Response to Reviewers and modifications to manuscript acp-2019-1001

Contents

Response to RC1

Review of "Analysis of 24 years of mesopause region OH rotational temperature observations at Davis, Antarctica. Part 1: Long---term trends" by French et al.

This manuscript presents the analysis of a very long dataset of OH temperatures over Antarctica. This is an extension of 8 years of the dataset presented by French and Klekociuk (2011). Indeed, results for the trends derived here coincide with their previous results and those for the solar response are only slightly different. In this work, the authors further identify and isolate a close---to---4---year period signal, which is to be studied
in the second part of this work. Even if the results in this paper may initially look as a mere update of previous results using an extended database, they are interesting and certainly worth publishing because they show the persistence of the trend and the consistency of the solar signal. Therefore, I suggest the publication of this paper in ACP, once the following concerns and comments are taken into account.

Line numbering corresponds to the latest version uploaded by the authors (acp---2019--1001---manuscript---version3.pdf).

Thank you for your considered and detailed review of acp-2019-1001. We address your comments and suggestions below.

Main comments

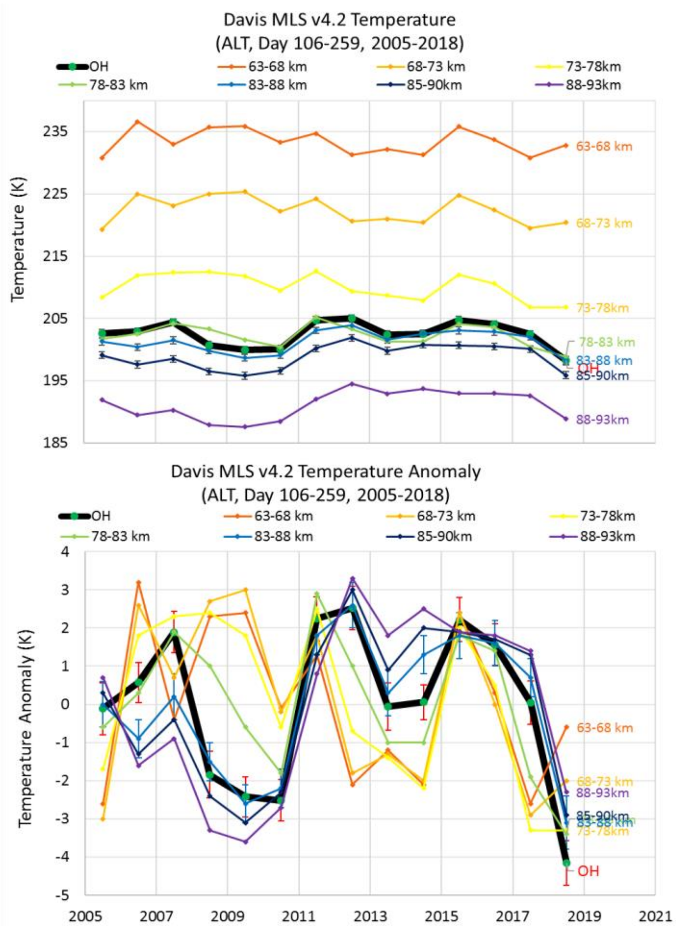
There is no discussion on the effect of MLS broad vertical resolution in the mesosphere and the potential impact on the comparisons shown in the manuscript. Indeed, it would be interesting to see comparisons with SABER, even with a smaller winter temporal--coverage. Additionally, SABER provides information on the altitude of the OH layer, potentially providing a more accurate approach. In the same context, the choice of a fixed pressure level in MLS data (seemingly done based just on a better agreement) is not very well justified, as the layer altitude varies.

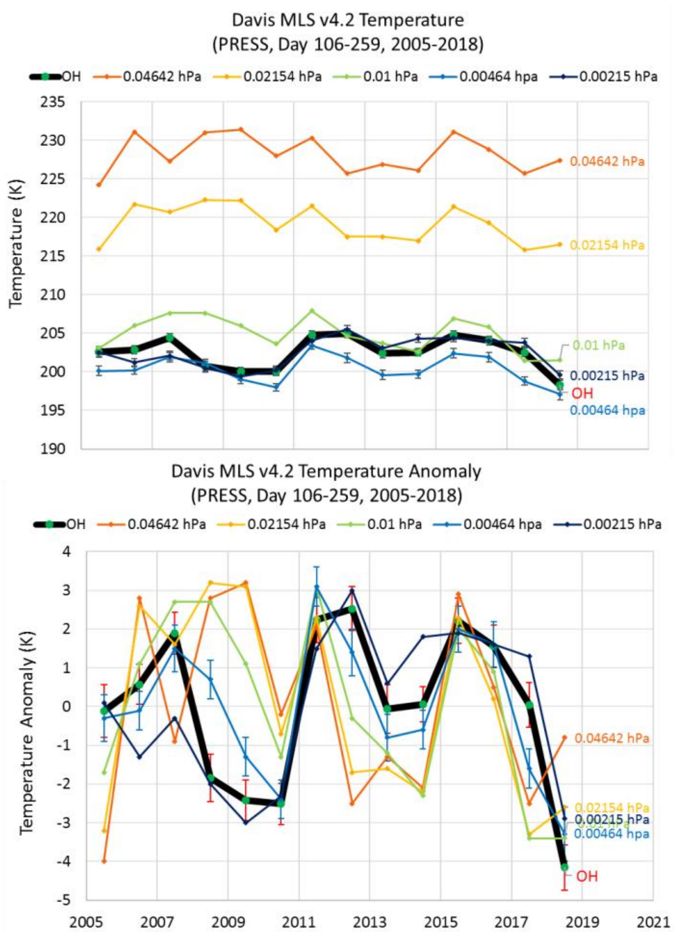
The broad vertical resolution (15km averaging kernel) of MLS profiles in the mesopause region is referenced on line 290. Its effect is to integrate temperatures from above and below the OH layer into the computed temperature for that pressure level. Bearing in mind that the OH profile is itself a broad layer (FWHM ~8km) and the rotational temperatures computed correspond to a similar integration over the width of the layer it is not unreasonable to compare OH with MLS temperatures. Indeed we find good agreement.

We have previously reported, and routinely compare our measured OH temperatures with both Aura/MLS and SABER profiles. In particular, French and Mulligan, 2010 examined biases between Davis OH temperatures and both Aura/MLS and SABER. A significant limitation of SABER for comparisons with Davis observations is the yaw cycle sampling of the satellite. Comparable observations over Davis are confined to two intervals (day-of-year 75–140 and 196–262) and days prior to 106 and after 259 are outside the OH winter averaging interval. Therefore only days 106-140 and 195-259 are comparable and a large part of the winter months is not sampled. As a consequence SABER winter averages do not fit the OH observations at Davis as well as MLS.

We agree that the geometrical altitude of the layer varies. Since the hydroxyl layer position is primarily controlled by collisional quenching with O_2 and N_2 on the bottom-side of the layer, and reaction with atomic oxygen on the top-side of the layer it is the concentration (density) of the reacting species that governs the layer position. Therefore it is reasonable to compare with MLS pressure (proportional to density) levels than on geometrical altitude levels. See also details of SABER altitude and pressure plots addresses in item 20 below.

This work is primarily concerned with mesopause region trends and absolute temperature biases are removed by subtraction of the climatological mean. We work here with anomalies (difference from the mean of all years) and residuals (solar cycle component removed). In selecting the 0.0046hPa level we compared the Davis OH winter average anomaly with MLS [AMJJAS average] anomaly over a range of altitude and pressure levels. (see plots below, the first are altitude ranges, the second pressure levels, both temperatures and the anomaly are shown).

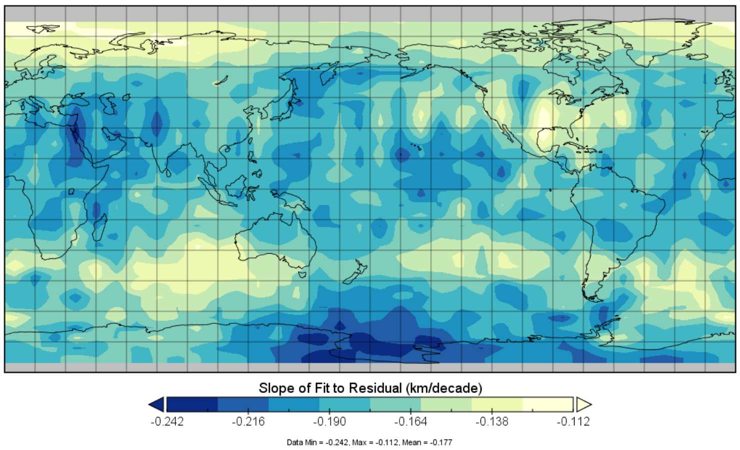

Davis MLS v4.2 Temperature
(ALT, Day 106-259, 2005-2018)
OH
63-68 km
68-73 km
73-78km
78-83 km
83-88 km
85-90km
88-93km
Temperature (K)
Davis MLS v4.2 Temperature Anomaly
(ALT, Day 106-259, 2005-2018)
Temperature Anomaly (K)



Altitude ranges 78-83km, 83-88km and 85-90km and pressure ranges 0.00215hPa and 0.00464hPa are all in reasonable agreement (<5K in absolute terms) with the OH temperatures, but we know there are biases with MLS (see French and Mulligan, 2010), and we know that the Davis OH temperatures are ~2 K high using LWR transition probabilities compared to those computed with the experimentally measured transition probability ratios determined in French et al 2000. These biases are removed by comparing anomalies. We calculate the Chi-Square goodness of fit parameter between the OH winter average *anomaly* with the Aura/MLS anomalies. The 0.0046hPa pressure level yields the smallest chi-sqr (14.8) compared to a layer centred on the traditional 87km altitude level (85-90km chisqr=18.8). This difference is small, but we prefer the pressure level comparison for the reason given above.

The relationship between pressure and geopotential height (GPH) is examined below using the MLS data set. Global decreases in GPH anomaly (between ~110 to 240 metres/decade) at the 0.0046hPa pressure level are consistent with a contraction of the underlying atmosphere and also consistent with the SABER trend in OH mean winter layer altitude for Davis shown in Fig 7. (200metres/decade) and discussed in the text.

MLS slope of fit to residual geopotential height at 0.0046hPa



There is also a lack of discussion on previously reported seasonal or latitudinal effects on trends that the authors mention but do not connect with their results. It would be useful to overplot these results from other authors on the corresponding figures in the manuscript.

We have modified figure 6 to indicate the solar cycle response and long-term trend coefficients available from other authors listed in Table 1.

A comparison of the seasonal variations in long-term trends has been previously published in a similar figure to Fig 5. in French and Klekociuk, 2011 (their figure 7; reproduced below). This included the results of Offermann et al 2010 at Wuppertal and Espy and Stegman (2002) from Stockholm. Since there has not been further updates to the seasonal variation in trend coefficients at either site we have not replicated this comparison. Perminov et al., 2018 offer seasonal variances in OH temperatures for the Russian sites of Zvenigorod and Tory but do not compute seasonal trend components.

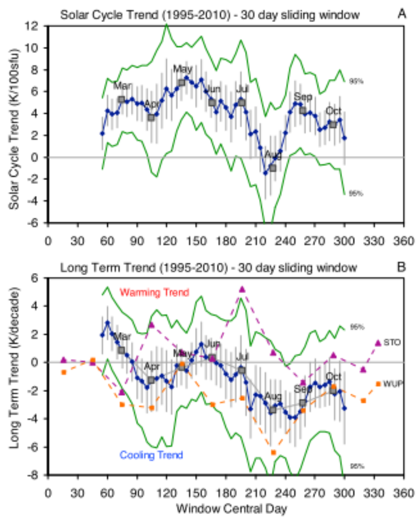


Figure 7. The 30 day sliding window (5 day step) evaluations of (a) solar cycle and (b) long-term trend coefficients at zero F10.7 lag (i.e., the vertical transect through Figure 6 (right-hand panels) at zero lag). One-sigma error bars and the 95% confidence limits (upper and lower traces) are as marked. The (true calendar) monthly evaluations are also plotted and labeled. Overlaid on Figure 7b are the equivalent monthly trend results by *Offermann et al.* [2010] from Wuppertal (labeled WUP) and *Espy and Stegman* [2002] from Stockholm (labeled STO). These are approximate values scaled off their respective seasonal trend plots and are offset by 6 months to match the Southern Hemisphere season.

The discussion section is too long. It is a good review but it is not easy to follow and, more importantly, to see how the results presented here fit on the discussion. I suggest revising the section, shortening it and putting the results of this paper into the context.

The discussion section has been substantially revised and some sections removed.

Other comments and suggestions are:

1. L42---45 Include in the abstract the result of the global MLS trend analysis.

We have included in the abstract the Aura/MLS solar cycle (3.39 ± 2.3 K/100 sfu) and long term trend (-1.3 ± 1.2 K/decade) coefficients at Davis for comparison. These are computed from anomalies derived from the AMJJAS means of all satellite observations within 500 km of Davis station over the 14 years of MLS observations (compared to 24 years of OH observations at Davis).

We note from Fig 6 that significant variability appears in both long term trend and solar cycle coefficients computed from MLS on a global scale. The global coefficients from MLS computed in the 5° latitude x 10° longitude grid boxes for AMJJAS averages at 0.0046hPa range from -2.3 to +2.3 K/decade (mean -0.01 K/dacade) for the long term trend and -0.2 to 8.8 K/100 sfu (mean 3.3 K/100sfu) for the solar cycle. It is not practical to include all the global MLS trend results in the abstract.

2. L148---L151. Is there any error associated to that interpolation?

Yes. The interpolation attempts to account for changes in the overall intensity of the OH emission during the course of the 7 minute scan. In some cases the intensity may not vary in a linear fashion, but in general interpolating intensities to a common time between consecutive scans provides a better estimate for varying intensity of the OH emission. Selection criteria limit extreme rates of change of intensities (<6%) between consecutive spectra.

The error assigned to each line intensity is the square root of the total number of counts, together with an error in estimating the background under each line. A standard deviation error is also derived from the 3 different ratios contributing to the weighted mean temperature for each pair of consecutive spectra. The process is described in detail in French and Burns (2004) in sections 2. Measurements and 3. Rotational temperature analysis.

French, W. J. R. and Burns, G. B.: (2004) "The influence of large-scale oscillations on long-term trend assessment in hydroxyl temperatures over Davis, Antarctica", J. Atmos. Sol. Terr. Phys., 66, 493–506. 634

3. L151. Is that 2% contribution independent of temperature? Or in other words, do potential uncertainties in the Q---line contribution incur into significant errors in the derived temperatures?

No, the $Q_1(5)$ is not temperature independent, but its contribution to $P_1(2)$ is computed in an iterative process using the final weighted mean temperature from three possible ratios from the $P_1(2)$, $P_1(4)$ and $P_1(5)$ lines. The contribution of the two lambda doubled components of $Q_1(5)$ and computed separately from the final weighted mean temperature. Approximately 98.0% of $Q_1(5)$e and 47.5% of $Q_1(5)$f contribute to the $P_1(2)$ emission intensity we measure, depending on the instrument line shape measured via frequency stabilised laser.

4. L157. Please, write a short sentence explaining why your choice is Langhoff et al. (1986).

We use Langhoff et al. (1986) transition probabilities because they are closest to the experimentally measured, temperature independent line ratios determined for the OH(6-2) band using the same instrument in French et al 2000.
Recent work by Noll et al (https://www.atmos-chem-phys-discuss.net/acp-2019-1102/acp-2019-1102.pdf) also show relatively small errors in the comparison of populations from P- and R- branch lines for the Langhoff et al (1986) coefficients, as well as van der Loo and Groenenboom (2008) and Brooke et al. (2016) coefficients. The latter two sets were not available at the time of that study.
The paragraph was modified to encompass this explanation.

5. Do you reach better agreement with satellites when using specific values?

This study is not an assessment of the bias between Davis OH temperatures derived with different transition probabilities and satellite measurements. We have previously examined this in French and Mulligan, 2010.

This work concentrates on the trends and variability inherent in the annual anomalies (any bias is removed by subtraction of the climatological mean) and thus is independent of the choice of transition probabilities (see further comments below). However, we do obtain good agreement of absolute temperatures using Langhoff et al (1986) probabilities (justified above) with Aura/MLS at the 0.0064hPa level and with SABER using a Gaussian fit to the OH-B channel VER (French and Mulligan, 2010), given the many assumptions made with regard to the layer height and shape, the width of the averaging kernel used for satellite retrievals.

6. L160. Did you explicitly test the effect on trends of the probabilities used? How much is "not significantly"? This is important in order to understand differences in trends between the different datasets.

From the rotational temperature equation for the temperature derived from the ratio of emission lines *m* and *n*

$$T_{rot} = \frac{E_m - E_n}{k\ln\left(\frac{I_n . A_m . (2J'_m + 1)}{I_m . A_n . (2J'_n + 1)}\right)}$$

Where E are the upper state energies, I are the measured intensities, A are the transition probabilities, J' are the upper state rotational quantum numbers and k is Boltzmann's constant.

