



Decoding long-term trends in the wet deposition of sulfate, nitrate and ammonium after reducing the perturbation from climate anomalies

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Abstract. Long-term trends of wet deposition of inorganic ions are affected by 1 2 multiple factors, among which emission changes and climate conditions are dominant ones. To assess the effectiveness of emission reductions on the wet deposition of 3 4 pollutants of interest, contributions from these factors to the long-term trends of wet deposition must be isolated. For this purpose, a two-step approach for preprocessing 5 6 wet deposition data is presented herein. This new approach aims to reduce the impact of climate anomalies on the trend analysis so that the impact of emission reductions on 7 the wet deposition can be revealed. This approach is applied to a two-decade wet 8 deposition dataset of sulfate (SO_4^{2-}) , nitrate (NO_3^{-}) and ammonium (NH_4^{+}) at rural 9 Canadian sites. Analysis results show that the approach allows for robustly identifying 10 inflection points on decreasing trends in the wet deposition fluxes of SO_4^{2-} and NO_3^{-} in 11 northern Ontario and Québec. The inflection points match well with the three-phase 12 mitigation of SO₂ emissions and two-phase mitigation of NOx emissions in Ontario. 13 Improved correlations between the wet deposition of ions and their precursors' 14 emissions were obtained after reducing the impact from climate anomalies. 15 Furthermore, decadal climate anomalies were identified as dominating the decreasing 16 trends in the wet deposition fluxes of SO42- and NO3- at a western coastal site. Long-17 term variations in NH₄⁺ wet deposition showed no clear trends due to the compensating 18 effects between NH3 emissions, climate anomalies, and chemistry associated with the 19 20 emission changes of sulfur and nitrogen.

21

22 1. Introduction

To assess the long-term impacts of acidifying pollutants on the environment, the wet deposition of sulfate (SO_4^{2-}), nitrate (NO_3^{-}) and ammonium (NH_4^+), among other inorganic ions, has been measured for several decades through monitoring networks





26	such as the European Monitoring and Evaluation Programme (EMEP) (Fowler et al.,
27	2005, 2007; Rogora et al., 2004, 2016), the National Atmospheric Deposition
28	Program/National Trends Network in the U.S. (Baumgardner et al., 2002; Lehmann et
29	al., 2007; Sickles & Shadwick, 2015), and the Canadian Air and Precipitation
30	Monitoring Network (CAPMoN) (Vet et al., 2014; Zbieranowski and Aherne, 2011).
31	The high-quality data collected from these networks have been widely used to quantify
32	the atmospheric deposition of acidifying pollutants (Lajtha & Jones, 2013; Lynch et al.,
33	2000; Pihl Karlsson et al., 2011; Strock et al., 2014; Vet et al., 2014). The data have
34	also been utilized to identify trends in the atmospheric deposition of reactive nitrogen
35	(Fagerli & Aas, 2008; Fowler et al., 2007; Lehmann et al., 2007; Zbieranowski and
36	Aherne, 2011) and to examine the impacts of acid rain and the perturbation of the
37	natural nitrogen cycle on sensitive ecosystems (Wright et al., 2018). The long-term data
38	can also be used for assessing the effectiveness of environmental policies (Butler et al.,
39	2005; Li et al., 2016; Lloret & Valiela, 2016).

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The wet deposition of SO42-, NO3- and NH4+ is affected by not only their gaseous 41 precursors' emissions (Butler et al., 2005; Fowler et al., 2007; Li et al., 2016) but also 42 complex atmospheric processes (Cheng & Zhang, 2017; Yao & Zhang, 2012; Zhang et 43 al., 2012). Those atmospheric processes sometimes lead to extremely high wet 44 45 deposition fluxes of ions during a precipitation event or even throughout a particular 46 month. Furthermore, climate anomalies can alter the relative contributions of local sources versus long-range transport to the total wet deposition amounts at reception 47 48 sites, thereby complicating the relationships between wet deposition and the emission of air pollutants of interest (Lloret & Valiela, 2016; Monteith et al., 2016; Wetherbee 49 & Mast, 2016). The emissions of SO2 and NOx have been decreasing substantially in 50





Europe and North America (Butler et al., 2005; Li et al., 2016; Pihl Karlsson et al.,
2011); coincidently, climate anomalies have also occurred more frequently in recent
decades (Burakowski et al., 2008; Lloret & Valiela, 2016; Wijngaard et al., 2003),
thereby leading to more complicated linkages between wet deposition and emission
trends on decadal scales.

