months. This holds a potential for forecasting in that time-scale. We investigate here if the strength of the stratospheric polar vortex at the start of the forecast  
15 could be used in improving the Extended-Range temperature Forecasts of the European Centre for Medium-Range Weather Forecasts (ECMWF) and in finding periods with higher prediction skill scores. For this, we developed a stratospheric wind indicator, *SWI*, based on the strength of the stratospheric polar vortex and the phase of the AO during the following weeks. We demonstrate that there was a statistically significant difference in the observed surface temperature in Northern Europe within the 3–6 weeks depending on the *SWI* at the start of the forecast.

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When our new *SWI* was applied in post-processing the ECMWF's two-week mean temperature reforecasts for weeks 3–4 and weeks 5–6 in Northern Europe during boreal winter, the skill scores of those weeks were slightly improved. This indicates there is some room to improve the Extended Range Forecasts, if the stratosphere-troposphere links were better captured in the modelling. In addition to this, we found that during the boreal winter, in cases the polar vortex was weak at the start of the  
25 forecast, the mean skill scores of the 3–6 weeks surface temperature forecasts were higher than average.

## 1 Introduction

Extended-range forecasts (ERF; lead time up to 46 days) by dynamical models have been developed since the 1990s with the aim to fill the gap between the medium-range weather forecasts and the seasonal forecasts. It is known that ERF skills are still rather modest in forecast weeks 3–6 especially in the Northern latitudes. If the skill of the forecasts improves, ERFs have the  
30 potential to become an essential element in climate services e.g., in the form of early warnings of climatic extremes. In an academic project CLIPS (CLimate services supporting Public mobility and Safety), climatic impact outlooks and early warnings of extremes (CLIPS forecasts) were developed by employing the ERF datasets (Ervasti et al. 2018). The CLIPS

forecasts were co-designed with the general public in Finland and experimented by a one year piloting phase. As many industries, e.g., energy and food production, as well as users from the general public considered they could use and would benefit from reliable ERFs (Ervasti et al. 2018), development of more skillful ERFs is clearly needed.

- 5 The European Centre for Medium-Range Weather Forecasts (ECMWF) has produced ERFs routinely since March 2002 (Vitart 2014). The verification results of the ECMWF model's ERF (Buizza and Leutbecher 2015; Vitart 2014) on a sub-continental and a regional scale (e.g., Monhart et al. 2018) demonstrated predictive skill beyond 2 weeks for temperature reforecasts over Northern Europe. ECMWF uses bias-correction of the mean in their automatic products, removing the mean bias computed from the reforecasts, depending on the time of the year (Buizza and Leutbecher 2015). We consider the bias over Northern  
10 Europe not to be dependent only on the time of the year but also on the prevailing weather pattern, and therefore, we aim to explore whether known teleconnections such as the strength of the stratospheric polar vortex and the phase of the Arctic Oscillation (AO) could be used in improving the forecasts.

The stratospheric polar vortex is an upper-level low-pressure area that forms over both the northern and southern poles during  
15 winter due to the growing temperature gradient between the pole and the tropics. Strong westerly winds circulate the polar vortex, isolating the gradually cooling polar cap air. The strength of the northern polar vortex varies from year to year and can be indicated by, e.g., the zonal mean zonal wind (ZMZW) at 60°N and 10 hPa or polar cap temperatures. The stronger the circumpolar winds and the colder the polar cap temperatures are, the stronger is the polar vortex. Planetary waves from the troposphere disturb the northern stratospheric polar vortex, leading to meandering and weakening of the westerlies and  
20 occasionally to reverse, i.e., easterly flow (Schoeberl, 1978). This weakening of the stratospheric polar vortex also leads to warming of the polar cap temperatures, sometimes even > 30–40 K within several days. A warming of this magnitude together with a reversal of the ZMZW at 60°N and 10 hPa is commonly defined as a major sudden stratospheric warming (SSW), albeit other definitions have also been used (Butler et al. 2015).

- 25 During boreal winter the strength of the polar vortex affects the phase of the AO, which characterizes air mass flow between the Arctic and the mid-latitudes. At the surface, the AO index is affected by the strength of the polar vortex with a time lag of about two to three weeks (Baldwin and Dunkerton 1999). A strong polar vortex is characterized by lower than average surface pressure in the Arctic, positive AO index, and strong westerly winds keeping the cold Arctic air locked in the polar region and bringing milder and wetter than average weather to Northern Europe (Limpasuvan et al. 2005). In contrast, a weak polar vortex  
30 is characterized by higher than average surface pressure in the Arctic, negative AO index, and the meandering and/or weakening of the polar jet stream and tropospheric jet stream enabling cold arctic/polar air outbreaks to Northern Europe (Thompson et al. 2002, Tomassini et al. 2012).

During boreal winters, the strength of the stratospheric polar vortex influences the surface weather in the Northern Hemisphere within weeks or months (Baldwin and Dunkerton 2001, Kidston et al. 2015), hence holding a potential for forecasting in that time scale. However, challenges related to the realistic modelling of the dynamical stratosphere-troposphere coupling have been adduced, e.g., by Shepherd et al. (2018) and Polichtchouk et al. (2018). Therefore, we investigate if the known  
5 stratospheric-tropospheric connection could be used to improve the ERFs by statistical post-processing.