Choice of a particular transition probability set only affects the ratio A_m/A_n and corresponds to an offset in T_{rot}. While this choice is important for comparisons of absolute temperature observations between sites, it is not important for studies of trends and variability so long as the same transition probability set has been used consistently for all years.

In this study, removal of the climatological mean, subtracts any offset due to differences in the transition probability ratio. The only conceivable differences between the temperatures derived using different sets is selection criteria boundary effects (whether individual measurements pass selection criteria on the extrema of the selection criteria limits). We believe (as for item 4 above) that Langhoff et al (1986) TP's are consistent with the experimentally measured, temperature independent ratio's examined in French et al 2000 and thus provide a reasonable estimate of the absolute temperature for comparison with SABER, Aura MLS, and other observations, however we make no assessment of bias against other observations here. Rather we assess trends and variability in the anomalies.
The paragraph has been modified to express these points.

7. L161---163. This sentence is somehow redundant. If biases due to the choice of probabilities can reach 12K, it is obvious that comparisons between different instruments depend on the choice. Please, delete.

Deleted.

8. L169. Specify selection criteria quantitatively.

Paragraph modified to read “Selection criteria limit extreme values of weighted standard deviation (< 20 K) and counting error (< 15 K), slope (< 0.06 counts/Å), magnitude (< 250 counts per second) and rate of change (< 3 counts per minute) of the backgrounds and the rate of change of branch line intensities (< 6%) between consecutive scans. Further details of the rotational temperature analysis procedure are available in Burns et al. (2003) and French and Burns (2004)”

9. L183 and Fig. 1. There are more recent versions for MSISE. Use latest version or at least show that it makes no difference.

Reference model updated to NRLMSISE-00. It makes little difference on the scale of the plot in fig 1.

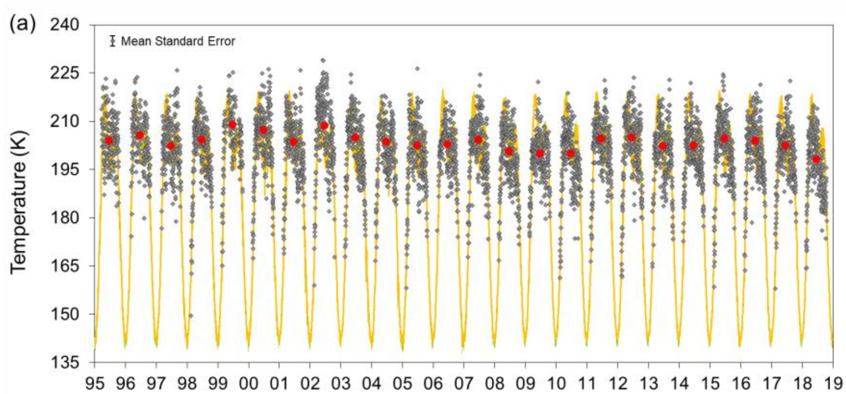


Fig1 caption modified accordingly and the reference updated to Picone et al 2002 [Picone, J. M., A. E. Hedin, D. P. Drob, and A. C. Aikin, NRLMSISE-00 empirical model of the atmosphere: Statistical comparisons and scientific issues, J. Geophys. Res., 107(A12), 1468, doi:10.1029/2002JA009430, 2002.]

The reference atmosphere (values for local midnight at Davis) has also been added to Figure 2 for comparison and text added to discuss the differences.

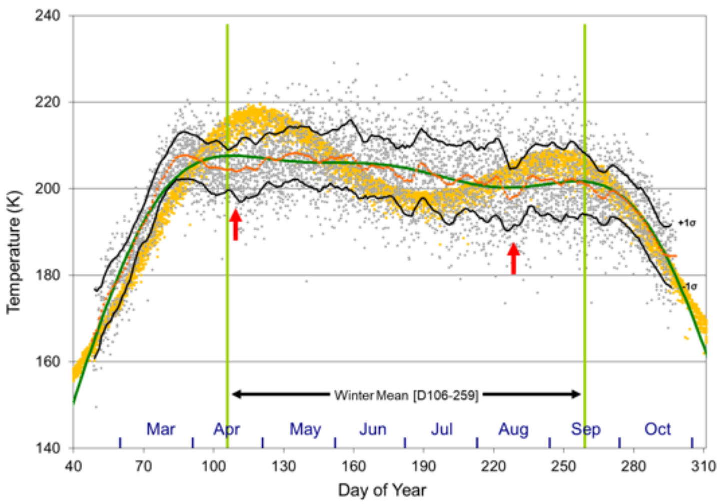


Figure 2. Superposed nightly mean temperatures from 1995 to 2018 [gray points] and a 5-day running mean which represents the climatological mean [orange line] with 1σ intervals [black lines]. The seasonal variation [green annual, semi-annual, ter-annual fit] and mid-April and mid-August dips [red arrows] are also indicated. Green vertical lines mark the calculation region for winter mean temperatures (inside the winter to summer transition intervals). The NRL-MSISE00 reference atmosphere (local midnight values for Davis) is also added for comparison [gold points]

10. L190. Introduce Fig. 2 at the beginning of this paragraph.

Inserted "(Fig. 2)" at the end of the first sentence and removed the sentence introducing Figure 2 midway through this paragraph.

11. L207---210. Remove this text from the caption. It is already in the text

Removed text and modified caption as for item 9 above.

12. L228. Are these MLS nightly means? Is there any time difference criterion used? If not, consider discussing possible sampling effects.

These are AMJJAS (6-month) means derived from all MLS observations within 500 km of Davis station as described in that line and in the paragraph from L282. There are only about 60 coincident samples per month (2 per night) within this range which gives little opportunity to apply a time restriction criterion.
Examination of nightly OH measurements show that tidal magnitudes are small (diurnal tide is <2K and semidiurnal <1K) and averaging over 6 months, and with the vertical averaging kernel of MLS at this altitude will average out tidal effects.

13. L246---252 and Fig. 4. This is a very nice and colorful figure but is it really essential or a short sentence would be enough?

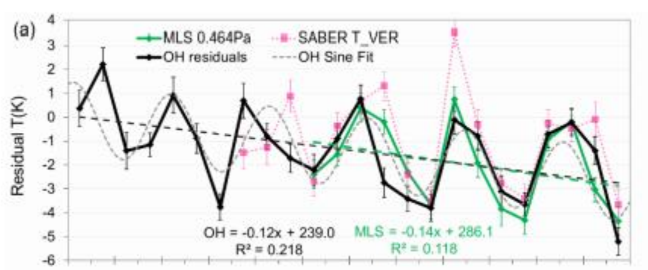
It was considered that the histograms were useful to show the variability and skewness in temperature distributions between years as a function of the mean annual winter temperature. It would be difficult to describe the range of distributions displayed in a short sentence.

14. L284. The choice of this altitude range sounds somehow arbitrary. The OH layer is not centered at 85km. Please, justify. What happens if you use 85---90km?

The altitude range selected was not arbitrary, but as the altitude range of the smallest chi-sqr fit to the OH temperatures. See discussion under first main comment.

15. L290. This is the main problem I see using MLS data for this study. Why don't you use SABER data, even with a smaller time---coverage during the winter months? It could do a good job in your 60---day running means solar and linear trends. Garcia---Comas et al. (2014) showed that MLS biases at these altitudes could be large on the South Pole.

We include the equivalent curve for SABER data in Part 2 (Figure 1) of this work on the QQO (available in discussions as acp-2019-1097). SABER's 60 day yaw cycle limits comparable observations over Davis to two intervals (day-of-year 75–140 and 196–262) and days prior to 106 and after 259 are outside the OH winter averaging interval. Therefore only days 106-140 and 195-259 are comparable and a large part of the winter months is not sampled. As a consequence SABER winter averages do not fit the OH observations at Davis as well as MLS.



SABER data and results are discussed extensively in the discussion and used to in section 5.3 to assess trends in the change in the height of the OH layer.

16. L293. The choice of this pressure level is not justified. You could be getting a good agreement due to a bias that could be masking a wrong selection of altitudes. Indeed, it is well known that the altitude of the OH---layer is variable (as you even mention in L582--584). Please, discuss this point.

We do not match absolute temperature but the variance in winter mean temperature anomalies over 14 years. The 0.0046hPa pressure level is selected as it yields the smallest chi-square of the pressure levels. See discussion above under main comments and item 20 below.

17. L299. Yes, you show this very nicely in Fig. 5, my favorite figure of the paper.

Yes, trends are not uniform, and the seasonal variation should be considered when comparing trends between observers. This is why we attempt to present some insight into seasonal and spatial trend variability of the mesopause region using the MLS dataset.

18. L302. Indicate the figure where this is shown.

Examination of the QQO feature is separated and undertaken in Part 2 of this work. The paper became too unwieldy to include both the trends analysis and QQO investigation in the one manuscript. This is brief statement on the seasonal variability of the QQO with the following sentence indicating it is discussed in more detail in part 2. The figure is provided as Figure 1b in Part 2. (reproduced below)

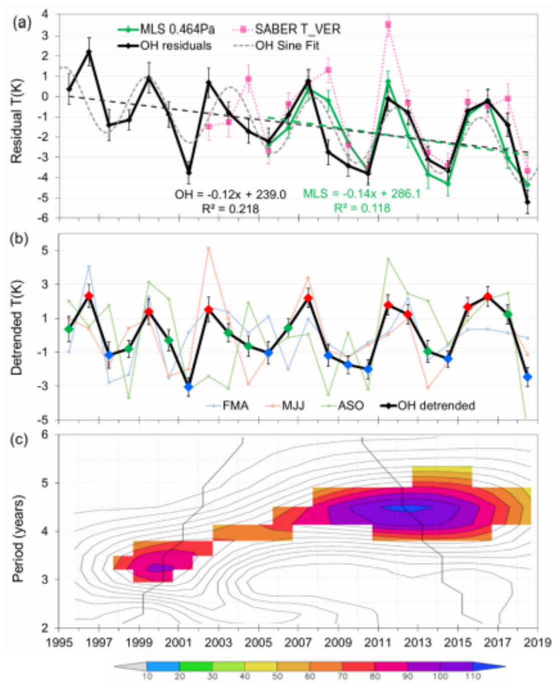


Figure 1. (a) Davis OH winter mean residual (solar response removed) temperatures (black line, standard error in mean error bars, dashed linear fit) compared with Aura/MLS [AMJJAS] mean residual temperatures for 0.0046 hPa (green line, standard error-in-mean error bars, dashed linear fit) and TIMED/SABER (pink dotted line, standard error-in-mean error bars). Gray dotted line is a sinusoid fit (peak-peak amplitude 3.0 K period 4.2 years). (b) Detrended Davis OH winter mean temperatures [AMJJAS] (black line, long-term linear fit removed) compared to FMA, MJJ and ASO monthly averages (red, green and blue points mark warm, mid and cold years for composite studies). (c) A Mortlet wavelet transform (order 6) of the detrended Davis OH winter mean temperatures. Coloured sections are power significant above 90% level as per colour bar. The black line indicates the cone of influence; points outside have been influenced by the boundaries of the time series.

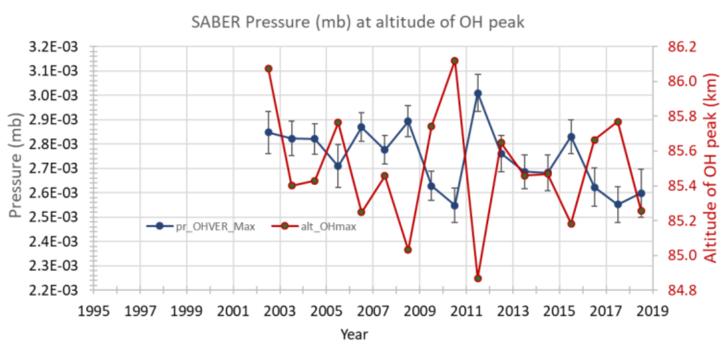
19. L304. Aura/MLS trends "at 0.00464 hPa"

Text added as suggested

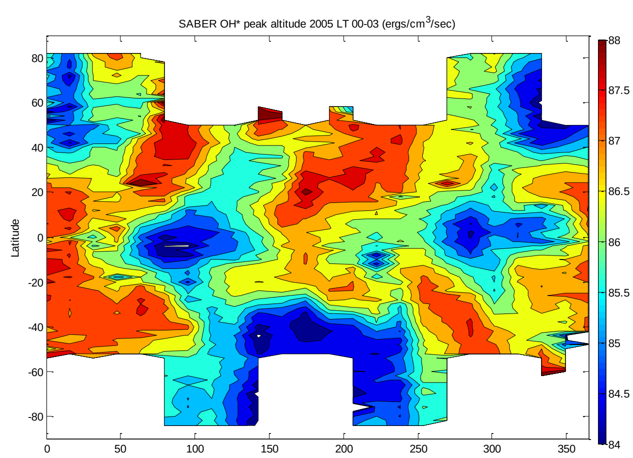
20. L307. How did you derive that this is the OH equivalent pressure level and that it does not change with latitude? Does SABER, measuring OH emission and temperature vs z and pressure, show a significant change of equivalent pressure with latitude?

As discussed above at the variation of the [AMJJAS] anomaly had the smallest chi-sqr compared to the measured OH winter average temperature anomaly at Davis.

The Figure below shows a comparison of SABER VER (altitude of peak) and corresponding pressure value (mb) for the years 2002-2018 (day 106 – 259 of each year) at Davis Station. The OH peak occurs at pressures in the range 0.00255 hPa to 0.003 hPa (2.55×10^{-3} - 3.0×10^{-3} mb). This value lies between two of the Aura MLS levels (0.00464 hPa and 0.00215 hPa) on which the averaging kernels are centered and is in excellent agreement with the Davis MLS 4.2 Temperature PRES plot (page 3 above) included in response to the first Main Comment above. An inverse relationship between altitude and pressure at the OH peak is clearly evident, and justifies the selection of a pressure level comparison for OH temperatures over an altitude level.



On the question of the altitude (or pressure) of OH peak as a function of latitude, the figure below shows the variation of the altitude of the OH peak as a function of latitude and day-of-year for the year 2005 from SABER data. The overall pattern shown here is repeated year after year with only minor changes in detail.



Based on the two figures above, the MLS averaging kernel centered on 0.00464 hPa would appear to be a good representative for the temperature of the OH layer.

21. L308. Monthly anomalies? 60---night running means? What are these?

These are AMJJAS (southern hemisphere winter months) averages as described on L309.

22. L316. What is the origin of these enhance bands?

To the best of our knowledge, this is the first time that these bands have been reported. This observation is discussed in the context of similar work (both observational and modelling) in section 5.4.

23. Section 5.1. I enjoyed reading the review but it could probably be shortened and better organized and your results should be put into the context you describe. Are they reasonable? Do they agree? Does the seasonality of your data agree with other results? Does it agree with the expected variations? I suggest extending the title of this section in order to include ozone.

Section 5.1 subtitle was modified to include ozone.

Section 5.1 does compare the present results with other reports in their context, e.g., in lines: 426-420, 459-460, 503-504, 510-511, 558-559 (line numbers refer to the original manuscript).

We have significantly shortened (from 114 to 84 lines) and reorganised the discussion in this section from line 427-487 – omitted the discussion of ACE-FTS CO2 rates of change and merged and modified the section from L436-477 to replace with the following -

"In a recent summary of progress in trends in the upper atmosphere, Laštovička (2017) identified greenhouse gases, particularly CO_2 as the primary driver of long-term trends there. The overall effect of greenhouse gases at mesospheric altitudes is radiative cooling. The important secondary trend drivers in the mesosphere and lower thermosphere (MLT) are stratospheric ozone, water vapour concentration and atmospheric dynamics. Temperature trends are predominantly negative, and recent progress in understanding the magnitude of the cooling have arisen from confirmation and quantification of the role of ozone. Lübken et al. (2013) present the results of trend studies in the mesosphere in the period 1961-2009 from the Leibniz-Institute Middle Atmosphere (LIMA) chemistry-transport model which is driven with European Centre for Medium–Range Weather Forecasts (ECMWF) reanalysis below 40 km, and observed variations of CO_2 and O_3. They find that CO_2 is the main driver of temperature change in the mesosphere, with O_3 contributing approximately one third to the trend. Linear temperature trends were found to vary substantially depending on the time period chosen primarily due to the influence of the complicated temporal variation of ozone.

Figure 3 of (Lübken et al., 2013) show a monotonically increasing trend on CO2 compared with a much more complicated temporal ozone variation (essentially constant until 1980, a rapid decrease from 1980-1995, followed by an increase since then. Trends in ozone vary as a function of both altitude and latitude, with positive trends dominating in the lower stratosphere and mesosphere. Increases in water vapour concentration are considered a secondary but non-negligible effect particularly in the lower thermosphere (Akmaev et al. 2006). The trend effect of dynamics was found to be very slightly negative in the mesosphere, but very small compared with the radiatively induced trends. At the mesopause, the trend due to dynamics was positive and significantly larger (~1 K/decade). These results were found to be in good agreement with observations from lidars, Stratospheric Sounding Units (SSU) (Randall et al., 2009) and radio reflection heights which have decreased by more than 1 km in the last 50 years due to shrinking in the stratosphere/lower mesosphere caused by cooling."