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Many trend analysis studies in the literature simply examined annual or seasonal values 57 as the data inputs for two popular trend analysis tools, i.e., the Mann-Kendall (M-K) 58 and linear regression (LR) methods (Marchetto et al., 2013; Waldner et al., 2014 and 59 references therein). These studies focused on the detection of statistically significant 60 trends; for example, Waldner et al. (2014) conducted a comprehensive analysis on the 61 applicability of the techniques to different choices of length and temporal resolutions 62 of a data series. Regarding the resolved trend results, these approaches are not well 63 suited to separating the impact of air pollutants' mitigation from the perturbation by 64 65 climate anomalies. Large uncertainties thus existed in the studies interpreting the major driving forces determining the extracted trends in the wet deposition of SO42-, NO3- and 66 NH4⁺. Regarding that air pollutant's emission mitigation targets often vary in different 67 phases of the entire study period, inflection points may exist in the trends in the wet 68 deposition of ions. The inflection points were rarely studied, despite their importance 69 70 for assessing the effectiveness of environmental policies. An alternative would be to 71 use high time resolution data in the Ensemble Empirical Mode Decomposition (EEMD) 72 method (Wu & Huang, 2009); however, this method still suffers from the end effect in 73 certain scenarios, whereby the extracted trends cannot be explained (Yao & Zhang, 74 2016).

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A new approach is presented herein that aims to reduce the perturbations from climate 76 77 anomalies on data inputs so that robust trends can be elucidated for evaluating the effectiveness of emission control policies. In this approach, raw data are preprocessed 78 79 to generate a new variable, which is then applied to M-K and LR methods. A piecewise linear regression (PLR) is also used to extract trends for cases in presence of inflection 80 81 points. The extracted trends in the wet deposition data on a decadal scale are then properly linked to major driving forces such as emission reductions and climate 82 anomalies. This new approach is first applied to the wet deposition data of SO_4^{2-} , NO_3^{-1} 83 and NH_4^+ in Canada, as an example to demonstrate its capability and advantages over 84 the traditional approaches. The extracted trends in the wet deposition of ions are further 85 studied through correlation analysis with known emission trends of their respective 86 gaseous precursors (SO₂, NOx and NH₃) in Canada and the U.S. Major driving forces 87 for the trends of ion wet deposition and how the wet deposition ions responded to their 88 precursors' emissions in Canada are then revealed. 89

90

91 2. Methodology

92 Wet deposition flux (F_{wet}) data were obtained from CAPMoN
93 (https://www.canada.ca/en/environment-climate-change/services/air-

pollution/monitoring-networks-data/canadian-air-precipitation.html). Data from four
sites have been collected for over twenty years and were chosen herein to illustrate the
novel trend analysis method (Table S1). Site 1 is an inland forest site at Chapais in
Québec. Site 2 is situated in a coastal forest area at Saturna in British Columbia. Sites
3 and 4 are two inland forest sites at the Chalk River and at Algoma, respectively, in
northern Ontario. Details on data sampling, chemical analysis and quality control can
be found in previous studies (Cheng & Zhang, 2017; Vet & Ro, 2008; Vet et al., 2014).





The emissions data of gaseous precursors were downloaded from the Air Pollutant Emission Inventory (APEI, https://pollution-waste.canada.ca/air-emission-inventory/) in Canada and from the USEPA National Emissions Inventory (NEI, https://www.epa.gov/air-emissions-inventories/air-emissions-sources) in the U.S. These data were demarcated at a provincial level in Canada and at a state level in the U.S. Data for the years of 1990 to 2011, which correspond to the period of selected F_{wet} data, were used in this study.

108

The M-K method is a popular nonparametric statistical procedure that can yield 109 qualitative trend results, such as "an increasing/decreasing trend with a P value of 110 <0.05," "a probable increasing/decreasing trend with a P value of 0.05-0.1," "a stable 111 trend with a P value of >0.1, as well as a ratio of <1.0 between the standard deviation 112 and the mean of the dataset," and "a no trend for P>0.1 with all other conditions" 113 (Kampata et al., 2008; Marchetto et al., 2013). The LR method has also been widely 114 used to extract trends (Marchetto et al., 2013; Waldner et al., 2014). Zbieranowski and 115 Aherne (2011) used LR to extract trends by separating different phases because of the 116 presence of inflection points in the entire study period, and the approach is same as PLR 117 (Vieth, 1989). In this study, the three methods were employed to compute the trends 118 of ion wet deposition using software downloaded from https://www.gsi-119 net.com/en/software/free-software/gsi-mann-kendall-toolkit.html and Excel 2016, first 120 121 using the annual F_{wet} directly as input data, then using a modified input data set, as 122 described below.

