Supplement material for the paper

# Long-term aerosol optical hygroscopicity study at the ACTRIS SIRTA observatory: synergy between ceilometer and in-situ measurements

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### 1 Effect of beta-attenuated water vapor correction on calculated $f_{\beta}(RH)$ and $\gamma$ parameter

The  $\beta^{att}$  signal presents a dependency on water vapor absorption as shown in Sec. 3.1 (Eq.8), associated with the wavelength emission. This dependency may cause direct effects over calculations retrieved by using  $\beta^{att}$ . Here, we present the correction applied to  $\beta^{att}$  and the effects on  $f_{\beta}(RH)$  and  $\gamma$  calculations. Figure S1 (leftpanels) represents the temporal-evolution of  $\beta^{att}$ , beta attenuated water vapor corrected ( $\beta^{att}_{wv}$ ) signals over 3 h time-window, together with the temporal-evolution of q. . Figure S1 (right-panels) shows the biases for beta (*bias*<sub> $\beta$ </sub>) and  $\Delta_q$  obtained.

The quantifications are performed by means of the bias<sub> $\beta$ </sub> ( $\beta^{att} - \beta^{att}_{wv}$ ) and  $\Delta_q$  ( $q(t) - q(t_d)$ ) calculations. Fig. S1ac presents two cases (case 1 and case 2) with low absolute-differences in q, which produces slight changes in  $\beta^{att}_{wv}$  signal respect to  $\beta^{att}$ . On the other hand, Fig. S1eg (case 3 and case 4) show that high absolute-differences in q are linked to high changes on  $\beta^{att}_{wv}$ . The right side of the panel (Fig. 1Sbd, case 1 and case 2), presents the bias quantification, showing that low bias<sub> $\beta$ </sub> are associated with low  $\Delta_q$  and, on the contrary, Fig. 1Sfh (case 3 and case 4), show that increases in  $\Delta_q$  makes that bias<sub> $\beta$ </sub> becomes higher. This analysis leads us to conclude that no-water vapor correction will produce an overestimation of the total  $\beta$  signal, being  $\Delta_q$  the parameter that rules the  $\beta$  correction. From now, we will use  $\beta$  instead of  $\beta^{att}$  for simplicity.

To see the effect of  $\beta_{wv}^{att}$  on  $f_{\beta}(RH)$ , we applied the Hänel parameterization (Eq. 7, Sec. 3.1 of the manuscript) to the same 4 cases studied above. Figure S2 presents the  $f_{\beta}(RH)$  and the enhancement factor water vapor corrected,  $f_{\beta_{wv}^{att}}(RH)$ . The results reinforce those obtained above (Fig S1), where low/high changes in  $\Delta_q$  are linked with low/high bias<sub> $\beta$ </sub> and, on this way, this would affect  $f_{\beta}(RH)$  calculation. Additionally, the water vapor correction tends to decrease  $\gamma$  ( $\beta_{wv}^{att}$  is lower than  $\beta$ ). Therefore, cases with lower bias<sub> $\beta$ </sub> and  $\Delta_q$  (case 1 and case 2), exhibits lower bias<sub> $\gamma$ </sub> (0.02 and 0.05, respectively), meanwhile on case 3 and case 4 instead bias<sub> $\beta$ </sub> increase, the bias<sub> $\gamma$ </sub> becomes higher (0.11 and 0.09, respectively).

Once we applied the phase 2 of the methodology (Sec. 5), we obtain 94 hygroscopic potential cases for 3h timewindow (Fig. S3a), 9 cases for 4h time-window (Fig. S3a) and 4 cases for 5h time-window (Fig. S3a), resulting in a total of 107 cases. To establish a bias error for this hygroscopic study, we have calculated the median of the bias<sub>β</sub> and  $\Delta_q$ , highlighting two main aspects: (i) The median bias<sub>β</sub> follows the median  $\Delta_q$  variability, remarking their dependency. This fact is well seen from the scatter plot (Fig. 3Sb), where these variables show a positive correlation, however the correlation coefficient is not too high (R<sup>2</sup>= 0.61), mainly because the data dispersion increases for bias<sub>β</sub> > 1.5 10<sup>-7</sup> m<sup>-1</sup>·sr<sup>-1</sup> and  $\Delta_q$  > 3.0 g/m<sup>3</sup>; (ii) The mean bias error calculated for  $\beta_{wv}^{att}$ over the 107 potential cases evaluated is lower than 2.5 · 10<sup>-7</sup> m<sup>-1</sup>·sr<sup>-1</sup> and the mean  $\Delta_q$  is lower than 5.5 g/m<sup>3</sup>.