The paragraph from lines 478-486 had also been omitted.

24. L440. Include reference.
25. L449. Include reference.

The reference for both lines is Laštovička (2017), which appears in the opening line of the paragraph containing those lines. The reference was included again at the end of the (modified) paragraph.

26. L465. This is already said, also mentioning the same reference.

Replaced with the modified paragraph as above (Item 23).

27. L478---486. It is not clear to me what this has to do with this work. Hervig et al. (2019) paper mainly deals with the paradox on the solar response of H2O. Perhaps mentioning only their result related to temperature makes more sense. By the way, do you actually see a change of solar response of temperature from 1995 to 2018?

The Hervig et al (2019) reference and following discussion has been omitted in the revision of section 5.1.

The period 1995-2018 spans only 2 solar cycles, assessing the response of the two cycles independently would not be constructive considering the uncertainties.

28. L488. Also Solomon et al. 2018

The work of Solomon et al. (2018) using WACCM-X is cited in lines 662-671. The two studies are discussed separately, since Solomon et al. use constant low solar activity conditions in an attempt to disentangle temperature changes arising from anthropogenic effects from solar induced variations.

29. L501---502 I do not understand "global averaged temperature (..) as a function of latitude". How does their month---to---month variability compare with the seasonal variability you derive? Please, overplot on Fig. 5 and discuss.

Corrected ".. zonal average ..". Sentence now reads " Qian et al. (2019) provide zonal averaged temperature trend values as a function of altitude (50-110 km) and latitude for each month (their Fig. 3) some of which are statistically significant."

30. L507. Please, write altitude

The word altitude *is* present in the sentence.

31. L559. Write "from 1995 to 2018". If you remove your QQO, don't you see such breaks? From your measurements, it seems that your trend is not monotonic. Quantify "no obvious sign".

Added “from 1995-2018” to the end of the sentence.

We are not certain how you propose we remove the QQO variation. In order that it be removed the process generating the QQO needs to be understood. We devote significant effort to examine this QQO signal in more detail in Part 2 of this work but are unable at this stage to isolate the mechanism, therefore have no index to model the QQO.

If there was a trend break in the period 1995-2018, one might expect to see a significant change in the trend, and in the solar response, when extending the period of the study from 16 years to 24 years. Such a change is not observed (as you note in your opening paragraph “the trends derived here coincide with their previous results and those for the solar response are only slightly different”. To quantify, the coefficients over 16 years (French and Klekociuk, 2011) were
4.30 ± 1.02 K/100sfu (95% confidence limits 2.2 K/100sfu < S < 6.4 K/100sfu)
-1.20 ± 0.51 K/decade (95% confidence limits -0.14 K/decade < L < -2.26)
and 24 years were
4.79 ± 1.02 K/100sfu (95% confidence limits 2.6 K/100sfu < S < 6.99 K/100sfu)
-1.18 ± 0.87 K/decade (95% confidence limits 0.71 K/decade < L < -3.06)
Neither coefficient has changed outside the uncertainty.

We are unsure how you deduce from our measurements that the “trend is not monotonic”?

32. L584. Please, also mention Liu et al. (2006).

Added Liu et al. (2006).

33. L592. Garcia---Comas et al. (2017) also estimated the trend and solar response of OH* altitude and temperature from SABER.

The following sentence has been added after line 580. “García-Comas et al. (2017) reported a slightly larger decrease of 40 m/decade in SABER OH volume emission rate weighted altitude at mid-latitudes which accompanied a 0.7%/decade increase in OH intensity and a 0.6K/decade decrease in OH equivalent temperature.”

34. P610. What is the expected change in MLS temperatures due to a change of 0.02---0.04 km? This may lead to a bias in the comparison between DAVIS and MLS. A way to test this could be done with SABER data by comparing temperature trends at a fixed altitude to temperature trends at a OH VER weighted temperature.

Since the averaging kernel of MLS temperatures at the OH altitude is of the order of 15km a change of 20 to 40 metres over 16 years (determined from SABER VER altitude 2002-2018, shown in Fig 7) is negligible!.

35. L657---661 Why are these results so different? Could the difference be due to sampling?

WACCM-X is a model, Aura/MLS is measured data. The overall long term trend in WACCM-X at 85km is -0.52 ± 0.64 K/decade (Fig 10 in Qian et al., 2019) although zonally averaged monthly trends are more variable (Fig 3 in Qian et al., 2019).

36. L720. According to a possible contribution of 30% by ozone, these values have (at least) a 30% error.

The clause in parenthesis in lines 721-722 acknowledges the fact that the trend quoted ignores possible contributions of stratospheric ozone to the trend.

37. L723---726. Could you provide the reason for this minimum in 2009?

The minimum of both the solar cycle and QQO cycle occur in 2009. (See Fig 3a.)

38. L747. Mention here also the results from other OH observations (those listed in Table 1).

We have re-written these sections with the request to shorten the discussion in section 5. Trend comparisons with other ground based observers listed in Table 1 is covered in section 4.4

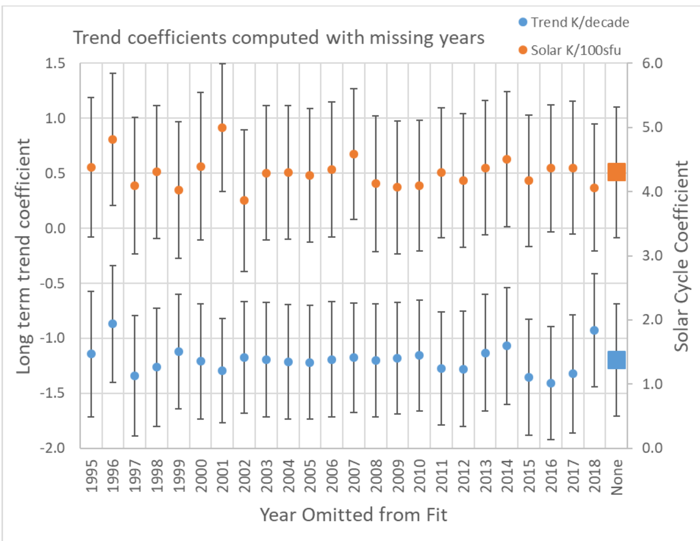
39. Fig1b. Do all years contain the same days of measurements from doy 106 to 259? If not, there could be a sample bias.

With the exception of 1999 when 2 intervals D095-126 and 213-249 were used to scan the OH(8-3) band and 1996 missing D176-202 all other years only have more than 85% nights
within the winter averaging window sampled. (ie 85% of the nights have a valid nightly average temperature with at least 10 measurements that pass selection criteria). A sample bias could be introduced in computing the anomalies if there was a significant departure from the climatological mean in those intervals.

year	1995	1996	1997	1998	1999	2000
of 153	133	111	131	151	81	145
% nights	86.9%	72.5%	85.6%	98.7%	52.9%	94.8%
year	2001	2002	2003	2004	2005	2006
of 153	148	149	138	133	130	139
% nights	96.7%	97.4%	90.2%	86.9%	85.0%	90.8%
year	2007	2008	2009	2010	2011	2012
of 153	146	143	146	137	150	150
% nights	95.4%	93.5%	95.4%	89.5%	98.0%	98.0%
year	2013	2014	2015	2016	2017	2018
of 153	146	149	142	152	150	142
% nights	95.4%	97.4%	92.8%	99.3%	98.0%	92.8%

The sliding window (seasonal variation in trend parameters) give some indication of the range of trends obtained by selecting different intervals compared to the winter mean interval.

We have also tested the effect on the derived coefficients by omitting individual years sequentially from the model fit computation. These show the range of coefficients if a data gap for the entire winter interval was missing. All coefficients derived from the omitted year computations remain within the uncertainty limits of the solar cycle and long-term trend coefficients when all years are included.



The following paragraph was added to section 4.1 to address this concern

"The stability of trend coefficients was tested for the presence of sampling gaps in the OH temperature record. With the exception of 1999 when 2 intervals D095-126 and 213-249 were used to scan the OH(8-3) band and 1996 missing D176-202 all other years only have more than 85% nights within the winter averaging window sampled. (ie 85% of the nights have a valid nightly average temperature with at least 10 measurements that pass selection criteria). A sample bias could be introduced in computing the anomalies if there was a significant departure from the climatological mean in those intervals. The test examined the effect on the derived coefficients by omitting individual years sequentially from the model fit computation. These show the range of L and S coefficients if a data gap for the entire winter interval was missing in a particular year. All coefficients derived from the omitted year computations remained within the uncertainty limits of the solar cycle and long-term trend coefficients when all years were included. "

40. Figure 2. Time coverage changes with doy. What is the effect of DW1?

Over the winter averaging window (D106-259) the diurnal time coverage varies from 13:13 hrs (D106, 15-Apr) to 19:00 hrs (D177, 21-Jun) and 10:45 hrs (D259, 15-Sep). From the OH nightly observations we observe that the amplitude of the diurnal tide is <2K and semidiurnal tide <1 K.

We sample the same hours on the same days each year, and average those over days 106-259 each year to derive the trends.

41. Fig 5.a. the minimum in solar trend is during the month when downwelling is maximum. This migh be an indirect compensation of the cooling due to the direct dependence of downwelling (warming) and solar flux (COMPROBAR!!!)

We agree that the (August) minimum in the solar trend may indeed be the result of indirect compensation of cooling by the warming from maximum downwelling at that time of the year. However, we consider that removal of the climatological mean calculated over two full solar cycles is the best that we can do to eliminate the substantial part of the seasonal trend, leaving the anomaly values as shown in Figure 1(b). The solar cycle and long-term trends were calculated simultaneously using linear regression on the anomaly. A non-linear effect of the seasonal behaviour of the OH layer on the solar cycle response, e.g., through downwelling of atomic oxygen rich air, which could lead to increase production of CO cannot be ruled out, but is beyond the scope of the present work.

42. Fig 5b. Perhaps you might be sounding different altitudes? What is the seasonal change of the altitude of the OH layer? Did you look at SABER data? Also, this might be connected to O3 trend seasonality or CO2 trend seasonality.

The altitude of the OH layer peak appears to have a substantial seasonal response at DAVIS with an altitude minimum in mid-winter. However, we believe that removal of the climatological mean over two full solar cycles together with fitting solar cycle and long-term trends simultaneously would take account of this variation. Part 2 of this work examines the seasonal and long-term relationships between observed trends in temperature, CO_2, O_3, and CO.

43. Figure 5. Define grey boxes and blue dots.

These *are* defined in the figure caption.

44. Figure 6. What is time sampling for MLS? Are you removing tides? Trends strongly depend on sampling (Rezac et al. 2018)

These are 6 month anomaly averages (AMJJAS and ONDJFM) in each grid box, as described in the caption. They are the same averaging intervals each year. Tides are small and average out over 6-month means.

45. Fig 6. Please, overplot trend at Davis on the 1d plot. Also indicate CAP and LEO position on the maps.

Have modified figure 6 to indicate the positions of all ground-based observations in table 1 where long-term trend and solar cycle coefficients have been provided for comparison. Note that the 1d plots are a zonal mean, and as the map plots show there is considerable spatial variability in the trends derived from MLS data.

46. Fig 6. b. The blue/red bands at 70N in the NH winter months look like the trend and solar response related to stationary PWs.

We are grateful for this interesting suggestion. In the interest of keeping the manuscript to a reasonable length, we decided to defer a detailed study of this point, and we are reluctant to speculate on it without supporting evidence.

47. Fig. 7. What is the realtionship between this plot and the temperature anomaly? Can it help to explain differences between DAVIS and MLS?

Section 5.3 describes the effect that a vertical shift in the altitude of the OH layer would have on the emission weighted temperature which is measured by a ground-based instrument like that at Davis. The purpose of Fig. 7 is to show that SABER data does not indicate a significant change in the altitude of the OH layer during the period 2002-2018. Therefore we can eliminate change in the altitude of the OH layer as a cause of temperature change detected by the spectrometer at Davis.

This Figure does not address any differences between Davis and MLS.

48. Table 1. Discuss these results in the text, particularly mention them in section 5.

The majority of these results are discussed in the text in section 4.4 Section 5 has been substantially re-written. See response to point 38 above.

49. Table 1. Include MLS results in this list

This table was constructed on the basis of ground-based measurements only (as described in the caption), with the result that Aura/MLS results are not included. The global results of MLS are provided in Figure 6.

Response to RC2

Interactive comment on "Analysis of 24 years of mesopause region OH rotational temperature observations at Davis, Antarctica – Part 1: Long-term trends" by W. John R. French et al. Anonymous Referee #1 Received and published: 7 February 2020

Reviewer Report on the manuscript acp-2019-1001 Analysis of 24 years of mesopause region OH rotational temperature observations at Davis, Antarctica – Part 1: Long-term trends by W.John R. French et al.

General Remarks

1. . .The paper presents 24 years of observations of OH temperatures, which is an intersesting extension of an earlier data set worth publishing.

Thank you.

2 The data were taken in Antarctica where such measurements cannot be performed in summer. This is a drawback for several interpretation aspects and must be carefully considered.

Observations of the hydroxyl nightglow cannot be made over the summer at this latitude and we understand this is a limitation of the observational program for long-term trends using this technique. We contribute a solar and long-term trend assessment of the mean winter temperatures at this high southern latitude site and make comparisons with satellite observations to place these observations into global context.

3 The data are discussed in the context of increasing CO2 mixing ratios. They are extensively compared to MLS and SABER satellite results, and to computer models (WACCM-X).

Yes, comparison with other observations and models place these measurements into context.

4 The paper gives a long term analysis and discusses possible trend breaks. These results are questionable because of the lack of winter data.

Winter data is provided. It is the southern hemisphere *summer* data that is lacking. We make this clear from the outset of the manuscript.

5 The authors see a quasi-quadrennial oscillation (QQO) in their data. They announce a detailed discussion in a second part of the paper. This should take into account recent work in the literature on 3 – 5 year oscillations.

The second part of this work is available in discussions as acp-2019-1097

6 The paper is well written, and is recommende for publication after some modifications.

Thank you.

Major Comments

Line 221 pp: Figure 3 indicates five oscillation periods. The approximate period lengths are 2x 3 yr, 1x 4 yr, and 2x 5 yr. It is not obvious that a mean can be taken. Superposition of a 3 yr and a 5 yr oscillation should be checked (see for instance Offermann et al., JASTP 135,1, 2015).

The quasi-quadrennial periodicity revealed in the residual temperatures (seasonal, solar cycle and long-term trend fits removed) is an interesting feature and forms the basis of part 2 of this work; available in discussions as **acp-2019-1097**. We examine many possible sources for this feature using correlation and composite analyses with other data sets. We cannot ascertain whether the oscillation is a superposition of 3 and 5 year periodicities.

Line 383 pp, 540: The paper Offermann et al., 2006, should not be used to demonstrate a trend break. It was outdated by Offermann et al., JGR 115, D18127, 2010, who show a longer data series.

The reference to Offermann et al., 2006 applied to a time period when trend breaks first began to appear in the literature in relation to mesopause region temperature trends. In that context the reference is still valid despite it being later updated. We have added the updated reference (Offermann et al., 2010) to the paragraph noting the continued occurrence of trend breaks.

Offermann, D., P. Hoffmann, P. Knieling, R. Koppmann, J. Oberheide, and W. Steinbrecht (2010), Long-term trends and solar cycle variations of mesospheric temperature and dynamics, J. Geophys. Res., 115, D18127, doi:10.1029/2009JD013363

Line 405 pp, Section 5.1, 5.4: In the discussion of the trend data it should be elaborated that the summer data at Davis are missing. Trend data are different in summer and winter as shown by the MLS data in your Fig. 6. They can also vary from month to month as shown in your Fig. 5, but variations could be much larger (see for instance Offermann et al., 2010, their Fig.9). Possibly the summer trends are larger than your 1.2 K/decade (by number), and so might be the trend of annual data. Hence, if you want to include Davis data to Tab.1 please use annual MLS data!

The fact that we derive a trend in the mean *winter* temperatures is explicitly stated in the first sentence of the discussion in section 5.1 (lines 407-408) and in the first sentence of section 5.4 (line 619) “southern hemisphere (SH) *winter* months (AMJJAS)”. We think it should be clear to the reader that hydroxyl temperatures at Davis cannot be obtained over the summer months as the sun does not go down.

We do explore the seasonal variation in trends where possible over the observing season at Davis (fig 5) and understand that the trends are variable.