In this paper, we first verify the raw and the mean bias-corrected surface temperature reforecasts of the ECMWF's ERFs for forecast weeks 1 to 6 over Northern Europe against ERA-Interim surface temperature re-analysis (Dee et al. 2011). After that, our aim is to find out which thresholds of the ZMW at 60°N and 10 hPa are followed by a statistically significantly weaker  
10 AO index. For this, we explore the observed daily AO index during boreal winters 1981–2016, 1–2 weeks, 3–4 weeks, and 5–6 weeks after different strengths of the observed the ZMW at 60°N and 10 hPa. According to the observed daily AO index, after different thresholds of the ZMW at 60°N and 10 hPa, we define a novel Stratospheric Wind Indicator (*SWI*). For a statistically significantly weaker mean AO index, the *SWI* is defined as *SWI<sub>neg</sub>*; otherwise, *SWI* is defined as *SWI<sub>plain</sub>*. Further, we study the mean surface temperature anomalies observed in Northern Europe 1–2 weeks, 3–4 weeks, and 5–6 weeks after  
15 *SWI<sub>neg</sub>* in comparison to *SWI<sub>plain</sub>* and we utilize these anomalies in post-processing the temperature forecasts of the ECMWF reforecasts. Finally, we compare the *SWI* based post-processed ECMWF reforecasts with the mean bias-corrected ECMWF reforecasts. Our paper is constructed as follows: First, we present the datasets and methods. Then, we present results about the selection of the *SWI<sub>neg</sub>* and *SWI<sub>plain</sub>* and the skill scores of the forecasts without post-processing and with post-processing. In the Discussions and Conclusions section, we present our view on our findings and the possible next steps.

## 20 **2 Datasets and Methods**

We verified and post-processed ERFs of the ECMWF's IFS Cycle 43r1 (Vitart, 2014), which belongs to the models of the Sub-seasonal to seasonal (S2S) prediction project of the World Weather Research Program/World Climate Research Program (Vitart et al. 2017). These forecasts are run twice a week, on Mondays and Thursdays, in a horizontal resolution of 0.4 degrees. We first studied the weekly mean temperatures of the Monday runs over Northern Europe (52° N to 71.2° N and 10° E to 33.2°  
25 E) with lead times of 1 to 6 weeks, here called forecast weeks 1 to 6. We verified the 20 years × 52 weeks = 1040 reforecasts (11 members ensemble) for 1997–2016 run for the same dates as the operational forecasts, i.e., as Mondays in 2017. The weekly averages of the raw, mean bias-corrected (Section 2.2), and post-processed (Sections 2.3 and 2.4) surface temperature forecasts over Northern Europe were verified against ERA-Interim 1981–2016 temperature re-analyses (Dee et al. 2011). Years 1981–2010 of the ERA-Interim data were used as the climatological reference period and as the statistical/climatological  
30 forecast.

## 2.1 Skill scores of the forecasts

A commonly used measure for the probabilistic forecasts is the continuous ranked probability score (CRPS, Hersbach 2000) calculated by the following Eq. (1):

$$CRPS = \int |F(y) - F_o(y)|^2 dx, \quad (1)$$

5 where  $F(y)$  and  $F_o(y)$  are the cumulative distribution functions of the forecast and the observation, respectively.

The CRPSs were calculated by the R package ‘ScoringRules’ (Jordan et al. 2018) for the ECMWF’s reforecast ( $CRPS_{rf}$ ) and the climatological forecasts (ERA-Interim weekly mean temperatures in 1981–2010), which were used as the reference ( $CRPS_{clim}$ ). As the ensemble size of the reforecasts,  $m$ , was only 11, and the ensemble size of the operational forecasts of the ECMWF’s IFS,  $M$ , was 51, the expected CRPS, the  $CRPS_{RF}$  of the ECMWF’s reforecast was calculated for 51 members using

10 equation 26 in Ferro et al. (2008):

$$CRPS_{RF} = \frac{m(M+1)}{M(m+1)} CRPS_{rf} \quad (2).$$

We calculated the annual means of the expected  $CRPS_{RF}$  across all weeks (1 to 52) of the years 1997–2016 reforecasts. These annual means were computed separately for lead times of 1 week, 2 weeks, 3 weeks, 4 weeks, 5 weeks, and 6 weeks, here called forecast week 1, forecast week 2, forecast week 3, forecast week 4, forecast week 5, and forecast week 6, respectively.

15 Further, the skill scores of the annual mean CPRSSs, the annual mean CRPSSs, for each lead time were calculated as follows:

$$CRPSS = 1 - \frac{CRPS_{RF}}{CRPS_{clim}} \quad (3).$$

The statistical significances of each forecast week’s annual mean CRPSS was determined for each grid point. The p-value with the null hypothesis that the CRPSS is zero was calculated by bootstrap resampling procedure with replacement and a sample size of 5000 for significance level 0.05.

## 20 2.2 Bias-correction of the ensemble mean

The mean bias-correction (as in Buizza and Leutbecher 2015, eq. 7a) removed the mean bias computed from the ensemble reforecasts for the 20 years (1997–2016) depending on the forecast week date. For the 1997–2016 reforecasts, the average bias was calculated considering  $19 \times 11 \times 5 = 1045$  ensemble reforecast members: 11 members’ reforecast with initial dates defined by five weeks centred on the forecast week date for the 19 years reforecasts (1997–2016 excluding the reforecast year). The

25 mean bias-corrected weekly mean temperatures were verified against the ERA-Interim data by calculating the annual mean CRPS separately for each lead time, i.e., forecast weeks 1 to 6. The skill scores of the mean bias-corrected forecasts and their statistical significance were calculated as explained in Section 2.1.

### 2.3 Definition of the stratospheric wind indicator (*SWI*)

As numerous observational and modelling studies have shown, the stratospheric polar vortex influences the weather in the Northern Hemisphere during boreal winter; strong polar vortex coincides more often with positive AO index and mild surface weather in Northern Europe, whereas weak polar vortex is more often followed by negative AO index and cold air outbreaks (Thompson and Wallace 1998, 2001, Kidston et al. 2015 and references therein). We aimed to find a stratospheric precursor for a statistically significantly weaker AO index available at the start of the forecast. The daily surface AO index were downloaded from the National Centers for Environmental Prediction (NCEP), Climate Prediction Center (CPC). This daily AO index from the NCEP CPC is produced by projecting the daily 1000 hPa geopotential height anomalies north of 20° N onto the loading pattern of AO, which is defined as the first leading mode from the Empirical Orthogonal Function (EOF) analysis of monthly mean 1000 hPa height anomalies poleward of 20° N during 1979–2000. As a precursor for the AO index we used the daily ZMW at 60°N and 10 hPa during 1981–2016 of the Modern-Era Retrospective analysis for Research and Applications, Version 2 (MERRA-2; Rienecker et al. 2011) reanalysis data provided by the National Aeronautics and Space Administration (NASA).

