

## Response to Referee #1 (Dr. Lucas Henneman)

### General comments:

Seo et al. present an evaluation of trends at various frequencies in Seoul, Korea. They use meteorological detrending techniques that apply the KZ filter and multiple linear regressions along with a simplified continuity formulation to attribute long term changes to local and transported sources.

Overall, the analysis is vigorous, and the conclusions appear relevant to future policy decisions. Other than a deeper discussion of the continuity approach and the minor suggestions below, I believe the overall analysis to be sound and an important contribution to the published literature.

We appreciate the reviewer for careful reading and helpful comments that improve the quality of the manuscript. As indicated in the following point-by-point responses, we have incorporated the reviewer's comments and suggestions into the revised manuscript. Each response to reviewer colored in blue and changes in the manuscript colored in red.

### Specific comments:

The meteorological detrending approach has been applied in similar fashion in previous applications, but I believe the continuity-based derivation of local/transported emissions on the long-term trend to be innovative. My major comments for this manuscript revolve around the development and discussion of this approach—I have listed questions here that I hope will inspire the authors to consider and discuss the approach in more detail. While the interpretation of the results appears to fit within scientific understanding of current atmospheric processes and emissions trends, I think the manuscript would be greatly improved by further development of this method.

1. Please clarify that local/long term transported emissions to long-term trends only are available to Seoul-wide data, not on individual sites.

As the reviewer pointed out, the evaluation method for the changes in local emissions ( $X_{LT}^{emis(L)}$ ) and the changes in transport of regional emissions ( $X_{LT}^{emis(T)}$ ) we suggested here is unavailable for individual air quality monitoring sites because the horizontal advection was obtained by using the winds at a weather station and the horizontal gradient of air pollutants measured in the surrounding area. To clarify this, we add the following sentence at the end of L24 on p.7.

Note that the above method to evaluate the changes in  $X_{LT}^{emis(T)}$  and  $X_{LT}^{emis(L)}$  is applicable not to data from an individual site but to data from the wide area, because of the requirement of horizontal gradient term in Eq. (10).

What are the implications of the distance scales, numbers of monitors available, and the choice of centering the Cartesian coordinates at the weather station?

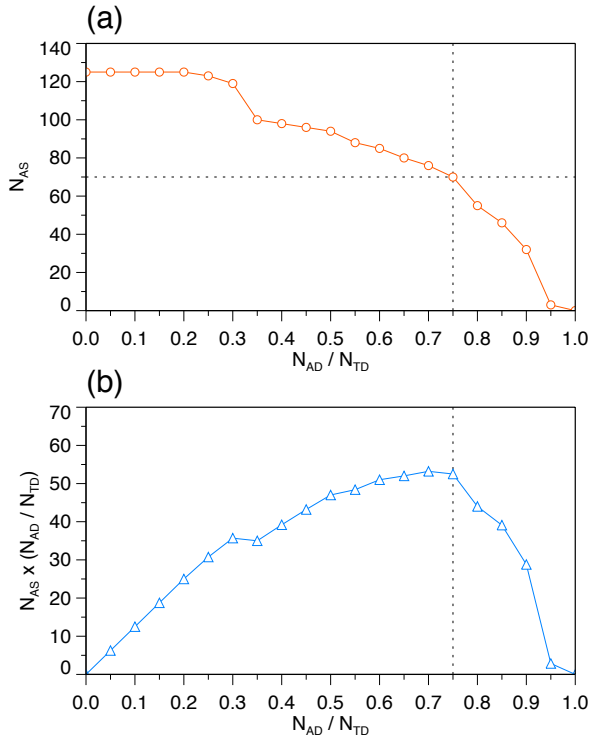
As we described in Sect. 2, we selected 18 air quality monitoring sites within the Seoul area and took average the daily data from the selected sites to produce the daily city-average data. Since the selected 18 monitoring sites in Seoul are located within the area of ~15 km radius from the weather station (37.571°N, 126.966°E; Fig. 1a), we can assume a 30 km grid box centered at the weather station, which is representative for the daily air pollution data in Seoul. If we try to calculate the horizontal advection of the long-term components ( $-\vec{V}_{LT} \cdot \nabla X_{LT}$ ) at the center of the coordinate (the  $(i, j)$ th grid point), the long-term component of  $u$ - and  $v$ -wind data at the  $(i, j)$ th grid point and  $X_{LT}$  at the  $(i-1, j)$ th,  $(i+1, j)$ th,  $(i, j-1)$ th, and  $(i, j+1)$ th grid points will be required as follows:

$$(-\vec{V}_{LT} \cdot \nabla X_{LT})_{(i,j)} = -u_{LT(i,j)} \frac{X_{LT(i+1,j)} - X_{LT(i-1,j)}}{2\Delta x} - v_{LT(i,j)} \frac{X_{LT(i,j+1)} - X_{LT(i,j-1)}}{2\Delta y}$$

where  $\Delta x$  and  $\Delta y$  are 30 km, and thus the total area we need for the calculation will be 90 km  $\times$  90 km (3 grid boxes each for each  $x$ - and  $y$ -direction). Although  $\nabla X_{LT}$  was obtained by using the linear regression method

in this study because of scattered monitoring sites over the Seoul Metropolitan Area (SMA), we partly adopted the finite difference concept to set up the center of coordinates (to the Seoul weather station) and specified the area for the available sites (as a 50 km radius range from the center of coordinates).

The number of monitoring sites used here was decided by setting a criterion for data availability at each site, which is a ratio of the available number of data ( $N_{AD}$ ) to the total number of data ( $N_{TD}$ ). As shown in Fig. S8a, the number of available sites ( $N_{AS}$ ) is rapidly decreased where the data availability is larger than 75% ( $N_{AD} / N_{TD} > 0.75$ ). Because both  $N_{AS}$  and  $N_{AD} / N_{TD}$  are important, we examined the multiplication of  $N_{AS}$  and  $N_{AD} / N_{TD}$  as a score and found that the score was high enough when the criterion of data availability was 75% (Fig. S8b). Fig. S8 was now added to the supplement.



**Figure S8. (a) Numbers of available air quality monitoring sites ( $N_{AS}$ ) within the area of 50 km radius from the Seoul weather station and (b)  $N_{AS}$  multiplied with ratios of the available number of data ( $N_{AD}$ ) to the total number of data ( $N_{TD}$ ). Vertical dotted lines at  $N_{AD} / N_{TD}$  of 0.75 show the data availability of 75% and a horizontal dotted line at  $N_{AS}$  of 70 represents the number of air quality monitoring sites used for obtaining the horizontal gradient of long-term components in this study.**

As the distance scale used here is on the order  $10^1$  km, can we assume the origin of the transported pollutants to be a certain distance away (e.g., on the order of  $10^2$  km)?