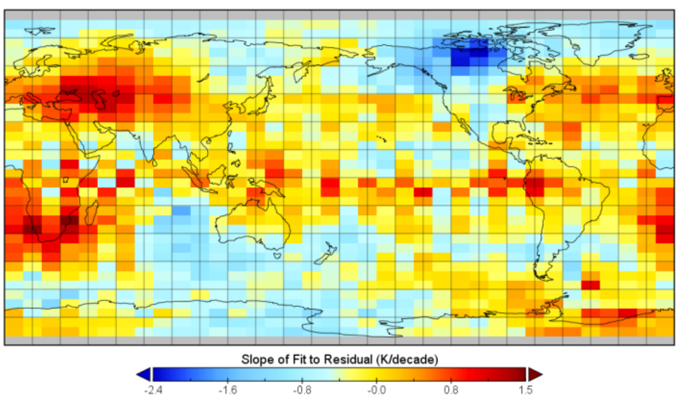
We also compare with MLS over winter and summer for each hemisphere (Figure 6) to show the seasonal difference in trends. MLS does not indicate that the summer trends [ONDJFM average trend plots in Fig 6] are larger than -1.2 K/decade for the grid box over Davis. The grid box over Davis yields a long-term trend of +0.02±0.08 K/decade for the summer months [ONDJFM] and an annual average value [JFMAMJJASOND] of -0.37±0.06 K/decade.

Table 1 is a list of ground based observations. It does not contain the MLS trends as they are globally and seasonally variable. Figure 6 shows the map of these coefficients for comparison with the ground based observations in Table 1.

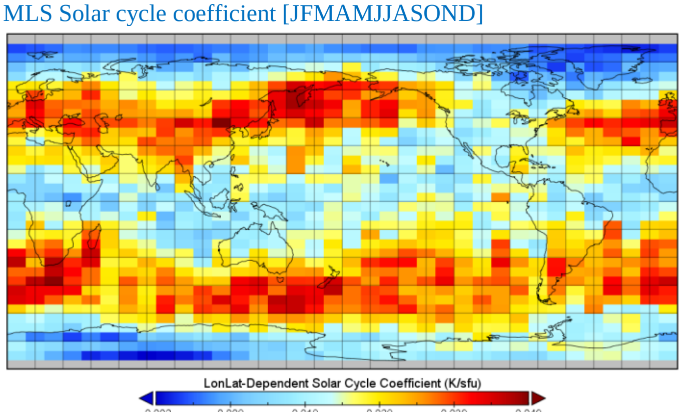
Line 558/559: . . .”no . . .sign of a discontinuity in the trend. . .” Kalicinsky et al., 2018, in their long-term analysis find that the summer data may be much more important than the winter data. Therefore please check your above statement by means of annual MLS data.

MLS values for the summer and winter seasons for each hemisphere are shown in Figure 6. It does not appear that the summer values are greater than winter values for the MLS trend coefficients globally (the two linear trend plots in figure 6 use the same colorbar scale). Coefficients maps (linear trend and solar cycle) for the whole year (all month MLS averages) are provided below for comparison with those in Fig 6. It merges the winter and summer features (as one would expect) and particularly highlights the mid-latitude maxima in solar cycle response.

MLS Linear trend coefficient [JFMAMJJASOND]



MLS Solar cycle coefficient [JFMAMJJASOND]



Minor Comments

Line 219: The error of your solar cycle response (1.02 K/100sfu) appears relatively large (see Tab.1). Do you know a reason?

The solar cycle response error is not unreasonably large compared to others in Table 1. (less than, or of the same order as 7 out of 11). We would suggest that the main reason for the error is the goodness-of-fit of the F10.7 and linear trend model. In particular that the model does not contain a quasi-quadrennial oscillation term that is evident in the residuals.

Line 295: Please give the error of the MLS trend.

MLS trend is 1.4±1.1 K/decade . added to text

Line 325 pp, Fig.6: a) Please show Panel numbers. b) Please show latitude scales. c) Part of the captions are difficult to read. d) Line 334: There is no Fig.1B. Do you mean Fig.3b?

Added panel numbers and latitude scales and modified the figure to improve readability. Yes, thank you we mean figure 3b for the comparison, but the caption has changed with the addition of all sites to the map for comparison as requested by another reviewer.

Line 353 pp, Table 1: Plese give the selection criteria for the sires shown.

We have updated the table from Table 2 in French and Klekociuk (2011) where updates were available since 2011.

Line 456 pp: ". . .peak altitudes.." Here and in the following it is sometimes unclear whether you mean the maximum of the peak or the geometric altitude. Please clarify.

We mean the altitude of the peak of the VER profile. This section has been extensively modified in response to another reviewer and this paragraph is no longer included.

Line 505 pp, 508: It is unclear whether you mean your Fig.5 or Qqian et al.. Please clarify.

This refers to the seasonal variation in the solar cycle coefficient in our Fig 5(a) and can be compared with Fig 4 of Qian et al (2019); this is clarified in the text.

Line 686: Fig.1a does not show this! Do you mean that you derived it from this Figure?

The R^2 value is shown in Fig 3(a) (Fit of solar cycle and long-term trend model to OH temperatures). This is corrected in the text.

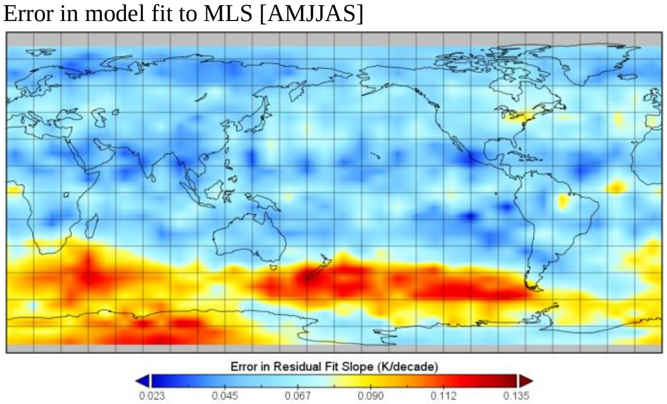
Line 693: Sentence difficult to understand.

Modified sentence to read "Although the altitude of the mesospheric cold point changes with season (e.g., Yuan et al., 2019) and tends to be higher than the centroid height of the OH* layer, the global solar response value obtained for T-CPM (4.89 ± 0.67 K/100 SFU) is in good agreement with the solar response coefficient derived from ground-based OH* observations."

Line 706 pp: Where can this be seen? The Supplementary Material was not available to me.

The original Fig 6. also contained a plot of the error in fitting the solar cycle and long term linear trend model (see below) and the point was made here that where the error was largest coincides with a strong QQO signal. However the QQO investigation is now discussed entirely in Part 2 of this work (acp-2019-1097).
The paragraph was modified to read "As a final comment on the global trends, it is noted that the largest errors in the linear trend fit for the SH winter understandably occur coincident with the regions positively or negatively correlated with the QQO (not shown here). The fit can be significantly improved if the QQO component can be understood and modelled. We investigate the QQO in detail in part 2 of this work. "

Error in model fit to MLS [AMJJAS]



Tracked Changes Manuscript

Analysis of 24 years of mesopause region OH rotational temperature observations at Davis, Antarctica. Part 1: Long-term trends.

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Abstract

The long term trend, solar cycle response and residual variability in 24 years of hydroxyl nightglow rotational temperatures above Davis Research Station, Antarctica (68° S, 78° E) is reported. Hydroxyl rotational temperatures are a layer-weighted proxy for kinetic temperatures near 87 km altitude and have been used for many decades to monitor trends in the mesopause region in response to increasing greenhouse gas emissions. Routine observations of the OH(6-2) band P-branch emission lines using a scanning spectrometer at Davis station have been made continuously over each winter season since 1995. Significant outcomes of this most recent analysis update are (a) a record low winter-average temperature of 198.3 K is obtained for 2018 (1.7 K below previous low in 2009) (b) a long term cooling trend of -1.2 ± 0.51~~1.2~~ K/decade persists, coupled with a solar cycle response of 4.3 ± 1.02 K/100 solar flux units and (c) we find evidence in the residual winter mean temperatures of an oscillation on a quasi-quadrennial (QQO) timescale which is investigated in detail in part 2 of this work.

Our observations and trend analyses are compared with satellite measurements from Aura/MLS version v4.2 level 2 data over the last 14 years and we find close agreement (a best fit to temperature anomalies) with the 0.00464 hPa pressure level values. The solar cycle response (3.4 ± 2.3 K/100sfu), long-term trend (-1.3 ± 1.2 K/decade) and underlying QQO residuals in Aura/MLS are consistent with the Davis observations. Consequently, we extend the Aura/MLS trend analysis to provide a global view of solar response and long term trend for southern and northern hemisphere winter seasons at the 0.00464 hPa pressure level to compare with other observers and models.

1. Introduction

Long-term monitoring of basic atmospheric parameters is fundamentally important to understand natural, periodic and episodic variability in atmospheric processes, to provide data to verify increasingly sophisticated atmospheric models and to resolve and quantify perturbations due to global change on decadal to century timescales. Dynamical processes, including gravity waves, tides, planetary waves, large scale circulation patterns and quasi-periodic teleconnections (such as the quasi-biennial oscillation (QBO), El Niño Southern Oscillation (ENSO), and the Pacific Decadal Oscillation (PDO)), changes to the chemical composition and radiative balance (particularly due to anthropogenic emissions of greenhouse and chlorofluorocarbon gasses) and external forcing such as the 27-day solar rotation and 11-year solar activity cycle, all play significant roles (directly and through interactions) in defining and perturbing the mean state of the atmosphere. Decades of well calibrated measurements are required to accurately quantify variations and trends on these timescales.

Meteorological reanalyses derived from assimilation of a vast number of surface observations provide time-series for useful trend analyses ~~for~~ in the lower atmosphere e.g. (Bengtsson et al., 2004). A few satellite based data sets are now also reaching multi-decadal timescales (e.g. the Thermosphere Ionosphere Mesosphere Energetics Dynamics satellite's Sounding of the Atmosphere using Broadband Emission Radiometry instrument (TIMED /SABER) (Mertens et al., 2003), and the Earth Observing System satellite Aura Microwave Limb Sounder (Aura/MLS) (Schwartz et al., 2008), that extend observations to the upper atmosphere. Of current and particular interest to climate science in the modern era are the atmospheric temperature trends in response to increasing global greenhouse gas emissions, principally from carbon dioxide (CO_2). Modelling studies over many years suggest that the sensitivity to CO_2 changes in the upper atmosphere, particularly at high latitudes, is much larger than in the lower atmosphere (e.g. Roble (2000), the Canadian Middle Atmosphere Model (CMAM) (Fomichev et al., 2007)) and the Hamburg Model of the Neutral and Ionized Atmosphere (HAMMONIA) (Schmidt et al., 2006)).

Above the stratosphere, the low collision frequency means that CO_2 preferentially radiates absorbed energy to space, resulting in a net cooling. Thus, the expected long-term temperature trends in the mesosphere and lower thermosphere due to CO_2 are negative. Ground based optical measurements of the Meinel emission bands of the hydroxyl (OH) molecule produced by the exothermic hydrogen (H) – ozone (O_3) reaction ($H + O_3 \rightarrow OH^*$ + 3.34 eV) have been used extensively over almost six decades as a method of measuring atmospheric temperature in the vicinity of the mesopause (Kvifte, 1961; Sivjee, 1992; Beig et al. 2003; Beig 2006; Beig et al. 2008; Beig 2011). The emission is centred about 87 km altitude and the rotational temperatures derived are representative of the kinetic temperatures, weighted by the shape and width of the layer (~8 km full-width at half-maximum (FWHM)). Temperatures thus obtained have always been considered ambiguous to the extent that they are dependent on the altitude of the emitting layer, and they are weighted by the altitude profile of that layer. In the case of the OH* layer, different vibrational bands are known to be weighted towards different altitude layers (von Savigny et al. 2012), and on short time scales, individual bands vary in altitude with diurnal, semi-diurnal, annual, semi-annual and solar cycle variations (García-Comas et al., 2017; Liu and Shepherd, 2006; Mulligan et al., 2009). Over long timescales (more than one solar cycle) however, recent studies using satellite data (Gao et al., 2016; von Savigny, 2015) and OH Chemistry-Dynamics (OHCD) models have shown that, the OH* layer altitude is remarkably insensitive to changes in CO_2 concentration or solar cycle variation. This makes these measurements very valuable for monitoring long term changes in the atmosphere.

This work provides an update on the solar cycle and long term trend analysis of the OH rotational temperature measurements taken through each winter season at Davis Research Station, Antarctica (68° S, 78° E). The dataset used here extends for 24 consecutive years and this analysis includes a further 8 years of measurements since the previously published trend assessment using these data (French and Klekociuk, 2011). Here we expand on the earlier analysis to provide a more detailed assessment of the solar response, trends and variability in the Davis record in comparison with v4.2 measurements from the Microwave Limb Sounder (MLS) on the Aura satellite (Aura/MLS) and a network of similar ground based observations (coordinated by the Network for Detection of Mesospheric Change, (NDMC), Reisin et al. 2014).

The outline of this paper is as follows. The instrumentation used and the acquired rotational temperature data collection are presented in Sections 2 and 3. Analysis of solar cycle response and the long-term linear trend is undertaken in Section 4 including comparisons with other ground-based observers and satellite measurements. Discussion of the results, summary and conclusions drawn are given in Sections 5 and 6, respectively.

We use the following terminology for the analysed temperature series in this manuscript. From the measured temperatures and their nightly, monthly, seasonal or winter means, *temperature anomalies* are produced by subtracting the climatological mean or monthly mean (we fit solar cycle and linear trend to the anomalies), *residual temperatures* additionally have the solar cycle component subtracted (used in discussion of long-term trends) and *detrended temperatures* additionally have the long term linear trend subtracted (used in discussion about remaining variability).

2. Instrumentation

A SPEX Industries Czerny-Turner grating spectrometer of 1.26 m focal length has been used to autonomously scan the OH(6-2) P-branch emission spectra (λ839-851 nm) at Davis (68.6° S, 78.0° E) each winter season over the last 24 years (1995-2018). Night-time observations (sun > 8° below the horizon) are only possible between mid-February (~day 048) and end of October (~day 300) at the latitude of Davis.

The spectrometer views the sky in the zenith with a 5.3° field-of-view and an instrument resolution of ~0.16 nm, sufficient to separate P_1 and P_2 branch lines but not to resolve their Lambda-doubling components. Observations are made regardless of cloud or moon conditions and take of the order of 7 minutes to acquire a complete spectrum.

Spectral response calibration has been maintained by reference to several tungsten filament Low Brightness Source units (a total of 4164 scans over the 24 years at Davis) which are in turn cross referenced to national standard lamps at the Australian National Measurement Institute (a total of 781 cross reference calibrations over 24 years). The response correction accounts mainly for the fall-off in response of the cooled gallium arsenide (GaAs) photomultiplier detector and amounts to 8.5% between the $P_1(2)$ and $P_1(5)$ of the OH(6-2) band. The total change in spectral response correction over 24 years is less than 0.3% (equates to less than 0.3 K for the $P_1(2) / P_1(5)$ ratio) despite changing the diffraction grating in 2006 and four changes of the GaAs photomultiplier detector which are carefully characterised over the years. The assigned annual calibration uncertainty is generally <0.3 K except for 1995 (1.8 K) due to calibration via a secondary calibration lamp and in 2002 (1.2 K) due to detector cooling problems. Further details of the instrument are contained in Greet et al. (1997) and French et al. (2000).

3. Davis 24 year rotational temperature dataset

We use the three possible ratios from the $P_1(2)$, $P_1(4)$ and $P_1(5)$ emission line intensities to derive a weighted mean temperature. Intensity values are interpolated to a common time between consecutive spectra to reduce errors uncertainty associated with the 7 minute acquisition cycle time. The weighting factor is the statistical counting error (based on the error in estimating each line intensity, taken as the square-root of the total number of counts for each line). $P_1(2)$ is corrected for the ~2% contribution by $Q_1(5)$, computed using the final weighted temperature. Line backgrounds are selected to balance the small auroral contribution of the $N_2$1PG and N_2^+ Meinel bands and solar Fraunhofer absorption for spectra acquired under moonlit conditions. Correction factors account for the difference in Lambda-doubling between the P-branch lines determined with knowledge of the instrument line shape from high-resolution scans of a frequency-stabilized laser.

Langhoff et al. (1986) transition probabilities are used to derive rotational temperatures as they are closest to the experimentally measured, temperature independent line ratios determined for the OH(6-2) band using the same instrument in French et al., 2000. Recent work by (Noll S., Winkler, H, Goussev, O, and Proxaufet al., (2020)Noll et al (2020), show that these remain a reasonable choice as the Langhoff et al (1986) coefficients show relatively small errors in the comparison of populations from P- and R-branch lines, as well as those of van der Loo and Groenenboom (2008) and Brooke et al. (2016). (see French et al., 2000). Other published sets (e.g., Mies, 1974; Turnbull and Lowe, 1989; van der Loo and Groenenboom, 2007; Brooke et al., 2016) can change offset the absolute temperatures derived by up to 12 K. While the choice is important for comparisons of absolute temperature between observers, it does , but does not significantlynot affect the trend analysis reported here (as long as the same transition probability set has been used consistently for all years)here as the offset is removed by

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subtracting the climatological mean (trends are derived from temperature *anomalies*). It should be noted however that comparison of absolute temperatures with other observations are significantly affected by different choices of transition probabilities. .

