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3	Top-down estimate of black carbon emissions for city cluster
4	using ground observations: A case study in southern Jiangsu,
5	China
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20 Abstract

21 We combined a chemistry transport model (CTM), a multiple regression model and 22 available ground observations, to derive top-down estimate of black carbon (BC) 23 emissions and to reduce deviations between simulations and observations for southern 24 Jiangsu city cluster, a typical developed region of eastern China. Scaled from a 25 high-resolution inventory for 2012 based on changes in activity levels, the BC 26 emissions in southern Jiangsu were calculated at 27.0 Gg/yr for 2015 (JS-prior). The 27 annual mean concentration of BC at Xianlin Campus of Nanjing University (NJU, a suburban site) was simulated at 3.4 μ g/m³, 11% lower than the observed 3.8 μ g/m³. In 28 contrast, it was simulated at 3.4 µg/m3 at Jiangsu Provincial Academy of 29 30 Environmental Science (PAES, an urban site), 36% higher than the observed 2.5 31 $\mu g/m^3$. The discrepancies at the two sites implied the uncertainty of the bottom-up 32 inventory of BC emissions. Assuming a near-linear response of BC concentrations to 33 emission changes, we applied a multiple regression model to fit the hourly surface 34 concentrations of BC at the two sites, based on the detailed source contributions to 35 ambient BC levels from brute-force simulation. Constrained with this top-down 36 method, BC emissions were estimated at 13.4 Gg/yr (JS-posterior), 50% smaller than 37 the bottom-up estimate, and stronger seasonal variations were found. Biases between 38 simulations and observations were reduced for most months at the two sites when 39 JS-posterior was applied. At PAES, in particular, the simulated annual mean was 40 elevated to 2.6 µg/m³ and the annual normalized mean error (NME) decreased from 72.0% to 57.6%. However, application of JS-posterior slightly enhanced NMEs in 41 42 July and October at NJU where simulated concentrations with JS-prior were lower 43 than observations, implying that reduction in total emissions could not correct CTM 44 underestimation. The effects of numbers and spatial representativeness of observation 45 sites on top-down estimate were further quantified. The best CTM performance was 46 obtained when observations of both sites were used with their difference in spatial 47 functions considered in emission constraining. Given the limited BC observation data 2/63





48 in the area, therefore, more measurements with better spatiotemporal coverage were 49 recommended for constraining BC emissions effectively. Top-down estimates derived 50 from JS-prior and the Multi-resolution Emission Inventory for China (MEIC) were 51 compared to test the sensitivity of the method to initial emission input. The 52 differences in emission levels, spatial distributions and CTM performances were 53 largely reduced after constraining, implying that the impact of initial inventory was 54 limited on top-down estimate. Sensitivity analysis proved the rationality of near 55 linearity assumption between emissions and concentrations, and the impact of wet 56 deposition on the multiple regression model was demonstrated moderate through data 57 screening based on simulated wet deposition and satellite-derived precipitation.

58 **1 Introduction**

59 Black carbon (BC), alternatively referred as elemental carbon (EC), is an crucial 60 component of atmospheric particle and comes mainly from incomplete combustion of 61 fossil fuels and biomass. BC has adverse effect on human health as it absorbs harmful 62 volatile organic compounds like polycyclic aromatic hydrocarbons (Dachs and 63 Eisenreich, 2000). Furthermore, BC contributes to global warming by intercepting and absorbing sunlight (Jacobson, 2001; Ramanathan and Carmichael, 2008). Bond et 64 al. (2013) assessed that the global average radiative forcing of BC was $+1.1 \text{ W/m}^2$ 65 (90% confidence interval: 0.17-2.1 W/m²), which was more than two-thirds of that 66 67 from CO₂ (+1.56 W/m²). Since BC remains for only a few days in the atmosphere, it is an effective way to mitigate climate warming in the short term by reducing BC 68 69 emissions. However, due to lack of sufficient understanding of major emission 70 sources, the effect of BC on regional climate was not fully quantified by models.

BC emission inventories are traditionally developed with the bottom-up method based on activity levels and emission factors. Previous studies of chemistry transport modeling (CTM) based on emission inventories found large discrepancies between simulated and observed BC concentrations. Koch et al. (2009) found that sixteen models applied in the AeroCom aerosol model inter-comparison project





76 underestimated surface BC levels by a factor of 2-3. Hu et al. (2016) found that CTM 77 significantly underestimated the peak surface concentrations of BC over northwestern 78 United States, likely due to missing strong local fire events in fire emissions. 79 Moreover, large differences existed in various bottom-up emission inventories, 80 particularly for China with large energy consumption, complicated emission source 81 categories, and fast changes in emission characteristics. BC emissions in China for 82 2001 and 2006 in the Regional Emission inventory in ASia (REAS 2.1, Kurokawa et 83 al., 2013) were smaller than those in the Intercontinental Chemical Transport 84 Experiment-Phase B (INTEX-B, Zhang et al., 2009), but the growth rate of BC 85 emissions in REAS 2.1 was larger than that in INTEX-B (30% versus 15%) for the 86 five years. Ohara et al. (2007) evaluated the inter-annual trend in China's BC emissions with constant emission factors, and found that the national emissions 87 continuously decreased by 23% from 1990 to 2000. In contrast, Lei et al. (2011) 88 89 suggested a much smaller inter-annual variability with the peak annual emissions found in 1996 for the same period. The differences resulted largely from the use of 90 91 activity levels from various data sources, especially for residential biofuel combustion. 92 The gaps between different studies implied potentially large uncertainties in BC 93 bottom-up emission inventories. The uncertainties of BC emission estimates for China 94 were reported at $\pm 484\%$, $\pm 208\%$, and $\pm 98\%$ by Streets et al. (2003), Zhang et al. 95 (2009), and Lu et al. (2011), respectively. Due to lack of sufficient local field tests, 96 emission factors were commonly taken from foreign studies with big variety 97 depending on fuel and combustion condition (Bond et al., 2004; Cao et al., 2006; Lei 98 et al., 2011; Qin and Xie, 2012; Streets et al., 2003; Streets et al., 2001; Zhang et al., 99 2009). It was also difficult to obtain accurate and detailed activity data, particularly 100 for the main sources of BC including small industries (e.g., coke and brick 101 production), off-road transportation, and residential solid fuel combustion. Besides the 102 large uncertainty in emission estimation, challenges existed as well in updating BC 103 inventories continuously (Hong et al., 2017; Lu et al., 2011; Xia et al., 2016; Zhao et





104 al., 2013). To beat severe air pollution, China has been conducting series of measures 105 in energy conservation and emission control, leading to dramatic changes in energy 106 structure, emission factors and removal rates of air pollutant control devices (Zhao et 107 al., 2014). Such changes could be partly tracked by continuous emission monitoring 108 system (CEMS) that was commonly installed at big industrial enterprises. Large 109 fractions of BC emissions, however, came from medium and small sources, and their 110 most recent improvements in manufacturing technologies and emission controls were 111 relatively difficult to be obtained timely and efficiently.

112 Given above limitations in bottom-up inventories, different top-down approaches 113 were applied to evaluate BC emissions. For example, Cohen and Wang (2014) 114 presented a Kalman filter technique to estimate the global BC emissions based on 115 satellite-derived radiances and surface concentrations from global and regional networks. The adjoint-based 4-D variational approach was also applied to constrain 116 117 the bottom-up BC emissions at the global or national scales (Zhang et al., 2015; Xu et 118 al., 2013; Guerrette et al., 2017). A near-linear response of BC concentrations to 119 emission changes was generally assumed at national (Fu et al., 2012; Kondo et al., 120 2011; Wang et al., 2013) and regional scales (Li et al., 2015; Wang et al., 2011), due to 121 its weak activity in atmospheric chemistry reaction. The ratio of observed to 122 simulated concentration can be used as a scaling factor to correct BC emissions. 123 Kondo et al. (2011) made continuous measurement of BC concentrations for a full 124 year on a remote island in East China Sea. With the data strongly affected by 125 emissions from China identified and those largely influenced by wet deposition excluded, they estimated China's annual anthropogenic BC emissions at 1.92 TgC/yr. 126 127 Wang et al. (2013) verified this linearity by conducting sensitivity simulation in which 128 emissions were increased by 50%. After excluding observation data of heavy 129 pollution and strong precipitation events at five Chinese sites, they calculated China's 130 annual BC emissions at 1.80 TgC/yr. The results of both studies were close to a 131 bottom-up estimate at 1.81 TgC/yr by Zhang et al. (2009). Based on observations at





132 10 Chinese background and rural sites, Fu et al. (2012) applied a multiple regression model and CTM to quantify China's BC emissions. They calculated the total 133 134 emissions at 3.05 TgC/yr, 59% larger than those by Zhang et al. (2009). Using similar 135 approach, Li et al. (2015) estimated BC emissions to be 34% larger than bottom-up 136 inventory in Pearl River Delta in south China by Zheng et al. (2012). Park et al. (2003) 137 used the multiple linear regression to fit the Interagency Monitoring of Protected 138 Visual Environments (IMPROVE) data and estimated that BC emissions from fossil 139 fuel and biofuel burning in the United States should be increased by 15%. Combining 140 a general circulation model simulation and the receptor modeling approach, Verma et 141 al. (2017) constrained BC emissions over India based on the scaling factor (the ratio 142 of simulated to observed BC concentration).

