

Response to the comments of Reviewer #1- Manuscript: *Seasonal stratospheric ozone trends over 2000-2018 derived from several merged data sets* by Szelag et. al.

We would like to thank Reviewer #1 for the comments and the effort that the reviewer has put into this paper. Please see our response below.

MAJOR COMMENTS

Reviewer: Pg. 04, Ln. 10: *“The two-step approach is equivalent to the one-step regression . . .” The two approaches are not equivalent. While the results of the fitting coefficients tend to be very close to the same (it depends on the nature of the data and the regression model), the truth is that the uncertainty analysis becomes less robust, because covariance information between dependence on the different proxies is lost with each step. The multistep procedure tends to result in smaller uncertainties than a single step procedure, but this is because those uncertainties are artificially biased low from the lost information, not because it is necessarily a better process. On that note, technically this is a three-step regression since it is performed on deseasonalized anomalies (i.e., the first step). In lieu of doing things different, I would merely state that the multi-step approach is meant to avoid the use of more complicated regression models applied to a reduced amount of data at the expense of potentially biasing the uncertainties low.*

Authors: In the revised version, we will write “nearly equivalent”. We would like to note that the autocorrelations of residuals are taken into account in our regression method also in the second step, therefore the method should not bias the uncertainties low.

In the revised version, we will add that autocorrelations are taken into account in both steps.

Reviewer: *The authors describe their “two-step approach” on page 4 (Lns. 15–33) where the data is fit with Eqn. 1 with more data to derive the natural cycles before removing them and then fitting two different trends over two different time periods. This is intended to offer sufficient data to properly fit the natural variability. Generally, the wording here is fine (e.g., “thus providing more accurate fitting of these proxies” and “thus providing smaller sensitivity to defining turnaround point”) but I would add a caveat. Namely, Eqn 1 allows for the use of the PWLT term to potentially alias into some of this natural variability such that its removal for step 2 will affect the residuals when the trends are fit.*

Authors: Of course, any method for detecting cycles includes assumptions about the trend function (linear, piecewise, smooth), which might affect the estimated amplitude of the cycles. This is not a caveat for our method, but for multiple linear regression in general.

Reviewer: Pg. 05, Ln. 15: *The paragraph that starts here (and goes onto the next page) discusses potential reasons why Method 1 would result in smaller uncertainties than Method 2. The seasonal dependence of the QBO, for example, is a likely culprit but multiple potential reasons exist. One is simply that using data from all seasons would require a more complicated QBO model in the regression and so the residuals are greater and potentially seasonally dependent. This is easily analyzed by looking at the amplitude of the QBO in the regression from the coefficients between the different Methods.*

Authors: Indeed, the QBO amplitude is different in different seasons.

Reviewer: *However, another is that the uncertainties in the trends computed in step 2 are done without any consideration of the uncertainties in the residuals from which they were fit.*

Authors: No, uncertainties at the 2nd step are estimated from the fit residuals. We will state this in the revised version (PG. 05, Ln. 35).

The uncertainties are smaller in method #1 because of better fitting the cycles (smaller residuals).

Reviewer: *Pg. 10, Lns. 35–41 discuss the hemispheric asymmetry of the summer trends in the middle stratosphere at mid-latitudes (25-35km). This made me think to compare this work with what was shown in the last ozone assessment, as both use the same data sets. While WMO (2018) does not attempt to look at the seasonal trends, it does look at the overall trends. As such, I would expect the black lines in Figs. 5 and S3 to agree with what is shown in Fig. 3-19 of the Assessment. While much of the results are in agreement, I see some clear offsets and discrepancies, particularly between CCI in the Southern/Northern Hemisphere and SOO in the Northern Hemisphere. This makes me wonder if the term “All” in the figures in this paper are using all of the data at once in the analysis or if they are instead some sort of average of the results from different seasons. If the former, I would expect better agreement between this work and the Assessment. If the latter, then each value should be weighted by its seasonal mean to better represent the former, otherwise it does not really carry any meaning. This does not necessarily mean that the results from the seasonal regressions are incorrect, but it does make me wonder if double checking that all of the results shown here are accurate is needed. If the data presented are accurate, then I would want to know why they disagree with those results from the Assessment.*

Authors:

In our analysis, yearly trend estimates (“all”) are obtained using all the data at once.

We would like to note that:

- (1) slightly different latitude zones are used in the WMO (2018) report;
- (2) different time periods are used in WMO assessment (2000-2016) and in our paper (2000-2018);
- (3) different regression methods are used in WMO-2018 report (independent linear trend regression) and in our work (two-step regression).

And, despite these differences, the trends in WMO-2018 report and yearly trends in our paper are close to each other. We compared directly the trends presented in Figure 3-19 of WMO-2018 report (V. Sofieva produced this figure) with the yearly trends from our analyses, they are shown in Figures 1 and 2 below.

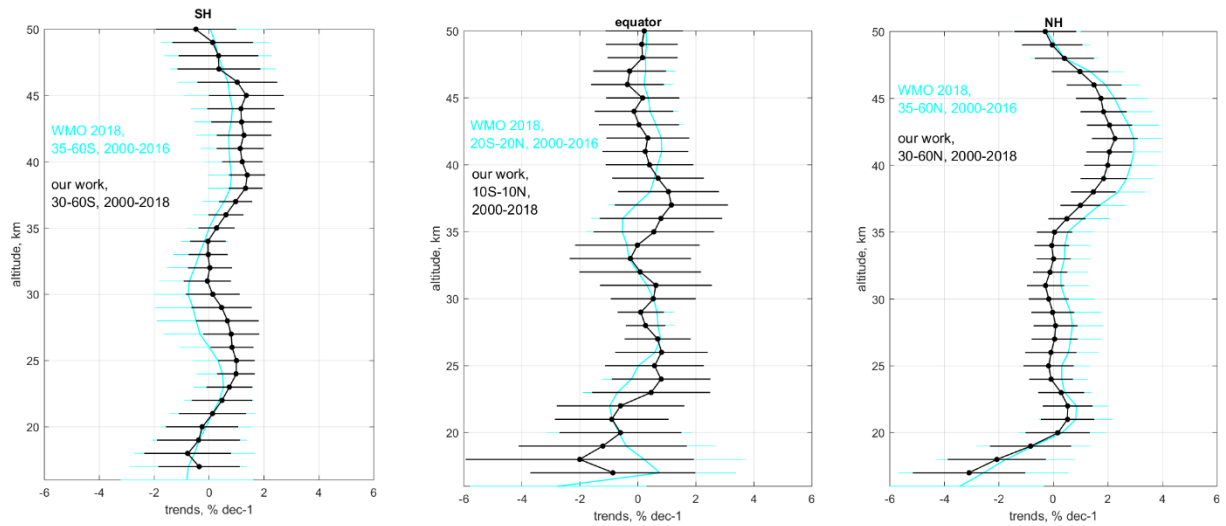


Figure 1 Comparison of post-2000 ozone trends for SAGE-CCI-OMPS datasets. Cyan lines: WMO 2018 assessment (2000-2016), black lines: yearly trends in our work (years 2000-2018).

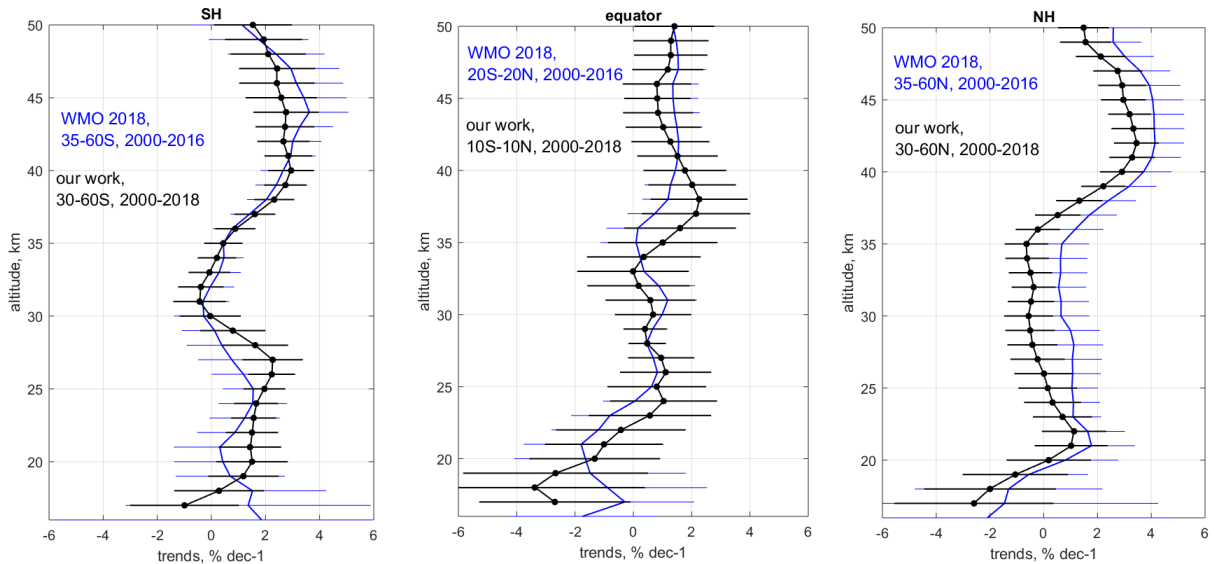


Figure 2. As Figure 1, but for SAGE-OSIRIS-OMPS dataset (SOO). Blue lines: trends from WMO-2018 assessment, black line: our analysis.