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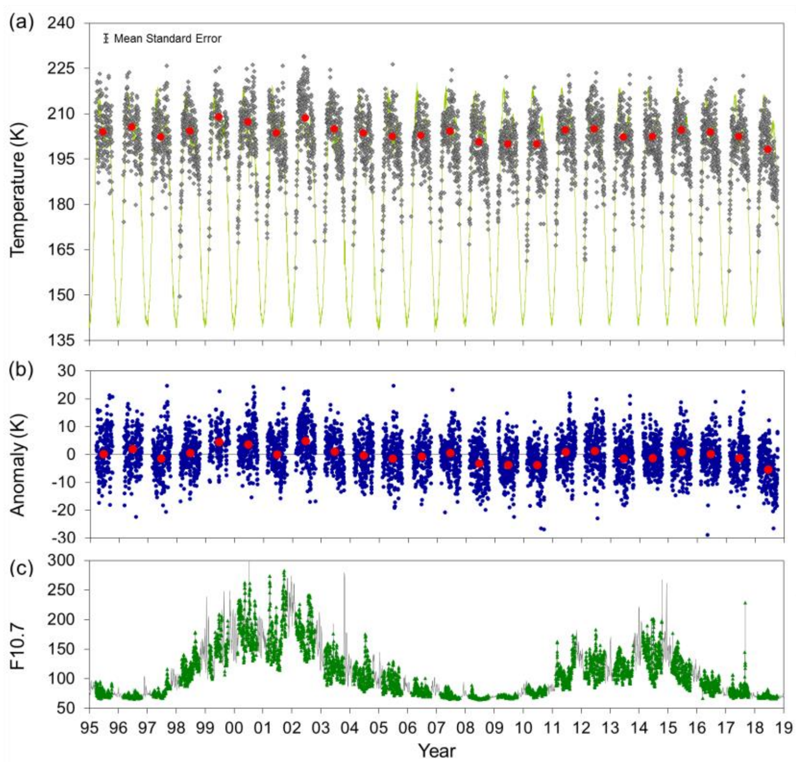
Selection criteria limit extreme values of weighted standard deviation (< 20 K) and counting error (< 15 K), slope (< 0.06 counts/Å), magnitude (< 250 counts per second) and rate of change (< 3 counts per minute) of the backgrounds and the rate of change of branch line intensities (< 6%) between consecutive scansSelection criteria limit extreme values of weighted standard deviation and counting error, slope and magnitude of the background and the rate of change of branch line intensities between consecutive scans. Further details of the rotational temperature analysis procedure are available in Burns et al. (2003) and French and Burns (2004).

Of over 624,000 measurements (typically ~26,000 profiles/year), 403,437 derived temperatures pass the reasonably tight selection criteria (many low signal-to-noise ratio profiles taken through thick cloud or high background profiles around full moon are rejected). These yield 5,309 nightly mean temperatures, where there are at least 10 valid samples that contribute within ±12 hours of local midnight (~1850 Universal Time (UT)). The time series spans two solar cycles (cycles 23 and 24) with peaks in 2001 and 2014. Annual mean temperatures show a dependence on solar activity (see French and Klekociuk (2011) for a comparison of different measures of solar activity with the Davis OH temperature data). We use the 10.7 -cm solar radio flux index (F10.7; 1 solar flux unit (sfu) = 10^{-22} W m^{-2} Hz^{-1}) as our preferred measure of solar activity (F10.7 is fitted and subtracted to examine residual variability). A plot of the nightly and winter mean temperatures with the F10.7 time series used in this work is provided in Fig. 1.



(a)
Mean Standard Error
Temperature (K)
240
225
210
195
180
165
150
135
(b)
Anomaly (K)
30
20
10
0
-10
-20
-30
(c)
F10.7
300
250
200
150
100
50
95 96 97 98 99 00 01 02 03 04 05 06 07 08 09 10 11 12 13 14 15 16 17 18 19
Year

1089

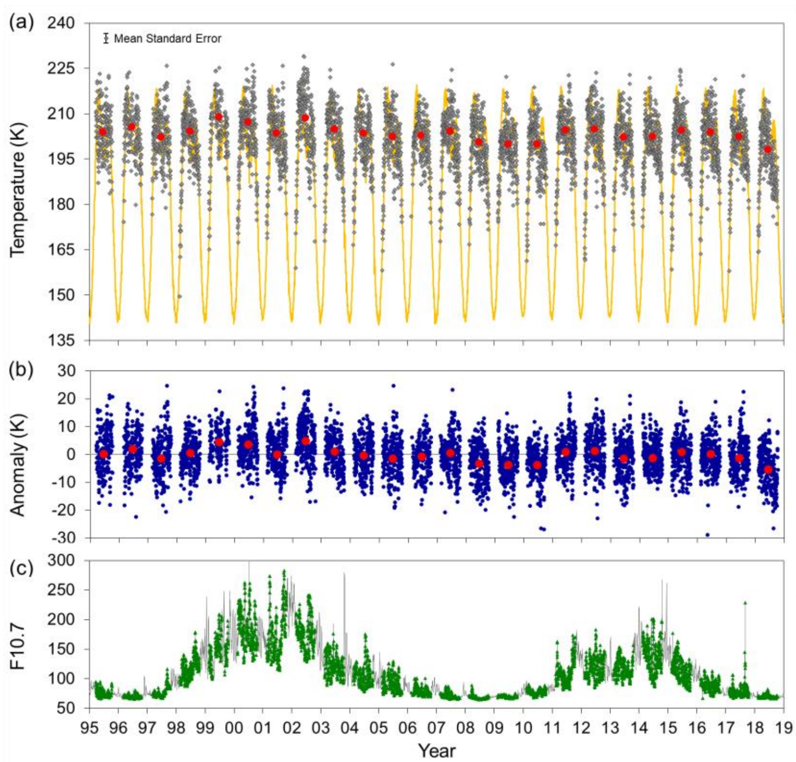
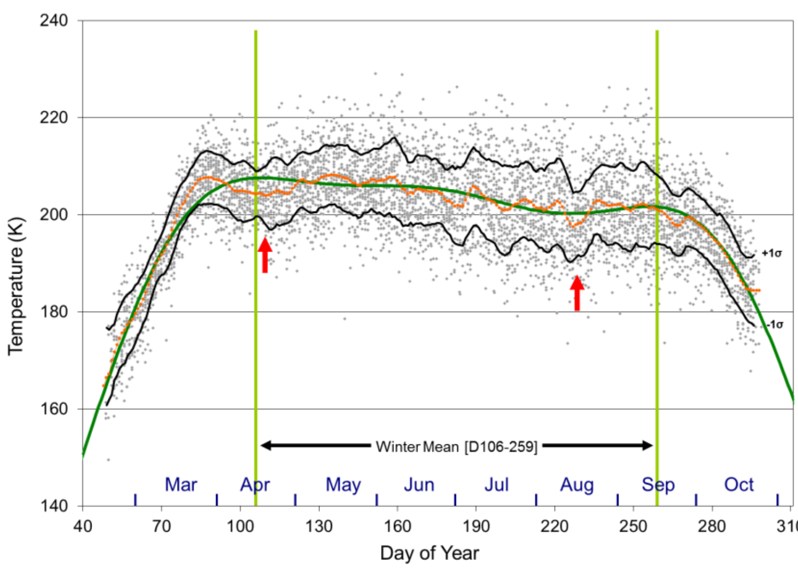


Figure 1 (a). Davis nightly mean temperatures (grey dots; 5309 samples) and winter mean temperatures (D106-259; red points) plotted over the MSISE90 NRL-MSISE00 model temperature for 68°S (87 km altitude, local midnight values) for seasonal reference ((Picone et al., 2002)Hedin, 1991Picone et al, 2002; gold line). (b). nightly mean and winter mean temperature anomalies derived by subtracting the climatological mean (see text) and (c). Daily mean F10.7 cm solar flux index (green points correspond to Davis OH temperature samples over the grey line which are all daily observations)

A climatological mean is derived from a fit to the superposition of nightly mean temperatures for all annual series (Fig. 2). The climatological mean is characterised by a rapid autumn transition (February-March) increasing at 1.2 K/day until a turn-over about 29 March (day of year D088), a slow winter decline (April-September) of -0.4 K/day that is punctuated by mid-April (~D113) and mid-August (~D227) dips corresponding to reversals in the mean meridional flow (Murphy et al., 2007), followed by a rapid spring transition (October-November) of -1.0 K/day. Figure 2 shows the superposed nightly

~~means for each year and the climatological mean fit.~~ Subtracting the climatological mean produces 5309 nightly mean temperature anomalies. Winter mean temperatures are calculated over the interval from 15 April (D106) to 15 September (D259) which avoids the winter to summer transition intervals and lower numbers of nightly observations due to the shorter night length in March and October. A seasonal fit (annual amplitude 41.9 K, semi-annual 23.0 K, ter-annual 7.5 K; green line) and the NRL-MSISE00 reference atmosphere (87 km altitude, local midnight values for Davis; gold points) are also added to Fig. 2 for reference and comparison. The model is limited in its representation of the seasonal cycle as only annual and semi-annual terms are modeled.



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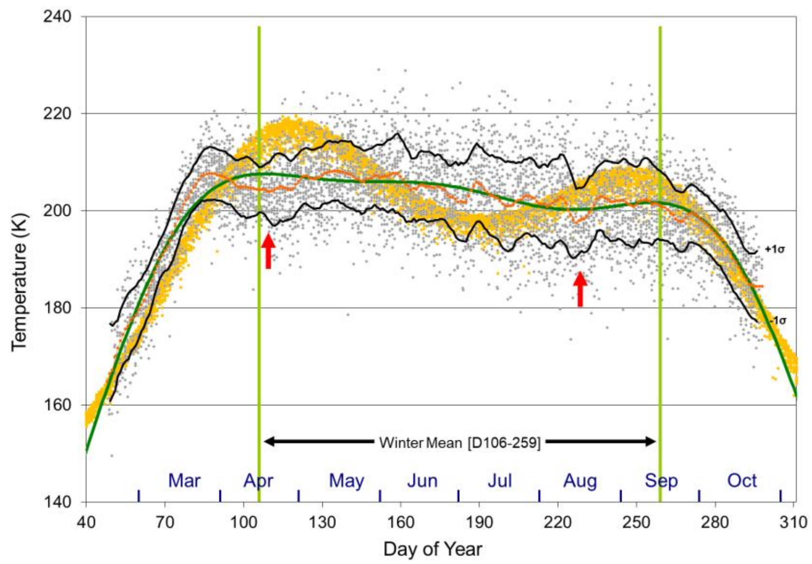

Figure 2. Superposed nightly mean temperatures from 1995 to 2018 [~~gray~~grey points] and a 5-day running mean which represents the climatological mean [orange line] with 1σ intervals [black lines]. The seasonal variation [~~green~~ annual, semi-annual, ter-annual fit; green line] ~~is characterised by a rapid autumn transition (Feb-Mar) increasing at 1.2 K/day until a turn-over about 29th March (day 088), a slow winter decline (Apr-Sep) of -0.4 K/day, punctuated by~~and mid-April and mid-August dips [~~indicated by~~ red arrows] are also indicated~~, followed by a rapid spring transition (Oct-Nov) of -1.0 K/day~~. Green vertical lines mark the calculation region for winter mean temperatures (~~outside~~ avoiding spring and autumn~~the winter to summer~~ transition intervals). The NRL-MSISE00 reference atmosphere (local midnight values for Davis) is also added for comparison [gold points]

4. Trend Assessment

4.1 Davis winter mean trends

Winter mean temperature anomalies over the 24 years of observations are plotted in Fig. 3a. The time series is fitted with a linear model containing a solar cycle term (F10.7) and long term linear trend. This model yields a solar cycle response coefficient (*S*) of 4.30 ± 1.02 K/100sfu (95% confidence limits 2.2 K/100sfu < *S* < 6.4 K/100sfu) and a long term linear trend (*L*) of -1.20 ± 0.51 K/decade (95% confidence limits -0.14 K/decade < *L* < -2.26 K/decade) and accounts for 58% of the temperature variability.

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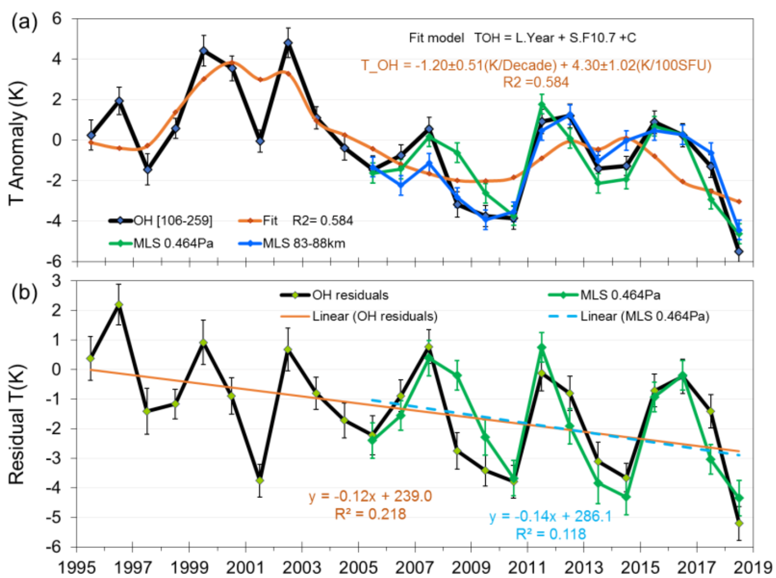

Figure 3 (a). Winter mean (D106-259) temperature anomalies (black line) for Davis station (68°S, 78°E) fitted with a linear model containing a solar cycle term (F10.7cm flux) and long term linear trend (orange line). Fit coefficients are 4.30±1.02 K/100sfu (95% confidence limits 2.18 to 6.42 K/100sfu) and -1.20±0.51 K/decade (95% confidence limits -0.14 to -2.26 K/decade) respectively and account for 58% of the temperature variability. Also plotted (from 2005) are Aura/MLS temperature anomalies derived from the AMJJAS means of all satellite observations within 500 km of Davis station. (b) As for (a), but with the solar cycle component removed to better reveal the long term trend and quasi-quadrennial oscillation (QQO). OH residuals (black line) are compared with Aura/MLS temperature residuals at the 0.00464 hPa level, corrected with the same solar cycle component as used for the Davis OH measurements.

The stability of trend coefficients were tested for the presence of sampling gaps in the OH temperature record. With the exception of 1999 when 2 intervals D095-126 and 213-249 were used to scan the OH(8-3) band and 1996 missing D176-202 all other years only have more than 85% nights within the winter averaging window sampled. (ie 85% of the nights have a valid nightly average temperature with at least 10 measurements that pass selection criteria). A sample bias could be introduced in computing the anomalies if there was a significant departure from the climatological mean in those intervals. The test examined the effect on the derived coefficients by omitting individual years sequentially from the model fit computation. These show the range of *L* and *S* coefficients if a data gap for the entire winter interval was missing in a particular year. All coefficients derived from the omitted year computations remained within the uncertainty limits of the solar cycle and long-term trend coefficients when all years were included.

We ~~note~~ report that a new record low winter-mean temperature of 198.3 K was set for the Davis measurements in 2018, ~~with a value of 198.3 K,~~ which is 1.7 K below the previous minimum recorded in 2009 (200.0 K). This is not entirely due to the low solar activity in 2018 (winter mean flux of 70.4 sfu) as both 2008 (66.9 sfu) and 2009 (69.1 sfu) had lower mean flux and comparable years 1996 (70.6 sfu) was 7.4 K warmer (205.7 K) and 2007 (71.9 sfu) was 6.1 K warmer (204.4 K).

Extracting the solar cycle contribution from the time series yields the long term linear trend and residual variability plotted in Fig. 3b. It is apparent from this plot that a significant oscillation on an approximately 4-year (quasi-quadrennial) timescale remains. A least-squares fit of a sinusoid~~al~~ function to the data yields a period of 4.2 years and peak-peak amplitude of ~3 K. This feature will be examined in detail in Part 2 of this work (French~~, W. J. R., Klekociuk, A. R., Mulligan~~ et al., ~~2019~~2020).