In this study, we investigated the changes in transport of regional emissions ( $\frac{\partial}{\partial t} X_{LT}^{emis(T)}$ ) using the data within the area of the radius of 50 km from the Seoul weather station. Therefore, the interpretation of the result should be limited within the analysis area. Considering that the interpretable area almost agrees with the Seoul Metropolitan Area (SMA), the origins of the transported pollutants from the outside of Seoul can not only be the outside of the SMA (such as the Chinese eastern coasts) but also be the inside of the SMA (such as industrial areas in the southwestern SMA).

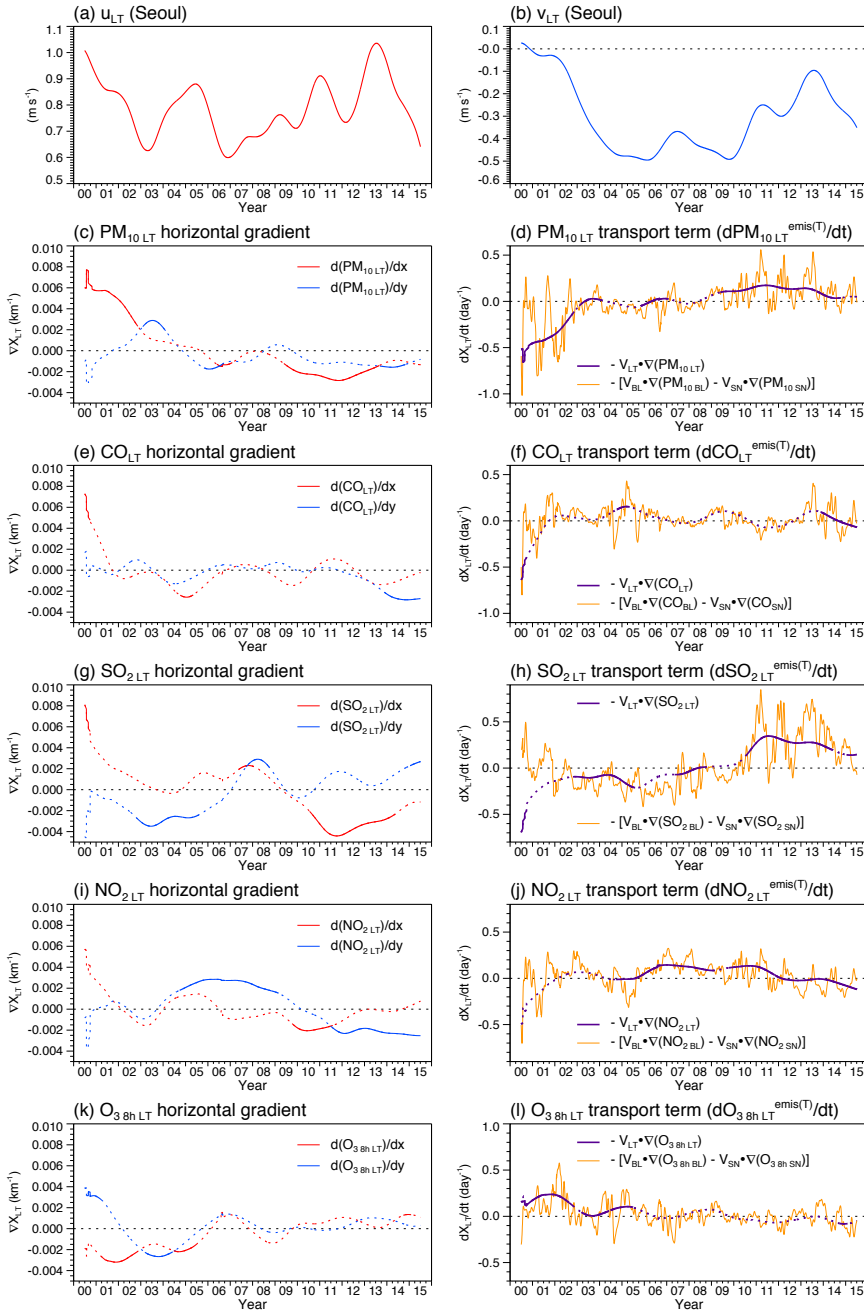
2. What is the interpretation of the high nonlinearities in the meridional gradients (Figure S5)?

The nonlinearity of the slope in the scatter plot of  $X_{BL}$  versus  $y$ -axis ( $r = -0.385$  and  $p < 0.001$ ; Fig. S5b) mainly arises from the overall zonal gradient of  $X_{BL}$  ( $\frac{\partial X_{BL}}{\partial x} = -0.0017 \text{ km}^{-1}$ ; Fig. S5c). Similarly, the nonlinearity of the slope in the scatter plot of  $X_{BL}$  versus  $x$ -axis ( $r = -0.202$  and  $p = 0.094$ ; Fig. S5c) mostly