We explored the mean AO index 1 to 6 weeks after the beginning of each week in November–February (1981–2016) and the minimum daily ZMW at 60°N and 10 hPa during the preceding 10 days to find a threshold for the ZMW at 60°N and 10 hPa to be followed by statistically significantly weaker AO index 1–2, 3–4, and 5–6 weeks later. The statistical significance of the difference between the AO index following the different thresholds of the ZMW at 60°N and 10 hPa, was determined using a two-sided Student's t-test with the null hypothesis that there is no difference. The threshold of the ZMW at 60°N and 10 hPa for statistically significantly weaker (at the 99% confidence level) AO index observed 1–2 weeks, 3–4 weeks, and 5–6 weeks later, was used to define the *SWI* as follows: below the threshold the *SWI* was defined negative,  $SWI_{neg}$ , and above the threshold the *SWI* was defined as plain,  $SWI_{plain}$ .

### 2.4 Utilizing the stratospheric winds indicator (*SWI*) in forecasting

In this section, we investigated the observed and reforecasted surface temperature anomalies 1–2 weeks, 3–4 weeks, and 5–6 weeks after  $SWI_{neg}/SWI_{plain}$  defined in Section 2.3. First, we calculated the two-week mean temperature anomalies of the ERA-Interim reanalyses (Dee et al. 2011) of the 1–2 weeks, the 3–4 weeks, and the 5–6 weeks from the beginning of each week in January, February, November, and December in 1981–2016 in Northern Europe. Subsequently, we divided the observed two-week mean temperature anomalies to sets of anomalies, representing  $SWI_{neg}$ , and  $SWI_{plain}$  according to the minimum ZMW at 60°N and 10 hPa during the preceding 10 days. Thereafter, we determined the statistical significance of the difference between the surface temperatures after  $SWI_{neg}$  and  $SWI_{plain}$  using a two-sided Student's t-test with the null hypothesis that there is no difference between  $SWI_{neg}$  and  $SWI_{plain}$ . This same procedure to define the difference between the surface temperatures after  $SWI_{neg}$  and  $SWI_{plain}$  was used for the ERA-Interim reanalyses for the period 1997–2016 to see how the selection of a

shorter period affects the temperature anomalies. Further, the mean surface temperature anomalies 1–2 weeks, 3–4 weeks, and 5–6 weeks after  $SWI_{neg}$  and  $SWI_{plain}$  in the ECMWF reforecasts run in the beginning of each week in November–February 1997–2016 were calculated to examine how the model reproduced the anomalies.

- 5 For post-processing the ECMWF reforecasts we calculated  $TA_{SWIneg}$  and  $TA_{SWIplain}$  representing mean temperature anomalies in November–February 1981–2016 after  $SWI_{neg}$  and  $SWI_{plain}$ , respectively. The  $TA_{SWIneg}$  and the  $TA_{SWIplain}$  were calculated separately for each  $0.4^\circ \times 0.4^\circ$  grid point over Northern Europe.

10 For the post-processing of the ECMWF reforecasts, we first defined the  $SWI$  either  $SWI_{neg}$  or  $SWI_{plain}$  at the start of the forecast according to the minimum ZMW at  $60^\circ\text{N}$  and 10 hPa during the preceding 10 days. According to the  $SWI$ , we added either  $TA_{SWIneg}$  or  $TA_{SWIplain}$  to the ERA-Interim mean temperature during 1981–2016, corresponding to forecast weeks 1–2, 3–4, and 5–6 to get a  $SWI_{neg}$  and  $SWI_{plain}$  based mean temperatures,  $T_{SWIneg}$  and  $T_{SWIplain}$ , for weeks 1–2, 3–4, and 5–6, respectively. The  $T_{SWIneg}$  and  $T_{SWIplain}$  were used in post-processing the ECMWF reforecasts’ mean bias-corrected ensemble members,  $T_{BC}$ , by calculating a weighted average,  $T_{SWI\_BC}$ , for  $SWI_{neg}$  as follows:

$$15 \quad T_{SWI\_BC} = (1 - k_{SWI}) * T_{BC} + k_{SWI} * T_{SWIneg} \quad (4)$$

And for  $SWI_{plain}$ ,

$$T_{SWI\_BC} = (1 - k_{SWI}) * T_{BC} + k_{SWI} * T_{SWIplain} \quad (5)$$

20 where  $T_{SWI\_BC}$  was a post-processed ensemble member.  $k_{SWI}$  was the weight of the  $T_{SWIneg}$  or  $T_{SWIplain}$ , which was tested between 0–1 and defined according to the best improvement in the skill scores of the post-processed forecast. By Eq. (4) and Eq. (5), we adjusted each ensemble member with the same weight, and hence, the original spread of the ECMWF reforecasts remained unchanged. The skill scores of the  $SWI$  based post-processed forecasts, and their statistical significance, were calculated as explained in Section 2.1.

### 3 Results

#### 3.1 Skill scores of the forecasts

- 25 The annual mean of the expected CRPSS and its 95% level of confidence of the raw and the mean bias-corrected (Section 2.2) weekly mean temperature of the ECMWF reforecasts for 1997–2016 are displayed in Figure 1. In grid points where the CRPSS was higher than zero and the confidence level was higher than 95% (dotted areas), the reforecasts were statistically significantly better than just the statistical forecast based on 1981–2010 climatology. Figure 1 illustrates that for forecast weeks 1–6 the mean bias-corrected ERF reforecasts were on average significantly better than climatology. The annual mean CRPSS values

show that in forecast weeks 1–3 the CRPSSs are for the most part above 0.1, whereas on in forecast weeks 4–6 they are mostly lower, between 0 and 0.1.