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Distributions of the nightly mean residual temperatures for each year are shown for comparison in Fig. 4. Histogram colour scale indicates the winter mean temperature from warmest year (1999; red) to coldest year (2018; blue). Distributions vary between years from sharp normal distributions (e.g., 1998, 2007, 2016), to broad flat distributions (e.g., 1996, 1997), to skewed or double peaked distributions (e.g., 2004, 2012, 2014, 2018). These differences can be attributed to the variability in large scale planetary wave activity from year to year (French and Klekociuk, 2011)

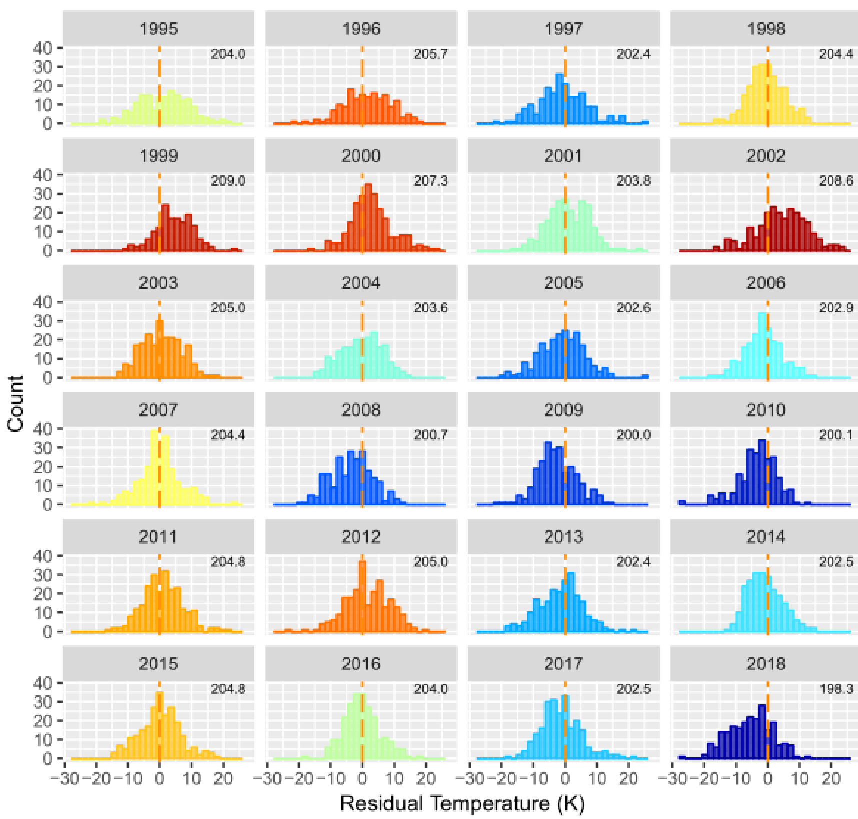

Figure 4. Histograms of nightly mean residual temperatures showing the distribution about the mean winter temperature (annotated in top right corner) coloured from red (warmest year: 1999) to blue (coldest year: 2018).

4.2 Seasonal variability in trends.

Seasonal trend coefficients are ~~also somewhat variable~~examined using a 60 day sliding window, and also from monthly average anomalies. Figure 5 shows the seasonal variability in solar cycle and long-term trend coefficients ~~derived using a 60 day sliding window, and as monthly trends~~as~~,~~ compared to the winter mean trends (D106-259, 154 day mean; red lines) derived for Fig. 3. Seasonal solar response shows a maximum in May-June (~5 K/100sfu) and minimum around August (~2 K/100sfu). Note that April and August temperatures are affected by the characteristic dips seen in the climatological mean during these months (see Fig. 2). Linear trend coefficients show maximum cooling responses in April-May (~ -1.3 K/decade) and in August-October (~ -2.5 K/decade). Virtually no long-term cooling trend is apparent for the midwinter months of June-July.

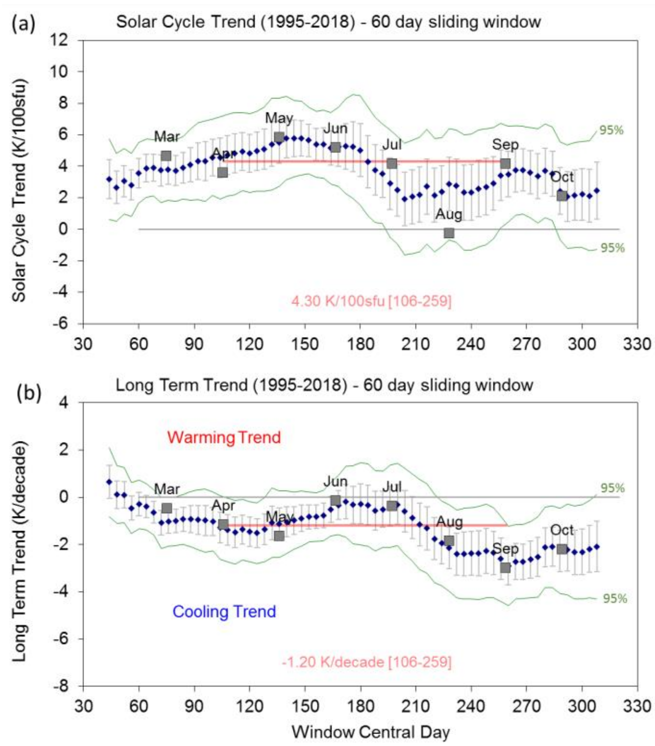

Figure 5. The seasonal variability in (a) solar cycle and (b) long-term trend coefficients derived using a 60 day sliding window (blue dots), and as monthly trends (grey boxes) compared to the winter mean trends (red lines) derived for Fig. 3. The green lines show the confidence limits (95%) for the trend coefficients.

4.3 Aura/MLS trend comparison

For comparison with the Davis trend measurements, we use version v4.2 level 2 data from the Microwave Limb Sounder (MLS) instrument on the Earth Observing System Aura satellite launched in July 2004 (Schwartz et al., 2008). Aura/MLS provides almost complete global coverage (82° S-82° N) of limb scanned vertical profiles (~5-100 km) of temperature and geopotential height derived from the thermal microwave emissions near the spectral lines 118 GHz O_2 and 234 GHz $O^{18}O$. Previous comparisons of these data with MLS v2.2 temperatures were conducted by French and Mulligan, 2010.

Over-plotted in Fig. 3a (extending from 2005) are the equivalent Aura/MLS mean temperature anomalies computed by averaging all observations within 500 km of Davis, for months April to September (AMJJAS) over altitudes 83-88 km (blue line, obtained from a linear interpolation of Aura/MLS geopotential height profiles to geometric height in 1 km steps) and at the 0.00464 hPa (native Aura/MLS retrieval) pressure level (green line). The Aura/MLS data were selected according to the quality control recommendations described in ~~(~~Livesey~~, Nathaniel J., William G. Read, Paul A. Wagner, Lucien Froidevaux~~ et al.~~,~~ (2018)~~Livesey et al. (2018)~~. Approximately 60 samples per ~~/~~month (~2 per day) are coincident within this range. We see very close agreement to both the pressure and interpolated altitude coordinates considering that ~~-~~at these altitudes the vertical resolution (FWHM of the averaging kernel) of Aura/MLS in approximately 15 km (Schwartz et al., 2008), compared to the ~8km FWHM integration of the hydroxyl layer temperatures. The Aura/MLS measurements closely follow the solar response, the long-term linear trend and the magnitude and period of the quasi-quadrennial oscillation (QQO)~~ and the underlying long-term linear trend~~.

We prefer the use of Aura/MLS pressure level data for the comparison with OH temperatures since it is the concentration (density) of reacting species that governs the hydroxyl layer position (primarily collisional quenching with O_2 and N_2 on the bottom-side of the layer, and reaction with atomic oxygen on the top-side of the layer; eg Xu et al., 2012). Statistically, (-from a chi-squared fit to the anomalies) the closest agreement is with the 0.00464 hPa pressure level and this is over-plotted on Fig. 3b corrected ~~for~~ using the same solar cycle response that was determined from the Davis OH measurements. The linear long-term trend fit for Aura/MLS over 14 years is -1.43 ± 1.1 K/decade which compares very well ~~(comparable~~ to ~~the~~ -1.2 ± 0.51 K/decade for the 24 years of Davis OH measurements, considering the seasonal variability shown above (section 4.2) and the ~~) but~~

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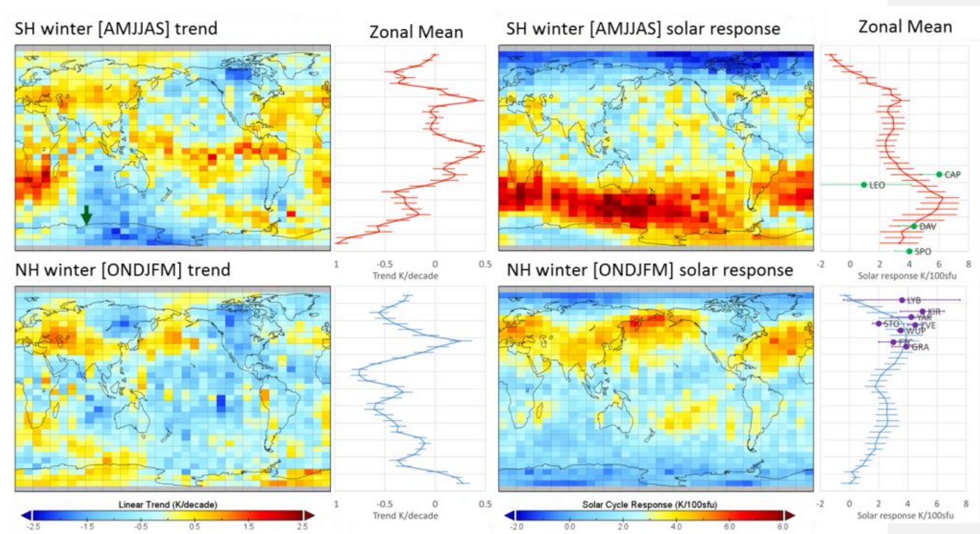
~~clearly the~~ underlying QQO ~~variability~~residual evident in both series, which has a significant effect on the fit over the different data spans ~~has a significant effect on the fit~~.~~.~~

~~It is important to note that the winter mean residual trend coefficients in Fig. 3b are derived as a mean across 6 months of significantly varying solar and long-term responses. Nevertheless, the residual QQO signature remains readily apparent in the 60-day sliding window means through April to July [AMJJ] although somewhat breaking down in August to October [ASO].~~

We examine the QQO feature in ~~more~~ greater detail in the second part of this work (French, ~~W. J. R., Klekociuk, A. R., Mulligan~~et al., ~~2019~~2020), but here, given the close agreement of Davis and Aura/MLS 0.00464 hPa trends in Fig. 3b, we apply the same model fit procedure to derive Aura/MLS solar cycle and linear long-term trend coefficients to obtain a global picture of trends at the hydroxyl layer equivalent pressure level (0.00464 hPa). Figure 6 shows global trends determined by averaging Aura/MLS pressure level 0.00464 hPa temperature anomalies into a 5° x 10° (latitude x longitude) grid, over Southern Hemisphere (SH) winter months (April-September; AMJJAS; panel a) trend; panel c) solar response~~; top panels~~) compared to Northern Hemisphere (NH) winter months (October-March; ONDJFM; panel b) trend; panel d) solar response~~; bottom panels~~). Each grid box has been corrected for the solar cycle response determined from a linear regression of temperature to F10.7 over the 14 years of Aura/MLS measurements. The long-term linear trend (~~left-hand panels~~panels a) and b)) and solar cycle response (~~right-hand panels~~panels c) and d)), for each grid box, together with their corresponding zonal means are presented. The maps contain some interesting features; enhanced bands of solar activity response occur at mid-latitudes in both winter hemispheres although strongest in the SH (colour scales are the same for each hemisphere). Minima in sensitivity to solar forcing occur over the equator and the poles. Long-term trends over the Aura/MLS era are not globally uniform. While the global mean trend for the SH winter [AMJJAS] is -0.31 K/decade, there are regions of warming, notably around the equator, southern Africa, Europe and the Atlantic ocean and strongest cooling over Antarctica and northern Canada. For the NH winter [ONDJFM] the global mean is -0.11 K/decade with generally global cooling, except for warming over Antarctica, Europe, southern Africa and the northern Pacific Ocean.



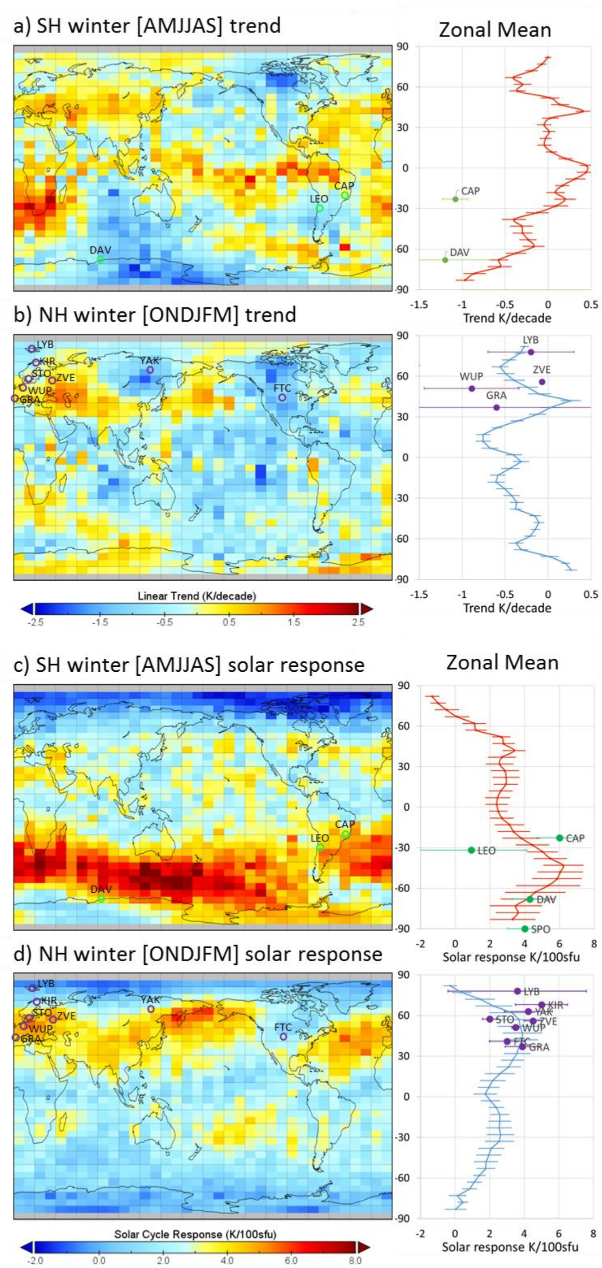


Figure 6. Global temperature trends (a. & b.) and solar cycle responses (c. & d.), together with -their corresponding zonal means determined from 14 years of MLS v4.2 pressure level 0.0046hPa (hydroxyl layer equivalent), averaged into 5° latitude x 10° longitude grid,

and over southern hemisphere winter months (AMJJAS; ~~top~~ panels a. & c.) compared to northern hemisphere winter months (ONDJFM; ~~bottom~~ panels b & d). The linear trend and solar cycle response coefficients have been derived individually for each grid box from Aura/MLS over 14 years with no lag. ~~The green arrow in panel 1~~Station locations indicate~~s~~ the Aura/MLS comparison with ~~Davis shown in Fig 1B~~ground based observations of long-term trend and solar response given in Table 1. ~~Solar response coefficients from other observers are indicated on the zonal solar response plots (see text for site information)~~

4.4 Trend comparisons with other ground based observations

It is useful to compare these Aura/MLS derived solar response and trend coefficients with other observations, carefully bearing in mind that these observations may span different time intervals than available in the Aura/MLS measurement epoch. At Davis the solar cycle response (indicated by the green ~~dot~~ point and label DAV in Fig. 6 c.) determined over 24 years matches well with the zonal mean at 68° S determined from the Aura/MLS measurements. Davis appears to be on the poleward boundary of the strong band of solar sensitivity (~40-70° S) in the SH winter. The long-term trend at Davis is marked by the green ~~arrow~~ point and label DAV on ~~the left-hand upper~~ panel a) in Fig 6, and as we have seen from Fig. 3, agrees well with Aura/MLS.

Table 1 summarises the data-span, derived long term trend, and solar cycle coefficients from a collection of ground-based observers. Where new results are available these have been updated from Table 2 in French and Klekociuk (2011) and as compiled in Beig et al. (2008). Solar cycle and long-term trend coefficients from these sites are also marked on Fig 6 where possible. The majority of these observations agree well (within error estimates) with the Aura/MLS zonal mean solar response and long term trends evaluated here, given the different measurement epochs and geographic variability in the trends coefficients shown by Aura/MLS.

~~it is a zonal mean response.~~

Site	Data Span	Trend K/decade	Solar response K/100sfu	Reference
Longyearbyen (LYB, 78°N, 16°E)	1983-2013	-0.2±0.5	3.6±4.0	Holmen et al. (2014)
Kiruna (KIR, 68°N, 21°E)	2003-2014	-2.6±1.5	5.0 ± 1.5	Kim et al. (2017)
Yakutia (YAK, 63°N, 129°E)	1999-2013	Not Significant	4.24±1.39	Ammosov et al. (2014)
Stockholm (STO, 57°N, 12°E)	1991, 1993-1998	Not Determined	2.0±0.4	Espy et al. (2011)
Zvenigorod (ZVE, 56°N, 37°E)	2000-2016	-0.07±0.03	4.5±0.5	Perminov et al. (2018)
Wuppertal (WUP, 51°N, 7°E)	1988-2015	-0.89±0.55	3.5±0.21	Kalicinsky et al. (2016)
Fort Collins (FTC, 41°N, 105°W)	1990-2018	-2.3±0.5	3.0±1.0	Yuan et al. (2019 in press)
Granada (GRA, 37°N,3°W)	2002-2015	-0.6±2.0	3.9±0.1	Garcia-Comas et al. (2017)
Cachoeira Paulista (CAP, 23°S, 45°W)	1987-2000	-1.08±0.15	6±1.3	Clemesha at al. (2005)
El Leoncito (LEO, 32°S, 69°W)	1998-2002	Not Determined	0.92±3.2	Scheer et al. (2005)
Davis (DAV, 68°S,78°E)	1995-2018	-1.20±0.51	4.30±1.02	This Work
South Pole (SPO, 90°S)	1994-2004	0.1±0.2	4.0±1.0	Azeem et al. (2007)

Table 1. A comparison of solar cycle response and temperature trend observations from the ground-based OH observer network with updates since 2011 where available.