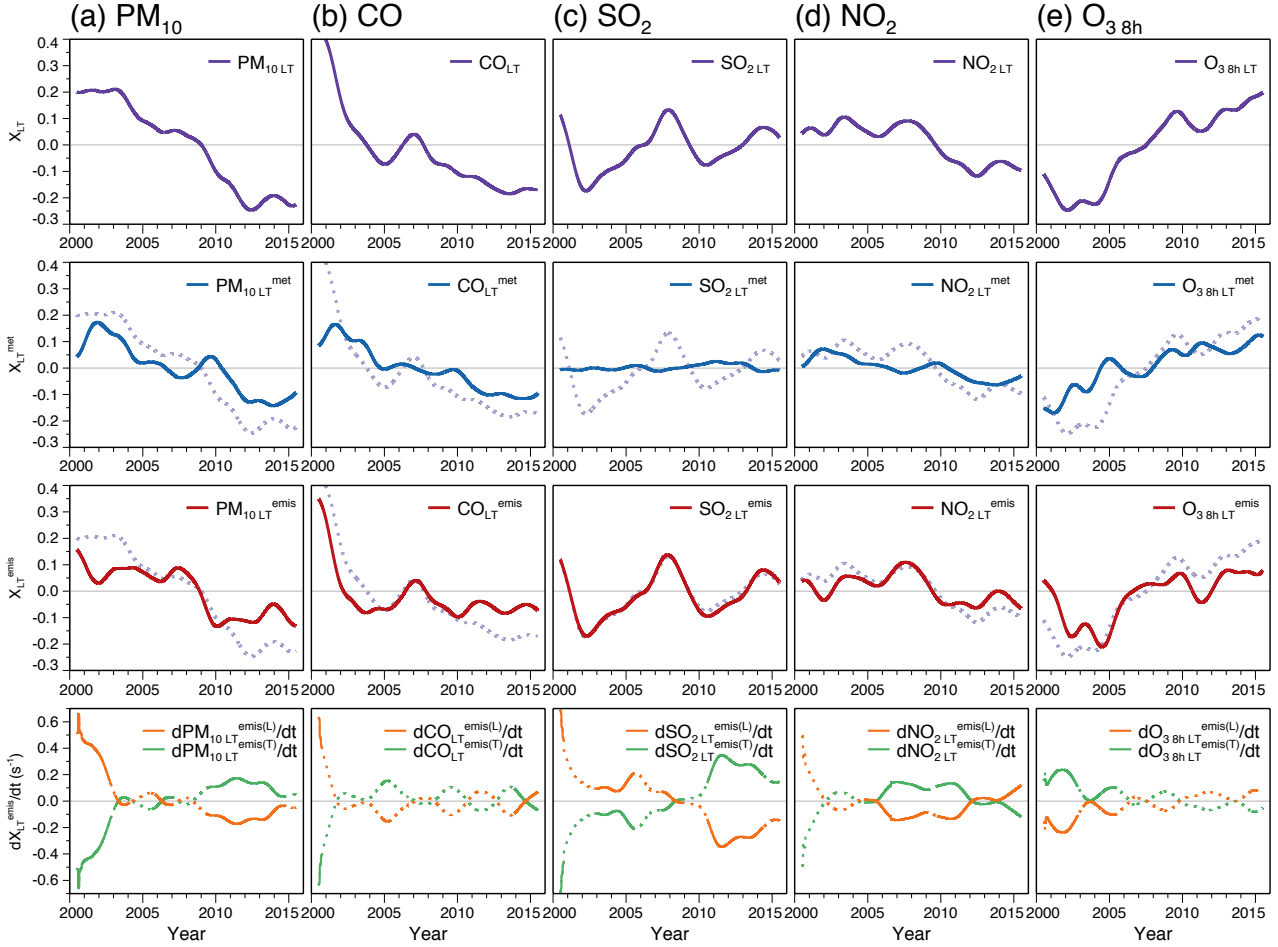
results from the meridional gradient of  $X_{BL}$  over the analysis area ( $\frac{\partial X_{BL}}{\partial y} = -0.0019 \text{ km}^{-1}$ ; Fig. S5b). Since  $\left| \frac{\partial X_{BL}}{\partial y} \right| > \left| \frac{\partial X_{BL}}{\partial x} \right|$ , the nonlinearity of the slope in Fig. S5c ( $X_{BL}$  vs.  $x$ ) is larger than that in Fig. S5b ( $X_{BL}$  vs.  $y$ ).

3. Are the meridional slopes generally statistically significantly different from zero?

The zonal and meridional slopes for the long-term components are statistically significant when the slope is high enough. We modified Fig. S6 and Fig. 5 to reveal the statistically significant ( $p < 0.05$ ) slope of each long-term component ( $X_{LT}$ ) and changes in local emissions ( $\frac{\partial}{\partial t} (X_{LT}^{\text{emis(L)}})$ ) and transport ( $\frac{\partial}{\partial t} (X_{LT}^{\text{emis(T)}})$ ).



**Figure S6.** Long-term component of (a) zonal wind ( $u_{LT}$ ) and (b) meridional wind ( $v_{LT}$ ) at the Seoul weather station. Zonal gradients ( $\partial X_{LT}/\partial x$ , red lines) and meridional gradients ( $\partial X_{LT}/\partial y$ , blue lines) of the long-term components and transport terms ( $-\vec{V}_{LT} \cdot \nabla X_{LT}$ , violet lines) by long-term components of horizontal winds ( $\vec{V}_{LT} = (u_{LT}, v_{LT})$ ) for (c–d)  $\text{PM}_{10}$ , (e–f)  $\text{CO}$ , (g–h)  $\text{SO}_2$ , (i–j)  $\text{NO}_2$ , (k–l)  $\text{O}_3$  8h. Solid lines in horizontal gradients and transport terms indicate that the gradients obtained by linear regression are statistically significant at the 95% level or higher ( $p < 0.05$ ).



**Figure 5: Long-term components those are unadjusted for the meteorological variables ( $X_{LT}$ ; violet lines), meteorology-related ( $X_{LT}^{met}$ ; blue lines) and emission-related ( $X_{LT}^{emis}$ ; red lines) long-term components, and contributions of local emissions ( $\frac{\partial}{\partial t} X_{LT}^{emis(L)}$ ; orange lines) and transport of regional emissions ( $\frac{\partial}{\partial t} X_{LT}^{emis(T)}$ ; green lines) to the long-term trends of (a)  $PM_{10}$ , (b)  $CO$ , (c)  $SO_2$ , (d)  $NO_2$ , and (e)  $O_3$  in Seoul. Solid lines in  $\frac{\partial}{\partial t} X_{LT}^{emis(L)}$  and  $\frac{\partial}{\partial t} X_{LT}^{emis(T)}$  show that the horizontal gradients of  $X_{LT}$  ( $-\vec{V}_{LT} \cdot \nabla X_{LT}$ ) obtained by linear regression are statistically significant at the 95% level or higher ( $p < 0.05$ ).**

Although the interpretable periods for each rate of changes in  $X_{LT}^{emis(L)}$  and  $X_{LT}^{emis(T)}$  are reduced based on the statistical significance test ( $p < 0.05$ ), important features we described in Sect. 4.3.2 still remain in the modified version of Fig. 5.

4. Are there limitations to this approach in regard to fewer available monitoring locations, spatial distributions of monitoring sites, etc.?

The most severe limitation of this approach arises from the high dependence on both number and spatial distribution of monitoring sites. If the number of sites is not enough for statistical analysis or the distribution of sites is biased from the center of the analysis area, the advection (transport) term of the tracer continuity equation must be unreliable. In addition, there must be wind data at the center of the analysis area.

High levels of missing data on certain days could severely impact calculated meridional slopes. Did the authors find evidence of this? If so, was anything done to correct for it?

In this study, we applied this approach to the long-term components of each air pollutant. Because of the iterative moving average process of the KZ-filter, together with the high data availability criterion for each site (75%), the number of missing data for the long-term components at each site was negligible. However, it

can be easily speculated that the data missing at several sites must affect the slope of data and cause abrupt changes in the advection term.

**Technical comments:**

Page 4, line 9: How were Asian Dust days identified?

We used daily records of the Asian Dust events, which were provided by the Korea Meteorological Administration (KMA) based on both naked eye observations (following the WMO recommendation) and PM<sub>10</sub> measurements (using the beta attenuation monitoring method). The KMA website is now cited in L9 on p.4 and included in the reference list as follows:

KMA (Korea Meteorological Administration): Asian Dust observation days, available at:  
<http://www.weather.go.kr/weather/asiandust/observday.jsp?type=2&stnId=108&year=2016&x=20&y=11>, (last access: 18 October 2018), 2018 (in Korean).