The most pronounced difference is the reduced trends in Northern Hemisphere. Most probably, the main impact is due to two-year extension. This reduction of trends due to 2 years extension is also observed using traditional trend analysis.

MINOR COMMENTS

Reviewer: Pg. 01, Ln. 35: “The Antarctic ozone hole is showing some signs of recovery, and the first signatures of global recovery . . .” I would recommend adding a reference to the first part of this sentence such as Solomon et al. (2016) (DOI: 10.1126/science.aae0061)

Authors: We now add the reference as suggested by the Reviewer.

Reviewer: Pg. 04, Ln. 01: *“The CCI and SOO datasets provide deseasonalized ozone anomalies. For GOZCARDS and SWOOSH, the deseasonalized anomalies were computed in the same way. The seasonal cycle was evaluated using the data from 2005-2011.” Was the seasonal cycle computed as averages of data in each month or using some sinusoidal fit? I am assuming the former but please specify.*

Authors: Yes, the seasonal cycle is computed via averaging the data in each month. We will clarify this in the revised version.

Reviewer: Pg. 04, Ln. 25: *“... 2-month lagged ENSO proxy ...” Why two months?*

Authors: The MEI ENSO index is defined as a 2-month average. It was used with 2-month lag in several trend analyses, for example, (Bourassa et al., 2014; Randel and Thompson, 2011; Sofieva et al., 2017). The reason for the lagging is somewhat smaller residuals in the stratosphere. However, the sensitivity of the fit to lagging by 1-2 months is small (Petropavlovskikh et al., 2019).

In the revised version, we clarify this.

Reviewer: *Figs. 3 and 6 are very busy (and large) figures and have the same information content. Wouldn't it make more sense to only include one of them in the paper for discussion? The ratio of figure space to text space is quite large.*

Authors: We think that Figure 6 is useful for discussion, and we would prefer to keep it in the manuscript.

References

Bourassa, A. E., Degenstein, D. A., Randel, W. J., Zawodny, J. M., Kyrölä, E., McLinden, C. A., Sioris, C. E. and Roth, C. Z.: Trends in stratospheric ozone derived from merged SAGE II and Odin-OSIRIS satellite observations, *Atmos. Chem. Phys.*, 14(13), 6983–6994, doi:10.5194/acp-14-6983-2014, 2014.

Randel, W. J. and Thompson, A. M.: Interannual variability and trends in tropical ozone derived from SAGE II satellite data and SHADOZ ozonesondes, *J. Geophys. Res. Atmos.*, 116(D07303), doi:10.1029/2010JD015195, 2011.

Sofieva, V. F., Kyrölä, E., Laine, M., Tamminen, J., Degenstein, D., Bourassa, A., Roth, C., Zawada, D., Weber, M., Rozanov, A., Rahpoe, N., Stiller, G., Laeng, A., von Clarmann, T., Walker, K. A., Sheese, P., Hubert, D., van Roozendaal, M., Zehner, C., Damadeo, R., Zawodny, J., Kramarova, N. and Bhartia, P. K.: Merged SAGE II, Ozone_cci and OMPS ozone profile dataset and evaluation of ozone trends in the stratosphere, *Atmos. Chem. Phys.*, 17(20), 12533–12552, doi:10.5194/acp-17-12533-2017, 2017.

Petropavlovskikh, I., Godin-Beekmann, S., Hubert, D., Damadeo, R., Hassler, B. and Sofieva, V.: SPARC/IO3C/GAW Report on Long-term Ozone Trends and Uncertainties in the Stratosphere, edited by M. Kennnertner and B. Ziegele, [online] Available from: <https://elib.dlr.de/126666/>, 2019.

Response to the comments of Reviewer #2- Manuscript: *Seasonal stratospheric ozone trends over 2000-2018 derived from several merged data sets by Szelag et. al.*

We would like to thank Reviewer #2 for the comments and the effort that the reviewer has put into this paper. Please see our response below.

MAJOR COMMENTS

Reviewer: *Of the four data sets, two have native units of volume mixing ratio on pressure, and two have native units of number density on altitude. If I understand correctly, each data set is analyzed and plotted in its native units. Could the authors say more about whether these should be directly comparable? They would not be if the pressure surfaces are changing in time relative to the altitude surfaces, that is, in the presence of a temperature trend. (see McLinden and Fioletov, doi: 10.1029/2010GL046012). I believe the answer is the temperature trends over the period are small, but this should be stated. Also, the seasonally-varying temperature trends mean this effect will differ by season. Do the authors see differences in the number density vs VMR trends that are consistent with seasonal temperature trends?*

Authors: We agree, the ozone trends in different representations and vertical coordinate can be different due to temperature trends. However, since the temperature trends after 2000 are small, we expect also a minor difference in ozone trends in different ozone representations. In the revised version, we note this and add the reference to the paper by McLinden and Fioletov, 2011.

We have plot the difference of seasonal trends for mean(CCI, SOO) minus mean(GOZCARDS, SWOOSH), similar way to Figure 4 in the manuscript. The difference in the upper stratosphere is consistent with the pattern on temperature trends and predictions by McLinden and Fioletov (see Figure 1 below). In the revised version, we include this figure into the Supplement.

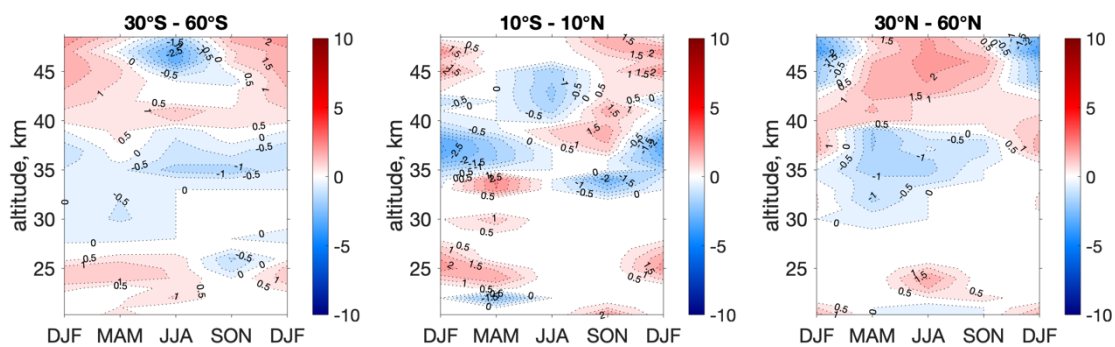


Figure 1. Altitude-seasons variation of the difference of the seasonal trends (mean[CCI, SOO] minus mean[GOZCARDS, SWOOSH]), calculated over 2000-2018 for three selected latitudinal bands.

Reviewer: *In the tropics, the authors show that both the trend and uncertainty can vary substantially between their method 1 (use seasonal time series only to fit and remove natural variability) and their method 2 (use full time series to remove variability). The authors state that fitting the natural variability proxies with insufficient number of points will cause larger uncertainties (P4 L17), and in P4 L30 that the two-step approach provides for a sufficient number of points. However, is this true for Method 1? In Method 1, only seasonal points are fit, and although three months of each year are fit, the proxies are all*

slowly varying such that they are highly correlated over the three-month segments of each year, so I question the assumption that there is a sufficient number of points to accurately fit to the proxies. The authors find that the uncertainty is generally less using method 1, but I'm not sure this is a true reduction in uncertainty. With the smaller number of points being fit, there may be higher correlation between proxies that is arbitrary (i.e. not physical). This would allow more cross-talk between the proxies in the regression, and thus a better fit, but not a physically meaningful fit. For example if there was correlation between the seasonal QBO and PWLT, the best fit in the regression might be a high coefficient on the QBO term, and an equally high but opposite sign coefficient on the PWLT, which when added allow the regression to better fit some short-term variability while the large-scale changes cancel. This would lead to smaller short-term variability in the residual (after QBO fit removed), but a larger PWLT term left behind, which then affects the trend segment fits in step 2. This is not necessarily the case, and the correspondence with the temperature trends lends support to the results being physical, but these caveats should be discussed. It would be ok to refer this to future work, but to investigate these possibilities, the authors might compare the derived ozone variability based on the full QBO proxy to that from the seasonal QBO proxies. Are the seasonal variations consistent with how we expect the ozone QBO to behave seasonally, and are they consistent in latitude and altitude, or generally noisy?

Authors:

In the revised version, we add caveats that the correlation between proxies can be different for Method 1 and Method 2, thus the reduction of uncertainty can be not fully physical. The detailed analyses of proxy correlations can be subject of future studies.

MINOR COMMENTS

Reviewer: *P1 L36 Is there a specific reference for the Antarctic ozone hole recovery (more recent than 2015) in addition to WMO 2018?*

Authors: We added the reference (Solomon et al., 2016).