As some observers have found, there is a ~~significant~~important question about a time delay in the OH layer temperature response to solar forcing via the various solar absorption mechanisms in the atmosphere. The major absorbers and altitude of solar extreme ultraviolet radiation are molecular oxygen (Schumann-Runge continuum, 80-130 km, Schumann-Runge electronic and vibrational bands, 40-95 km, Herzberg continuum, below 50 km) and ozone (Hartley-Huggins bands, below 50 km).

We have previously found a lag of around 160 days (F10.7 leads temperature) is best fit to the linear model (French and Klekociuk, 2011), others find shorter: 80 days at Longyearbyen, Svalbard (Holmen et al. 2014), or larger lags: 25 months at Maimaga station, Yakutia (Ammosov et al. 2014; Reisin et al. 2014). Recalculating the long term trends for Aura/MLS assuming a uniform global solar response (as for Davis), or with a 160 day lag and zonal mean solar response ~~(see supplementary material)~~ does not significantly change the warming and cooling patterns shown in Fig. 2, but the lag does reduce the cooling trend (on average by 0.16 K/decade for the southern hemisphere (SH) winter and 0.11 K/decade for the northern hemisphere (NH) winter) and increases the fit error.

Beig (2011a, 2011b) in their reviews of long-term trends in the temperature of the mesosphere and lower thermosphere (MLT), highlight the difficulty of distinguishing between the anthropogenic and solar cycle influences. In their results, mesopause region temperature trends were found to be either slightly negative or zero. At that time, it was believed that the solar response becomes stronger with increasing latitude in the mesosphere with typical values in the range of a few degrees per 100 solar flux units in the lower part of the mesosphere but reaching 4-5 K/100 sfu near the mesopause. More recent studies using longer data sets (Ammosov et al. 2014; Holmen et al. 2014; Perminov et al. 2018) and satellite data (Tang et al. 2016) have reinforced that view.

Trend breaks began to appear in mesopause region temperatures in 2006 (Offermann et al., 2006, 2010), and these continue until now in certain locations (e.g., Jacobi et al., 2015; Kalicinsky et al., 2018; Yuan et al., 2019). These can be quite varied from site to site, ranging from -10 K/decade to +5 K/decade. Some of these estimates simply suffer from lack of observations (measurement spans less than a solar cycle). Few are longer than 2 solar cycles, but those of note are included in Table 1. OH temperature trend studies in the southern hemisphere are less common. Reid et al. (2017) report MLT-region nightglow intensities, temperatures and emission heights near Adelaide (35° S, 138° E), Australia. Five years (2001-2006) of spectrometer measurements using OH(6-2) and O_2(0-1) temperature are compared with 2 years of Aura/MLS data and 4.5 years of SABER data. Venturini et al. (2018) report mesopause region temperature variability and its trend in southern Brazil (Santa Maria, 30° S, 54° W), based on SABER data over the period 2003-2014. Nath and Sridharan (2014) examined the response of the middle atmosphere temperature to variations in solar cycle, QBO and ENSO in the altitude range 20-100 km and 10-15° N latitude using monthly averaged zonal mean SABER observations for the years 2002-2012. They found cooling trends in most of the stratosphere and the mesosphere (40–90 km). In the mesosphere, they found the temperature response to the solar cycle to be increasingly positive above 40 km. The temperature response to ENSO was found to be negative in the middle stratosphere and positive in the lower and upper stratosphere, whereas it appeared largely negative in the height range 60–80 km and positive above 80 km.

5. Discussion

5.1 Relationship between Davis trends and CO_2 -and O_3 change.

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Our updated trend assessment over 24 years yields a cooling rate of -1.20±0.51 K/decade for the mean winter [D106-259] temperatures in the hydroxyl layer above Davis. A slightly greater rate of -1.32±0.45 K/decade is derived if the full year [D040-310] of observations are included in the annual means. Over the same period, annual mean surface CO_2 volume mixing ratios (VMRs) increased from 360.82 ppm [1995] to 408.52 ppm [2018] (Mauna Loa values from Global Greenhouse Gas Reference Network www.esrl.noaa.gov/gmd/ccgg/trends/), an increase of 47.7 ppm or 13.2% (19.9 ppm per decade or 5.5% per decade). Qian et al. (2019) quote a CO_2 trend figure of 5.2%/decade (or 5.1 % if the seasonal variation is removed before the linear trend calculated) based on measurements made by TIMED/SABER from 2002-2015. If the primary factor for the observed temperature trend is considered to be CO_2 radiative cooling, a coefficient of -0.06 K/ppmCO_2 or -0.22 K/%CO_2 is implied. This is approximately twice the value obtained by (Huang, 2018) (her Figure 2) who employed a linear scaling of the result of a doubling of CO_2 concentration by (Roble and Dickinson, 1989). A CO_2 increase of 26.5% from 1960 to 2015 was accompanied by a temperature decrease of 1.4% at an altitude of 89.4 km near Salt Lake city, Utah (18° N, 290° E).

CO_2 is well mixed through the lower atmosphere with a constant VMR up to about 80 km. Above this height, diffusion and photolysis processes begin to have an effect, reducing the VMR (Garcia et al., 2014) but these processes vary with latitude and season (Rezac et al. 2015; López-Puertas et al., 2017).

~~Several studies of CO_2 VMR using profiles from the Atmosphere Chemistry Experiment Fourier Transform Spectrometer (ACE-FTS) and Sounding of the Atmosphere using Broadband Emission Radiometry (SABER) satellite instruments, reported~~

considerably larger rates of change of CO_2 in the upper atmosphere, increasing from about 5% per decade at 80 km to 12% per decade at 110 km (Emmert et al., 2012; Garcia et al., 2016; Yue et al., 2015). However, more recent analysis of the ACE-FTS and SABER CO_2 data with different deseasonalizing procedures have shown an average rate of 5.5% per decade in the 80-110 km region, consistent with surface rates (Qian et al., 2019; Rezac et al., 2018).

In a recent summary of progress in trends in the upper atmosphere, ((Laštovička and Jan, (2017)Laštovička (2017) identified greenhouse gases, particularly CO_2 as the primary driver of long-term trends there. The overall effect of greenhouse gases at mesospheric altitudes is radiative cooling. The important secondary trend drivers in the mesosphere and lower thermosphere (MLT) are stratospheric ozone, water vapour concentration and atmospheric dynamics. Temperature trends are predominantly negative, and recent progress in understanding the magnitude of the cooling have arisen from confirmation and quantification of the role of ozone. Lübken et al. (2013) present the results of trend studies in the mesosphere in the period 1961-2009 from the Leibniz-Institute Middle Atmosphere (LIMA) chemistry-transport model which is driven with European Centre for Medium–Range Weather Forecasts (ECMWF) reanalysis below 40 km, and observed variations of CO_2 and O_3. They find that CO_2 is the main driver of temperature change in the mesosphere, with O_3 contributing approximately one third to the trend. Linear temperature trends were found to vary substantially depending on the time period chosen primarily due to the influence of the complicated temporal variation of ozone. Figure 3 of (Lübken et al., 2013) show a monotonically increasing trend in CO_2 compared with a much more complicated temporal ozone variation (essentially constant until 1980, a rapid decrease from 1980-1995, followed by an increase since then. Trends in ozone vary as a function of both altitude and latitude, with positive trends

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dominating in the lower stratosphere and mesosphere (Laštovička, 2017). Increases in water vapour concentration are considered a secondary but non-negligible effect particularly in the lower thermosphere (Akmaev et al., 2006)~~(Akmaev et al. 2006)~~. The trend effect of dynamics was found to be very slightly negative in the mesosphere, but very small compared with the radiatively induced trends. At the mesopause, the trend due to dynamics was positive and significantly larger (~1 K/decade). These results were found to be in good agreement with observations from lidars, Stratospheric Sounding Units (SSU) (Randall et al., 2009) and radio reflection heights which have decreased by more than 1 km in the last 50 years due to shrinking in the stratosphere/lower mesosphere caused by cooling~~In a recent summary of progress in trends in the upper atmosphere, Laštovička (2017) identified greenhouse gases, particularly CO_2 as the primary driver of long-term trends there. The important secondary trend drivers in the mesosphere and lower thermosphere (MLT) are stratospheric ozone, water vapour concentration and atmospheric dynamics. The overall effect of greenhouse gases at mesospheric altitudes is radiative cooling. Temperature trends are predominantly negative, and recent progress in understanding the magnitude of the cooling have arisen from confirmation and quantification of the role of ozone. In the mesopause region, about two thirds of the cooling is attributed to increases in CO_2 concentration and one third to changing concentration of ozone in the stratosphere (Lübken et al., 2013). Increases in water vapour concentration are considered a secondary but non-negligible effect particularly in the lower thermosphere (Akmaev et al. 2006). Trends in ozone vary as a function of both altitude and latitude, with positive trends dominating in the lower stratosphere and mesosphere.~~

~~Huang (2018) examined the influence of CO_2 increase, solar cycle variation and geomagnetic activity on airglow from 1960 to 2015 using two airglow chemistry~~

dynamics models (OHCD – OH chemistry dynamics, and MACD – multiple airglow chemistry dynamics). As expected, the results showed that airglow intensity and peak volume emission rate (VER) are in phase and have a linear relationship with F10.7 values, whereas CO_2 increase leads to a slowly decreasing trend in OH(8-3) airglow intensity. OH(8-3) peak altitudes of the VER are unaffected by increases in CO_2 concentration, and are only slightly affected by the F10.7 cycle, with slightly lower peak altitudes when F10.7 is <100 SFU. Surprisingly, OH VER peak heights showed a significant inverse relationship with geomagnetic activity as measured by the Ap index. We find no significant correlation of the *T* residual from Davis with the Ap index for the months of AMJJAS.

Lübken et al. (2013) present the results of trend studies in the mesosphere in the period 1961-2009 from the Leibniz-Institute Middle Atmosphere (LIMA) chemistry-transport model which is driven with European Centre for Medium- Range Weather Forecasts (ECMWF) reanalysis below 40 km, and observed variations of CO_2 and O_3. They find that CO_2 is the main driver of temperature change in the mesosphere, with O_3 contributing approximately one third to the trend. Linear temperature trends were found to vary substantially depending on the time period chosen primarily due to the influence of the complicated temporal variation of ozone. The trend effect of dynamics was found to be very slightly negative in the mesosphere, but very small compared with the radiatively induced trends. At the mesopause, the trend due to dynamics was positive and significantly larger (~1 K/decade). These results were found to be in good agreement with observations from lidars, Stratospheric Sounding Units (SSU) (Randall et al., 2009) and radio reflection heights which have decreased by more than 1 km in the last 50 years due to shrinking in the stratosphere/lower mesosphere caused by cooling. Figure 3 of (Lübken et al., 2013) show a monotonically increasing trend on CO_2 compared with a much more complicated temporal ozone variation (essentially constant until 1980, a rapid decrease from 1980-1995, followed by an increase since then.

A recent paper by Hervig et al. (2019) report on the absence of a solar signal correlated response in polar mesospheric clouds (PMCs) in the summer mesopause following 2002. PMCs are controlled by temperature and water vapour. At solar maximum, temperatures are expected to be higher and water vapour lower, thereby leading to less PMCs at solar maximum. This anti-correlation was evident in satellite data until 2002, but has been absent since then. The main cause for the diminished solar cycle in PMCs at 68° N and 68° S appears to be the dramatic suppression of the solar cycle response in water vapour. The solar cycle response of temperature also decreases after 2002, but has a much lower effect on PMCs than the water vapour.

The Whole Atmosphere Community Climate Model (WACCM) extended into thermosphere (upper boundary ~700 km) (WACCM-X) was used by Qian et al. (2019) (with the lower atmosphere constrained by reanalysis data) to investigate temperature trends and the effect of solar irradiance on temperature trends on the mesosphere during the period 1980-2014. The overall temperature trend in the mesopause region at 85 km was statistically insignificant at -0.46 ± 0.60 K/decade. Solar irradiance effects on the global average temperature are positive and decrease monotonically with decreasing altitude from a value of ~3 K/100 sfu in the lower thermosphere to ~1 K/100 SFU at 55 km. This is readily explained by the decreasing external energy from the Sun with reducing altitude. A monthly mean global average trend of 2.46 K/100 sfu is quoted for the mesopause near 85 km. The mesosphere is affected by solar irradiance directly from local heating through absorption of radiation, and indirectly through dynamics by its effects on the geostrophic winds which control the upward propagation of gravity waves and planetary waves generated in the troposphere. Zonal mean temperatures show significant variability as a function of altitude, latitude and season. Qian et al. (2019) provide ~~globally~~ zonal averaged temperature trend values as a function of altitude (50-110 km) and latitude for each month (their Fig. 3) some of which are statistically significant. Solar cycle effects on temperature are in reasonable agreement with the ~~OH(6-2) temperatures~~Davis coefficients (shown in Fig~~ure~~ 5(a)) with positive values ranging from ~3-5 K/100 sfu, the largest values occurring in July and October (compare Qian et al. 2019 Fig 4.). The long-term trend is predominantly negative with values in the range -1 to -3 K/decade with the largest cooling occurring in March and September at the latitude and altitude of the OH temperatures measured at Davis Station. WACCM-X shows slightly positive trend values in the months of February, November and December at Davis Station, but OH(6-2) temperature data are not available in these months. The September maximum in cooling is in reasonable agreement with the Davis measurements shown in Figure 5 of this work.

More recent results from Garcia et al. (2019) using WACCMv4 free-running (coupled ocean) simulations for the period 1955-2100 using IPCC RCP 6.0 attribute the changes in the trends of the temperature profile to monotonic increases in CO_2 concentration together with a decrease in O_3 until 1995 followed by subsequent increase. Garcia et al. (2019) assign half of the stratopause negative temperature trend to ozone depleting substances. At the mesopause, the global mean trend in temperature is approximately -0.6 K/decade. Solar cycle signals at the mesopause are in the range 2-3 K/100 sfu with slightly higher values in the southern polar cap. Very large seasonal trends in temperature at all altitudes are associated with the development of the Antarctic ozone hole. Trends are largest in the November-December period, and teleconnections are made with the upper mesosphere via GW filtering by the zonal wind anomaly in the southern polar cap.

5.2 Trend breaks.

When analysing long-term trends, several authors (Lübken et al., 2013; Qian et al., 2019) emphasise the importance of specifying the length of the time period, as well as the beginning and end of the period, because trend drivers can be different for different periods (e.g., Yuan et al., 2019). Yuan et al. (2019) report long-term trends of the nocturnal mesopause temperature and altitude from LIDAR observations at mid-latitude (41-42° N, 105-112° W) in the period 1990-2018. They divided their observations into two categories, the high mesopause (HM) above 97 km during the non-summer months, mainly formed by radiative cooling, and the low mesopause (LM) below 92 km during the non-winter months generated by mostly by adiabatic cooling. This idea of the mesopause at two different altitudes is well established (e.g., von Zahn et al., 1996; Xu et al., 2007; Thulasiraman and Nee, 2002). Although Yuan et al. (2019) obtained a cooling trend of more than 2 K/decade in the mesopause temperature along with a decreasing trend in mesopause height since 1990, the temperature trend is statistically insignificant since 2000.

Trend breaks have been reported at other mid-latitude stations (Offermann et al., 2006, 2010) where a discontinuity was found in the overall trend in the year 2001/2002. Using some of the same data as Offermann et al. (2006), Kalicinsky et al. (2016) reported a trend break in the middle of 2008. Before the break point, there is a clear negative trend reported to be -2.4 ± 0.7 K/decade, whereas after 2008, a large positive trend of 6.4 ± 3.3 K/decade is ~~deciphered~~determined. Two possible explanations are suggested for the trend break: the first is that it is the result of a combination of the solar cycle and a long period oscillation such as the 22-year Hale cycle of the Sun. A second possible explanation of the very substantial change in the trend at 2008 is a combination of the solar flux with a sensitivity of 4.1 ± 0.8 K /100 SFU together with a long period oscillation 24-26 years with an amplitude of about 2K. Kalicinsky et al. (2018) find support for this idea in the identification of a quasi-decadal oscillation in the summer mesopause over Western Europe in plasma scale height observations (near 80 km altitude) which are in anti-correlation with the potential oscillation in temperature from OH* measurements. The anti-correlation in the two data sets is explained on the basis of the fact that they originate below (plasma scale height data) and above (OH* temperature data) the temperature minimum in the mesopause region in summer. Jacobi et al. (2015) find that the long-term behavior of both meridional and zonal winds at 90-95 km in northern mid-latitude stations exhibit trend breaks in summer near 1999, although the winter data are well described by a single linear trend over the years 1980- 2015. We find no obvious sign of a discontinuity in the trend obtained in the Davis data from 1995-2018. There is no significant change in the long-term trend or solar response when extending the period of study from 16 years (2005-2010 coefficients 4.79 ± 1.02 K/100sfu and -1.18 ± 0.87 K/decade; French and Klekociuk, 2011) to 24 years (this work). Neither coefficient has changed outside the uncertainty.