143 To our knowledge, limitations remained in the assessment of BC emissions based 144 on the top-down approach. Current available studies focused mainly on global or 145 national scale, and few evaluations could be found for city clusters. In aims of examining emission control policies and quantifying impacts of BC on local climate 146 147 and air quality, there was a strong need for studies at city cluster scale that require 148 ground observation and emission inventory with improved details. Regarding 149 measurement data, monthly or annual means were commonly used in previous studies, 150 and information of heavy-polluted events were lost when targeting a local scale. In 151 general, observations at a higher temporal resolution were considered as an important 152 means to effectively reduce uncertainties (Matsui et al., 2013; Wang et al., 2013; 153 Gilardoni et al., 2011). Moreover, it was somewhat arbitrary to differentiate emissions 154 by sector in previous top-down estimates, attributed to lack of detailed information on 155 source categories from bottom-up inventories. The method was thus insufficient to 156 make substantial improvement on emission evaluation by sector, or to clearly stress 157 the direction of further revisions on bottom-up inventories.

158 In this work, therefore, we integrated CTM, multiple regression model and 159 available hourly ground observations to provide top-down constraint of BC emissions





160 and to reduce deviations between simulations and observations at city cluster scale. 161 We selected southern Jiangsu city cluster including cities of Suzhou, Wuxi, 162 Changzhou, Zhenjiang, and Nanjing, a typical region with large population and 163 economy in Yangtze River Delta (YRD), China (see the geographic location and cities 164 in Figure S1 in the supplement). Given its intensive industry and energy consumption, 165 the city cluster was regarded as one of the largest BC emission sources in eastern 166 China and BC emissions from this region accounted for nearly half of the total 167 emissions in Jiangsu (Zhou et al., 2017). The heavy air pollution was found in the 168 region: the annual averages of fine particle (PM_{2.5}) concentrations in all the cities 169 exceeded the National Ambient Air Quality Standard (NAAQS, 35 μ g/m³) in 2012. 170 Under the pressure of air quality improvement, Jiangsu conducted aggressive actions 171 of emission control, leading to 20% reduction in the annual average of PM_{25} concentration from 2013 to 2015. Based on a provincial bottom-up emission inventory, 172 173 we estimated the contributions to BC concentrations by sector at two ground observation sites through the brute-force method in CTM. The results, together with 174 175 observed ambient BC concentrations, were incorporated in a multiple regression 176 model to derive the top-down estimate of BC emissions for southern Jiangsu city 177 cluster. The advantage of top-down estimate against bottom-up inventory was then 178 judged by CTM and ground observations. The factors that would potentially influence 179 the top-down estimate were also evaluated, including number and spatial 180 representativeness of observation sites, and initial bottom-up emission input. The 181 uncertainties of the multiple regression model were finally evaluated including the 182 influence of precipitation and the near-linear assumption between BC emissions and 183 concentrations.

184 2 Data and method

185 **2.1 Bottom-up inventories of BC emissions**

186 Two bottom-up emission inventories at different spatial scales were used in this





187 work. At the national scale, the Multi-resolution Emission Inventory for China (MEIC, 188 http://www.meicmodel.org/) was developed by Tsinghua University, with an original horizontal resolution at $0.25^{\circ} \times 0.25^{\circ}$. At the provincial scale, Zhou et al. (2017) 189 190 collected the best available information of industrial sources in Jiangsu and developed 191 an inventory with higher resolution at 3×3 km. The latter was proved to be more 192 supportive in air quality simulation at city cluster scale (Zhou et al., 2017; Zhao et al., 193 2017). In both inventories, anthropogenic BC emissions for 2012 came from four 194 major sectors: power generation, industry, residential sources and transportation. The 195 national and provincial inventories for 2015 (mentioned respectively as MEIC-prior 196 and JS-prior hereinafter) were obtained using a simple scaling method based mainly 197 on changes in activity levels (energy consumption and industrial production, etc) 198 between the four years. Table S1 in the supplement summarizes the data sources of 199 activity levels and the scaling factors by sector in JS-prior. As MEIC-prior includes 200 only four major sectors, the scaling factor for each sector was calculated as the 201 average of those for subcategories within the sector. Potential changes in BC emission 202 factors from 2012 to 2015, e.g., those attributed to varied manufacturing technologies 203 and/or penetrations of emission control devices, were not considered in the calculation. 204 The implication and uncertainty from that simplified emission scaling method will be 205 further discussed in Section 4.3. The temporal distribution of the emissions was 206 dependent on that of activity levels by source category. Such information was investigated by Zhou et al. (2017) according to the official statistics of the country 207 208 (http://data.stats.gov.cn/) and directly adopted in this work.

209 2.2 Top-down emission estimation with multiple regression model

The top-down emissions of BC in southern Jiangsu (mentioned as JS-posterior hereinafter) were estimated with a multiple regression model using ground observations as constraint. The regression model matched BC contributions by sector (calculated through CTM) against measured ambient hourly BC concentrations:





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$$c_{obs} = \beta_1 c_{power} + \beta_2 c_{industry} + \beta_3 c_{residential} + \beta_4 c_{transportation} + \varepsilon$$
(1)

where c_{obs} is the vector of observed hourly BC concentrations; c_{power} , $c_{industry}$, $c_{residential}$, and $c_{transportation}$ are the vectors of BC concentrations contributed by power generation, industry, residential sources and transportation, respectively, and they were simulated using the brute-force method as described in Section 2.3; β_1 - β_4 are the domain-wide scaling factors obtained by sector in the multiple regression model to best match observations; and ε is the error vector of the model.

221 As BC is not one of the six regulated air pollutants in the NAAQS, it was a big 222 challenge to obtain observation data with high temporal resolution in most cities of 223 southern Jiangsu. For the whole year 2015, hourly ambient BC concentrations were available at two sites in Nanjing, the capital of Jiangsu. As illustrated in Figure 1, one 224 225 is a suburban site located in the Xianlin Campus of Nanjing University in northeast 226 Nanjing (NJU), and the other is an urban site in Jiangsu Provincial Academy of 227 Environmental Science (PAES). At both sites, BC was sampled and analyzed hourly 228 with semi-continuous carbon analyzer (Model-4, Sunset Lab, USA). Details of the 229 measurement approach were described in Chen et al. (2017). The statistics of 230 observed ambient BC concentrations at the two sites are shown in Figure S2 in the 231 supplement. The annual average BC concentrations (calculated as the mean of January, April, July and October) were 3.83 and 2.47 μ g/m³ at NJU and PAES, respectively. 232 233 The hourly average BC observations ranged 0.06-17.65 µg/m³ and 0.22-19.76 µg/m³ at NJU and PAES, respectively. The values were similar to those observed in the 234 235 Guanzhong basin (0.4-23.1 μ g/m³), the Pearl River Delta region (1-13 μ g/m³) and the 236 Beijing-Tianjin-Hebei region (2-32 µg/m³) (Li et al., 2016). Much higher BC 237 concentrations were observed in autumn and winter at both sites, with the monthly means at 3.96 and 5.44 μ g/m³ at NJU and 3.62 and 2.80 μ g/m³ at PAES, respectively. 238 239 The scaling factors derived from Eq. (1) were used to constrain BC emissions in