Reviewer: *P3 Table 1 Under Ozone Profile Representation, are the first two entries (deseason- alized anomalies) in units of number density (or concentration) or percent? I suggest adding the units to each column.*

Authors: We now add the units to each column in Table 1.

Reviewer: *P4 L2-3 Suggest "For GOZCARDS and SWOOSH, the deseasonalized anomalies were computed relative to their 2005-2011 mean seasonal cycle." [These were not necessarily computed the same way as those initially provided as anomalies. I suspect that for those, the seasonal cycle was subtracted from each instrument individually before merging, which, if there are seasonal biases in the individual records, is different than using the seasonal cycle of the final merged record.]*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L9-10 Suggest re-wording slightly, “The two-step approach allows us to avoid fitting over the period when the ozone trends transition from negative to positive and are not well-represented by a linear function.” I want to get across the idea that it is not only that we don’t know the exact turn-around time (and this varies with latitude and altitude) but that the ozone change is not linear over this period anyway, so there may not be a well-defined turn-around time.*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L12-14 Suggest dropping the last sentence, as using dynamical linear modelling is not otherwise discussed.*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L27 certain period -> certain season*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L31-33 Similar to above, suggest slight re-wording of the last sentence. The authors do fit near the “trend turnaround point in 1997” if the decline segment fit is 1984-1997. It is not so much sensitivity to the turnaround time as it is that the ozone change in time does not look like the hockey stick representation (it is a curve rather than two intersecting slopes). What about “In the second step, fitting is only done during periods when the ozone change is approximately linear, thus avoiding the problem of how to properly model the ozone change in the trend turnaround period (such as sensitivity to trend turnaround time when using a hockey-stick representation).”*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P5 L15 Related to above comments, with fewer points, the individual proxies may be more correlated, giving the regression more room to “play” (more degrees of freedom) and get a better fit, but that fit might not be physical.*

Authors: Please see our replies above, in the major comments.

Reviewer: *P8 L1 Just clarifying, the analysis is redone for ozone averaged in the broad latitude bands, as opposed to averaging the trends in the smaller bands. This is how I read the text, I just want to be sure.*

Authors: That is correct.

EDITORIAL COMMENTS/TYPOS

Reviewer: *P1 L13, add comma after SWOOSH*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P2 L11 -1 K dec-1 (add space)*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P3 Table 1 Suggest some re-wording on the Merging Method: Median value of deseasonalized anomalies Average value of deseasonalized anomalies referenced to SAGE II Average value of original values referenced to SAGE II Average value of original values referenced to SAGE II*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L4 Trend analyses are usually performed using a multiple linear regression. . .*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L8 . . . two steps, by detecting and removing natural cycles in the first step, and estimating bulk changes over specific periods in the second step . . .*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L30 In the two-step approach, *a* sufficient number . . .*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L 34 Figure 2 illustrates each step of our analysis . . .*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P4 L35 Seasonal data (3-months, method #1)*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P5 L1 estimated for the period 2000-2008*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P5 L9 like in -> as in*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P6 Figure 2 Caption Last sentence, Right panel shows. . .*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P6 L16 Figure 2 -> Figure 3*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P7 Figure 3 GOZCARDS and SWOOSH titles give altitude range, but shouldn't these give pressure ranges. The caption also says pressure bands. I would suggest changing the word bands to ranges in the caption, but also include the pressure range in the plot titles rather than the altitude range.*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P9 L8 are dominating over -> dominate in*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P10 L2 Figure S4 does not match Figure 6.*

Authors: The part with Supplementary Figure S4 is now moved to correct place in the manuscript (Description of the Figure 4). Note, that the order of Supplementary Figures has changed.

Reviewer: *P10 L10 larger positive ozone trends*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P10 L14 due to greenhouse gas (remove 'the')*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P10 L30 so do -> as are*

Authors: Corrected as suggested by the Reviewer.

Reviewer: *P10 L31 additional confirmation of our hypothesis*

Authors: Corrected as suggested by the Reviewer.

Seasonal stratospheric ozone trends over 2000-2018 derived from several merged data sets.

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Abstract

In this work, we analyse the seasonal dependence of ozone trends in the stratosphere using four long-term merged datasets: SAGE-CCI-OMPS, SAGE-OSIRIS-OMPS, GOZCARDS and SWOOSH, which provide more than 30 years of monthly zonal mean ozone profiles in the stratosphere. We focus here on trends between 2000 and 2018. All datasets show similar results, although some discrepancies are observed. In the upper stratosphere, the trends are positive throughout all seasons and the majority of latitudes. The largest upper stratospheric ozone trends are observed during local winter (up to 6% dec⁻¹) and equinox (up to 3% dec⁻¹) at mid-latitudes. In the equatorial region, we find a very strong seasonal dependence of ozone trends at all altitudes: the trends vary from positive to negative, with the sign of transition depending on altitude and season. The trends are negative in the upper stratospheric winter (-1 to -2% dec⁻¹) and in the lower stratospheric spring (-2 to -4% dec⁻¹), but positive (2-3% dec⁻¹) at 30-35 km in spring, while the opposite pattern is observed in summer. The tropical trends below 25 km are negative and maximize during summer (up to -2 % dec⁻¹) and spring (up to -3% dec⁻¹). In the lower mid-latitude stratosphere, our analysis points to a hemispheric asymmetry hemispheric asymmetry: during local summers and equinoxes, positive trends are observed in the South (+1 to +2% dec⁻¹) while negative trends are observed in the North (-1 to -2% dec⁻¹).

We compare the seasonal dependence of ozone trends with available analyses of the seasonal dependence of stratospheric temperature trends. We find that ozone and temperature trends show positive correlation in the dynamically controlled lower stratosphere, and negative correlation above 30 km, where photochemistry dominates.

Seasonal trend analysis gives information beyond that contained in annual mean trends, which can be helpful in order to better understand the role of dynamical variability in short-term trends and future ozone recovery predictions.

1 Introduction

The stratospheric ozone layer plays an important role in the energy budget and dynamics of the middle atmosphere by absorbing a great part of harmful ultraviolet solar radiation. Its changes in the stratosphere contribute to ground-level climate variability. International efforts by scientists and politicians resulted in the Montreal Protocol agreement and its amendments, which regulate ozone-depleting substances (ODSs). Upper stratospheric ozone declined by about 4-8 % dec⁻¹ in the upper stratosphere from 1980 to the late 1990s (Steinbrecht et al., 2017a; WMO, 2018). The Antarctic ozone hole is showing some signs of recovery (Solomon et al., 2016), and the first signatures of global recovery have now been observed in the upper

stratosphere (Bourassa et al., 2014; Kyrölä et al., 2013; Newchurch et al., 2003; Tummon et al., 2015; WMO, 2018). However, no significant trend has been detected in global total column ozone and even though upper stratospheric ozone is recovering, negative trends have been reported in the lower stratosphere (Ball et al., 2018; WMO, 2018). Because of the strong linkage and feedbacks between ozone depletion and climate change, continuous monitoring of ozone is very important for understanding its variability and its role in climate change. Furthermore, stratospheric ozone depletion impacts stratospheric temperature, which is also a key indicator of global climate variability. The majority of ozone trend studies have assumed that the trends can be described as a simple piecewise-linear function (Bourassa et al., 2014; Kyrölä et al., 2013; Nair et al., 2015; Neu et al., 2014; Sofieva et al., 2017; Steinbrecht et al., 2017). On the other hand, more comprehensive analyses of stratospheric temperature trends have revealed large seasonal variability (Funatsu et al., 2016; Khaykin et al., 2017; Randel et al., 2016). Funatsu et al. (2016) reported cooling in the stratosphere over the period 2002–2014, at a rate of about 0.5 K dec⁻¹ above 25 km. They have also shown large seasonal variability at midlatitudes, with significant negative trends of -0.6 to -1 K dec⁻¹ during summer and autumn. Khaykin et al. (2017) revealed a middle stratospheric cooling over 2002–2016 at an average rate of -0.1 to -0.3 K dec⁻¹, with seasonal trend patterns indicating changes in the stratospheric circulation. Randel et al. (2016) have pointed to a stratospheric cooling varying from -0.1 to -0.6 K dec⁻¹ over 1979–2015, with larger cooling during 1979–1997 compared to 1998–2015, mirroring the differences in upper-stratospheric ozone trends.

This paper is dedicated to an investigation of the seasonal dependence of ozone trends in the stratosphere for the post-2000 time period. For our analysis, we have used several merged satellite datasets that cover more than three decades (from 1984 to 2018). Ozone trends are estimated based on two-step multiple linear regression. The paper is organized as follows. Section 2 describes the merged ozone datasets. Section 3 describes the methods used for seasonal trends estimation. Section 4 describes our results regarding the seasonal dependence of ozone trends in the stratosphere. Our conclusions are summarized in Section 5.

2 Data

In order to analyze seasonal stratospheric ozone trends, four long-term merged data sets of ozone profiles have been used.