Throughout the manuscript (and especially in the Data section), please clarify the monitoring station being referred to, or whether the data is an average of all monitoring stations.

Because we aimed to analyzed city-scale air pollution variability in Seoul, we averaged daily data for selected 18 sites within the city and used in this study. To clarify, we added the following sentence to the first paragraph of Sect. 2 (L5 on p.4).

We averaged daily concentration data from the selected 18 sites in Seoul and utilized in this study.

Eqs. 2 & 4: I recommend changing the syntax of these equations slightly to improve clarification. The current form of, e.g.,  $KZ_{(m,p)}X(t)$  looks like the KZ term is being multiplied by  $X(t)$ . I recommend changing to  $KZ_{(m,p)}[X(t)]$  or similar.

We modified Eqs. (2), (4), and (7), and L2 on p.6, following the reviewer's suggestion.

Eqs. 10 & 11: Please include section (possibly in the supplement) with more detailed derivation steps.

We added detailed derivation steps for Eqs. (10) and (11) as Appendix S1 in the supplement.

Page 8, Line 15: I believe there is a typo in this sentence. Possibly it should be "...not balanced *against* each other in *the* short-term timescale"

Thanks for the correction. It was now corrected.

Page 8, Section 4: Please provide explanation for why the variances described by each of the trends do not sum to 100%.

Explained variances in Table 1 are identical to coefficients of determination ( $R^2$ ) between the original time series ( $X$ ) and each decomposed component ( $X_{ST}$ ,  $X_{SN}$ , and  $X_{LT}$ ), which represent how much of the variability of  $X$  is accounted for by the variability of each component. The explained variances of the decomposed time series should sum to 100%, if  $X_{ST}$ ,  $X_{SN}$ , and  $X_{LT}$  were completely independent of each other. However, because the decomposed components by the KZ filter method still have minor correlations among them albeit very weak, the explained variance of each component ( $R^2$  with  $X$  in percentage) are contributed not only by the variance itself but also by the covariances with other components.

In the revised version, to clarify this, we replaced the explained variances in the original version of Table 1 with the proportions of each variance and covariance to the total variances. The sum of variances of and covariances among the short-term, seasonal, and long-term components are exactly the same as the total variances of the original time series (see the formula in the revised Table 1).

Following the modified Table 1, “(~50–70%)” in L19 on p.8 is now “(~46–68%).” Also, the original version of Table 1 is modified to the  $R^2$  values and is now added to the supplement as Table S2.

**Table 1: Total variances ( $Var(X)$ ) of log-scale times series for Seoul average concentrations of  $PM_{10}$ ,  $SO_2$ ,  $NO_2$ ,  $CO$ , and  $O_3$  8h, and relative contributions (%) of variances of and covariances ( $Cov$ ) among each component to  $Var(X)$ . Daily data for the period of July 2000 to Jun 2015 that  $X_{LT}$  data is available were used.**

	$PM_{10}$	$SO_2$	$NO_2$	$CO$	$O_3$ 8h
$Var(X)^a$	0.2998	0.1345	0.1400	0.1540	0.4068
$Var(X_{ST})$	63.98%	48.92%	67.94%	46.19%	42.96%
$Var(X_{SN})$	22.07%	40.58%	23.86%	34.66%	47.06%
$Var(X_{LT})$	8.74%	4.50%	3.64%	14.32%	5.00%
$Cov(X_{ST}, X_{SN})$	2.91%	2.53%	2.20%	2.00%	2.17%
$Cov(X_{SN}, X_{LT})$	-0.29%	0.47%	0.09%	0.42%	0.32%
$Cov(X_{ST}, X_{LT})$	-0.01%	0.00%	-0.01%	0.00%	0.01%
$Var(X_{LT}^{emis})$	2.49%	4.88%	1.80%	4.97%	1.90%
$Var(X_{LT}^{met})$	2.85%	0.09%	1.08%	4.45%	1.52%
$Cov(X_{LT}^{emis}, X_{LT}^{met})$	1.70%	-0.23%	0.38%	2.45%	0.79%

<sup>a</sup> Values of variances of each log-scale concentration time series

$$Var(X) = Var(X_{ST}) + Var(X_{SN}) + Var(X_{LT}) + 2[Cov(X_{ST}, X_{SN}) + Cov(X_{SN}, X_{LT}) + Cov(X_{ST}, X_{LT})]$$

**Table S2. Coefficients of determination ( $R^2$ ) between each component and original time series ( $X$ ) for Seoul average concentrations of  $PM_{10}$ ,  $SO_2$ ,  $NO_2$ ,  $CO$ , and  $O_3$  8h. Daily data for the period of July 2000 to Jun 2015 that  $X_{LT}$  data is available were used.**

Components	$PM_{10}$	$SO_2$	$NO_2$	$CO$	$O_3$ 8h	Notes
$X_{ST}$	0.699	0.541	0.724	0.503	0.474	Short-term components
$X_{SN}$	0.276	0.468	0.287	0.397	0.522	Seasonal components
$X_{LT}$	0.081	0.055	0.038	0.152	0.057	Long-term components
$X_{LT}^{emis}$	0.064	0.054	0.030	0.124	0.043	Emission-related $X_{LT}$
$X_{LT}^{met}$	0.069	0.003	0.019	0.107	0.040	Meteorology-related $X_{LT}$

Page 9, line 22, suggest remove final word (“were”).

Thanks. It was now corrected.

Page 12, Line 9: Why do you use  $p < 0.1$  here, and 0.05 elsewhere?

Here “a significant ( $p < 0.1$ ) linear trend” was intended not to emphasize that “ $p < 0.1$ ” is the significant threshold but to describe that the linear trend of  $WS_{LT}$  is statistically significant at 90% confidence level. Although the  $p$ -value of  $WS_{LT}$  is slightly larger than 0.05 ( $p = 0.065$ ), this is much lower than those of other long-term meteorological components (Table 4). In general, the long-term trends have large autocorrelations for long lag-periods and thus very small numbers of independent data points. For example, although the correlation coefficient of the slope of  $WS_{LT}$  is very high (0.935), its degree of freedom is only 2 (Table S1) probably related to the two data points of the lowest  $WS_{LT}$  in 2002 and the highest  $WS_{LT}$  in 2013, and therefore the  $p$ -value of the slope is not much small ( $p = 0.065$ ).

Page 12, Line 29: I advise referencing the relevant emissions changes in Figure 7 to more fully describe impacts potential impacts of the Legislation.