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124 The modified input data set was produced in two steps. The first step was an effort to 125 reduce the perturbation from the monthly climate anomalies to the input data. This was





done by creating a new variable that was defined as the slopes of the regression 126 equations of a series of study years against a climatology (base) year using monthly 127 Fwet data. Note that the monthly Fwet data were aggregated from daily raw data before 128 129 the regression analysis. To ensure the presence of enough data points in each regression equation, the data corresponding to two-year periods (or 24 monthly Fwet values) were 130 131 grouped together, as detailed below. At a selected site and for a given chemical 132 component, monthly Fwet data were generated for the first two years and were grouped together and rearranged from the smallest to the largest values to form an array of data 133 134 with 24 data points, i.e., A(i) with i=1 to 24. Repeating the above procedure for the subsequent years using a two-year interval to eventually obtain a series of data arrays, 135 A(i) now becomes A(i, j) with i=1 to 24 and j=1 to N, where N is the total number of 136 data arrays. The climatology data array (CA(i)) was then defined as the average of all 137 of the arrays as follows: 138

139
$$CA(i) = \frac{1}{N} \sum_{j=1}^{N} A(i,j), i = 1 \text{ to } 24.$$

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LR with zero interception was applied for each individual data array against the climatology data array. In cases where the maximum monthly deposition flux deviated greatly from the general regression curve, the slopes (m-values) were calculated after excluding the maximum monthly deposition flux, which is an approach that reduced the perturbation to the m-values from the monthly scale climate anomalies. The second step was to screen out the outliers in m-values, which reduced the perturbation to the m-values from the annual-scale climate anomalies.

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An analysis of Site 1 is used to illustrate the new approach and demonstrate itsadvantages against the existing common approaches used in the literature. Twelve four-





151	year periods of data (1988-1989, 1990-1991, etc.) are available from this site. The
152	regression of each data set against the climatology data set was first performed using
153	all of the monthly values to obtain an m-value (the slope) (Fig. 1a-d). For eight out of
154	the 12 data sets, the m-values were recalculated after excluding the maximum monthly
155	value of $F_{\text{wet}},$ which appeared to be an apparent outlier of the linear regression. The R^2
156	values, which are conventionally used in LR, were then significantly increased for these
157	eight sets, e.g., from the original values of 0.79-0.94 to the improved values of 0.92-
158	0.98. To demonstrate that the excluded maximum value was an outlier, the case of the
159	1990-1991 data set was taken as an example. The new regression equation (y=1.47x,
160	R^2 =0.98, Fig. 1a) predicted a maximum value in the range of 330-368 mg m ⁻² month ⁻¹
161	using three times the standard deviation (± 3 SD, 0.08) at a 99% confidence level. The
162	actual observed maximum value of 532 mg m ⁻² month ⁻¹ was much larger than the upper
163	range of the predicted value and was thus believed to be caused by monthly scale
164	climate anomalies. The maximum value was treated as an outlier and excluded for
165	analysis. Using the similar procedure, all outliers in this study were identified. The
166	exclusion of the observed maximum value greatly reduced the perturbation of the short-
167	term climate anomalies to the calculated m-value in this two-year period, i.e., the m-
168	value decreased from 1.67 to 1.47, which in turn increased the relative contribution of
169	the air pollutants' emissions to the calculated m-value. In summary, this new approach
170	meets the objective of identifying outlier data points by applying the criteria of being
171	outside the boundaries of ± 3 times the standard deviation of the general trend.

172

The revised m-values were further scrutinized by eliminating the outliers caused by the
annual-scale climate anomalies. For example, the m-value of 1.31 in 1998-1999 greatly
deviated from other m-values, narrowly oscillating approximately 0.96±0.07 (average





176	\pm 1 SD) during the period of 1994-2005, even with the ± 3 SD being considered (Fig.
177	1a-d). Using the value of 0.96 as the reference, climate anomalies likely increased the
178	F_{wet} of SO ₄ ²⁻ by 37% in 1998-1999. The m-values were then calculated by shifting one
179	year in time to 1997-1998 (1.07) and to 1999-2000 (1.24). The $F_{wet}\xspace$ in 1998 was less
180	affected by climate anomalies than that in 1999. Thus, the m-value in 1997-1998 was
181	within 0.96 \pm 0.21 (average \pm 3 SD) and used to replace the m-value in 1998-1999 for
182	the trend analysis. Similar to the first step discussed above, this approach meets the
183	objective of identifying outlier m-values by applying the criteria of being outside the
184	range of ± 3 SD plus the average m-value during a decade or a longer period.