240 JS-prior from a top-down perspective by assuming a near-linear relation between

241 changes in BC concentrations and emissions:





242
$$E_{JS-posterior} = \beta_1 E_{power} + \beta_2 E_{industry} + \beta_3 E_{residential} + \beta_4 E_{transportation}$$
(2)

- 243 where *E_{JS-posterior}* is the vector of the total BC emissions from the top-down approach;
- 244 E_{power}, E_{industry}, E_{residential} and E_{transportation} are the vectors of BC emissions from power
- 245 generation, industry, residential sources and transportation, respectively, in JS-prior.
- 246 **2.3 Air quality simulation**

247 We used the Models-3 Community Multi-scale Air Quality (CMAQ) version 248 4.7.1 to simulate ambient BC concentrations. As shown in Figure 1, three nested 249 domains were applied with horizontal resolutions of 27, 9, and 3 km, respectively, on 250 a Lambert Conformal Conic projection centered at (110°E, 34°N). The mother domain 251 (D1, 177×127 cells) covered most parts of China and other surrounding countries. The 252 second domain (D2, 118×121 cells) covered Jiangsu, Anhui, Zhejiang, Shanghai, and 253 parts of other provinces in China. The third domain (D3, 133×73 cells) covered 254 Shanghai, part of Anhui province and the city cluster in southern Jiangsu. There were 255 27 vertical levels from the ground surface up to 50 hPa on terrain-following 256 coordinated. The simulations were conducted for January, April, July and October to 257 represent four typical seasons in 2015. A 5-day spin-up period of each month was 258 applied to minimize the influence of initial conditions in the simulations.

259 Meteorological fields were simulated by the Weather Research and Forecasting Model (WRF) version 3.4 and the carbon bond gas-phase mechanism (CB05) and 260 261 AERO5 aerosol module were adopted in CMAQ. Relevant details of model 262 configuration can be found in Zhou et al. (2017). Statistical indicators including 263 averages of simulations and observations, bias, normalized mean bias (NMB), normalized mean error (NME), root mean squared error (RMSE) and index of 264 agreement (IOA) were applied to evaluate the modeling performance of WRF (Baker 265 266 et al, 2004; Zhang et al., 2006). Ground observation data at 1 or 3 h interval at 267 meteorological stations including Lukou, Hongqiao and Liyang stations in the third 268 domain (labeled in Figure 1) were taken from National Climatic Data Center (NCDC).





The statistical indicators for temperature at 2 m (T2) and relative humidity at 2 m (RH2), wind speed and direction at 10 m (WS10 and WD10) for the four typical months in 2015 are summarized in Table S2 in the supplement. Discrepancies between ground observations and WRF modeling were within acceptable range (Emery et al., 2001).

274 To make it applicable in our CTM, MEIC-prior was downscaled into grid 275 systems of each modeling domain, based on the spatial distributions of gross domestic 276 product (GDP, for power generation and industrial emissions) and population (for 277 residential and transportation emissions) at a horizontal resolution of 1×1 km. The 278 downscaled MEIC-prior was used for the first, the second domains and the regions 279 outside Jiangsu of the third domains, while JS-prior was applied for the Jiangsu region 280 of the third domain. Brute-force method was applied to estimate contributions to ambient BC concentrations by sector. Five scenarios were designed in this study: 281 282 Scenario B (the base scenario) in which emissions from all sources in the third 283 domain were included, and Scenarios S1, S2, S3, and S4 in which BC emissions from 284 power generation, industry, residential sources and transportation were zeroed out, 285 respectively. We compared simulated BC concentrations in S1, S2, S3 and S4 with 286 those in Scenario B in four months, and the contributions from four major emission 287 sectors to ambient BC levels were determined as the differences in simulated 288 concentrations between Scenarios B and S.

289 **3 Results**

290 **3.1 Bottom-up emission estimate**

The total annual BC emissions of JS-prior were estimated at 26.99 Gg for southern Jiangsu city cluster in 2015, including 0.18 Gg from power generation, 17.67 Gg from industry, 3.80 Gg from residential sources and 5.33 Gg from transportation, as shown in Figure 2. Accounting for 66% of total annual emissions, industry was identified as the dominant contributor to BC, followed by transportation (20%) and





296 residential sources (14%). Although the policies of energy conservation and emission 297 control have been conducted for years, there were still a number of small facilities 298 with low operation temperatures and combustion efficiencies in southern Jiangsu, 299 leading to a large amount of BC from incomplete combustion. When scaling 300 emissions from 2012 to 2015, in addition, improvements in emission controls were 301 not taken into account, such as elevated combustion technologies and enhanced use of 302 dust collectors. The potential reductions in net emission factors for major factories, 303 therefore, were not well quantified, and the emissions from industry could be 304 overestimated. Emissions from power generation were few, resulting from relatively 305 high combustion efficiency of pulverized boilers and large penetrations and removal 306 rates of dust collectors. Besides the annual total, the emissions of four months 307 (January, April, July and October) were also estimated and limited seasonal 308 differences were found as shown in Figure 2.

309 Figure S3 in the supplement shows the spatial distribution of annual BC emissions in JS-prior. For power generation and industry sectors, latitude and 310 longitude of each plant were applied to allocate BC emissions, and the outstandingly 311 312 high emissions shown in the map indicated the existence of big industrial plants. For 313 residential sources, large emissions were found in the regions with intensive 314 population. Emissions from transportation were mainly distributed along the road net 315 and downtown regions in southern Jiangsu cities (see the geographic locations of downtowns in Figure S1 in the supplement), slightly overlapping with those from 316 317 residential sources.

318 **3.2 Top-down emission estimate**

The time series of BC concentrations contributed by various sectors (*c* in Eq. (1)) were simulated with CTM and illustrated in Figures S4 and S5 in the supplement for NJU and PAES, respectively. Among all the sectors, the largest seasonal variation in BC contribution was found for residential sources. The average concentrations contributed by this sector in January reached 0.76 and 0.94 μ g/m³ at NJU and PAES,





324 respectively, approximately double of those in another three months. The 325 concentrations contributed by industry were significantly enhanced in certain periods (e.g., January 20th, April 9th-11th, and July 15th-17th), and industrial emissions were 326 327 expected to be an important reason for the overestimation in BC concentrations 328 through CTM (see the model evaluation in Section 3.3). Table S3 in the supplement 329 summarizes the monthly and annual mean BC contributions by sector. The annual 330 contributions of industry at the two sites were close to each other (21.0% and 21.9% 331 at NJU and PAES respectively). Contributions of residential sources and 332 transportation were higher at PAES resulting from large population and heavy traffic 333 in the urban area. Minor contribution of power generation to BC concentrations was 334 found at both sites (the annual means were less than 1%), attributed to its very limited 335 emissions.

Summarized in Table 1 are the scaling factors β_1 - β_4 estimated from multiple 336 337 regression model (Eq. (1)) by season, together with the statistical indicators including 338 the values of t, Sig. (or p) and variance inflation factor (VIF). The values of t and Sig. 339 indicate statistical significance with a threshold of 2 and 0.05, respectively. VIF is a 340 test for multicollinearity and the model is reasonable with VIF smaller than 10. Since 341 the emissions from power generation were small and they contributed very little to 342 ambient BC concentrations, inclusion of power generation component would not 343 significantly improve the regression model. In this study, therefore, we assumed that the simulated BC concentrations from power generation were correct by setting β_l at 344 345 1 and further subtracted them from the observations. Most statistical indicators in Table 1 met the criteria (t>2, Sig.<0.05, VIF<10) and the overall significance was 346 347 0.00 in four months, implying acceptable robustness of the multiple regression model. 348 However, the results were not statistically significant indicated by t and p values for 349 some months and sectors (e.g., industry in April and residential in April and July), 350 implying that the constrained emissions for those month/sectors need to be cautiously 351 analyzed.