We appreciate the reviewer's comment. To describe the "Special Act on the Improvement of Air Quality in Seoul Metropolitan Area" and its potential impact on the emission trends in Seoul, following sentences were now added to the end of L2 on p.13.

Although the act includes the introduction of the emission cap-and-trade system and strengthening VOCs management, most of the budget has been allocated for expanding of DPF/DOC usage for old diesel vehicles, which was considered to be effective for reduction of  $\text{NO}_x$  and primary PM emissions (Kim and Lee, 2018). The impact of such policy on the decrease in  $\text{PM}_{10}$  and  $\text{NO}_x$  emissions (Fig. 7a) is not easily distinguishable from the influence of the decrease in diesel consumption (Fig. 7c). However, the generalized use of DPF/DOC in diesel vehicles may be one reason for relatively stable trends of  $\text{PM}_{10}$  and  $\text{NO}_x$  emissions despite the rapid increase in diesel consumption mainly by vehicles since 2012.

Table 4: Relatively higher (positive or negative) correlation between some met variable ST trends (SI and RH, for example) may affect the interpretation of

We agree with the reviewer's points. In fact, the correlations among the short-term meteorological components are closely related to the midlatitude synoptic weather system. The positive correlation between P and SI and the negative correlations of RH with P and SI implies clear-sky and dry conditions in the high-pressure system and cloudy and wet conditions in the low-pressure system. In Sect. 4.2, we already considered the interpretation of the correlations among the short-term meteorological components to depict the impact of the synoptic meteorological conditions on the short-term variability of pollutants.

Figure S3: Please clarify which monitoring site this data is from, or if it is an average. This clarification should be made throughout the manuscript.

In this study, we used daily pollutant concentrations averaged for the selected 18 sites in Seoul. We modified captions of Fig. 2, Fig. S1, and Fig. S3 as follows:

Figure 2: Schematic flowchart of temporal decomposition of air pollution time series (Seoul average daily  $\text{PM}_{10}$  concentration) into short-term, seasonal, and emission-related and meteorology-related long-term components.

Figure S1. (a) Numbers of available air quality monitoring sites in Seoul, of which missing data are less than 10% of the total. (b) Average and (c) standard deviation of  $\text{PM}_{10}$  concentrations of the available 18 sites in Seoul. Asian dust events those were excluded from the  $\text{PM}_{10}$  analysis are marked with orange color.

Figure S3. Decompositions of log-scale daily time series of (a)  $\text{PM}_{10}$  and (b)  $\text{O}_3$  those averaged for 18 sites in Seoul.

Figure S4: I recommend stating that the gray line in each subfigure represents the raw spectrum.

We added following sentence at the end of the caption of Fig. S4.

The power spectrum of the original time series in (a) is represented with gray lines in (b-d).

## Response to Anonymous Referee #2

In the manuscript, the authors used various statistical tools such as the K-Z filter and multiple linear regression method to the 18 year long-term data of the criteria air pollutants and meteorological variables. They could separate short term variations and long-term variation. Further, out of the long-term trend, they could separate the meteorological and emission driven trends. In addition they calculated local emission driven and transported components from the emission driven part based on the continuity equation approach.

It is a well organized manuscript and the results are of great importance since this study result can be complementary to the 3-dimensional chemical transport modeling results and, thus, it can provide a scientific background for effective policy development in South Korea.

However, there are several points that should be improved and clarified. Thus, I recommend the manuscript be accepted for the publication in the Journal with minor revisions. Specific points are:

We appreciate the reviewer's careful reading, valuable comments, and constructive suggestions. We modified the original manuscript as following the reviewer's comments and suggestions. Each response to the reviewer colored in blue and changes in the manuscript colored in red.

1. Abstract needs revision to further emphasize the scientific significances of the study.

We revised the last sentence of the abstract (L27–29 on p.1) as follows:

The present results not only reveal an important role of synoptic meteorological conditions on the episodic air pollution events but also give insights into the practical effects of environmental policies and regulations on the long-term air pollution trends. As a complementary approach to the chemical transport modeling, this study will provide a scientific background for developing and improving effective air quality management strategy in Seoul and its metropolitan area.

2. It would make the manuscript more valuable to compare the results with other the study results for South Korea in which the same statistical tools have been used (for example, Shin et al., AAQR, 12, 93, 2012). Also, it would be nice to refer and compare the study results on the influence of meteorological parameters on the air pollutant level (for example, Lim et al., JKOSAE, 28, 325, 2012).

Thanks for the suggestion. To add a brief comparison with the previous results from Shin et al. (2012), we modified L2–4 on p.14 as follows:

In terms of  $O_3$ , positive  $\frac{\partial}{\partial t}(O_{3\ 8hLT}^{emis(T)})$  during the 2000s (Fig. 5e) implies that the transport of regional background  $O_3$  played an important role in the meteorologically-adjusted long-term  $O_{3\ 8hLT}^{emis}$  trend ( $O_{3\ 8hLT}^{emis}$ ) during the period. This result is consistent with the previous study on  $O_3$  in Seoul for 2002–2006 using the OZone Isopleth Plotting Package for Research (OZIPR) and the KZ filter (Shin et al., 2012). However, because  $O_{3\ 8hLT}^{emis(L)}$  has been gradually changed from the decreasing phase to the increasing phase over the analysis period (Fig. 5e), the recent changes in  $O_{3\ 8hLT}^{emis}$  in the 2010s are probably more related to the local secondary production rather than the background  $O_3$  transport.

Also, a short discussion of comparison with the dispersion effect of winds on  $PM_{10}$  and  $NO_2$  levels reported by Lim et al. (2012) is inserted after L11 on p.12 as follows:

The long-term component of wind speed ( $WS_{LT}$ ) in Seoul was increased by  $\sim 0.9\ m\ s^{-1}$  from 2002 to 2012 (Fig. 6a). Previous statistical research on the dispersion effects of winds reported that the fraction of concentrations reduced by wind speed rising of  $2\ m\ s^{-1}$  to  $3\ m\ s^{-1}$  was  $\sim 24\%$  for  $PM_{10}$  and  $\sim 9\%$  for  $NO_2$  (Lim et al., 2012). These ratios are comparable to the decrease of meteorology-related long-term concentrations



$(\chi_{LT}^{met})$  in PM<sub>10</sub> and NO<sub>2</sub> by ~30% and ~10% for the period, respectively ( $\Delta\chi_{LT}^{met}$  ( $= \Delta\chi_{LT}^{met}/\chi_{LT}^{met}$ ) is  $-0.3$  for PM<sub>10</sub> and  $-0.1$  for NO<sub>2</sub>; Fig. 5a and d).

3. In Table 1, add a part on explaining why the sum of the variances is not 100%. Also, in Table 4, it would better to make the sum of the trends of ‘emis’ and ‘met’ be equal to the total.