### 3.2 The stratospheric observations and the thereafter observed AO index and surface temperature

Figure 2 shows boxplots of the observed mean of the daily AO index 1–2 weeks, 3–4 weeks, and 5–6 weeks after different strengths of the ZMZW at 60°N and 10 hPa. In Fig. 2 the first box (brown) represents the mean AO index after all the cases in 1981–2016 November–February, i.e., 36 years \* 17 weeks=612 cases. The blue, yellow, and red boxes in Fig. 2 show the mean AO index after cases in which the daily ZMZW at 60°N and 10 hPa was during the preceding 10 days below its 10<sup>th</sup> (2.5ms<sup>-1</sup>), 15<sup>th</sup> (6.7ms<sup>-1</sup>), and 20<sup>th</sup> (10ms<sup>-1</sup>) percentile, respectively. The observed mean AO index was statistically significantly weaker at the 99% confidence level 1–2 weeks, 3–4 weeks, and 5–6 weeks after the daily ZMZW at 60°N and 10 hPa had been below its overall wintertime 15th percentile, 6.7ms<sup>-1</sup>. Based on this, we defined the *SWI* to be negative (positive) and to indicate statistically significantly lower (higher) AO index in cases the minimum ZWZW at 60°N and 10 hPa was below (above) its 15<sup>th</sup> percentile, 6.7ms<sup>-1</sup>, during the preceding 10 days.

Figure 3 shows the ERA-Interim (1981–2016 and 1997–2016) and model forecasted mean temperature anomalies 1–6 weeks after *SWI<sub>neg</sub>* and *SWI<sub>plain</sub>*. Cases with ZMZW at 60°N and 10 hPa weaker than 6.7ms<sup>-1</sup>, i.e., *SWI<sub>neg</sub>*, in Fig. 3a–3c and Fig 3g–3i (stronger than 6.7ms<sup>-1</sup>, i.e., *SWI<sub>plain</sub>*, in Fig. 3d–3f and Fig. 3j–l) were on average followed by colder (warmer) than average mean temperature. The ECMWF reforecasts (Fig. 3m–r) capture these mean anomalies clearly, in some areas even too strong in comparison to the ERA-Interim 1997–2016 (Fig. 3g–l).

### 3.3 The *SWI* and the forecasted mean temperatures

The mean temperature anomalies in Fig. 3(a–f) for Northern Europe were used for the *SWI* based post-processing as described in Section 2.4. Figure 4 shows how the post-processing based on the *SWI* affected the forecasting skill scores, in cases of *SWI<sub>neg</sub>* and *SWI<sub>plain</sub>*. The CRPSSs of the mean temperatures of the forecast weeks 3–4 and 5–6 were improved by the *SWI* based post-processing and the best median CRPSS was achieved by  $k_{SWI}=0.3$ , for forecast weeks 3–4 in cases of *SWI<sub>plain</sub>*, and by  $k_{SWI}=0.6$  for all the other cases.

Figure 5 shows the forecast skill of the mean bias-corrected mean temperature reforecasts of forecast weeks 3–4 and 5–6 in all cases (Fig. 5a and 5c), in cases the ZMZW at 60°N and 10 hPa was below 6.7 ms<sup>-1</sup> at the start of the forecast (*SWI<sub>neg</sub>*, Fig. 5e and g), and in cases the ZMZW at 60°N and 10 hPa was above 6.7 ms<sup>-1</sup> at the start of the forecast (*SWI<sub>plain</sub>*, Fig. 5i and 5k). In cases of weak ZMZW at 60°N and 10 hPa at the start of the forecast (Fig. 5e and g) the CRPSSs of forecast weeks 3–4 and 5–

6 reached even higher than 0.4 values in some areas, indicating there higher predictability in comparison to cases in which the ZMZW at 60°N and 10 hPa was stronger than 6.7 ms<sup>-1</sup> at the start of the forecast (Fig. 5i and 5k).

Figure 5 depicts also the mean CRPSS of the mean bias-corrected and *SWI* post-processed reforecasts in all cases (Fig. 5b and 5d), in cases the ZMZW at 60°N and 10 hPa was below 6.7 ms<sup>-1</sup> (*SWI<sub>neg</sub>*) at the start of the forecast ( $k_{SWI}=0.6$  and  $k_{SWI}=0.6$  in Fig. 5e and Fig. 5g, respectively), and in cases the ZMZW at 60°N and 10 hPa was above 6.7 ms<sup>-1</sup>, (*SWI<sub>plain</sub>*) at the start of the forecast ( $k_{SWI}=0.3$  and  $k_{SWI}=0.6$  in Fig. 5j and Fig. 5l, respectively). In comparison to the only mean bias corrected ECMWF reforecasts (see Fig. 5a, 5c, 5e, 5g, 5i, 5k), adding the *SWI* based post-processing to the ECMWF reforecasts (see Fig. 5b, 5d, 5f, 5h, 5j), the CRPSSs for weeks 3–4 and weeks 5–6 were slightly improved and the area of these forecasts being significantly better than just the climatological forecast was expanded.

#### 4 Discussion and Conclusions

Based on ECMWF's extended-range reforecasts for the period 1997–2016, we found that the weekly mean surface temperature forecasts over Northern Europe were on average significantly better than just the climatological forecast in weeks 1–6, however, in weeks 4–6, the CRPSSs were quite low, mostly between 0 and 0.1.

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We studied the mean AO index after different thresholds of ZMZW at 60°N and 10 hPa. We found that the mean AO index was statistically significantly weaker 1–2 weeks, 3–4 weeks, and 5–6 weeks after the daily ZMZW at 60°N and 10 hPa had been below its November-February 15th percentile, 6.7ms<sup>-1</sup>.