352 By applying β_1 - β_4 in Eq. (2), the top-down estimates of BC emissions 353 (JS-posterior) were estimated and illustrated in Figure 2. The total BC emissions for 354 southern Jiangsu city cluster were calculated at 13.4 Gg, 50% smaller than those of 355 JS-prior. The scaling factors of emissions from industry and transportation (β_2 and β_4) ranged from 0.22 to 0.42 and from 0.55 to 0.79 for different months, respectively. 356 357 Accordingly, the emissions from industry and transportation in JS-posterior were 358 estimated 67% and 32% smaller than those in JS-prior, respectively. As mentioned 359 above, the emissions in JS-prior 2015 were simply scaled from those in 2012 360 according to activity data, and changes in emission factors were not considered. In the 361 actual fact, however, a series of measures in industry and transportation were 362 conducted to improve energy efficiency and to reduce emissions over recent years. 363 Issued in 2013, for example, the Air Pollution Control Planning for the Key Regions for the 12th Five-Year Plan period (2010-2015) aimed to achieve 7% and 15% 364 365 reductions in the annual average concentration and industrial emissions of fine particles in Jiangsu province from 2010 to 2015, respectively (Qian, 2013). The 366 367 measures included eliminating old and energy-inefficient plants of heavy-polluted 368 industries (thermal power generation and steel/building material production), and 369 optimizing the energy structure through application of sustainable energy. Meanwhile, 370 the enhanced use of cleaner gasoline and diesel products (National stage V standard) 371 in transportation could lead to reduced vehicle emissions. The government efforts in 372 emissions controls proved effective, indicated by the scaling factors much smaller 373 than 1 (β_2 and β_4 in Table 1) and the reduced emissions of JS-posterior. For residential 374 sources, the emissions in JS-posterior were 3% smaller than those in JS-prior, 375 indicating limited difference in the annual total emissions between the two inventories. 376 However, the scaling factors (β_3) in January and October were 1.31 and 1.52 377 respectively, showing a stronger enhancement in BC emissions in winter and autumn 378 in JS-posterior than those in JS-prior. It thus implied that there were missing sources 379 likely associated with low-quality fossil fuels or biofuel used for heating in winter and





380 crop waste burning in autumn in JS-prior. For the capital city of Jiangsu Province, 381 Nanjing, Huang et al. (in preparation) conducted detailed analysis on the changes in 382 operation activities and emission control technologies of individual sources based on 383 annually updated official environmental statistics and pollution census. With the 384 bottom-up approach, the annual BC emissions in the city were estimated to decrease 385 by 60% from 2012 to 2015 as shown in Figure S6 in the supplement. The relative 386 change in annual emissions was close to that between JS-prior and JS-posterior, and 387 the validity of the two methods (the bottom-up approach by Huang et al. and the 388 top-down approach in this work) could be verified.

389 Figure 3 presents the seasonal variations in BC emissions of JS-prior, 390 JS-posterior and MEIC-prior by sector, and stronger variations were generally found 391 in JS-posterior. As shown in Figure 3a, the largest difference among the three 392 inventories existed in the residential sources, and the ratio of maximum to minimum 393 monthly emissions was 4.33 in JS-posterior, close to that in MEIC-prior at 4.00 and 394 nearly 4 times of that in JS-prior at 1.13. The analogue ratio for industry was 2.05 in 395 JS-posterior, nearly twice of those in JS-prior at 1.14 and MEIC-prior at 1.12. The 396 smallest difference was found for transportation among the three inventories. 397 Seasonal variations in total emissions were a combination of those by sector weighted 398 by the contribution of each sector to total emissions. The ratios of maximum to 399 minimum monthly emissions were 1.13, 1.83 and 1.29 for JS-prior, JS-posterior and 400 MEIC-prior, respectively (Figure 3b). The value for JS-posterior was closer to 2.1 for 401 an anthropogenic BC emission inventory in China by Lu et al. (2011) that considered 402 enhanced use of fossil fuels for residential heating in winter in northern China. The 403 comparison thus implied again that current bottom-up inventories might 404 underestimate the emissions of residential solid fuel burning in winter in southern 405 Jiangsu. As central household heating was not conducted in the area in winter, the 406 official energy statistics on which bottom-up inventories were based may not fully 407 capture the elevated fuel burning by disperse households. Spatial distribution of BC





408 emissions in JS-posterior was illustrated in Figure S3 in the supplement. Compared to

409 JS-prior, BC emissions from industry and transportation were greatly reduced in

410 downtown regions in southern Jiangsu city cluster.

411 **3.3 Evaluation of the top-down emission estimate**

The simulated BC concentrations based on bottom-up (JS-prior) and top-down estimation in emissions (JS-posterior) were compared with observations to evaluate the two inventories, and the results were illustrated in Figures 4 and 5 for NJU and PAES sites, respectively. Statistical indicators including mean concentrations from simulations and observations, NMB and NME, as well as the regression correlation (R) were calculated to evaluate the modeling performance, as summarized in Table 2.

418 In general, CTM based on JS-prior reproduced well the temporal variations of 419 the observed BC concentrations at the two sites. The highest and lowest 420 concentrations were respectively simulated in winter and summer, consistent with 421 observations with an exception at PAES where the observed monthly mean in January $(2.80 \ \mu\text{g/m}^3)$ was lower than that in October $(3.62 \ \mu\text{g/m}^3)$. The seasonal variation of 422 423 BC concentrations at NJU was larger than that at PAES, suggesting bigger impact of 424 household solid fuel use on the suburban and rural regions. Though the model was able to capture the seasonal variability, discrepancies between simulations and 425 426 observations existed, and CTM commonly underestimated BC concentrations at the 427 suburban site NJU and overestimated those at the urban site PAES. With the monthly 428 means ranged 1.99-5.97 µg/m³ at NJU, the annual average of BC concentration 429 (calculated as the mean of January, April, July and October) was simulated at 3.44 430 $\mu g/m^3$, smaller than the observed 3.83 $\mu g/m^3$. With the monthly means ranged 2.61-6.46 μ g/m³, in contrast, the annual concentration at PAES was simulated at 3.39 431 $\mu g/m^3$, larger than the observed 2.48 $\mu g/m^3$. Better correlation between observation 432 433 and simulation was found at NJU, indicated by the larger R. The annual mean NMBs 434 were calculated at -10.16% and 36.67%, and the NMEs were 41.15% and 72.00% at 435 NJU and PAES, respectively. The discrepancy suggested that JS-prior used in CTM





436 might misrepresent the spatial pattern of emissions. Population and economy densities 437 were applied to allocate BC emissions, leading to overestimation in emissions and 438 thereby simulated concentrations in urban areas with more population and economic 439 activity. Besides, the model overestimated the peak surface concentrations at both 440 sites particularly when the contribution from industry sector was enhanced as 441 mentioned in Section 3.2 (e.g., January 9th-11th and April 9th-10th at NJU, and April 442 9th-12th, the second half of July, and October 20th at PAES).

443 Application of JS-posterior in CTM effectively corrected large biases between 444 simulations and observations at the two sites. As shown in Table 2, NMEs were 445 reduced for most months while effects of applying JS-posterior in CTM varied at two 446 sites. At PAES, the annual average NME declined from 72.00% to 57.55% and the 447 annual mean of BC concentration was simulated at 2.57 µg/m³, in better agreement with the observed 2.48 μ g/m³ than the simulated 3.39 μ g/m³ using JS-prior. The 448 449 largest reductions in NMEs were found in April and July, from 73.18% to 42.87% and 450 from 92.74% to 42.37%, respectively. Moreover the overestimation in peak 451 concentrations using JS-prior were partly corrected when JS-posterior was applied, 452 resulting mainly from the reduced emissions from industry and transportation. 453 Although simulation of peak concentrations at NJU were improved as well, the annual 454 average NME at NJU slightly increased from 41.15% to 44.16% and the annual mean of BC concentration was simulated at 2.82 μ g/m³, smaller than the simulated 3.44 455 456 µg/m³ using JS-prior. Bigger bias was found in July and October at NJU, since the 457 reduced emission estimates in JS-posterior led to further underestimation in simulated ambient BC levels compared to JS-prior. Limitation of current multiple regression 458 459 model was thus indicated that overestimation and underestimation in concentrations at 460 different sites could hardly be corrected simultaneously without further improvement 461 in spatial distribution of emissions.





463 **4 Discussions**

We selected April to evaluate the sensitivity of observation and bottom-up 464 465 emission input to top-down constraint. Observation site number, spatial representativeness of sites, and initial bottom-up inventory were changed separately in 466 467 the constraining approach, and various top-down estimates could be derived and 468 compared with each other. The statistical indicators of modeling performances based 469 on different bottom-up and top-down emission estimates in April are summarized in 470 Table 3. Furthermore, we evaluated the uncertainty of the multiple regression model, 471 including the assumption of near linearity between emissions and concentrations and 472 the impact of precipitation. Details were described as below.