Explained variances in Table 1 are identical to coefficients of determination ( $R^2$ ) between the original time series ( $X$ ) and each decomposed component ( $X_{ST}$ ,  $X_{SN}$ , and  $X_{LT}$ ), which represent how much of the variability of  $X$  is accounted for by the variability of each component. The explained variances of the decomposed time series should sum to 100%, if  $X_{ST}$ ,  $X_{SN}$ , and  $X_{LT}$  were completely independent of each other. However, because the components decomposed by the KZ filter still have some minor correlations among them albeit very weak, the explained variance of each component ( $R^2$  with  $X$  in percentage) are contributed not only by the variance itself but also by the covariances with other components.

In the revised version, to clarify this, we replaced the explained variances in the original Table 1 with the proportions of each variance and covariance to the total variances. The sum of variances and covariances among the short-term, seasonal, and long-term components are exactly the same as the total variances of the original time series (see the formula in the revised Table 1).

Following the modified Table 1, “(~50–70%)” in L19 on p.8 is now “(~46–68%).” Also, the original version of Table 1 is modified to the  $R^2$  values and is now added to the supplement as Table S2.

**Table 1: Total variances ( $Var(X)$ ) of log-scale times series for Seoul average concentrations of PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3 8h</sub>, and relative contributions (%) of variances and covariances ( $Cov$ ) among each component to  $Var(X)$ . Daily data for the period of July 2000 to Jun 2015 that  $X_{LT}$  data is available were used.**

	PM <sub>10</sub>	SO <sub>2</sub>	NO <sub>2</sub>	CO	O <sub>3 8h</sub>
$Var(X)^a$	0.2998	0.1345	0.1400	0.1540	0.4068
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$Var(X_{LT}^{emis})$	2.49%	4.88%	1.80%	4.97%	1.90%
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<sup>a</sup> Values of variances of each log-scale concentration time series

$$Var(X) = Var(X_{ST}) + Var(X_{SN}) + Var(X_{LT}) + 2[Cov(X_{ST}, X_{SN}) + Cov(X_{SN}, X_{LT}) + Cov(X_{ST}, X_{LT})]$$

**Table S2. Coefficients of determination ( $R^2$ ) between each component and original time series ( $X$ ) for Seoul average concentrations of PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3 8h</sub>. Daily data for the period of July 2000 to Jun 2015 that  $X_{LT}$  data is available were used.**

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$X_{ST}$	0.699	0.541	0.724	0.503	0.474	Short-term components
$X_{SN}$	0.276	0.468	0.287	0.397	0.522	Seasonal components
$X_{LT}$	0.081	0.055	0.038	0.152	0.057	Long-term components
$X_{LT}^{emis}$	0.064	0.054	0.030	0.124	0.043	Emission-related $X_{LT}$
$X_{LT}^{met}$	0.069	0.003	0.019	0.107	0.040	Meteorology-related $X_{LT}$

In Table 4, the sum of the linear trends of  $X_{LT}^{emis}$  and  $X_{LT}^{met}$  should be exactly the same as the  $X_{LT}$  trend. However, the rounding after the calculation that we used here can induce tiny difference of  $\pm 1$  in the last decimal place of both  $\frac{\partial X_{LT}}{\partial t}$  and the sum of  $\frac{\partial}{\partial t} X_{LT}^{emis}$  and  $\frac{\partial}{\partial t} X_{LT}^{met}$ . For example,  $PM_{10LT}^{emis}$  trend ( $-1.6883300\% \text{ yr}^{-1}$ ) and  $PM_{10LT}^{met}$  trend ( $-1.9456153\% \text{ yr}^{-1}$ ) exactly sum to  $PM_{10LT}$  trend ( $-3.6339453\% \text{ yr}^{-1}$ ). However, the trends rounded to the second decimal place are  $-1.69\% \text{ yr}^{-1}$  for  $PM_{10LT}^{emis}$  and  $-1.95\% \text{ yr}^{-1}$  for  $PM_{10LT}^{met}$ , and thus their sum is  $0.01\% \text{ yr}^{-1}$  larger than the rounded-off  $PM_{10LT}$  trend ( $-3.63\% \text{ yr}^{-1}$ ). For the result value, we think that the rounding after the calculation is more accurate than the rounding before the calculation. Therefore, we left Table 4 as the original version.

4. In eq. (9), should it be a residual term or  $S_{LT}$  contains all terms except advection?

By Eq. (9), the term  $S_{LT}$  should contain all effects on the change rate of long-term trend ( $\frac{\partial X_{LT}}{\partial t}$ ), except the changes of inflow and outflow by horizontal advection ( $-\vec{V}_{LT} \cdot \nabla X_{LT}$ ). The residual effects are source (surface emissions and atmospheric secondary production related to  $X_{LT}^{emis(L)}$  and partly  $X_{LT}^{met}$ ), sink (physical and chemical scavenging affected by  $X_{LT}^{met}$ ), and diffusion (accumulation and dissipation), as well as the vertical advection term ( $-w_{LT} \frac{\partial X_{LT}}{\partial z}$ ). The term  $-w_{LT} \frac{\partial X_{LT}}{\partial z}$  is negligible because the long-term trend of vertical motion ( $w_{LT}$ ) is close to zero. Therefore,  $S_{LT}$  can be regarded as the sum of source term (directly related to  $\frac{\partial}{\partial t} X_{LT}^{emis(L)}$ ) and sink and diffusion terms (affected by the long-term meteorological effect;  $\frac{\partial}{\partial t} X_{LT}^{met}$ ).

5. What would be the main reason for  $CO(LT, Met)$  and  $SO_2(LT, Met)$  to show different trends though both of them are primary air pollutants?

The local meteorological factors such as wind speed, relative humidity, insolation, and temperature are not directly related to the regional transport of air pollutants, and thus can affect mainly on the contribution of local source pollutants rather than transported pollutants. In other words, if the local emissions at a local place are assumed to be zero, the local meteorological effects on accumulation and secondary production ( $X_{LT}^{met}$ ) in the long-term air pollution trend ( $X_{LT}$ ) will be negligible. As we wrote in the third paragraph on p.11, the emission intensity of CO in Seoul ( $215 \text{ t/km}^2$ ) is higher than those in the upwind source areas (e.g.,  $98 \text{ t/km}^2$  for the Chinese eastern coast), while that of  $SO_2$  in Seoul ( $7 \text{ t/km}^2$ ) is much lower compared to those in the upwind source areas (e.g.,  $14 \text{ t/km}^2$  for the Chinese eastern coast). Therefore, the long-term  $SO_2$  trend ( $SO_{2LT}$ ) is much less contributed by the changes in meteorological effects on its local emission ( $SO_{2LT}^{met}$ ) than by the changes in its regional emissions ( $SO_{2LT}^{emis}$ ). As we describe in L9–13 on p.13, the interannual variation of satellite-based  $SO_2$  level over China is similar to  $SO_{2LT}$  (and  $SO_{2LT}^{emis}$ ) in this study (Fig. 5c) and supports the large influence of the regional background  $SO_2$  on the long-term  $SO_2$  trend in Seoul.