20 Cases preceded by weaker (stronger) than 6.7 ms<sup>-1</sup> ZMZW at 60°N and 10 hPa were defined as *SWI<sub>neg</sub>* (*SWI<sub>plain</sub>*). As negative AO index enables cold air outbreaks to Northern Europe (Thompson et al. 2002, Tomassini et al. 2012) and positive AO index tends to bring milder and wetter than average weather to Northern Europe (Limpasuvan et al. 2005), we investigated how the mean surface temperatures were in November-February (1981–2016) in Northern Europe 1–6 weeks after *SWI<sub>neg</sub>*/*SWI<sub>plain</sub>*. We found that the mean surface temperature anomalies in Northern Europe in November–February in 1981–2016 after *SWI<sub>neg</sub>* and *SWI<sub>plain</sub>* were in many places statistically significantly different, with anomalously cold surface temperatures more common 1–6 weeks after *SWI<sub>neg</sub>*. The mean temperature anomalies corresponding to *SWI<sub>neg</sub>*/*SWI<sub>plain</sub>* were used in post processing the ECMWF's mean temperature reforecast for weeks 3–4 and 5–6 in Northern Europe during boreal winter, and thereby, those weeks' forecast skills were slightly improved.

30 We also investigated the forecast skill in cases of ZMZW at 60°N and 10 hPa below or above the threshold of 6.7 ms<sup>-1</sup>. We found that cases of weaker than 6.7 ms<sup>-1</sup> ZMZW at 60°N and 10 hPa at the start of the forecast were followed by higher than average forecasting skill scores of mean surface temperature for forecast weeks 3–4 and 5–6. Also earlier studies have reported



enhanced forecast skill during periods of negative AO, e.g., in 500 hPa geopotential height forecasts in the northern mid-latitudes in both medium-range (Langland and Mauw 2012) extended range (Minami and Takaya 2020).

In future the *SWI* based post-processing method introduced in this paper could also be tested for other northern areas affected by the polar vortex and to precipitation and windiness forecasts, and it could be further developed by, e.g., the Madden-Julian-Oscillation (Madden and Julian 1994; Zhang 2005; Jiang et al. 2017; Vitart 2017; Vitart and Molteni 2010; Robertson et al. 2018, Cassou 2008) and the quasi-biennial oscillation (Watson and Gray 2014; Scaife et al. 2014; Garfinkel et al. 2018; Gray et al. 2018). In this study, the effect of global warming was not filtered from the temperature anomalies used for statistical post-processing. In future work, the impact of filtering the effect of global warming could be tested. Moreover, the next step would be looking for the stratospheric signal from the forecast model.

*Data availability.* ERA-Interim data available at <https://apps.ecmwf.int/datasets/data/interim-full-daily/levtype=sfc/> (last accessed 24 June 2019). ECMWF reforecasts data available at <https://apps.ecmwf.int/mars-catalogue/> (last accessed 28 June 2019). AO index data available at [https://www.cpc.ncep.noaa.gov/products/precip/CWlink/daily\\_ao\\_index/ao.shtml](https://www.cpc.ncep.noaa.gov/products/precip/CWlink/daily_ao_index/ao.shtml) (last accessed 24 June 2019). The daily ZMW at 60°N and 10 hPa data available at [https://acd-ext.gsfc.nasa.gov/Data\\_services/met/ann\\_data.html](https://acd-ext.gsfc.nasa.gov/Data_services/met/ann_data.html) (last accessed 24 June 2019). The data of Figures 1–5 available at <https://etsin.fairdata.fi/dataset/34d0f8b3-a593-46aa-8fcf-358d72f6cac1>.

*Competing interests.* The authors declare that they have no conflict of interest.

*Author contributions.* NK designed the study, analysed the results and prepared the manuscript with contributions from all co-authors. OH participated in the study design and analysing the results. MK contributed to the discussions and fine-tuned the experiments. DSR contributed to the discussions and to the interpretation of the results. HJ provided supervision during the experiments and writing. HG contributed to the study design and was in charge of the management and the acquisition of the financial support for the CLIPS-project leading to this publication.

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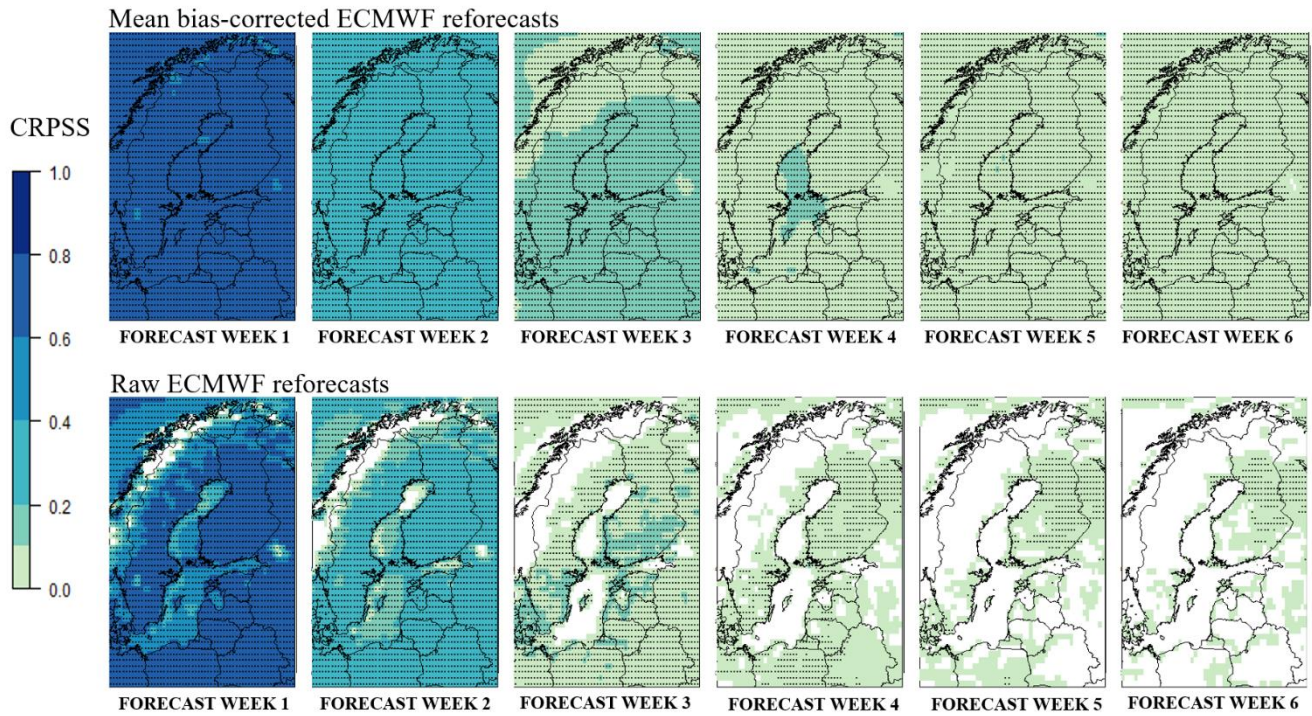