473 **4.1 The effect of observation site number**

A major challenge in understanding the sources and distributions of BC in China 474 475 was lack of a consistent and stable measurement network with good spatiotemporal 476 coverage, such as the IMPROVE network in the United States (Malm et al., 1994). 477 Uncertainty existed in the top-down estimates in this work, as hourly measurements 478 on BC concentrations were only available at two sites in southern Jiangsu. Therefore, 479 besides JS-posterior derived from observations at both sites as described in Section 480 3.2 (mentioned as Case 1 hereinafter), we conducted a Case 2 in which observation 481 data at only one site (NJU) was used in the top-down approach, to analyze the effect 482 of the site number on emission estimates. The scaling factors of emissions from industry, residential sources and transportation were recalculated at 0.42, 0.95 and 483 484 0.65, respectively. Compared with Scenario B, the NMEs of Case 2 decreased from 485 42.31% to 32.47% and from 73.18% to 61.59% at NJU and PAES, respectively, implying the benefits of ground measurements (even available only at one site) on 486 487 emission constraint. The NME in Case 2 was slightly smaller than that in Case 1 at 488 NJU, suggesting that application of measurement data at one single site could 489 improve model performance moderately at that site. At PAES, in contrast, much larger





490 NME was found in Case 2. Much better model performance in Case 1 at PAES

491 indicated that inclusion of more measurements with better spatiotemporal coverage

492 could constrain BC emissions at city cluster level more effectively.

493 **4.2** The effect of spatial representativeness of observation sites

494 Spatial representativeness of observation sites was identified and its impact on 495 top-down emission constraint was evaluated. Considering the prevailing winds from 496 northeast and southeast, on one hand, NJU located upwind Nanjing is hardly 497 influenced by the emissions from the downtown of the city. Besides the site is 498 downwind of the Yangtze River Delta region (YRD) including the 499 Suzhou-Wuxi-Changzhou-Zhenjiang city cluster (Chen et al., 2017), thus it is more 500 representative for the western YRD emissions through regional transport. On the other 501 hand, PAES is located at urban Nanjing and its air quality is commonly influenced by 502 surrounding transportation, residential, and commercial sources, thus the site is 503 representative for the local emissions of Nanjing. In contrast to previous top-down 504 studies that did not distinguish influence of local emissions and transport on air 505 quality in sub-regions of the research domain (Wang et al., 2011; Fu et al., 2012), the 506 spatial representativeness of the two observation sites were taken into account to 507 improve the top-down approach and the result of constraining BC emissions in 508 southern Jiangsu city cluster. Through the brute-force method described in Section 2.3, 509 we zeroed out the emissions from Nanjing and Suzhou-Wuxi-Changzhou-Zhenjiang 510 city cluster in CTM, respectively, and compared the simulated concentrations with those in Scenario B to analyze the contributions of the two regions to ambient BC 511 512 concentrations at NJU and PAES sites. As shown in Figure S7 in the supplement, the 513 contribution of emissions from Nanjing to PAES was greater than that to NJU in 82% 514 of the modeling period, and the analogue number was 81% for the contribution of 515 Suzhou-Wuxi-Changzhou-Zhenjiang city cluster to NJU greater than that to PAES. 516 We thus concluded that emissions from Nanjing contributed significantly to PAES 517 while those from Suzhou-Wuxi-Changzhou-Zhenjiang city cluster contributed





518 significantly to NJU. We then developed a new case of top-down emission estimate in 519 southern Jiangsu (Case 3), in which observation data at PAES and NJU were applied 520 to constrain emissions from Nanjing and Suzhou–Wuxi–Changzhou-Zhenjiang city 521 cluster, respectively.

522 The scaling factors in Case 3 are provided in Table 4. To avoid the collinearity in 523 the multiple regression model, we expected that the relative changes in emissions 524 from transportation in Nanjing and Suzhou-Wuxi-Changzhou-Zhenjiang city cluster 525 were similar for recent years, resulting from the same progress of emission standard implementation (National Standard Stage IV) in southern Jiangsu and the frequent 526 527 circulation of vehicles among the cities. Therefore a same scaling factor was assumed 528 for transportation in the two regions. As shown in Table 4, all the scaling factors at 529 PAES were smaller than those at NJU, implying that implementation of emission 530 controls in Nanjing were more stringent than that in 531 Suzhou-Wuxi-Changzhou-Zhenjiang city cluster from 2012 to 2015. As the host city of the 2nd Asian Youth Games in 2013 and the 2nd Youth Olympic Games in 2014, 532 533 Nanjing was undertaking series of restrictions on air pollutant emissions. The city 534 conducted emission control action on small coal-fired boilers since 2013 and over 535 1200 coal-fired boilers had been shut down by the end of 2014. In addition, central 536 heating units were largely applied to replace the coal with electricity, natural gas or 537 biofuel. As shown in Table 3, the NMEs in Case 3 were the smallest at both sites 538 among all the cases with an exception: the NME at NJU in Case 3 was 32.64%, 539 slightly larger than that in Case 2 at 32.47%. The result implied that inclusion of more 540 measurement data with their spatial representativeness considered could improve the 541 top-down approach in terms of spatial distribution of emissions and could reduce the 542 deviation between observations and simulations.

543 Summarized in Table 5 are BC emissions from Nanjing and 544 Suzhou–Wuxi–Changzhou-Zhenjiang city cluster estimated in different cases. All the 545 top-down estimates were approximately half of the bottom-up estimate and the





546 estimate in Case 1 was the smallest among all the cases. The same scaling factors 547 were generated and applied in Cases 2 and 3 to calculate BC emissions from 548 Suzhou-Wuxi-Changzhou-Zhenjiang city cluster which accounted for 80% of the 549 total emissions in southern Jiangsu, resulting in similar top-down emission estimates 550 between the two cases.

551 4.3 The effect of initial bottom-up emission input

552 Given the large uncertainty in JS-prior that was simply developed based on the 553 changes of activity levels in recent years, we applied MEIC-prior as well to explore 554 the effect of initial emission inventory on top-down BC constraints.

555 Figures 6 and 7a compare the total amount and spatial distribution of emissions 556 between JS-prior and MEIC-prior in April, respectively. The total BC emissions of 557 southern Jiangsu city cluster in JS-prior were 21% lower than those in MEIC-prior. In 558 JS-prior, as shown in Figure 7a, the emissions from some industrial plants were 559 extremely larger than those in MEIC-prior, while the emissions in urban areas were 560 found smaller. Both inventories indicated extremely small contribution from power 561 generation. BC emissions from industry sector were calculated at 1.34 Gg in JS-prior, 0.22 Gg smaller than MEIC-prior. Emissions from industry in MEIC-prior were 562 563 calculated based on regional average of emission factors and allocated according to 564 spatial distribution of GDP. The method would possibly result in underestimation in 565 emissions from big industrial plants but overestimation in urban areas. Emissions 566 from residential sources in JS-prior were close to those in MEIC-prior as similar 567 methodology was applied for the sector in the two inventories. BC emissions from transportation in MEIC-prior (0.85 Gg) were twice of those in JS-prior (0.42 Gg) 568 569 attributable probably to the application of different emission factors. For on-road 570 transportation, the emission factors in JS-prior were calculated with CORPERT model 571 (EEA, 2012; Zhou et al., 2017) while they were obtained from available domestic 572 measurements in MEIC-prior.