Fig. S4 quantifies the effect of the  $\beta_{wv}^{att}$  over  $f_{\beta}(RH)$  and  $\gamma$  hygroscopic properties, by means of the median  $\operatorname{bias}_{f_{\beta}(RH)}(f_{\beta}(RH) - f_{\beta_{wv}^{att}}(RH))$  and the bias  $\gamma(\gamma_{\beta}att - \gamma_{\beta_{wv}^{att}})$ . Figure S4 reveals the no-correlation between  $\operatorname{bias}_{\beta}$  and  $\operatorname{bias}_{f_{\beta}(RH)}$ . However, combining the results from Fig. S3 and Fig. S4, it is possible to establish that  $\operatorname{bias}_{\beta} > 1 \text{ m}^{-1} \cdot \operatorname{sr}^{-1}$  would cause an increment of  $\operatorname{bias}_{f_{\beta}(RH)}$  above 0.2, increasing the error on  $f_{\beta}(RH)$ . Finally, it was obtained that  $\operatorname{bias}_{f_{\beta}(RH)}$  and  $\operatorname{bias}_{\gamma}$  for the whole study were lower than 0.3.

## 2 Results

### 2.1 Methodology applied to eight aerosol hygroscopic growth cases

The 8 hygroscopic growth cases reported in this study (Table 1, Sec. 5.2) were found at daytime in the early morning, with  $RH_{ref}$  around 50 % and the maximum RH reached was up to 90% over 3h time-window. For cases 1, 2, 4, 5, 6 and 7 the perceptual composition was dominated by OA with values up to 50 %, except on case 7 where OA decreased up to 38%. The BC concentration was relatively low almost for all cases found (close to 6 %). The concentration of inorganic compounds were dominated by  $SO_4^{2-}$  (lower than 21 %) and NH<sub>4</sub><sup>4</sup>(lower than 19 %), however NO<sub>3</sub><sup>-</sup> reached values of 21% on case 6. The air masses that come mainly from W-NW direction are related to case 1 ( $\Delta u$ =3.6 m/s W), case 2 ( $\Delta u$ =23.0 m/s NW) and case 6 ( $\Delta u$ =4.4 m/s W), with speed variability up to 14.2 %, 20.7 % and 18.5 %, respectively; and the air masses that coming from W-SW direction at low wind direction variability are associated with case 4 ( $\Delta u$ =2.7 m/s W), case 5 ( $\Delta u$ =2.4 m/s SW), and case 7 ( $\Delta u$ =1.7 m/s SW), and wind speed variability about 15.4 m/s, 20.4 m/s and 10.9 m/s, respectively. All these cases fulfilled the threshold established for  $\frac{\Delta f_{PM1}(RH)}{\Delta f_{\beta}(RH)} < 0.5$  indicating that increases/decreases in  $f_{\beta}(RH)$  are not related with advected aerosol into the atmospheric volume studied. The Hänel parameterization is calculated for both  $f_{\beta}(RH)$  and  $f_{PM1}(RH)$  (see panel Fig. S4 to S9 b.). The hygroscopicity properties of the 6 cases presented here were evaluated and compared against literature in the Sec. 5.2 of the article.

# Figures



**Figure S1.** Time evolution of  $\beta^{att}$  (blue line),  $\beta_c^{att}$  (red line) and q (orange line)[left panels (a,c,e,g *bias*<sub> $\beta$ </sub> (black line) and  $\Delta_q$  (green line) [right panel (b,d,f,h)].



**Figure S2.** Experimental data points (blue/red dots) and Hänel parameterization (blue/red lines). Case 1 and Case 2 show the effect of the lower bias<sub> $\beta$ </sub> and  $\Delta_q$  differences over  $f_{\beta_{wv}^{att}}(RH)$  and  $f_{\beta}(RH)$ . Case 3 and Case 4 present the effect of the higher bias<sub> $\beta$ </sub> and  $\Delta_q$  differences over  $f_{\beta_{wv}^{att}}(RH)$  and  $f_{\beta}(RH)$ .



**Figure S3.** Median of  $\text{bias}_{\beta}$  and  $\Delta_q$  for all potential cases of hygroscopic growth found from 2012 to 2016 at the ACTRS SIRTA observatory: (a median of  $\text{bias}_{\beta}$  and  $\Delta_q$  to 3h time-window analysis (green bars), 4h time-window analysis (orange bars) and 5h time-window analysis (blue bars); (b) scatter plot correlating median of  $\text{bias}_{\beta}$  and  $\Delta_q$  for whole time-windows.



**Figure S4.** bias<sub>*f*<sub>β</sub>(*RH*)</sub> and bias  $_{\gamma}$  for all potential cases of hygroscopic growth found from 2012 to 2016 at the ACTRIS SIRTA observatory. bias<sub>*f*<sub>β</sub>(*RH*)</sub> and bias  $_{\gamma}$  for 3h time-window analysis (green bars), 4h time-window analysis (orange bars) and 5h time-window analysis (blue bars).





Figure S5. Criterion for data selection: case 1 on 29 July 2012 from 06:30 to 09:30 UTC.



Figure S6. Criterion for data selection: case 2 on 02 September 2012 from 10:30 to 13:30 UTC.



Figure S7. Criterion for data selection: case 4 on 28 July 2014 from 09:10 to 12:10 UTC.



Figure S8. Criterion for data selection: case 5 on 17 August 2014 from 06:40 to 09:40 UTC.

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Figure S9. Criterion for data selection: case 6 on 21 May 2015 from 06:15 to 09:15 UTC.



Figure S10. Criterion for data selection: case 7 on 15 April 2016 from 07:05 to 10:05 UTC.