6. The different trends of  $NO_2$  and  $NO_x$  might be caused not only by the reasons explained in the manuscript but also with changing oxidative potential of the atmosphere (for example, Kim and Lee, AAQR, in press).

We appreciate the reviewer's comment. L8–L12 on p.11 was now modified as follows:

Interestingly,  $NO_2$  level has been stabilized despite increasing of the number of vehicles in Seoul from 2.3 million in 1999 to 3.1 million in 2016 probably owing to implementation of natural gas vehicles and low emission diesel engines, and  $NO_x$  ( $= NO + NO_2$ ) level has been even decreased from  $\sim 70$  ppb to  $\sim 50$  ppb for the same period (Shon and Kim, 2011; Kim and Lee, 2018). Such an increase of  $NO_2$  to  $NO_x$  ratio implies that additional conversion of NO to  $NO_2$  occurs somewhere before the emission (e.g., exhaust line of the vehicle) or in the atmosphere. Although further evidence is required, this can be attributable to expanding of diesel particulate filter (DPF) and diesel oxidation catalyst (DOC) usage for diesel vehicles or increase of the atmospheric oxidative potential in the SMA (Alvarez et al, 2008; Kim and Lee, 2018).

# Effects of meteorology and emissions on urban air quality: a quantitative statistical approach to long-term records (1999–2016) in Seoul, South Korea

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**Abstract.** Together with emissions of air pollutants and precursors, meteorological conditions play important roles in local air quality through accumulation or ventilation, regional transport, and atmospheric chemistry. In this study, we extensively investigated multi-timescale meteorological effects on the urban air pollution using the long-term measurements data of PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3</sub> and meteorological variables over the period of 1999–2016 in Seoul, South Korea. The long-term air quality data were decomposed into trend-free short-term components and long-term trends by the Kolmogorov-Zurbenko filter, and the effects of meteorology and emissions were quantitatively isolated using a multiple linear regression with meteorological variables. In terms of short-term variability, intercorrelations among the pollutants and meteorological variables and composite analysis of synoptic meteorological fields exhibited that the warm and stagnant conditions in the migratory high-pressure system are related to the high PM<sub>10</sub> and primary pollutant while the strong irradiance and low NO<sub>2</sub> by high winds at the rear of cyclone are related to the high O<sub>3</sub>. In terms of long-term trends, decrease in PM<sub>10</sub> ( $-1.75 \mu\text{g m}^{-3} \text{yr}^{-1}$ ) and increase in O<sub>3</sub> ( $+0.88 \text{ppb yr}^{-1}$ ) in Seoul were largely contributed by the meteorology-related trends ( $-0.94 \mu\text{g m}^{-3} \text{yr}^{-1}$  for PM<sub>10</sub> and  $+0.47 \text{ppb yr}^{-1}$  for O<sub>3</sub>), which were attributable to the subregional scale wind speeds increase. Comparisons with estimated local emissions and socioeconomic indices like GDP growth and fuel consumptions indicate probable influences of the 2008 global economic recession as well as the enforced regulations from the mid-2000s on the emission-related trends of PM<sub>10</sub> and other primary pollutants. Change rates of local emissions and transport term of long-term components calculated by the tracer continuity equation revealed a decrease in contributions of local emissions to the primary pollutants including PM<sub>10</sub> and an increase in contributions of local secondary productions to O<sub>3</sub>. This study shows The present results not only reveal an important role of synoptic meteorological conditions related to on the episodic air pollution events in short term and provides but also give insights into the current practical effects of environmental policies and regulations by isolation of emission-related trends in on the long-term air pollution trends. As a complementary approach to the chemical transport modeling, this study will provide a scientific background for developing and improving effective air quality management strategy in Seoul and its metropolitan area.

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## 1 Introduction

Over the past few decades, rapid urbanization, population and economic growth, and increase in energy consumption have exacerbated air pollution in the developing countries of South and East Asia (Sun et al., 2016; Liu et al., 2017; Shi et al., 2018). In response to growing concerns about air quality deterioration, several East Asian countries have implemented strict regulations on emissions in recent decades and achieved some degree of improvement in the primary air pollutants. For example, ambient concentrations of carbon monoxide (CO), sulfur dioxide (SO<sub>2</sub>), and nitrogen oxides (NO<sub>x</sub>) have decreased after the 1990s in South Korea and after the mid-2000s in China owing to various efforts to reduce their emissions (e.g., desulfurization on coal-fired power plants and industries, installation of selective catalytic reduction equipment, use of low-sulfur fuels and natural gas, and enhancement of vehicle emission standards; Shon and Kim, 2011; Klimont et al., 2013; Ray and Kim, 2014; van der A et al., 2017; C. Li et al., 2017; Kim and Lee, 2018). Despite the efforts to reduce the primary pollutants and secondary precursors, however, the East Asian countries still suffer from frequent severe haze pollutions (Huang et al., 2014) and experience continuous increasing of ozone (O<sub>3</sub>) levels (Seo et al., 2014; Sun et al., 2016).

Such a discrepancy between the past reduction of emissions and the current continuous air pollution can be minimized by considering effects of meteorological conditions on the air quality. The meteorological conditions often play important roles in local air quality through accumulation or ventilation of pollutants, regional transport of polluted or clean air, and atmospheric chemistry for the formation of secondary species (Zheng et al., 2015; Seo et al., 2017). For example, during the Chinese severe haze episode in January 2013, active secondary aerosol formation in the stagnant surface conditions and shallow boundary layer over the North China Plain induced the extremely high particulate matter (PM) concentrations without abrupt changes in emissions (Huang et al., 2014; Zheng et al., 2015; Cai et al., 2017; Zou et al., 2017). Meteorologically-adjusted long-term PM trend in Beijing shows that unfavorable meteorological condition has reduced efficiency of recent emission control (Zhang et al., 2018). Although the meteorological effects on the current increase of O<sub>3</sub> in China are unclear (Ma et al., 2016), the O<sub>3</sub> pollution is expected to become a more important issue in the future warmer climate (Jacob and Winner, 2009; Lin et al., 2008; Schnell et al., 2016) due to its high dependence on temperature (Sillman and Samson, 1995; Lin et al., 2001).