573 Simulation Case 4 was determined using MEIC-prior in CTM. As shown in

21 / 63





574 Table 3, the hourly average of BC concentrations at NJU was simulated at 2.49 μ g/m³ for April 2015 in Case 4, close to 2.38 µg/m³ simulated with JS-prior (Scenario B). At 575 PAES, however, application of MEIC-prior in CTM resulted in much larger 576 577 concentration than JS-prior (5.13 versus 2.98 μ g/m³), indicating again that 578 MEIC-prior would overestimate the emissions in urban area. Following the top-down 579 approach described in Section 2.2, we developed Case 5, using MEIC-prior instead of 580 JS-prior as the initial input of emission data in CTM. The scaling factors of emissions 581 from industry, residential sources and transportation were respectively calculated at 582 0.15, 1.30 and 0.25 through multiple regression model, and the top-down estimate in 583 BC emissions (mentioned as MEIC-posterior hereafter) were calculated at 0.75 Gg in 584 April 2015, close to 0.78 Gg in the JS-posterior (Figure 6). The differences in the 585 emissions from industry and transportation between JS-posterior and MEIC-posterior were 0.06 and 0.07 Gg, respectively, much smaller than those between JS-prior and 586 587 MEIC-prior. Besides the total amount, differences in spatial distribution in industry plants and urban area between the top-down estimates (JS-posterior and 588 589 MEIC-posterior) were also significantly reduced compared to those between 590 bottom-up estimates (JS-prior and MEIC-prior), as shown in Figure 7b. Figure 8 591 illustrates the scatterplots of the simulated BC concentrations from bottom-up and 592 top-down inventories at NJU (Figure 8a) and PAES (Figure 8b). Using two bottom-up 593 inventories in CTM, bigger difference in simulated BC concentrations was found at 594 PAES compared to that at NJU, indicated by the slope (1.10) closer to 1 at NJU in Figure 8a. The correlation coefficients (R^2) between simulated BC concentrations 595 596 using JS-prior and MEIC-prior were 0.81 at NJU and 0.40 at PAES respectively. 597 Using two top-down estimates, the difference between simulated concentrations at 598 PAES was significantly reduced and the slope got much closer to 1 in Figure 8b. The correlation coefficients (R²) were enhanced to 0.94 and 0.87 at NJU and PAES, 599 600 respectively. To summarize, similar results from top-down constraint approach could 601 be obtained in emission level, spatial distribution, and CTM performance, even clear





602 difference existed in the initial bottom-up inventories. In other word, limited effect of

603 initial emission input was evaluated on the top-down estimate from the multiple

604 regression model.

605 4.4 Uncertainty analysis of the multiple regression model

606 As mentioned in Section 2.2, the assumption of near linearity between emissions 607 and concentrations is a principle of the multiple regression model, given the weak 608 chemistry reactivity of BC. The principle has been applied in previous studies to 609 constrain BC emissions (Fu et al., 2012; Kondo et al., 2011; Wang et al., 2013; Park et 610 al., 2003; Verma et al., 2017). In the actual fact, however, processes other than 611 chemical reaction, e.g., precipitation or wet deposition, impact the linearity. Therefore, 612 the near-linear assumption needs to be justified, and the uncertainty of the 613 methodology could then be evaluated.

614 Sensitivity analysis was conducted to assess the rationality of brute-force method 615 described in Section 2.3, in which emissions of given sector were zeroed out to 616 determine their contribution to the ambient concentrations. As summarized in Table 617 S4 in the supplement, we first calculated the ratio of simulated wet deposition to 618 emissions by month for NJU, PAES and the whole southern Jiangsu city cluster with 619 JS-prior (Scenario B) and JS-posterior (Case 1), respectively. July and October were 620 identified as the months with the most and least impact from precipitation, suggested 621 by the largest and smallest ratio, respectively. Two sensitivity simulations were then 622 conducted for the selected two months, in which doubled and halved emissions (i.e., 623 200% and 50% of emissions in JS-prior, respectively) were used in CTM, and the simulated concentrations were then compared to those with JS-prior (i.e., Scenario B). 624 625 Figures 9 and 10 illustrate the linear correlations of the simulated concentrations in 626 these two sensitivity cases and the base scenario (Scenario B) at NJU and PAES, 627 respectively. As can be seen in all the panels, the fraction of change in simulated 628 monthly average concentration ($F_{conc.}$) was close to that of emission change ($F_{emis.}$), 629 i.e., the ratio of $F_{emis.}$ to $F_{conc.}$ was around 1.0, within a range of ±10%. Similar ratio of





change in emissions ($\triangle E$) to that in simulated average concentration ($\triangle C$) was obtained for each month and site as well. The results thus suggested that the impact of non-linearity between emissions and concentrations was limited, no matter the precipitation was strong or not. As the top-down constrained emissions (JS-posterior) were 50% smaller than the bottom-up estimates (JS-prior), the relative change was far beyond the uncertainty from non-linearity ($\pm 10\%$), implying the improvement of the top-down approach on emission estimation.

637 Many studies have reported the difficulty in precipitation simulation with WRF 638 (Annor et al., 2017; Liu et al., 2018; Yu et al., 2011; Yang et al., 2014; Kaewmesri, 639 2018). In this study, the observed ground precipitation at Lukou, Liyang and Shanghai 640 stations (see Figure 1 for locations) was compared with the simulated one to evaluate 641 the WRF performance for precipitation modeling. As shown in Figures S8-11 in the supplement, the model could capture the dates of precipitation, but it generally 642 643 overestimated the amount. Similar results were found in previous studies that WRF overestimated precipitation at fine spatial resolution (Politi et al., 2018; Kotlarski et 644 645 al., 2014; García-Díez et al., 2015). Improvement in physics parameterization 646 schemes in WRF will help better understanding the wet deposition of BC through 647 simulation. To further evaluate the effect of wet deposition on emission constraining, 648 we conducted an extra Case 6, in which the data influenced by simulated wet 649 deposition (i.e., the periods with simulated wet deposition at hourly basis) were excluded in the top-down approach. The new scaling factors $\beta_1 - \beta_4$ estimated from the 650 651 multiple regression model were summarized in Table 6. By applying β_1 '- β_4 ' in Eq. (2), the top-down estimates of BC emissions in Case 6 were calculated at 13.7 Gg, and the 652 653 emissions by sector and month were illustrated in Table 7, together with the relative 654 deviation (RD) compared to emissions in Case 1(JS-posterior). The relative deviations 655 of monthly total emissions between Case 6 and Case 1 were less than 5%, with an 656 exception of July at 14%, and that for annual total was 2.6%. Larger relative 657 deviations were found for given sources, e.g., residential in January and transportation





658 in July. The deviations, therefore, were much smaller than that between the emissions 659 in JS-prior and JS-posterior. We consequently applied CTM to evaluate the model 660 performance with the emissions in Case 6 for July. Illustrated in Table 8 were the simulated BC concentrations and the statistic indicators obtained through comparisons 661 with observation at the two sites. As suggested by the NME and R values, little 662 663 improvement on CTM performance was achieved with the emissions in Case 6, 664 compared to those with Case 1 (Table 2). The impact of simulated wet deposition on 665 the top-down approach was thus expected to be moderate in this work.

As the simulated wet deposition varied from the reality to some extent and the 666 667 impact of precipitation along the transport was not excluded in Case 6, we selected 668 July to conduct a Case 7, in which the data influenced by accumulative precipitation along the back trajectories at the two sites were excluded in the multiple regression 669 model. The merged high-quality precipitation measured by the Tropical Rainfall 670 671 Measuring Mission (TRMM) satellite instrument was adopted for wet deposition screening, with a temporal resolution of 3 h and a spatial resolution of $0.25^{\circ} \times 0.25^{\circ}$. 672 673 We used the Hybrid Single Particle Lagrangian Integrated Trajectory (HYSPLIT, 674 version 4.9) model (http://www.ready.noaa.gov) to calculate the 48 h back trajectories 675 of the air masses arriving at NJU and PAES. The back trajectories were calculated 676 every 3 hour for July with the simulated layer heights of 50, 100 and 500 m above the 677 ground and the time step of 3 h (the same as the temporal resolution of TRMM). The 678 hourly accumulative precipitation along the 48 h back trajectories at two sites were 679 then calculated to determine the BC-CO data pairs influenced by precipitation, given 680 the little effect of precipitation on CO. Figure 11 illustrates the changes in the $\triangle BC/$ 681 \triangle CO ratio observed at two sites for different accumulated precipitation intervals. At 682 NJU, the $\triangle BC / \triangle CO$ ratio of air masses receiving less than 3 mm accumulated 683 precipitation was significantly larger than that of air masses receiving more than 3 684 mm, and the analogue number was 5 mm at PAES. In Case 7, therefore, we excluded 685 the BC-CO data pairs receiving more than 3 mm and 5 mm accumulated precipitation





686 along their trajectories within the last 48 h at NJU and PAES, respectively, in the 687 multiple regression model. It minimized the effect of wet deposition while retained 688 sufficient data points for the statistical significance. Figure 12 shows the simulated 689 wet deposition in Case 6 and the accumulated precipitation in Case 7 for July to 690 compare the data selection in the two cases. In Case 6, the number of data points were reduced to 65% of Case 1 after data screening, and over 500 samples at the two sites 691 692 were available for the multiple regression model. In Case 7, only 31% of data points 693 remained. The periods excluded in Case 7 contained those in Case 6, implying a 694 stricter data screening to eliminate the effect of precipitation.