These datasets are:

- SAGE-CCI-OMPS dataset developed in the framework of ESA Climate Change Initiative (Sofieva et al., 2017), referred to as CCI hereafter
- SAGE-OSIRIS-OMPS developed at the University of Saskatchewan (Bourassa et al., 2014), referred to as SOO hereafter
- The Global Ozone Chemistry And Related trace gas Data records for the Stratosphere (GOZCARDS) (Froidevaux et al., 2015, 2019)
- The Stratospheric Water and Ozone Satellite Homogenized (SWOOSH) created by NOAA/ESRL Davis et al. (2016).

General information about the merged datasets is collected in Table 1. The basic information about individual ozone profile datasets used in the merged datasets is collected in Table 2. The collection of merged datasets represents two main groups categorized according to ozone profile representation: number density on altitude grid (CCI and SOO) and volume mixing ratio (vmr) on pressure grid (GOZCARDS and SWOOSH). **Note, that the ozone trends in different representations and vertical coordinates can be different due to temperature trends. However, since the temperature trends after 2000 are small, we expect also a minor difference in ozone trends in different ozone representations (McLinden and Fioletov, 2011).** The vmr-based group uses MLS/Aura as the main dataset after August 2004, while the number-density group uses other satellite instruments, with retrievals on altitude grid, after 2005. The merged datasets use different collections of limb-profiling

instruments and use different merging methods (see Table 1 and references for more details). All merged datasets are constructed from datasets having a good vertical resolution of 2-4 km. More detailed information about the merged datasets can be found in the publications mentioned above and in Table 1. All the merged datasets have been extended until December 2018. These long-term monthly zonal mean datasets cover all seasons.

5 **Table 1. General information about the merged datasets**

Dataset	SAGE-CCI-OMPS	SAGE-OSIRIS-OMPS	GOZCARDS (v2.20)	SWOOSH (v2.6)
Time coverage	Oct 1984- Dec 2018	Oct 1984- Dec 2018	Jan 1979-Dec 2018	Oct 1984- Dec 2018
Latitude coverage and representation	90°S-90°N 10° zones	60°S-60°N 10° zones	90°S-90°N 10° zones	90°S-90°N 10° zones (also 5°, 2.5 °)
Vertical coverage and sampling	10-50 km, 1-km grid	10-50 km, 1-km grid	215-0.2 hPa 12 levels dec ⁻¹ (~ 3 km)	316- 1 hPa 12 levels dec ⁻¹ (~ 3 km)
Included instruments	SAGE II OSIRIS OMPS-LP GOMOS MIPAS SCIAMACHY ACE-FTS	SAGE II OSIRIS OMPS-LP	SAGE I SAGE II HALOE Aura MLS	SAGE II HALOE UARS MLS Aura MLS SAGE III
Merging method	median value of deseasonalized anomalies	average value of deseasonalized anomalies referenced to SAGE II	average value of original values referenced to SAGE II, bias corrections for non-SAGE II using monthly mean differences	average value of original values referenced to SAGE II, bias corrections for non-SAGE II instruments using collocated data
Ozone profile representation	deseasonalized anomalies (%) and ozone concentrations (mol m⁻³)	deseasonalized anomalies, relative	vmr	vmr
Reference	Sofieva et al. (2017)	Bourassa et al. (2014), Zawada et al. (2017)	Froidevaux et al., (2015, 2019)	Davis et al. (2016)

Table 2 Basic information about individual satellite datasets used in the merged datasets

Instrument/platform	Acronym explanation	Used period	Data processor
SAGE I/AEM-B	Stratospheric Aerosol and Gases Experiment I	1979-1981	V5.9 rev
SAGE II/ERBS	Stratospheric Aerosol and Gases Experiment II	1984-2005	V7
SAGE III/Meteor-3M	Stratospheric Aerosol and Gases Experiment III	2002-2005	V4
HALOE/UARS	HALogen Occultation Experiment	1991-2005	V19
MLS/UARS	Microwave Limb Sounder	1991-1997	V5
MLS/Aura	Microwave Limb Sounder	2004-2018	V4.2
OSIRIS/Odin	Optical Spectrograph and InfraRed Imaging System	2001-2018	v5.10
MIPAS/Envisat	Michelson Interferometer for Passive Atmospheric Sounding	2005-2012	IMK/IAA v7
SCIAMACHY/Envisat	SCanning Imaging Absorption spectroMeter for Atmospheric ChartographY	2003-2012	UB v3.5
GOMOS/Envisat	Global Ozone Monitoring by Occultation of Stars	2002-2011	ALGOM2s v1
ACE-FTS/SCISAT	Atmospheric Chemistry Experiment Fourier Transform Spectrometer	2004-2018	v3.5/3.6

OMPS-LP/Suomi-NPP	Ozone Monitor Profiling Suite-Limb Profiler	2012-2018	Usask2D v1.1.0
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3 Methods

The seasonal trend analysis was performed on monthly deseasonalized anomalies for all merged data sets. The CCI and SOO datasets provide deseasonalized ozone anomalies. **The seasonal cycle is computed by averaging the data in each month.**

5 **For GOZCARDS and SWOOSH, the deseasonalized anomalies were computed relative to their 2005-2011 mean seasonal cycle.**

The trend analyses are usually performed using a multiple linear regression technique, in order to separate the natural variability and long-term trends. The known main sources of ozone variability, which are characterized by corresponding proxies, are solar cycle, QBO and ENSO. In many studies, the regression is performed in one step by assuming an approximation of ozone trends by a piecewise linear function (e.g., Bourassa et al., 2014; Harris et al., 2015; Kyrölä et al., 2013; Sofieva et al., 2017).
10 Alternatively, the regression can be performed in two steps, by detecting and removing natural cycles in the first step, and estimating bulk changes over specific periods in the second step (e.g., Steinbrecht et al., 2017; WMO, 2018). **This two-step approach allows us to avoid fitting over the period when the ozone trends transition from negative to positive and are not well-represented by a piecewise linear function. The two-step approach is nearly equivalent to a one-step regression,**
15 **in which a piecewise linear function with three segments is fitted to the time series, as done in the SPARC LOTUS report (Petrovavlovskikh et al., 2019, the independent-linear-trend (ILT) method).**

Analyses of the seasonal dependence of trends are based on smaller number of data points, compared to the traditional trend analyses, in which data from all months are used. For example, if one analyzes trends in a given month, a 30-yr long dataset will have only 30 data points. The fitting of all proxies in the standard multiple linear regression with insufficient data points
20 will result in substantial uncertainty of estimated parameters. Therefore, we propose the following two-step multiple regression for seasonal trend analysis. In the first step, natural cycles are estimated from the data using the traditional multiple linear regression:

$$O_3(t) = PWLT(t, t_0) + q_1 QBO_{30}(t) + q_2 QBO_{50}(t) + sF_{10.7}(t) + dENSO(t), \quad (1)$$

where ozone trends are approximated with a piecewise linear function ($PWLT(t, t_0)$) with the turnaround point in 1997,
25 $QBO_{30}(t_s)$ and $QBO_{50}(t_s)$ are equatorial winds at 30 and 50 hPa, respectively (<http://www.cpc.ncep.noaa.gov/data/indices/>), $F_{10.7}(t_s)$ is the monthly average 10.7 cm solar radio flux (ftp://ftp.geolab.nrcan.gc.ca/data/solar_flux/monthly_averages/), and $ENSO(t_s)$ is the ENSO proxy (<http://www.esrl.noaa.gov/psd/enso/mei/table.html>). **The ENSO index is used with a 2 months lag, as done in several previous ozone trend analyses (e.g., Bourassa et al., 2014; Randel and Thompson, 2011; Sofieva et al., 2017).** Then the fitted natural cycles are removed from the data, so only smooth variations remain in the resulting
30 time series. The regression Eq. (1) can be performed using the data only from **a certain season**, usually 3 months (referred to as method #1 hereafter) or using data from all seasons (we will call this method #2). In the second step, the trends are estimated separately for the decline (1984-1997) and the recovery period (2000-2018) using a simple linear regression. In the two-step approach, a sufficient number of data points are available for detection of natural cycles - solar, QBO and ENSO, thus providing more accurate fitting of these proxies. **In the second step, fitting is only done during periods when the ozone change is**
35 **approximately linear, thus avoiding the problem of modelling the ozone change in the turnaround period (such as the**

sensitivity to trend turnaround time when using a hockey-stick representation). Uncertainties in the second step are estimated from the fit residuals. Autocorrelations are removed in both steps using the Cochrane–Orcutt transformation (Cochrane and Orcutt, 1949).

Figure 1 illustrates **each step** of our analysis for the CCI dataset and for two selected periods (MAM, SON) in the tropics (10°S-10°N) at altitudes between 30-35 km. **Seasonal data** (3-months, method #1) are used in the first step (Eq.(1)). The original ozone anomalies (top panel of Figure 1) show a variability of about $\pm 10\%$ during 1984-2018 for both seasons. The anomalies are reduced to about $\pm 6\%$ after the first step of the regression analysis where all proxies (QBO_{30} , QBO_{50} , $F_{10.7}$, $ENSO$) are subtracted from the original data (middle panel of Figure 1). Finally, the bottom panel of Figure 1 illustrates the second step of the regression analysis, i.e., linear trends are estimated for the period 2000-2018 in April (trend of $+2.3\% \text{ dec}^{-1}$) and October (trend of $-2.1\% \text{ dec}^{-1}$).