185

More justification of the new approach can be found in the Supporting Information, 186 including Figs. S1-3, wherein the statistical comparison between this and other 187 approaches was presented. Theoretically, the extracted trend using the data 188 189 preprocessed with the new approach is determined by the local emissions of air pollutants, the regional transport of air pollutants, and a small portion of climate 190 anomalies that are unable to be removed by the new approach. It is assumed that the 191 192 extracted trend is less affected by microphysical/chemical processes, since two-year data were used together to calculate the m-value. 193

194

In theory, if the data from different sites in the same region are grouped together for trend analysis, the results may be better linked to the trends of the regional emissions of related air pollutants. In the following sections, trend analysis results from individual sites as well as those from grouped sites are discussed. Sites 1, 3 and 4 showed similar trends in the wet deposition of SO_4^{2-} and NO_3^{-} , and these three sites were grouped together.





201 3. Results and discussion

202 3.1 Trends at Site 1 after reducing perturbations from climate anomalies

Trends in the m-values shown in Fig. 2 represent the trends after removing the perturbations from climate anomalies at Site 1 in northern Québec from 1988 to 2011. SO_4^{2-} and NO_3^{-} showed decreasing trends from a LR analysis, with R² values of 0.81 and 0.71, respectively, and P values <0.01. The decreasing trends were also confirmed by the M-K method analysis. NH_4^+ exhibited a stable trend from M-K analysis, as well as no significant trend with P value >0.05 from LR analysis.

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The m-values of SO42- and NO3- also allowed for the identification of trends in different 210 phases. The inflection point for each phase is critical to a) link the annual F_{wet} of ions 211 and the emissions of the corresponding precursors and b) assess the effectiveness of 212 environmental policies. For example, the trends in the m-values of SO42- can be clearly 213 classified into three phases. The m-values oscillated approximately 1.38±0.08 during 214 Phase 1 (1988 to 1993) and approximately 1.02±0.08 during Phase 2 (1994 to 2005), 215 with a significant difference between the two phases under the t-test (P value <0.01), 216 217 thereby implying an abrupt decrease of approximately 30% at the inflection point between the two phases. In contrast, the m-values linearly decreased by approximately 218 219 20% every two years, starting from the end of Phase 2 to Phase 3 (2006-2011). Again, a significant difference existed between Phase 2 and Phase 3 under the t-test (P value 220 <0.01). Overall, PLR should be applied separately for the different phases in the 221 presence of the inflection points, rather than LR for the entire period, and the result is 222 223 presented as:

224
$$\begin{cases} m - value = 1.38, 1988 \le x < 1994 \\ m - value = 1.02, 1994 \le x \le 2004 \\ m - value = -0.185 * \left(\frac{x}{2} - 1001\right) + 1.15, 2004 < x \le 2010 \end{cases}$$





- where x represent the calendar year from 1988 to 2010.
- 226

The trend in the m-values of NO_3^- can be classified into two phases, with the inflection point at 2003, according to the t-test result, i.e., the values oscillated approximately 1.09±0.09 during the period from 1988 to 2003 and then exhibited a significant decrease of approximately 50% overall afterwards, with P value <0.01. The PRL result is expressed as below:

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$$\begin{cases} m - value = 1.09, 1988 \le x < 2004 \\ m - vlaue = -0.128 * \left(\frac{x}{2} - 1001\right) + 1.08, 2004 \le x \le 2010 \end{cases}$$

233 The m-value of NO_3 in 1998-1999 was approximately 30% larger than the mean value in 1988-2003 and exceeded the mean value plus 3 SD in 1998-2003, and thus was not 234 235 included in the trend analysis. The sharp increase in F_{wet} of NO_3^{-} occurred mainly in 1999, which was probably caused by a large perturbation from climate anomalies. 236 Moreover, the monthly F_{wet} values of NO₃⁻ in March, April, July and August 1999 were 237 actually lower than the corresponding long-term averages in 1988-2003 (excluding 238 239 1999) (Fig. S4a). This outcome indicates that the large increase in annual Fwet of NO3in 1999 was unlikely to have been determined by the emissions of its gaseous 240 precursors. The same can be said for the large increase in Fwet of SO42- in 1999 (Fig. 2a, 241 242 S4b).