695 Table 9 shows the scaling factors estimated from the multiple regression model 696 in Case 7, and no big changes were found compared to the scaling factors for July in 697 Case 6 (Table 6). Consequently, the emissions by sector and total emissions in Case 7 698 were close to those in Case 6 (Table 7). The relative deviation of total emissions in 699 July between Case 7 and Case 1 (RD in Table 9) was 13%, and those for residential 700 and transportation were larger. The influence of precipitation was again indicated 701 insignificant, as the deviation was much smaller than that between the estimates 702 obtained from the bottom-up and top-down methods. Moreover, the CTM 703 performance based on Case 7, indicated by NMB and NME, was found similar to that 704 based on Case 6, implying the small effect of precipitation screening on simulation. 705 Even excluding the influence of precipitation along the back trajectories, the Sig. for 706 residential sources in Case 7 was still much larger than 0.05 (Table 9), suggesting 707 more efforts on quantification of emissions for this highly uncertain source category.

708 5 Conclusions

Monthly top-down estimates of BC emissions were derived from a multiple regression model that integrated CTM and hourly BC concentrations from two ground observation sites in southern Jiangsu city cluster. The annual emissions from top-down approach (JS-posterior) were estimated at 13.4 Gg for 2015, 50.3% smaller than those in bottom-up emission inventory that did not include the improved





714 emission controls in recent years (JS-prior), implying the effectiveness of air pollution 715 prevention measures on emission abatement. Application of JS-posterior in CTM 716 reduced the deviations between simulations and observations at two ground sites 717 effectively, especially at the urban site PAES. To evaluate the effects of observation 718 data on top-down estimate, two more cases in which observation data of only one site 719 (NJU) and observation data at both sites with their spatial representativeness 720 differentiated were applied to constrain the emissions, respectively. Best CTM 721 performance was found for the third case, indicating that inclusion of more ground 722 measurements with better spatiotemporal coverage in the city cluster would improve 723 the understanding of spatial distributions of BC emissions. In addition, top-down 724 estimates were derived from various bottom-up inventories, and the differences in 725 emission amount, spatial distribution and CTM performance between the constrained emission estimates were significantly reduced compared to those between the 726 727 bottom-up inventories. The results implied that changes in initial emission input in the 728 regression model and CTM had limited effect on the top-down estimation. Finally, the 729 assumption of near-linearity between emissions and concentrations was justified, and 730 the influence of wet deposition on the estimated emissions was evaluated to be 731 moderate. This work demonstrated that top-down approach based on ground 732 observations and CTM could capture the fast changes in BC emissions attributed to 733 tightened pollution control policy at a city cluster scale. To further reduce uncertainty of the approach, more ground measurements with sufficient temporal resolution 734 735 would be recommended at other regions in the city cluster. Data from other sources, such as aerosol optical depth from satellite observation, could also be included to 736 737 improve the spatial and temporal distributions of emission estimates.

738

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945 **Figure captions**

- 946 Figure 1. Modeling domain and locations of two observation sites and three
- 947 meteorological stations.
- 948 Figure 2. The monthly (left axis) and annual emissions (right axis) by sector for
- 949 southern Jiangsu 2015 in JS-prior and JS-posterior (unit: Gg).
- 950 Figure 3. The seasonal variation of BC emissions by source (a) and total emissions (b)
- 951 in JS-prior, JS-posterior and MEIC-prior.
- 952 Figure 4. The observed and simulated hourly BC concentrations at NJU using JS-prior
- 953 and JS-posterior for January (a), April (b), July (c) and October (d) in 2015.
- 954 Figure 5. The same as Figure 4 but at PAES.
- 955 Figure 6. BC emission estimates by source of JS-prior, MEIC-prior, JS-posterior, and
- 956 MEIC-posterior in April 2015 in southern Jiangsu.
- 957 Figure 7. The spatial distributions of the deviations (JS-MEIC) between JS-prior and
- 958 MEIC-prior (a) and those between JS-posterior and MEIC-posterior (b).
- 959 Figure 8. The scatter plots of the simulated BC concentrations using JS inventories
- 960 versus those using MEIC at NJU (a) and PAES (b).
- Figure 9. The correlation between the simulated BC concentrations with JS-prior and those with doubled (a and c) or halved emissions in JS-prior (b and d) in July (a and b) and October (c and d) at NJU. F_{emis} and $F_{conc.}$ indicate respectively the fraction of changed emissions and that of changed simulated monthly average concentrations between sensitivity and base simulation (Scenario B). $\triangle E$ and $\triangle C$ indicated the change in emissions and that in simulated monthly average concentrations, respectively.
- 968 Figure 10. The same as Figure 9 but at PAES.





- 969 Figure 11. The $\triangle BC / \triangle CO$ ratio at NJU (a) and PAES (b) separated by different
- 970 accumulated precipitation along the back trajectories during 48 h. The number of
- 971 remaining data points is also given.
- 972 Figure 12. The wet deposition in Case 6 and accumulated precipitation in Case 7 at
- 973 NJU (a) and PAES (b). The number of remaining data points is also given.





974 Tables

Table 1. The scaling factors and statistical indicators from the multiple regression model for estimation of JS-posterior.

Month	Sector	Scaling factor	t ^a	Sig. ^b	VIF ^c	Sig. ^d
	Industry (β_2)	0.42	2.65	0.01	1.76	
January	Residential (β_3)	1.31	3.67	0.00	2.37	0.00
	Transportation (β_4)	0.79	2.23	0.03	2.72	
	Industry (β_2)	0.22	0.96	0.34	2.65	
April	Residential (β_3)	0.58	1.63	0.11	4.62	0.00
	Transportation (β_4)	0.67	2.21	0.03	4.19	
	Industry (β_2)	0.35	3.09	0.00	2.09	
July	Residential (β_3)	0.39	0.95	0.34	2.95	0.00
-	Transportation (β_4)	0.55	2.20	0.03	3.46	
	Industry (β_2)	0.34	1.92	0.06	1.53	
October	Residential (β_3)	1.52	4.12	0.00	2.20	0.00
_	Transportation (β_4)	0.74	2.80	0.01	2.65	

977 Note: The criteria for the statistical significance of the model: a: t>2, b: Sig.<0.05, and

978 c: VIF<10, d: the overall significance.





0:10	Deremotor	Ja	nuary	Ł	April	-	July	ŏ	stober	-A1	nnual
olle	raiaincici	JS-prior	JS-posterior								
	Average SIM (µg/m ³)	5.97	5.50	2.38	1.82	1.99	1.29	2.80	2.42	3.44	2.82
	Average OBS (µg/m ³)	5.44	5.44	2.69	2.69	2.65	2.65	3.96	3.96	3.83	3.83
NJU	NMB (%)	8.35	-0.08	-16.02	-32.40	-23.09	-51.32	-29.20	-39.01	-10.16	26.43
	NME (%)	37.83	35.54	42.31	38.61	49.62	57.49	40.52	43.06	41.15	44.16
	R	0.67	0.66	0.34	0.43	0.36	0.31	0.42	0.48	0.67	0.69
	Average SIM (μg/m ³)	6.46	5.91	2.98	1.95	2.61	1.63	3.19	2.88	3.39	2.57
	Average OBS (µg/m ³)	2.80	2.80	1.70	1.70	1.51	1.51	3.62	3.62	2.48	2.48
PAES	NMB (%)	151.93	134.59	61.57	14.73	72.17	8.28	-12.01	-20.48	36.67	3.54
	NME (%)	155.53	139.50	73.18	42.87	92.74	42.37	43.10	40.80	72.00	57.55
	R	0.38	0.38	0.64	0.53	0.35	0.37	0.57	0.72	0.38	0.45

Q allu respectively. INMID UUSCI VALIUII, ITUILI SIIIIUIAUUII AIIU rinsol Note: SIM and UBS indicated the 900

equations (P and O indicated the results from modeling prediction and observation, respectively): 981

982
$$NMB = \frac{\sum_{i=1}^{n} (P_i - O_i)}{\sum_{i=1}^{n} O_i} \times 100\%; . NME = \frac{\sum_{i=1}^{n} |P_i - O_i|}{\sum_{i=1}^{n} O_i} \times 100\%$$

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983	Table 3. Statistical indicators for observed and simulated BC concentrations in
984	different cases in April 2015 at NJU and PAES.