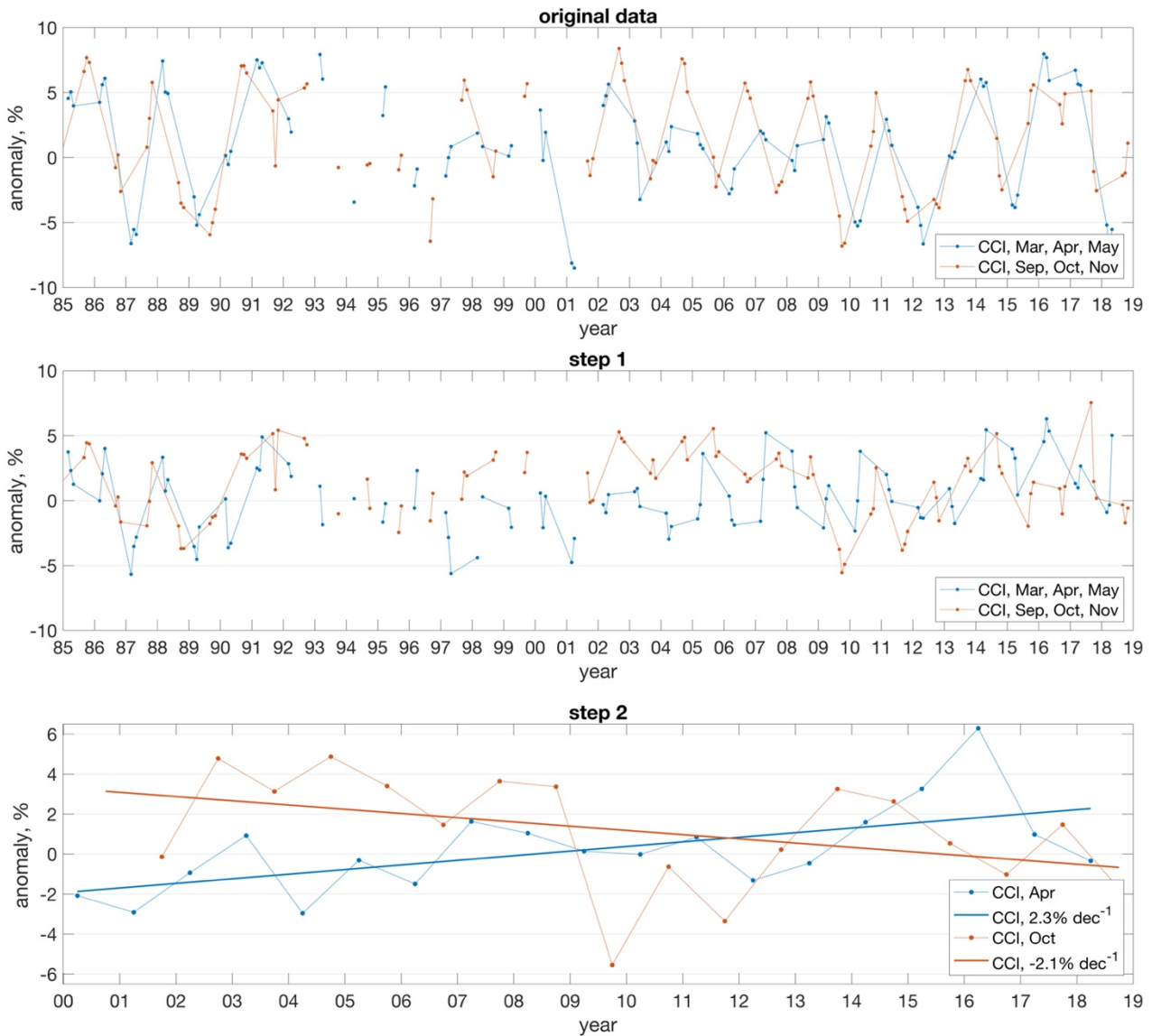


Figure 1. Two-step multiple linear regression example from CCI for 10°S-10°N and 30-35 km. Top panel: original ozone anomalies. Middle panel: anomalies with cycles removed. Bottom panel: linear trends estimated for two selected months.

As described above, the amplitude of the natural cycles (solar, QBO, ENSO), which will be removed from the time series, can be estimated in two ways in the first step of our regression method, i.e., for each season separately (method #1) or using the

data from all months, as in the traditional trend analysis (method #2). We have investigated both methods for estimating the natural cycles and we found that in general, uncertainties of method #1 are smaller. This is illustrated in Figure 2, which shows the ozone trends at latitudinal bands 10°S-10°N (top panels) and 30°N-60°N (bottom panels) in April for these two methods (left and middle panels). The right panel of Figure 2 shows the differences in trends and uncertainties between both methods.

At mid-latitudes, the difference in uncertainties is negligible (less than 0.3%), but method#1 provides some smaller trend uncertainties, especially around ~35 km.

Such an observation is counter-intuitive: one would expect better fitting using more data points. The reason for this might be the seasonal dependence of natural cycles, particularly the QBO (Gabis et al., 2018). Assuming this, one would expect larger differences between the two methods in the tropical region, where the QBO dominates, and this is observed in Figure 2.

Although our observations – smaller residuals when fitting natural cycles using the data from a 3-months season – seem to support the hypothesis of seasonal dependence of natural cycles, this discussion is beyond the scope of our paper. **From another point of view, the correlation between proxies can be different for methods #1 and #2, thus the uncertainty reduction may not be fully realistic. More detailed analyses of proxy correlations can be the subject of future studies.**

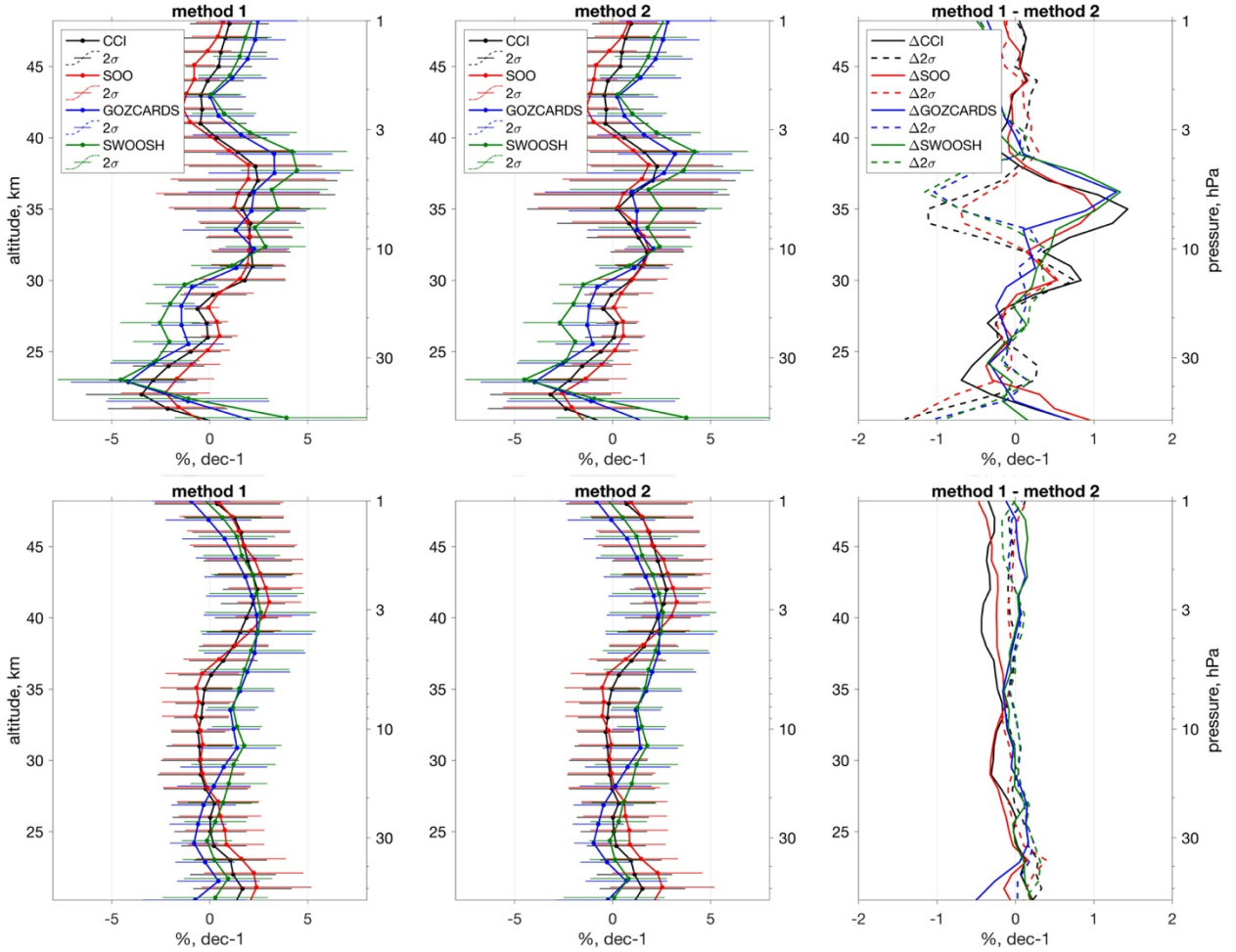


Figure 2. Vertical profiles of ozone trends in 2000-2018 from CCI (black), SOO (red), GOZCARDS (blue) and SWOOSH (green). The results are shown for April, latitudinal bands 10°S-10°N (top panels) and 30°N-60°N (bottom panels) for two separate methods (see description in the text). Error bars are 2-sigma uncertainties. Data are presented on their natural vertical coordinate: altitude grid (left axis) for CCI and SOO and pressure grid (right axis) for GOZCARDS and SWOOSH. Right panel shows the difference between methods for trends (solid lines) and uncertainties (dashed lines).