243

The trends in the m-value of SO_4^{2-} at Site 1 (Fig. 2a) were clearly different from those of the SO₂ emissions in Québec (Fig. 2c) but matched well to those in Ontario (Fig. 2c), which is also supported by their Pearson correlation coefficients, e.g., no significant correlation (r = 0.46 and P value >0.05) for the former case and a good correlation (r = 0.96 and P value <0.01) for the latter case. Note that r instead of R² is conventionally





- used in correlation analysis and is therefore used here. Zhang et al. (2008) reported that
 this remote area can receive the long-range transport of air pollutants from Ontario but
 that transport is less likely from the intensive emission sources in Québec.
- 252

The trends in NOx emissions during 1990-2003 had similar bell-shape patterns in 253 254 Québec and Ontario, although with different magnitudes of emissions (Fig. 2f). A different trend pattern was seen for the m-value of NO3⁻ at Site 1 than for the 255 abovementioned provincial emissions during the same period (Fig. 2d), and there was 256 no significant correlation (r<0.41, with P value >0.05) between the m-value of NO₃-257 and the emissions of NOx in Québec or Ontario. Different results were found for the 258 period of 2002-2011 than those of 1990-2003 discussed above. In 2002-2011, the m-259 value of NO₃⁻ decreased by ~50% and the NOx emissions decreased by ~40% in 260 Québec and Ontario; also, good correlations (r = 0.94-0.95 with P values < 0.01) were 261 observed between m-values and emissions. The contrasting results between the two 262 263 different periods discussed above implied one possibility, i.e., that the perturbation from climate anomalies, which was unable to be removed by the new approach, 264 overwhelmed the effect of NOx emissions on trends in m-values of NO₃⁻ in 1990-2003, 265 while the reverse was true in 2002-2011. However, other possibilities cannot be 266 excluded. Fwet of NO3⁻ and precipitation depth exhibited only a weakly significant 267 correlation, with r = 0.58 and P<0.05 in 1988-2003 (the values in 1999 were excluded). 268 269 Annual precipitation varied by only ~20% during the fifteen years, and this factor alone 270 was unlikely to explain the ~100% interannual variation of Fwet of NO3⁻ during that period. 271

272

273 The m-values of NH_4^+ at Site 1 had no significant correlation (r = 0.21 and P value





274	>0.05) with the emissions of NH ₃ in Québec but exhibited a weakly significant
275	correlation (r = 0.60 and P value <0.05) with the emissions of NH_3 in Ontario. Nearly
276	all of the $\rm NH_{4^+}$ was associated with $\rm SO_{4^{2-}}$ and $\rm NO_{3^-}$ in the atmosphere (Cheng and
277	Zhang, 2017; Teng et al., 2017; Tost et al., 2007; Zhang et al., 2012), and the trends in
278	the m-value of NH_4^+ could be affected by many other factors besides NH_3 emissions
279	and climate anomalies.

280

LR analysis of the annual Fwet of these ions revealed decreasing trends for SO42- and 281 NO3⁻ (second row in Fig. 2). The M-K method analysis also confirmed the decreasing 282 trends with annual Fwet as input. However, the three-phase trends in Fwet of SO42- and 283 the two-phase trends in Fwet of NO3- and related inflection points, identified using the 284 m-values discussed above, were not identified by the t-test when simply using annual 285 Fwet data as input. The correlations between annual Fwet and emissions were 0.89 for 286 SO₄²⁻ vs. SO₂ in Ontario and 0.74-0.76 for NO₃⁻ vs. NO_x in Québec and Ontario (P 287 values <0.01), while the corresponding r values were as high as 0.95 and 0.84-0.85 288 between m-values and emissions. After reducing the perturbations from climatic factors 289 to the annual Fwet, stronger correlations were obtained between Fwet and emissions. 290

291

The stable trend in annual F_{wet} of NH₄⁺ and the decreasing trend in annual F_{wet} of NO₃⁻ gradually increased the relative contributions of reduced nitrogen in the total nitrogen wet deposition budget, e.g., from 40% in 1998-1999 to 52% in 2010-2011. A similar trend has also been recently reported in the U.S. (Li et al., 2016). Such a trend was mostly due to the mitigation of NOx rather than climate anomalies.

297

298 *3.2 Decadal climate anomalies drove trends at Site 2*





299	Fig. 3 shows the trend analysis results at Site 2. An obvious shift in the m-values and
300	annual F_{wet} occurred during 2001-2002, as detected by the t-test, i.e., the m-values of
301	$\mathrm{SO_4^{2\text{-}}}$ oscillated approximately 1.15±0.11 in 1990-2001 and 0.76±0.02 in 2002-2011
302	(or 0.83 ± 0.12 if the value in 2006-2007 was included), but with a significant difference
303	between the two periods with P value <0.01. The annual $F_{\rm wet}$ of $SO_4{}^{2\text{-}}$ oscillated
304	approximately 632 \pm 63 mg m ⁻² in 1990-2001 and 452 \pm 74 mg m ⁻² in 2002-2011, and the
305	values between the two periods showed significant differences. The shift led to the m-
306	values and annual F_{wet} of SO ₄ ²⁻ exhibiting a consistent decreasing trend by ~40% overall
307	from 1990 to 2011 using the LR and the M-K method.