5.3 Effect of changes in the OH*-layer height

There is widespread acceptance that cooling of the middle atmosphere due to increases in CO_2 concentration has resulted in shrinking of the middle atmosphere (e.g., (Grygalashvyly et al., 2014; Sonnemann et al., 2015). This does raise the question however of whether the OH* layer is fixed to a constant pressure level rather than a constant altitude. There are mixed reports on this topic. In a long-term study of the effects of chemistry, greenhouse gases, and the solar modulation on OH* layer trends using the Leibniz Institute Middle Atmosphere (LIMA) chemistry-transport model covering the period 1969 to 2009, Grygalashvyly et al. (2014) reported a downward shift in the OH*-layer by about 0.3 km/decade in all seasons due to shrinking of the middle atmosphere resulting from radiative cooling by increasing CO_2 concentrations. Wüst et al. (2017) report a descent in the mean altitude of the OH* layer of 0.02 km/ year from 14 years of SABER data (2002-2015) in the alpine region of southern Europe (44–48° N, 6–12° E). They refer to a paper by Bremer and Peters (2008) which reports low frequency reflection heights (ca. 80-83 km) between 1959 and 2006 and derive a figure of 0.032 km/year.

Sivakandan et al. (2016) have published a long-term variation paper on OH peak emission altitude and volume emission rate over Indian low latitudes using SABER data. A weak decreasing trend of 19.56 m/year was reported for the peak emission altitude of the night-time OH*-layer. (García-Comas et al., (2017)García-Comas et al. (2017) reported a slightly larger decrease of 40 m/decade in SABER OH volume emission rate weighted altitude at mid-latitudes which accompanied a 0.7%/decade increase in OH intensity and a 0.6K/decade decrease in OH equivalent temperature.

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A vertical shift of the OH* layer either upward or downward gives rise to a change in the emission weighted temperature which is measured by ground-based optical instruments (French and Mulligan, 2010; Liu and Shepherd, 2006; von Savigny, 2015). Von Savigny (2015) reported no apparent trend or solar cycle in OH emission altitude at the local time of the SCIAMACHY nighttime observations in the period 2003-2011. However, Teiser and von Savigny (2017) found evidence of an 11-year solar cycle in the vertically integrated emission rate and in the centroid emission altitude of both the OH(3-1) and OH(6-2) bands in SCIAMACHY data. Gao et al. (2016) found no evidence that the OH* peak heights are affected by solar cycle in 13 years of TIMED/SABER data, and deduced that the solar cycle variation of temperature obtained from ground-based OH nightglow observations were essentially immune from the OH emission altitude variations. Huang (2018) found no systematic response of airglow $O(^1S)$ green line, O_2(0,-1), or OH(8-

3)) VER peak heights with the F10.7 solar cycle using two airglow models OHCD and MACD-90. The Huang (2018) result is supported by Gao et al. (2016) using TIMED/SABER data and by von Savigny (2015) using SCIAMACHY data. These confirmations of the remarkable long-term stability of the peak altitude of the OH*-layer in an atmosphere with increasing CO_2 concentration and changing solar radiation are essential for the use of long-term studies of mesopause region temperatures derived from ground-based OH* optical measurements.

We have examined the altitude variation of the OH* layer over Davis during the period 2002-2018 using the OH-B channel volume mission rate (VER) from TIMED/SABER (version 2.0) sensitive in the wavelength range 1.56-1.72 µm, which includes mostly the OH(4-2) and OH(5-3) bands. All VER altitude profiles between day 105 and day 259 that satisfied the selection criteria (tangent point within 500 km of Davis and solar zenith angle > 97°), employed by French and Mulligan (2010) were used to determine the altitude of the layer. The altitude of the peak was obtained from a Gaussian profile fitted to the VER profile (for more details, see French and Mulligan, 2010). The slope of the best fit line to the winter annual average peak altitude was -0.02 ± 0.02 km/ year as shown in Figure 7, i.e., no significant change in altitude of the layer over the period in agreement with the result of Gao et al. (2016).

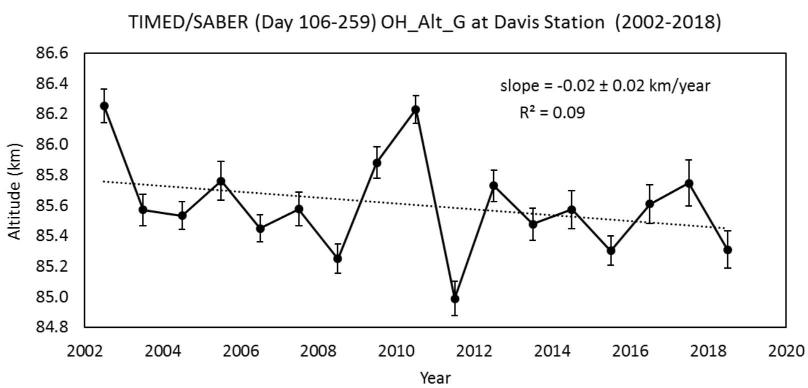

Figure 7. The trend in the mean winter OH layer altitude, derived from TIMED/SABER (version 2.0) OH-B channel volume emission rate. The slope of the best fit line is -0.02 ± 0.02 km/year, i.e., no significant change in altitude of the layer over this interval.

5.4 Global solar cycle and ~~linear~~ long-term trends

The ~~trend~~ long-term trend measured at Davis is well matched with the result from Aura/MLS over 14 years for the southern hemisphere ~~(SH)~~ winter months (AMJJAS) at the 0.00464 hPa level. Clearly though, applying the same analysis to the global temperature field reveals that trends are far from~~not~~ globally uniform (Fig 6) ~~(Fig 6)~~. In the SH winter the most significant cooling trends are seen over the southern polar cap and northern Canada, with warming trends over southern Africa, around the equator and over Europe and Russia. NH winter cooling trends are strongest over eastern Russia and North America, but warming trends remain over Europe.

There are a number of limitations and assumptions made for these derived trends: i) there are only 14 years from which to extract a solar cycle component, ii) a solar cycle component is computed for each grid box. The zonal means calculated are generally within 2 K/100 sfu of other reported solar response coefficients, but there is a strong latitudinal and seasonal dependence (strongest solar flux response in mid-latitude winter hemisphere – near zero response in high latitude summer), iii) we have assumed no lag between solar flux variations and the temperature response, whereas previous work for the Davis response for example indicates a ~160 day lag is optimal at least for Davis (French and Klekociuk, 2011) and iv) for comparison with other hydroxyl temperature long-term trends we assume the global OH layer height is well matched with the Aura/MLS 0.004664 hPa level.

To address uncertainties about the solar response coefficient (item ii above) we have recalculated the global trends assuming a fixed response for each grid box (4.2 K/100 sfu as derived from the Davis observations) and also as zonal means but for a lag of 160 days (F10.7 leads T) as previously found for Davis. These plotsis analysis are available in the supplementary material and determinesshow that, by and large, the warming and cooling patterns observed in Figure 6 do not change significantly for the different solar cycle response componentscomputations.

While the WACCM-X results presented by Qian et al. (2019) are in reasonable agreement with the OH temperature behaviour measured at Davis Station, the zonally averaged pattern of solar cycle response and linear trend obtained from WACCM-X differs considerably from that obtained from an analysis of the Aura/MLS data at the 0.00464 hPa level shown in Figure 6. In the Aura/MLS results, the solar response in both hemispheres in winter show a great deal more variation as a function of latitude than is evident in the WACCM-X results at 87 km (Figure 4 of Qian et al., 2019). The zonally averaged Aura/MLS pattern shows maxima (~6 K/100sfu) in southern mid-latitudes in the Southern Hemisphere (SH)SH winter, and similarly awhile the maximum (although a smaller peak ~4 K/100sfu compared to the SH response) is in northern mid-latitudes in the NHNorthern Hemisphere (NH) winter. The solar cycle response is essentially zero at 82° north and south during the NH winter months, but it is of the order of 3 K/100sfudecade at 82° south in SH winter. The southern hemisphereSH winter months have the largest variation with a pronounced maximum in the latitude range ~10° S to 40° S. (The maximum also shows longitudinal structure with a much broader maximum between 90° east and 90° west which is centred at higher southern latitudes.) Several authors (Perminov et al., 2014; Pertsev and Perminov, 2008) have reported that winter OH* temperatures are more sensitive to the solar flux variation than summer temperatures and t~~. T~~his agrees with the Aura/MLS variation shown here~~in Figure 6.~~

The ~~WACCM-X~~ long term trend modelled by WACCM-X is predominantly negative or zero at the altitude of the OH layer (87 km) at all latitudes and in all months apart from February, November and December, when a positive trend of up to ~3 K/decade is present at high southern latitudes (see Fig 3. in Qian et al., 2019). Aura/MLS results also show a predominantly slight negative trend ~0.5-1 K/decade, except at the equator, and at mid-latitudes in the SH winter months.

Solomon et al. (2018) simulated the anthropogenic global change through the entire atmosphere using WACCM-X in a free-running mode (i.e., lower atmosphere below 50 km not constrained by ECMWF reanalysis data) using constant low solar activity conditions. They find substantial cooling in the mesosphere of the order of -1 K/decade, increasing to -2.8K/decade in the thermosphere. Temperature decreases were small near the mesopause compared with the variation in the annual mean thus making trends there somewhat uncertain. Solomon et al. (2018) conclude that inconsistent observational results in the mesopause region, together with little or no global mean trends is due to the dominance of dynamical processes in controlling mesopause temperature, which exhibits significant inter-annual variability, even without variable solar forcing.

The SABER dataset (2002-2015) was used by Tang et al. (2016) to study the response of the cold-point temperature of the mesopause (T-CPM) to solar activity. The results showed that the T-CPM is significantly correlated to solar activity at all latitudes, and the solar response becomes stronger with increasing latitude. The solar-cycle dependence of the mesopause cold point temperature (T-CPM) is due to the relative importance of CO_2 and NO infrared cooling (Tang et al., 2016). NO density at solar max is about three times that at solar minimum. Consequently, CO_2 cooling is relatively less important at solar maximum, but is the dominant cooling mechanism during solar minimum.

Values of the solar response of T-CPM reported by Tang et al. (2016) increased from 2.82 ± 0.73 K/100 sfu at 0-10° S to 6.35 ± 1.16 K/100 sfu at 60-70° S (see their Fig 5(a)). Correlation coefficients of mesopause temperature with F10.7 cm solar irradiance data were higher for mid-latitudes (> 0.9) than at the equator (~0.7) and at higher latitude (see their Fig 5(b)). The ~~value~~ correlation coefficient found for 70° S (~0.8) is consistent with the ~~correlation coefficient~~value obtained for the OH* temperatures (Figure ~~1~~3(a) (R^2 = 0.584~~)½ — ~~~ or R = 0.76) obtained in ~~this~~ this work. At low latitudes, one would expect the QBO and ENSO to be significant factors ~~there~~ (see e.g., Nath and Sridharan, 2014), but at high latitudes, gravity wave activity is a candidate for the missing variance. Inter-annual variations of GWs at high latitudes are correlated with the strength of the polar vortex. A stronger polar vortex filters out more eastward propagating GWs, thus leading to more westward GW drag, which drives stronger meridional circulation (Karlsson and Shepherd, 2018).

Although the altitude of the mesospheric cold point changes with season (e.g., Yuan et al., 2019) and~~it~~ tends to be higher than the centroid height of the OH* layer, t~~, t~~he global solar response value obtained for T-CPM (4.89 ± 0.67 K/100 SFU) is in good agreement with the solar response coefficient derived from ground-based OH* observations.

The solar response of the T-CPM in Tang et al. (2016) shows some significant differences from the results in Figure 6 (zonal mean cycle from Aura/MLS) of this work. The solar response of the T-CPM increases more or less monotonically with latitude,

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whereas the solar response ~~registered~~ observed by Aura/MLS maximises at higher mid-latitudes. Of course the height of the T-CPM is some 7 km higher on average as indicated in Figure 9 (b) of Tang et al. (2016).

~~Several authors (Perminov et al., 2014; Pertsev and Perminov, 2008) have reported that winter OH* temperatures are more sensitive to the solar flux variation than summer temperatures. This agrees with the Aura/MLS variation shown in Figure 6.~~

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As a final comment on the global trends, it is noted that the largest errors in the linear trend fit for the SH winter understandably occur coincident with the regions positively or negatively correlated with the QQO (not shown here). The fit can be significantly improved if the QQO component can be understood and modelled. We investigate the QQO in detail in part 2 of this work.

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~~(cf. figure 3. i.e., eastern Antarctic polar cap, southern Pacific and southern Indian oceans). This is understandable if there is a significant QQO signal superposed on the underlying long-term linear trend.~~

6. Summary and Conclusions

We provide updates for the long-term trend and solar cycle response derived from 24 years of spectrometer observations of hydroxyl airglow at Davis Research Station, Antarctica (68° S, 78° E). A cooling trend in the mean winter temperatures [D106-259] of -1.20 ± 0.51 K/decade (95% confidence limits -0.14 K/decade < L < -2.26 K/decade) is obtained coupled with a solar cycle response coefficient of 4.30 ± 1.02 K/100sfu (95% confidence limits 2.2 K/100sfu < S < 6.4 K/100sfu). The observed cooling is consistent with radiative cooling due to increasing CO_2 concentrations and a rate of -0.06 K/ppmCO_2 or -0.22 K/%CO_2 is implied (ignoring possible contributions of stratospheric ozone change to the trend). A significant note is that a new record low winter-mean temperature was set for the Davis measurements in 2018, with a value of 198.3 K, which is 1.7 K below the previous minimum recorded in 2009 (200.0 K). An examination of the seasonal variation in the trend fit parameters reveals very little (no significant) long-term trend occurs over the ~~2~~ two midwinter months of June and July, but 95% significant trends of -1.5 to -2.6 K/decade during the April-May and August-October intervals. From examination of TIMED/SABER VER profiles we see no evidence that the trend results obtained can be significantly attributed to a change in the height of the OH layer.

We do not see evidence of a trend break or a change in the nature of the underlying trend after accounting for the solar cycle response in the Davis OH temperatures, however, this simple solar-cycle and linear trend model fit accounts for only 58% of the temperature variability. The remaining variability reveals evidence of a temperature oscillation on a quasi-quadrennial (~4 year period) timescale.

We compare our observations with Aura/MLS version v4.2 level 2 data over the last 14 years when these satellite data are available and find close agreement (a best fit to the variance in mean winter anomaly) with the 0.00464 hPa (native Aura/MLS retrieval) pressure level values. The solar cycle response, long-term trend and underlying QQO residuals are consistent with the Davis observations. Consequently, we derive global maps of Aura/MLS trend and solar response coefficients for the SH and NH winter periods to compare with other observers and models. Significant patterns for the zonally averaged solar cycle response are maxima in southern mid-latitudes in the ~~Southern Hemisphere (SH)~~SH winter and in northern mid-latitudes in the ~~Northern Hemisphere (NH)~~NH winter. Long term trends are ~~a~~ predominantly slight negative (~0.5-1 K/decade), except at the equator, and at mid-latitudes in the SH winter months. Comparisons are also made with the WACCM-X model and mesopause cold point temperature versus solar activity study using

TIMED/SABER data of Tang et al. (2016), both of which reveal significant differences in the zonally averaged patterns of solar cycle response and linear trend compared to the Aura/MLS data at 0.00464 hPa.

Further analysis using the datasets~~s~~ described here ~~are~~ is undertaken to ~~examine~~ explore the ~~residual~~ QQO signal ~~that this analysis has revealed~~revealed in the residual temperatures. A second part of this paper "Analysis of 24 years of mesopause region OH rotational temperature observations at Davis, Antarctica. Part 2: Evidence of a quasi-quadrennial oscillation (QQO) in the polar mesosphere." concerns this observation.

Data Availability

All Davis hydroxyl rotational data described in this manuscript are available through the Australian Antarctic Data Centre website (ref project AAS4157) via the following link https://data.aad.gov.au/metadata/records/Davis_OH_airglow . The satellite data used in this paper were obtained from the Aura/MLS data centre (see https://mls.jpl.nasa.gov), the SABER data centre (see http://saber.gats-inc.com/data.php) and are publicly available.

Author Contribution

WJRF managed data collection, performed data analysis, prepared manuscript with contributions from all co-authors

FJM analysis of SABER data, manuscript editing, figures, references

ARK analysis of Aura/MLS satellite data, manuscript editing.

Competing Interests

The authors declare that they have no conflict of interest.

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The satellite data used in this paper were obtained from the Aura/MLS data centre (see https://mls.jpl.nasa.gov), the SABER data centre (see http://saber.gats-inc.com/data.php) and are publicly available. We thank those teams and acknowledge the use of these data sets.

This work contributes to the understanding of mesospheric change processes coordinated through the Network for Detection of Mesospheric Change (see https://ndmc.dlr.de/)

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