Figure 1: Annual mean of the expected CRPSS of the weekly mean temperature of the mean bias-corrected (upper row) and raw  
5 (lower row) ECMWF reforecasts for years 1997–2016 using ERA-Interim climatology of 1981–2010 as the reference. The dotted areas represent the 95% level of confidence that the CRPSS is above zero.

Mean AO index after different thresholds of the Zonal Mean Zonal Wind at 60°N and 10 hPa (ZMZW)

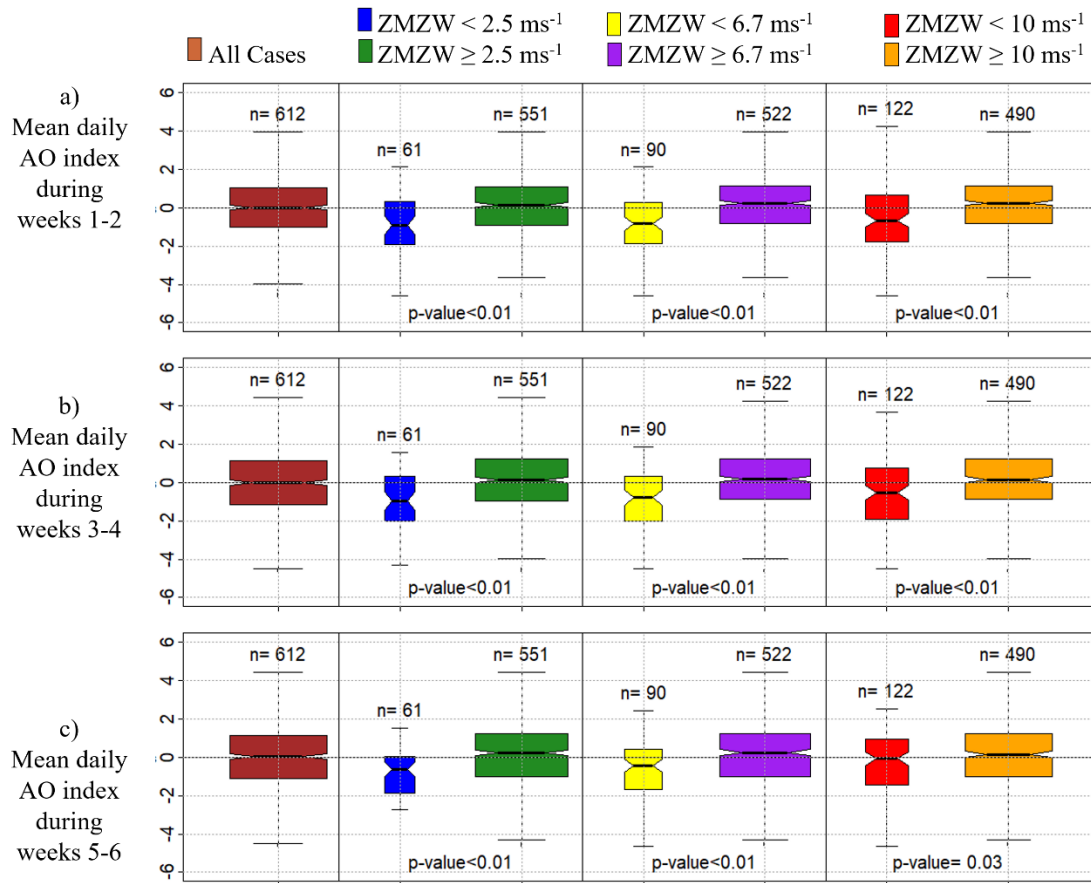


Figure 2: Observed mean AO index in November-March (1981-2016) a) 1–2, b) 3–4 and c) 5–6 weeks after different thresholds of the zonal mean zonal wind at 60°N and 10 hPa (ZMZW). The horizontal line dividing each box into two parts shows the median of the data, the ends of the box show the lower and upper quartiles, and the whiskers represent the highest and the lowest values excluding outliers. The n written above each box indicates the number of observations in each group. The p-value written below each boxplot pair indicates the likelihood of such a pair of distributions arising from a random sampling of a single distribution as given by a Student's t-test, i.e., p-values less than 0.01 indicate that the means of the data sets differ significantly at the 99% level of confidence. The notches of each side of the boxes were calculated by R boxplot.stats. If the notches of two plots do not overlap, this is 'strong evidence' that the two medians differ (Chambers et al., 1983, p. 62).