308

The emissions of SO₂ oscillated approximately 1.13±0.07 in 1990-2001 and 1.06±0.03 309 310 in 2002-2011 in British Columbia, which did not support the large decrease of approximately 40% in wet deposition of SO42- in 2002-2011. Statistically, no 311 correlation existed between annual Fwet of SO42- and the emissions of SO2 in British 312 313 Columbia, with r = 0.52 and P value >0.05. Although the transboundary transport of air pollutants from the U.S. cannot be excluded, the almost constant m-values from 2002 314 to 2011 (excluding 2006-2007) at Site 2 were inconsistent with the approximately 70% 315 decrease in emissions of SO₂ in the state of Washington in the U.S. during that period 316 (not shown). Precipitation cannot explain the jump in wet deposition either, because 317 there was no corresponding jump in precipitation during 2001-2002 (Fig. 3b). 318

319

van Donkelaar et al. (2008) analyzed aircraft and satellite measurements from the
Intercontinental Chemical Transport Experiment and proposed the long-range transport
of sulfur from East Asia to the west coast of Canada. The average wind fields in 19902011 at different altitudes also showed air masses primarily originating from the Pacific





324	Ocean in the west (Fig. 4a). However, the climate anomalies of wind fields in 1990-
325	2001 compared against 1990-2011 clearly showed a counterclockwise pattern in the
326	corresponding coastal area, including Site 2, while a clockwise pattern existed in 2002-
327	2011 against 1990-2011 (Fig. 4b, c). The decadal climate anomalies of wind fields in
328	2002-2011 very likely caused a large decrease in the contribution of air pollutants from
329	the eastern coast to Site 2, resulting in a distinct demarcation at 2002. This hypothesis
330	was also supported by a large rebound of the m-value in 2006-2007, due to the increase
331	in F_{wet} of SO_4^{2-} in 2007. The climate anomalies of wind fields in 2007 against 1990-
332	2011 showed a counterclockwise pattern in the north, while the clockwise pattern was
333	pushed to the south (Fig. 4d). A greater contribution of air pollutants from the eastern
334	coast to Site 2 might have led to the large increase in $F_{wet} of SO_4{}^{2\text{-}}$ during a few month-
335	long periods in 2007.

336

The present study is the first one to identify the decreasing trend in the annual F_{wet} of SO₄²⁻ as being very likely caused by decadal climate anomalies rather than by the emission reductions of SO₂. The decadal climate anomalies may substantially alter the long-range transport of air pollutants to the reception site. Note that the causes for the decadal climate anomalies in this region are beyond the scope of the present study, but some information can be found in the literature (Bond et al., 2003; Coopersmith et al., 2014; Deng et al., 2014).

344

For the wet deposition of NO₃⁻, the m-values also showed a clear shift, i.e., the m-values oscillated approximately 1.09 ± 0.14 in 1990-2001 and 0.88 ± 0.06 in 2002-2011, with a significant difference between the two periods under the t-test with P value <0.01. The annual F_{wet} of NO₃⁻ varied substantially, and the shift could not be identified





349	statistically. However, the annual Fwet of NO3 ⁻ exhibited a decreasing trend by M-K
350	method analysis. Similar to the case of SO_4^{2-} , no significant correlation (r = 0.49, P
351	value >0.05) existed between the annual F_{wet} of NO ₃ ⁻ and the emissions of NOx in
352	British Columbia.

353

354 In addition to decadal climate anomalies, the interannual climate variability also affected the trends in m-values and annual Fwet of NO3-. The perturbations from 355 interannual climate variability cannot be completely removed by the new approach, and 356 357 they complicate the relationship between the F_{wet} of NO₃⁻ and the emissions of NO_x in British Columbia. For example, the m-values in 1990-1991, 1996-1997, 1998-1999 and 358 359 2000-2001 were nearly constant at 1.17 ± 0.03 . However, the NOx emissions in British 360 Columbia in 1998-1999 were 26% greater than those in 1990-1991. There was a sharp decrease in the NOx emissions (by ~30%) from 2002 to 2011 in British Columbia. 361 362 However, the m-values oscillated approximately 0.88 ± 0.06 and showed no clear trend based on either the M-K method or LR analysis. The interannual climate variability 363 apparently negated the impact of reduced emissions during these periods. 364

365

The m-values and the annual F_{wet} of NH_4^+ oscillated approximately 0.99 ± 0.13 and 81±16 mg m⁻³, respectively, in the period of 1990-2011, and showed no trend (Fig. 3). Neither the m-values nor annual F_{wet} of NH_4^+ showed the two-period distribution pattern or had any significant correlation with the emissions of NH_3 in British Columbia at a 95% confidence level. Similarly to Site 1, the annual variation in F_{wet} of NH_4^+ at Site 2 cannot be simply explained by known emission trends.