Site	Parameter	Scenario B	Case1	Case2	Case3	Case4	Case5
	Average SIM (µg/m ³)	2.38	1.82	2.27	2.06	2.49	1.78
	Average OBS (µg/m ³)	2.69	2.69	2.69	2.69	2.69	2.69
NJU	NMB (%)	-16.02	-32.40	-21.59	-23.50	-7.46	-33.95
	NME (%)	42.31	38.61	32.47	32.64	41.58	38.94
	R	0.34	0.43	0.49	0.49	0.40	0.46
	Average SIM ($\mu g/m^3$)	2.98	1.95	2.45	2.01	5.13	2.29
	Average OBS (µg/m ³)	1.70	1.70	1.70	1.70	1.70	1.70
PAES	NMB (%)	61.57	14.73	49.86	18.02	201.35	34.71
	NME (%)	73.18	42.87	61.59	39.62	201.56	47.73
	R	0.64	0.53	0.63	0.66	0.65	0.59





Site	Sector	Scaling factor	t	Sig.	VIF
	Industry (β_2)	0.42	1.71	0.09	2.03
NJU	Residential (β_3)	0.95	2.50	0.01	2.52
	Transportation (β_4)	0.65	2.13	0.03	2.66
	Industry (β_2)	0.19	3.46	0.00	1.44
PAES	Residential (β_3)	0.36	1.89	0.06	1.44
	Transportation(β_4)	0.65	-	-	-

Table 4. The scaling factors and statistical indicators from the multiple regression model in Case 3.





Table 5. BC emissions from Nanjing and Suzhou–Wuxi–Changzhou-Zhenjiang
city cluster in different cases in April 2015 (Gg).

Casa	C a at a m	Nauliua	Suzhou-Wuxi-Changzhou	Southern
Case	Sector	Nanjing	-Zhenjiang	Jiangsu
	Power	0	0.01	0.01
	Industry	0.21	1.13	1.34
Scenario B	Residential	0.08	0.24	0.32
	Transportation	0.12	0.30	0.42
	Total	0.41	1.68	2.09
	Power	0	0.01	0.01
Casa 1	Industry	0.05	0.25	0.30
Case I	Residential	0.04	0.14	0.19
	Transportation	0.08	0.20	0.28
	Total	0.17	0.60	0.78
	Power	0	0.01	0.01
	Industry	0.09	0.47	0.56
Case 2	Residential	0.07	0.23	0.30
	Transportation	0.08	0.20	0.27
	Total	0.24	0.91	1.14
	Power	0	0.01	0.01
	Industry	0.04	0.47	0.51
Case 3	Residential	0.03	0.23	0.26
	Transportation	0.08	0.20	0.27
	Total	0.15	0.90	1.05





989	Table	6.	The	sca	ling	factors	and	statistical	indicators	from	the	multiple
000					~	-						

990 regression model in Case 6.

Month	Sector	Scaling factor	ť	Sig. ^b	VIF ^c	Sig. ^d
	Industry (β_2 ')	0.41	2.17	0.03	1.71	
January	Residential (β_3 ')	1.53	3.48	0.00	2.29	0.00
	Transportation (β_4 ')	0.73	1.65	0.10	2.66	
	Industry (B2')	0.24	0.92	0.36	1 91	
April	Residential (β_3)	0.51	1.32	0.19	3.29	0.00
1	Transportation (β_4 ')	0.70	2.12	0.03	3.03	
	Industry (β_2 ')	0.38	4.43	0.00	1.43	
July	Residential (β_3')	0.34	0.82	0.41	2.52	0.00
-	Transportation (β_4 ')	0.74	3.55	0.00	2.25	
	Industry (B. ')	0.33	1.00	0.32	1 44	
October	Residential (β_2)	1.36	2.61	0.02	1.44	0.00
	Transportation (β_4 ')	0.72	1.89	0.01	2.02	0.00

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992 Note: The criteria for the statistical significance of the model: a: t>2, b: Sig.<0.05, and

993 c: VIF<10, d: the overall significance.

> -0.6% 5.4% 2.6%

> 3.6 3.9 13.5

-10.2% -3.0% -4.2%

0.4 0.3 1.2

-13.7% 34.4% 13.6%

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-13.1% 4.3% 2.3%

0.2 0.3 0.8

16.7% -8.2% 2.4%

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Fransportation

Sum





nonthly and annual emissions by se 1 (RD: Case 6–Case 1)/Case 1). January January Case 6 RD 0.0 0.0% 0.0 y 0.6 -2.4% 0.3 ial 0.5 16.7% 0.2	sector for southern Jiangsu 2015 in Case 6 (unit: Gg) and the relative $d\varepsilon$	April July October Annual	RD Case 6 RD Case 6 RD Case 6 RD	0.0% 0.0 0.0% 0.0 0.0% 0.0 0.0%	9.9% 0.6 9.2% 0.5 -0.3% 6.0 3.1%	-13.1% 0.1 $-13.7%$ 0.4 $-10.2%$ 3.6 $-0.6%$
se 1 (RD: Case 6–Cas January Case 6 RI Case 6 RI 0.0 0.0 y 0.6 -2.4 ial 0.5 16.7	emissions by sector e 1)/Case 1).	April	D Case 6 1	% 0.0 0	!% 0.3 9	7% 0.2 -1
	nonthly and annual ese 1 (RD: Case 6–Cas	January	Case 6 RI	0.0 0.0	y 0.6 -2.4	ial 0.5 16.7

996 997 999 44 / 63









1002 Table 8. Statistical indicators for the observed and simulated BC concentrations

1003 in July 2015 at NJU and PAES in Case 6 and Case 7.

	Parameter	Case 6	Case 7
	Average SIM ($\mu g/m^3$)	1.40	1.41
	Average OBS $(\mu g/m^3)$	2.65	2.65
NJU	NMB (%)	-47.41	-46.72
	NME (%)	54.88	54.44
	R	0.33	0.33
	Average SIM ($\mu g/m^3$)	1.76	1.76
	Average OBS $(\mu g/m^3)$	1.51	1.51
PAES	NMB (%)	16.87	16.65
	NME (%)	44.46	42.71
	R	0.36	0.39

1005 Note: SIM and OBS indicated the results from simulation and observation,
1006 respectively. NMB and NME were calculated using following equations (P and O
1007 indicated the results from modeling prediction and observation, respectively):

1008
$$NMB = \frac{\sum_{i=1}^{n} (P_i - O_i)}{\sum_{i=1}^{n} O_i} \times 100\%$$
 $NME = \frac{\sum_{i=1}^{n} |P_i - O_i|}{\sum_{i=1}^{n} O_i} \times 100\%$





1027	Table 9. The scaling factors and statistical indicators from the multiple
1028	regression model in Case 7. The emissions by sector for southern Jiangsu 2015
1029	July in Case 7 (unit: Gg) and the relative deviations (RD) compared to Case 1

1030 (**RD: Case 7–Case 1**)/Case 1) are also shown in table.

Sector	Scaling factor	ť	Sig. ^b	VIF ^c	Sig. ^d	Emissions	RD
Power						0.0	0.0%
Industry (β_2 ')	0.38	2.38	0.02	1.31		0.5	9.5%
Residential (β_3 ')	0.31	0.31	0.75	2.31	0.00	0.1	-20.6%
Transportation (β_4 ')	0.75	1.8	0.07	1.95		0.4	36.4%
Sum						1.0	13.4%
)21							

1031

1032 Note: The criteria for the statistical significance of the model: a: t>2, b: Sig.<0.05, and

1033 c: VIF<10, d: the overall significance.

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1035





1037 Figure 1











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1040

Figure 2





1050 Figure 3











1051 Figure 4

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1058 Figure 8







1060 Figure 9







































1073 Figure 12