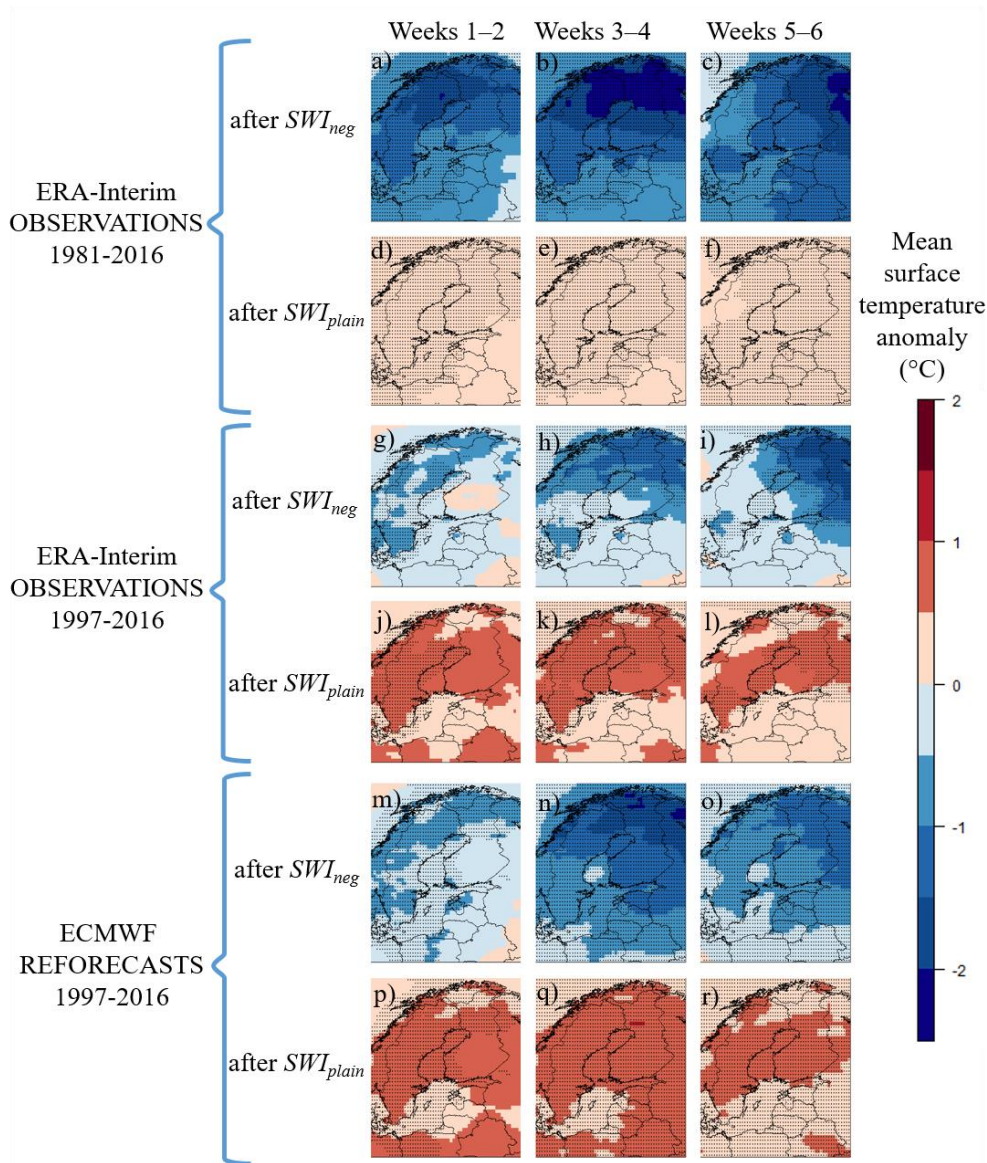


Figure 3. ERA-Interim observed (a-l) and ECMWF reforecasted (m-r) mean temperature anomalies in comparison to the 1981-2016 mean during boreal winters (November-February) in cases the previous week's SWI was negative ( $SWI_{neg}$ , covering about 17% of the winter weeks) or plain ( $SWI_{plain}$ , covering about 83% of the winter weeks). The dotted areas represent the 95% level of confidence where the means of surface temperature anomalies after  $SWI_{neg}$  and  $SWI_{plain}$  differ significantly.