372

373 In summary, decadal climate anomalies overwhelmingly determined the long-term





trends in the wet deposition of SO_4^{2-} and NO_3^{-} , with the perturbation from monthly and annual climate anomalies removed at Site 2. The interannual climate variability further complicated the trends, resulting in undetectable influences of the emission trends on the deposition trends. Since the decrease in F_{wet} of NO_3^{-} appeared to be primarily caused by climate anomalies, the relative contributions of NH_4^+ and NO_3^- in the total N wet deposition varied little, i.e., 33% versus 67% in 2010-2011 and 31% versus 69% in 1990-1991.

381

382 *3.3 Regional trends in wet deposition in northern Ontario* and Québec

Trends in the m-values or annual Fwet of ions at Sites 3 and 4 in the northern regions of 383 Ontario were generally similar to those found at Site 1 (Fig. S5 and S6). For example, 384 the three-phase trend in m-values of SO4²⁻ and the two-phase trend in m-values of NO3⁻ 385 were also obtained at Sites 3 and 4 after excluding a few m-values that were caused by 386 large perturbations from climate anomalies. Thus, Sites 1, 3 and 4 were combined 387 388 together to study regional trends in the northern areas of Ontario and Québec (Fig. 5ac). Similar to those found at the individual sites, the temporal profile of regional m-389 values of SO₄²⁻ can be clearly classified into three phases (Fig. 5a) as follows: Phase 1 390 from 1988 to 1993 with m-values oscillating approximately 1.31±0.08, Phase 2 from 391 1994 to 2003 with near-constant m-values of 1.05±0.04, and Phase 3 for 2004 onward 392 with a decreasing trend by an overall ~50%. Significant differences of m-values existed 393 between any two of the three phases, based on the t-test results (P value <0.01). The 394 395 PRL result is expressed as below:

396
$$\begin{cases} m - value = 1.31, 1988 \le x < 1994 \\ m - value = 1.05, 1994 \le x < 2004 \\ m - value = -0.129 * \left(\frac{x}{2} - 1001\right) + 1.03, 2004 \le x \le 2010 \end{cases}$$

397 The three-phase pattern of m-values matched well with the three-phase emission profile





of SO₂ in Ontario. Statistically, an ~70% decrease in m-value and an ~70% decrease in emissions were found from 1990 to 2011, with a correlation of r = 0.95 (P value <0.01).

The profile of the regional m-values of NO_3^- also clearly exhibited two phases, according to the following t-test results between them: Phase 1 from 1988 to 2003, with m-values narrowly varying approximately 1.11 ± 0.05 and Phase 2 from 2004 to 2011, with a decreasing trend by an overall ~40% against that in 2002-2003 (Fig. 5b). The PRL result is expressed as below:

406
$$\begin{cases} m - value = 1.11, 1988 \le x < 2004 \\ m - value = -0.11 * \left(\frac{x}{2} - 1001\right) + 1.03, 2004 \le x \le 2010 \end{cases}$$

From 2002 to 2011, the m-value had a moderately good correlation with the NOx 407 emission in Ontario (r = 0.91, P<0.01), and the two variables decreased by 30-40% in 408 this period. From 1990 to 2003, the near constant m-value was, however, inconsistent 409 with the bell-shape profile of the NOx emissions in Ontario and Québec, which could 410 be due to either the perturbation from climate anomalies or unrealistic emissions 411 inventory. Considering that the first possibility was minimal over a large regional scale, 412 especially when the consistency was determined in a different time frame (2002-2011) 413 in the same region, it is thus doubtful that the bell-shape profile of the NOx emissions 414 in 1990-2003 was realistic. 415

416

The regional m-values of NH_{4^+} largely oscillated from 1988 to 2003 (Fig. 5c). The mvalues of NH_{4^+} , however, decreased by ~30% from 2002 to 2011, leading to a probable decreasing trend in m-value from 1988 to 2011. No correlation was found between the m-values of NH_{4^+} and the emissions of NH_3 in Ontario, which is consistent with the findings at the individual sites discussed above.





Since the decrease in F_{wet} values of NO_3^- at Sites 3 and 4 were very likely due to the mitigation of NOx in Ontario, the decrease also changed the relative contributions between NH_4^+ and NO_3^- in the total N wet deposition budget. For example, NH_4^+ and NO_3^- contributed 52% and 48%, respectively, to the total budget in 2010-2011 and 34% and 66%, respectively, in 1984-1985 at Site 3. The corresponding numbers at Site 4 were 58% and 42% in 2010-2011 and 47% and 53% in 1985-1986.