In the analyses shown below, we have used method #1 for the evaluation and removal of natural cycles. We would like to emphasize, that the results are similar if method #2 is used (see text below and Supplementary information for details). In this work, we will focus and discuss post-2000 trends only.

5 4 Results

The seasonal variation of ozone trends over the 2000–2018 period as a function of latitude for 5 selected altitude regions is shown in Figure 3 (colored contours). The red/blue shading (positive/negative trends) in Figure 3 denote trends that are statistically significant at the 95% level.

In the tropical region (20°S-20°N), a strong seasonal dependence of ozone trends is observed. In the lower stratosphere (19-23 km), the pattern of statistically significant negative trends of about -2 to -3% dec⁻¹ is present in all merged data sets during the spring and summer months (MAM and JJA), while they are less pronounced in other months. At altitudes 24-28 km, the trends change from negative (-1 to -2% dec⁻¹, statistically significant for all except SOO) in spring (MAM) to positive (+1 to +2% dec⁻¹, statistically significant for CCI and SOO) in autumn (SON). At 31-35 km in the tropics, the trends are opposite: positive trends in MAM (+2 to +3% dec⁻¹, statistically significant) and negative in SON (-1 to -2% dec⁻¹, not statistically significant).

Above 40 km, trends are in general positive throughout the latitudes and months, with the largest trends observed at mid-latitudes (30°N/S-60°N/S) during the local winters and spring/autumn seasons (2-4% dec⁻¹). Upper stratospheric recovery for all latitude bands has been obtained by others (e.g., Petropavlovskikh et al., 2019; Steinbrecht et al., 2017; WMO, 2018), nonetheless, for seasonal analysis, negative trends of -1 to -2% dec⁻¹ during the late winter (DJF) are present in the tropics, although with statistical significance only in the CCI datasets.

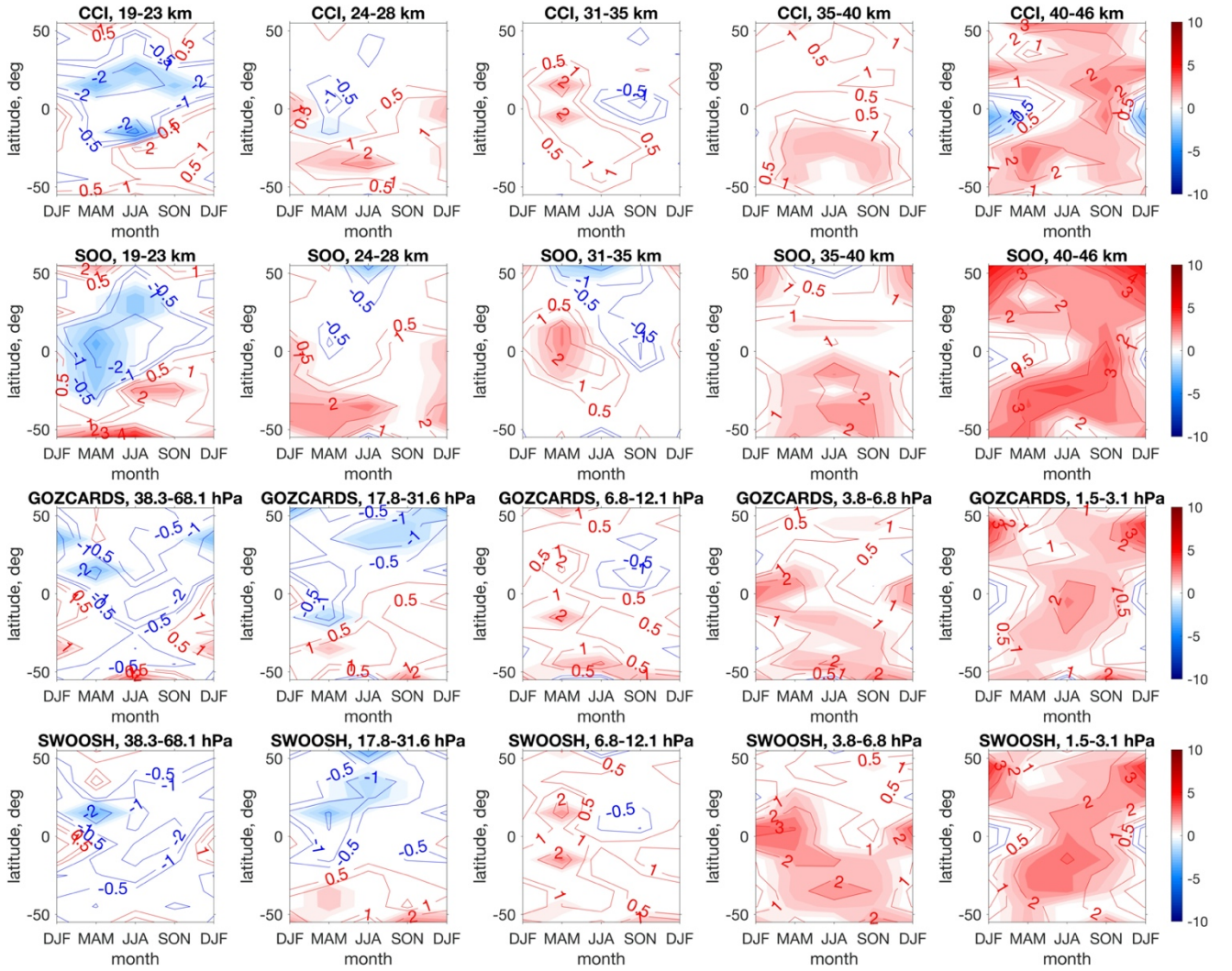


Figure 3. Latitude-seasons variation of linear trends in ozone for each of the merged data sets calculated over 2000-2018 for five selected altitude/pressure ranges. The shading denotes trends that are significant at the 95% level. Pressure ranges correspond approximately to altitude ranges.

5

Analysis results from broader latitude bands are presented in Figure 4 (method #1) and Supplementary Figure S1 (method #2). Both methods as well as all datasets generally show very similar results. The uncertainties of method #1 are usually smaller than for method #2, which makes the observed patterns more statistically significant. The basic structure and patterns of seasonal trends are apparent in the monthly data as well (Supplementary Figure S2 and Figure S3) but the magnitude of the calculated trends is higher for each single month.

Figure 4 reveal negative trends ranging from -1 to -2% dec^{-1} in the lower/middle stratosphere at 30°N-60°N during the summer (JJA). Negative trends are present in all datasets and are statistically significant for all except the CCI dataset. In contrast, in the lower/middle southern stratosphere (30°S-60°S), trends are positive (1-2% dec^{-1}) and statistically significant for all datasets.

15 In the equatorial region, all datasets show pronounced, statistically significant and very similar seasonal dependence of ozone trends. In the upper tropical stratosphere above 40 km, trends are negative in DJF (-1 to -3% dec^{-1} , significant for all except SOO), and positive in August-October (2-3 %, statistically significant for all datasets). At 30-35 km, trends are positive in MAM at altitudes 30-35 km (2 to 3 % dec^{-1} , significant for all) and negative in SON (~ -1 % dec^{-1} , statistically significant for

CCI and SOO). In the lower stratosphere, the negative trends are in MAM (-2 to -3% dec⁻¹, significant for all) and positive in DJF (1 to 2% dec⁻¹, significant for CCI and SOO).

Upper stratospheric trends at mid-latitudes are most pronounced during the local winters and equinoxes, varying from 3 to 4% dec⁻¹ in the North and from 2 to 3% dec⁻¹ in the South.