Sensitivity of the CRPSS to the  $k_{SWI}$  in the post-processing

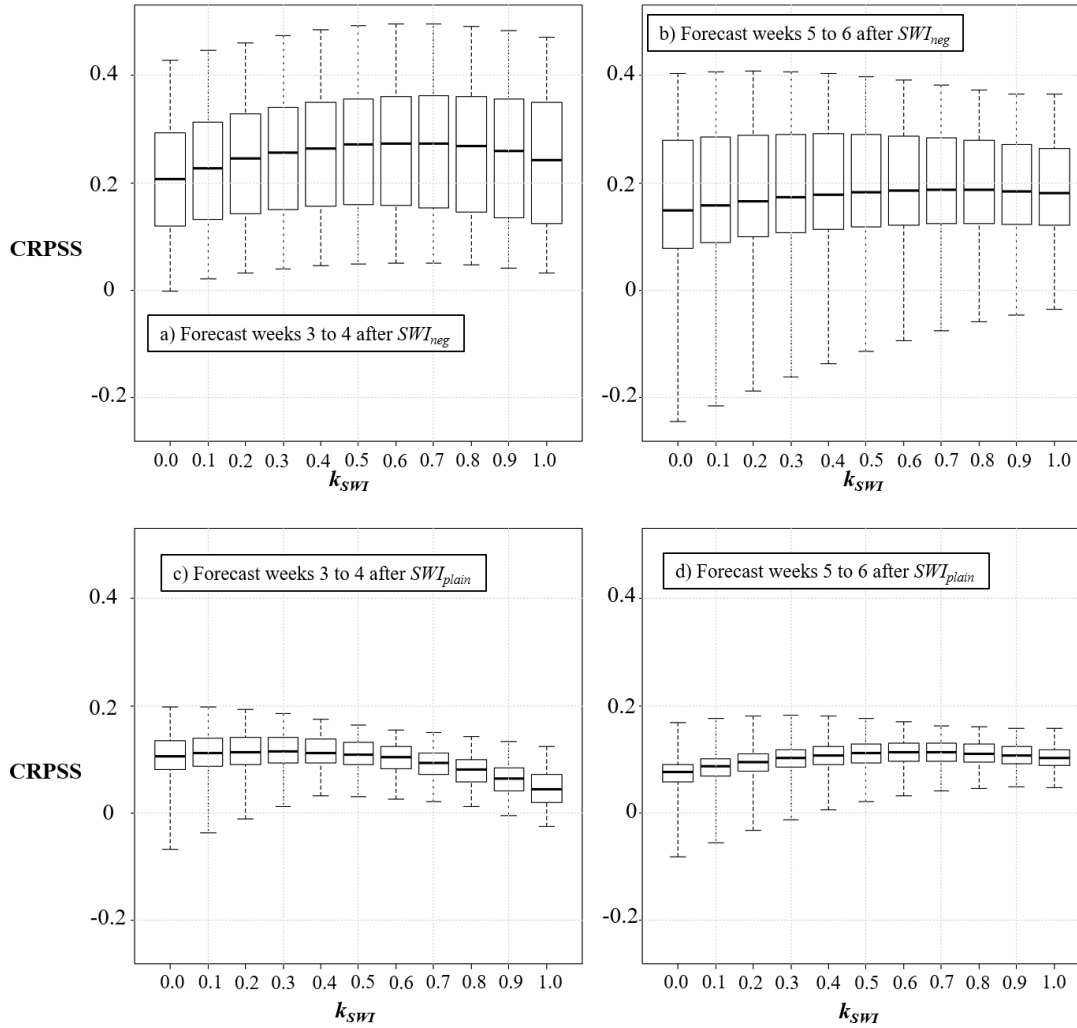


Figure 4. Sensitivity of the expected CRPSS of the post-processed ECMWF surface temperature reforecasts to the  $k_{SWI}$  ranging from 0.0 to 1.0 in forecast weeks 3–4 (a and c) and 5–6 (b and d) in cases of  $SWI_{neg}$  (upper row) and  $SWI_{plain}$  (bottom row). The black boxes show the lower and upper quartiles, and the whiskers illustrate the extremes of the November-February mean CRPSSs of all the grid points in Northern Europe.



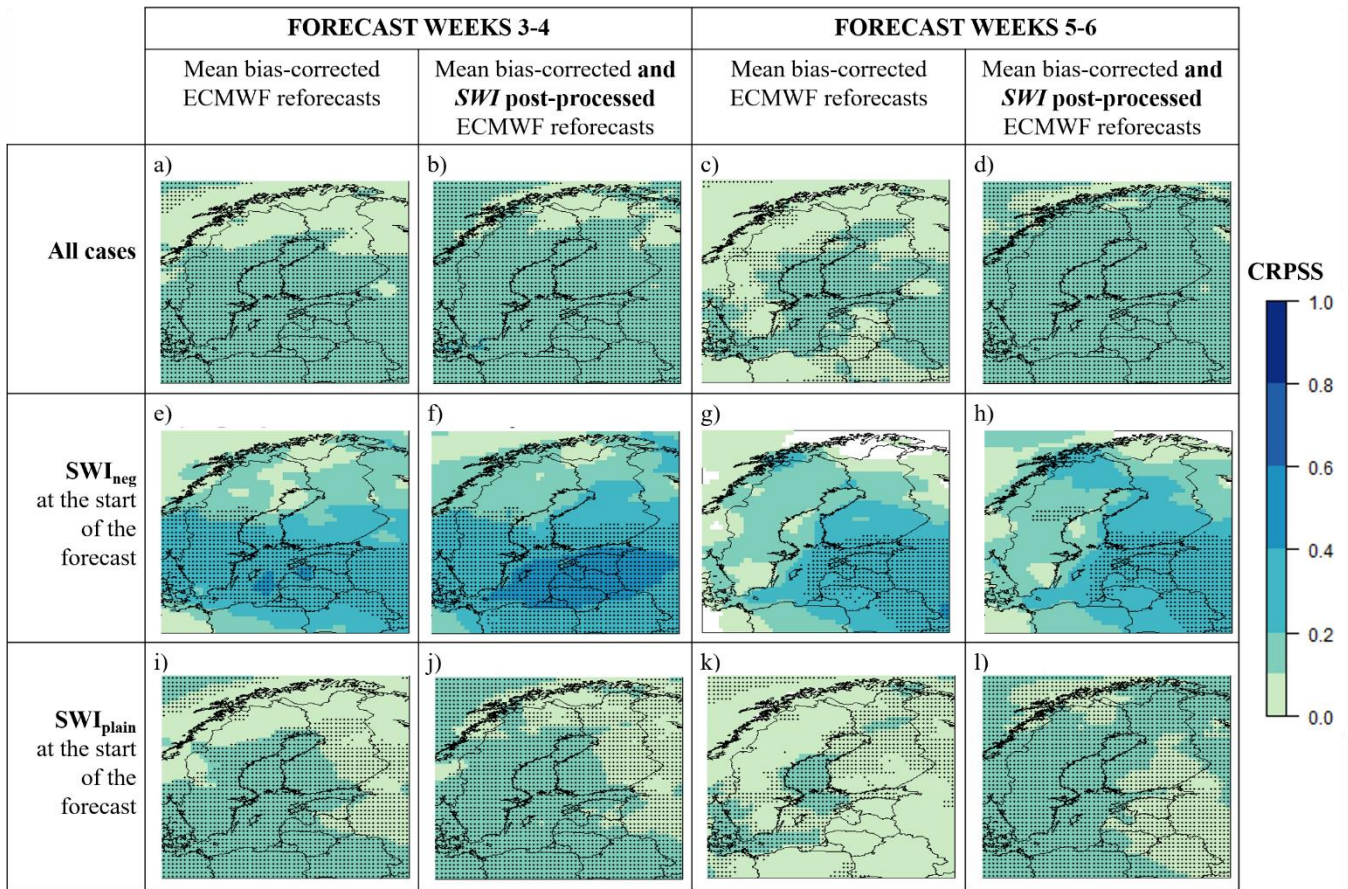


Figure 5. Expected CRPSS of forecast weeks 3–4 and 5–6 of the ECMWF’s two-weeks mean temperature reforecasts for November–February 1997–2016 in all cases (upper row), after *SWI*<sub>neg</sub> (middle row), and after *SWI*<sub>plain</sub> (bottom row) with mean bias-correction only (a, c, e, g, i, and k) and with both mean bias-correction and *SWI* based post-processing (b, d, f, h, j, and l). ERA-Interim climatology of 1981–2010 was used as the reference. The dotted areas represent the 95% level of confidence that the CRPSS is above zero.