428

429 4 Conclusions

Climate anomalies during the two-decade period resulted in annual Fwet of SO42- and/or 430 NO_3^- deviating from the normal value by up to ~40% at the rural Canadian sites. The 431 new approach of rearranging and screening Fwet data can largely reduce the impact of 432 climate anomalies when used for generating the decadal trends of Fwet. With the climate 433 perturbation being reduced, Fwet of SO42- exhibited a three-phase decreasing trend at 434 every individual site, as well as on a regional scale in northern Ontario and Québec. 435 The three-phase pattern of the decreasing trend in Fwet of SO42- matches well with the 436 emission trends of SO₂ in Ontario, as supported by the good correlation between wet 437 deposition and emission, with $r \ge 0.95$ and P<0.01. F_{wet} of NO₃⁻ exhibited a two-phase 438 decreasing trend, but only during the second phase Fwet of NO3-, and the emissions of 439 NOx in Ontario and Québec matched well, with a good correlation of r ≥0.91 and 440 P<0.01. Compared to the results obtained without applying the new approach, it is 441 442 concluded that, after reducing the perturbation from climate anomalies, 1) better 443 correlation was obtained between Fwet of ions and the emission of the corresponding gaseous precursors in northern Ontario and Québec, and 2) the inflection points in the 444 decreasing trends of F_{wet} of SO₄²⁻ and NO₃⁻ were clearly identified. 445

446





However, the new approach cannot completely remove the perturbations from climate 447 anomalies, especially when this is the dominant factor and/or on long timescales, as 448 was the case at a coastal site of Saturna in British Columbia. At this location, the 449 decreasing trends in Fwet of SO42- and NO3- were caused by the decadal climate 450 anomalies, as well as being affected by interannual climate variability, which 451 452 overwhelmed the impact of the emission changes of the gaseous precursors in this province. This is the first study that has identified that decadal climate anomalies can 453 dominate trends in Fwet of SO42- and NO3-. 454

455

The long-term variations in Fwet of NH4⁺ generally showed no clear long-term trends. 456 Moreover, no apparent cause-effect relationships were found between the wet 457 deposition of NH₄⁺ and the emission of NH₃. This outcome is not surprising because 458 additional key factors besides those discussed in this study also impact the trends of 459 F_{wet} of NH₄⁺. For example, NH₄⁺ may be more greatly impacted by changes in SO₂ and 460 461 NOx than are NH₃ emissions in NH₃-rich scenarios. It should be noted that F_{wet} of N via NH4⁺ exceeded those via NO3⁻ in 2010 and 2011 in northern Ontario and Québec, 462 where the decrease in F_{wet} of NO₃⁻ was associated with decreasing NOx emissions. In 463 contrast, F_{wet} of NH_4^+ did not exceed F_{wet} of NO_3^- in 2010 and 2011 in the coastal area 464 in British Columbia, where the decreasing trends of Fwet of NO3⁻ were determined to 465 result mainly from the perturbation by climate anomalies. 466

467

- 468 Data availability. Data used in this study are available from the corresponding authors.
- 469 *Supplement*. The supplement materials are available online.
- 470 Author contribution. X. Y. and L. Z. designed the study, analyzed he data and prepared the manuscript.
- 471 *Competing interests.* The authors declare that they have no conflict of interest.
- 472 Acknowledgments. X.Y. is supported by the National Key Research and Development Program in





473	China (No. 2016YFC0200500), and L.Z	. by the Ai	r Pollutants	program	of Environment	and	Climate
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474 Change Canada.

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- **Figure 2.** m-values and annual F_{wet} extracted trends at Site 1, and the annual emissions of air pollutants in Québec and Ontario, Canada. Full and empty markers in blue in (a), (b) and (g) represent the calculation of m-values without and with the outlier, respectively. Empty markers in red represent the outliers in m-values and are excluded for trend analysis, as detailed in Section 2. R² reflects the coefficient of determination of a variable against the calendar year from LR analysis, and the fitted lines represent the LR function.
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Figure 5. Regional m-values at Sites 1, 3 and 4: (a): SO₄²⁻, (b): NO₃⁻, and (c): NH₄⁺.





excluded for LR analysis. Fitted lines represent the LR function with zero interception.



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Empty markers in red represent the outliers in m-values and are excluded for trend analysis, as detailed in Section 2. R² reflects the Figure 2. m-values and annual F_{wet} extracted trends at Site 1, and the annual emissions of air pollutants in Québec and Ontario, Canada. Full and empty markers in blue in (a), (b) and (g) represent the calculation of m-values without and with the outlier, respectively. coefficient of determination of a variable against the calendar year from LR analysis, and the fitted lines represent the LR function.