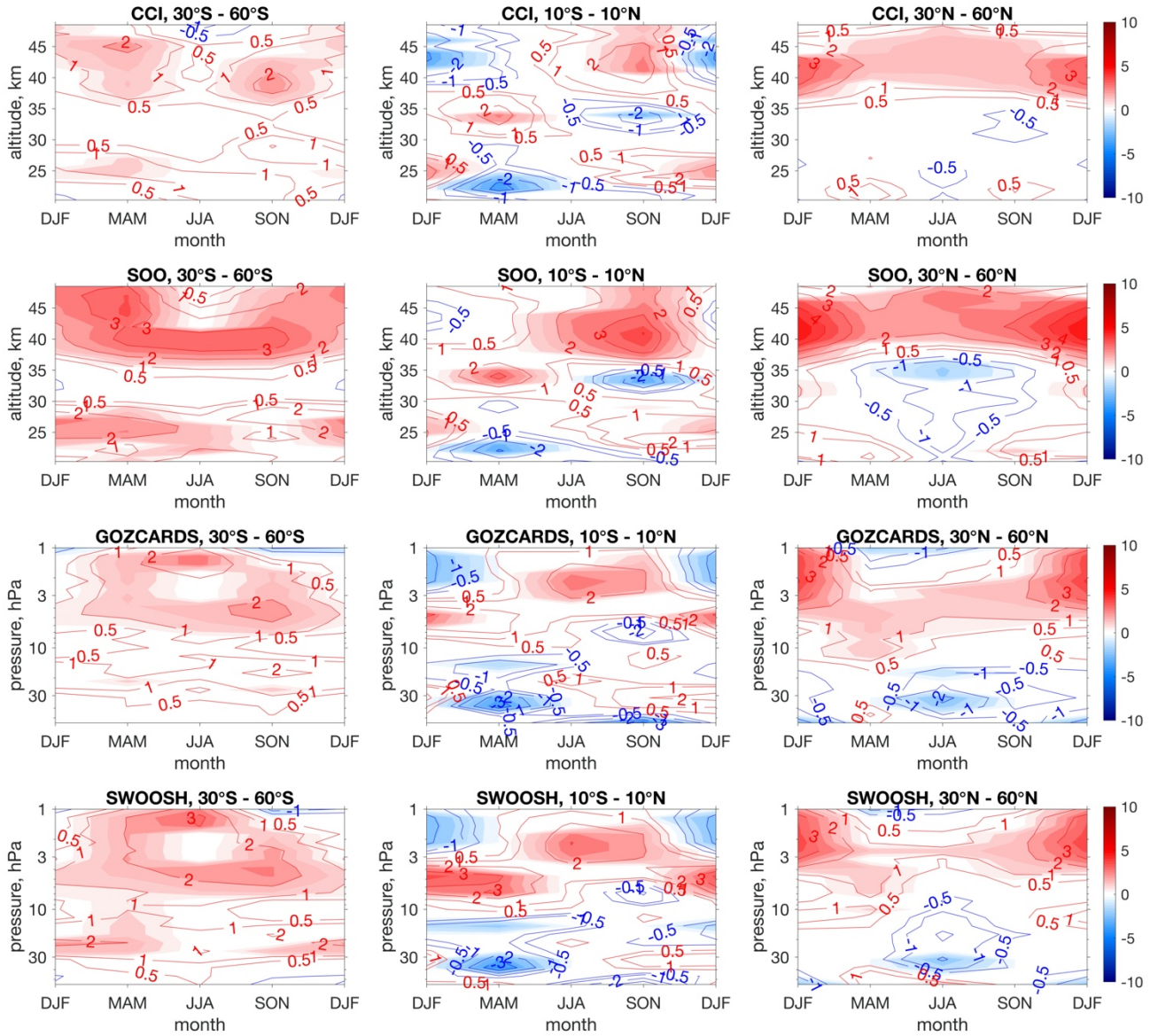


Figure 4. Altitude-seasons variation of linear trends in ozone for each of the merged data sets calculated over 2000-2018 for three selected latitudinal bands. Data are presented on their natural vertical coordinate: altitude grid for CCI and SOO and pressure grid for GOZCARDS and SWOOSH. The shading denotes trends that are significant at the 95% level.

All merged ozone datasets show a similar seasonal dependence of the ozone trends. Some difference in ozone trends can also be due to the differing vertical grids used in the various data sets. To estimate this, we have calculated the seasonal trends for “number density on altitude grid” (CCI, SOO) and also for “vmr on pressure grid” (GOZCARDS, SWOOSH), as shown in Supplementary Figure S4. The estimated difference is consistent with the pattern of temperature trends and predictions by McLinden and Fioletov (2011).

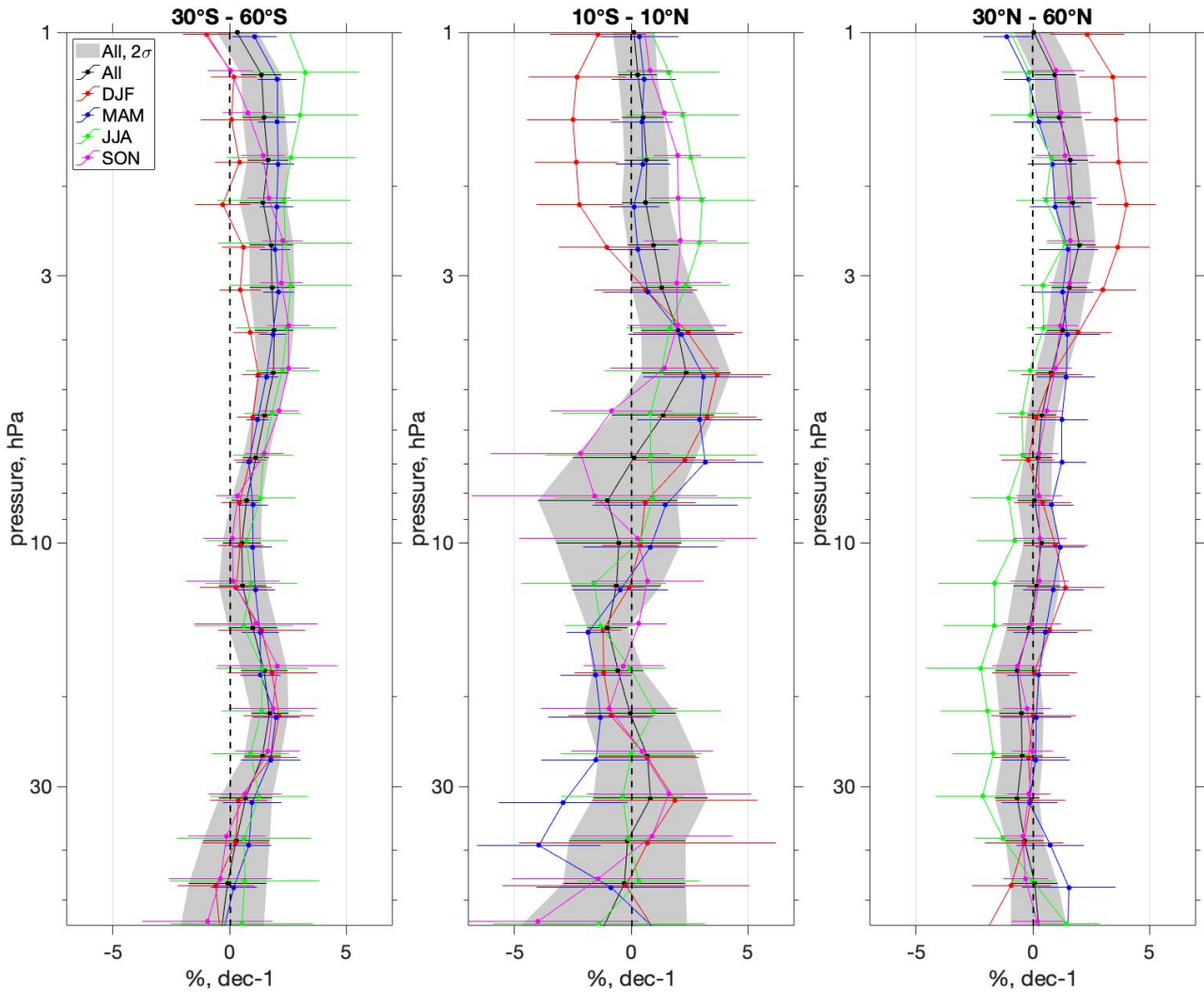


Figure 5. Vertical profiles of seasonal (red, blue, green, magenta) and yearly (black) ozone trends in 2000-2018 from SWOOSH. The results are shown for three selected latitude bands. Error bars and shaded area (gray) are 2-sigma uncertainties.

Figure 5 shows the seasonal ozone trends (color) together with yearly trend (black) plotted in the vertical distribution for three selected latitudinal bands. It is clear that at mid-latitudes, positive trends in the upper stratosphere are dominate in the local winters (up to 4% dec⁻¹) and are much higher than the yearly trend (up to 2% dec⁻¹). In the tropics, the main features observed during different seasons (i.e. negative-positive patterns observed throughout the seasons and vertical levels) mostly cancel out in the yearly trend. Negative winter trends (DJF) in the upper stratosphere (-3% dec⁻¹), negative spring trends (MAM) in the lower stratosphere (-3 to -4% dec⁻¹), and positive spring trends (MAM) in the middle stratosphere (2% dec⁻¹) are all counterbalanced by opposite or smaller trends during the remaining seasons. As a result, the yearly ozone trend in the tropics is much smaller than the seasonal trends. The other datasets show consistently similar features (Supplementary Figure S5). The main discrepancies are observed between number density- based group (SOO and CCI) and vmr-based group (GOZCARDS and SWOOSH) during summer (JJA) in the upper stratosphere in the South.

5 Discussion

To summarize our analysis, variations of ozone trends over the period 2000–2018 for each latitude and vertical level are plotted for each season separately in Figure 6.

In the upper stratosphere, trends are positive throughout all seasons and the majority of latitudes. One of the most pronounced features is that the mid-latitude upper stratospheric ozone trends are larger in local winters. As discussed in Sect 1., ozone and temperature trends are inter-related, as ozone and temperature are connected via photochemical reactions (effective above ~25–30 km, (Brasseur and Solomon, 2005). Randel et al. (2016) reported weaker upper-stratospheric cooling in local winter at mid- and high latitudes. This is fully consistent with our observations of larger positive ozone trends at these locations in winter. A hypothesized explanation of this feature might be the acceleration of the upper branch of the Brewer-Dobson circulation (or in general mean residual circulation), which controls the meridional transport of trace gases from the tropical region to the poles (Brewer, 1949; Dobson, 1956). The Brewer-Dobson circulation is most effective during winter season (Chipperfield, M. P., Jones, 1999). Several modelling studies have shown that due to greenhouse gas concentration (GHG) increases, the winter time BDC will strengthen and accelerate the expected ozone recovery (Butchart et al., 2006; Garcia and Randel, 2008; Gettelman et al., 2010; Li et al., 2008; Schnadt et al., 2002; Sigmond et al., 2004). Also, observational studies have shown an accelerated BDC over the tropical region (Thompson and Solomon, 2009) as well as at high latitudes (Hu and Fu, 2009). Increased speed of the BDC would have an effect on the transport of ozone and ODSs. While the main reason for positive ozone trends in the upper mid-latitude stratosphere is due to the decrease of ozone-depleting substances, the seasonal variations of the ozone trends can be due to dynamics. Seasonal dependence of both temperature and ozone trends support this hypothesis. However, the investigation of the mechanisms that controls the seasonality of ozone and temperature trends is beyond the scope of our paper; it can be the subject of future modelling and observational studies.

In the tropics, our analysis has shown a very strong seasonal dependence of ozone trends observed at all altitudes. The trends change from positive to negative, with the phase changing with altitude. In the tropical lower stratosphere below 25 km, strong negative trends are observed during boreal spring and summer, which are statistically significant for all datasets. Khaykin et al., (2017) also found an altitude-dependent pattern in temperature trends in the tropical region, with a strong seasonality. One can notice that the changes of phases in ozone and temperature trends are very similar (temperature trends are evaluated below 35 km in Khaykin et al., (2017)). In the lower tropical stratosphere, in the dynamically controlled region, ozone and temperature variations are positively correlated (Hauchecorne et al., 2010), as are the ozone and temperature trends (compare our Fig. 8 with Fig. 5 in Khaykin et al., 2017). This is rather expected and can serve as additional confirmation of our hypotheses explaining this seasonality. Khaykin et al. (2017) hypothesized that the observed trends structure might be related to seasonal variations in the Brewer–Dobson circulation.

The third interesting feature is hemispheric asymmetry of summer-time ozone trend patterns below 35 km: they are negative in the NH and positive in the SH. No similar analyses exist for temperature, and this can be the subject of future work. Hemispheric asymmetry was reported in recent studies (Froidevaux et al., 2019, Ball et al., 2019), even if not broken down by season. Froidevaux et al. (2019) showed that ozone trends derived from Aura Microwave Limb Sounder (Aura/MLS) data over a shorter period (2005–2018) have a tendency towards slightly positive values in the SH. Additionally, this asymmetry might be related to the hydrogen chloride (HCl) abundances and trends (Mahieu et al., 2014, Han et al., 2019). At the moment, we can only speculate that this might also be a contributing factor to the observed ozone negative trends in that region.

6 Summary

Using four long-term merged data sets of ozone profiles we have studied the seasonal dependence of ozone trends in the stratosphere. The results of our analysis, based on two-step multiple linear regression can be summarized as follows:

- The upper stratospheric ozone is recovering, and the recovery maximizes during local winters and equinoxes reaching up to 3-4% dec^{-1} , which is fully consistent with weaker upper-stratospheric cooling in local winter at mid- and high latitudes.
- In the tropics, there is very strong seasonal dependence of ozone trends at all altitudes. The trends are changing from positive to negative, with the sign of transition depending on altitude and season.
- Below 25 km in the tropical region, strong negative trends are observed during spring and summer, which are statistically significant for all datasets and consistent with the seasonal pattern of temperature trends in this region.
- In the lower and middle stratosphere, there is hemispheric asymmetry during the local summers and equinoxes at mid-latitudes with negative trend in the North and positive trend in the South.

Despite some discrepancies, the general coherence in trends derived from four different merged data sets gives us confidence in the validity and robustness of the results. We compared the seasonal dependence of ozone trends with available analyses of the seasonal dependence of stratospheric temperature trends and found a clear inter-relation of the trend patterns.

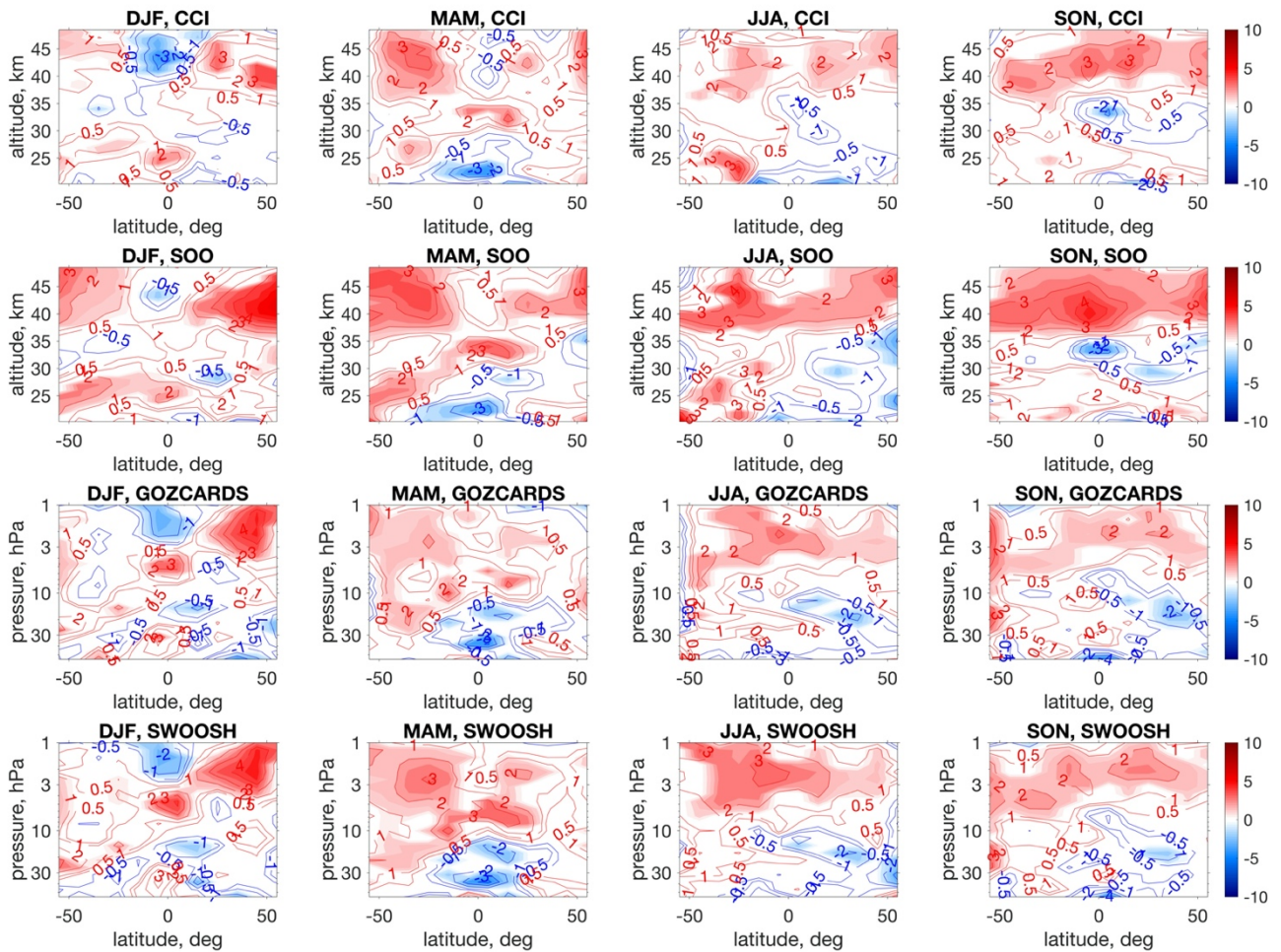


Figure 6. Altitude-latitude variation of linear trends in ozone calculated over 2000-2018 for each season. Data are presented on their natural vertical coordinate: altitude grid for CCI and SOO and pressure grid for GOZCARDS and SWOOSH. The shading denotes trends that are significant at the 95% level.

Data availability

The SAGE-CCI-OMPS dataset is available from CCI website (<http://www.esa-ozone-cci.org>). The SAGEII-OSIRIS-OMPS dataset is available from University of Saskatchewan ftp-site. GOZCARDS ozone data updates (version 2.20) are available by contacting Lucien Froidevaux; this data version will also be updated on the public Goddard Earth Sciences Data and Information Services Center (GES DISC) website (<https://disc.gsfc.nasa.gov>) in the near future.

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