Considering the important role of meteorological conditions on the local air pollution levels, long-term urban air quality data must be interpreted carefully with examining both meteorological effects and changes in local and regional emissions. The meteorological conditions contribute to the various timescale changes in urban air quality by synoptic scale weather in short-term (Seo et al., 2017) and climatic variabilities in long-term (Cai et al., 2017; Zou et al., 2017; Oh et al., 2018), as well as its inherent seasonality (Kim et al., 2018). On the other hand, the changes in emissions mostly occur in long-term timescale due to the implementation of policy and regulations (van der A et al., 2017) and economic boom or recession (Vrekoussis et al., 2013; Tong et al., 2016). Thus, temporal decomposition of air pollution measurement data into different timescales provides useful information on the meteorological effects on both day-to-day fluctuations, seasonal characteristics, and long-term trend of the air pollution levels. In addition, the effectiveness of the past regulations on emissions and the influences of

socioeconomic changes can be more reasonably assessed by considering and isolating the meteorological effects on air pollution.

As one of the highly populated megacities in East Asia, the Seoul metropolitan area (SMA), which has a population of 25 million and 9 million vehicles, is an apparent large emission source of various air pollutant species (NIER, [2016](#)[2018](#)). To mitigate air pollution in the SMA by regulation on primary emissions, the South Korean government enacted the Special Act on the Improvement of Air Quality in Seoul Metropolitan Area in 2003 and enforced the act in 2005 (Kim and Lee, 2018). In fact, long-term measurements in Seoul shows that concentrations of PM<sub>10</sub> (PM with aerodynamic diameter  $\leq 10 \mu\text{m}$ ) and other primary pollutants have been decreased over the decades, and these long-term decreasing trends have been regarded as results from the enhanced environmental regulations (Ghim et al., 2015; Kim and Lee, 2018). However, considering that the air quality of Seoul is largely affected by the transport of regional air pollutants and the synoptic meteorological conditions (Seo et al., 2017), the efficiency of the current emission control policy should be evaluated after understanding and quantitative isolating of the meteorological effects on the long-term measurement data. Using regional air quality simulation, Kim et al. (2017) recently reported that interannual variability in winds could fluctuate the PM<sub>10</sub> levels in Seoul despite the emission control efforts. However, quantification of the meteorological effects based on a direct comparison between measured air pollutants and meteorological variables are further needed to demonstrate the effects of the emission controls.

In this study, we aimed to extensively investigate the role of meteorological conditions on the air pollution in Seoul, South Korea, and quantitatively examine the contributions of meteorology and emissions using the statistical approach to 18-yr (1999–2016) long records of ground PM<sub>10</sub>, NO<sub>2</sub>, SO<sub>2</sub>, CO, and O<sub>3</sub> together, as well as various local meteorological variables. To decompose the air pollution time series into different timescales and isolate the meteorological effects on the air pollution, we employed a simple statistical concept using the Kolmogorov-Zurbenko (KZ) filter and multiple linear regression model with meteorological factors, which has been used in previous studies (Wise and Comrie, 2005; Seo et al., 2014; Henneman et al., 2015; Ma et al., 2016; P. Li et al., 2017; Zhang et al., 2018). Synoptic meteorological conditions related to episodic air pollution events were investigated using composite analyses of the trend-free short-term components, and ~~Isolation~~[isolation](#) of the meteorology- and emission-related components from the long-term air pollution trends were conducted. We further explored the long-term changes in contributions of local emissions and transport to the air pollution in Seoul using a simplified tracer continuity equation. Finally, the identified emission-related trends of each pollutant species were compared with the estimated emission inventories, as well as socioeconomic indices that could affect local emissions.

## 2 Data

The Korean Ministry of Environment has measured PM<sub>10</sub> and four gaseous air pollutants of SO<sub>2</sub>, nitrogen dioxide (NO<sub>2</sub>), CO, and O<sub>3</sub> at 264 urban air quality monitoring sites in South Korea and provides their 1 h average concentrations (NIER, 2017). Daily average concentrations for PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, and CO and daily maximum 8-h average concentration for O<sub>3</sub> (O<sub>3 8h</sub>) at each site were derived from the hourly data. We choose 18 sites located in Seoul to obtain daily air pollutant concentrations

representative for air quality of Seoul (Figs. 1a and S1a), based on data availability of more than 90% for 1999–2016. Although the selected 18 sites are spread over Seoul, concentrations from individual sites show fair spatial homogeneity (e.g., coefficient of variations for PM<sub>10</sub> at the selected sites are ~0.17; Fig. S1b and c). Since Seoul is located in a basin surrounded by mountainous terrain (higher than ~500 m above sea level) except its western exit of the Han River toward the Yellow Sea, air pollutants from both emissions and transport can be stagnated and mixed in the basin in the prevailing westerlies. We averaged daily concentration data from the selected 18 sites in Seoul and utilized in this study.

The most noticeable features in the air pollutants data in Seoul are decreasing of PM<sub>10</sub> and increasing of O<sub>3</sub> levels in the long-term timescale. Annual average PM<sub>10</sub> concentrations and exceedance days of the South Korean air quality standard (AQS; 100 μg m<sup>-3</sup>) were decreased from 69 μg m<sup>-3</sup> to 41 μg m<sup>-3</sup> and from 64 days to 2 days between 1999 and 2015 (Fig. 1b). Note that we excluded episodic Asian dust (AD) days (183 days for the analysis period; [KMA, 2018](#)) from the PM<sub>10</sub> analysis to focus on the anthropogenic sources, although AD did not much affect the long-term PM<sub>10</sub> trend. On the other hand, annual average O<sub>3</sub> 8h levels and exceedance days of the AQS (60 ppbv) have been increased from 25 ppbv to 39 ppbv and from 5 days to 58 days between 2002 and 2016, respectively (Fig. 1c).

In terms of meteorological characteristics in Seoul, we used daily averages of temperature (°C), sea level pressure (hPa), relative humidity (%), wind speed (m s<sup>-1</sup>), and solar irradiance (W m<sup>-2</sup>) at the Seoul weather station (Fig. 1a) managed by the Korea Meteorological Administration. Note that the air quality monitoring sites in Seoul are spread over the area within a radius of ~15 km from the weather station, and such spatial size is enough to resolve the influence of synoptic conditions on the local meteorological factors. To investigate the synoptic meteorological conditions, the geopotential height and wind fields at 850 hPa, 10 m wind fields, total cloud cover, and surface solar radiation were derived from the European Centre for Medium-Range Weather Forecasts Reanalysis Interim (ERA-Interim) data.

We additionally employed estimated local emission data and several socioeconomic indices that could affect the emissions to compare with the separated emission-related air pollution trends. The annual estimated emissions of sulfur oxides (SO<sub>x</sub>), NO<sub>x</sub>, CO, volatile organic compounds (VOCs), and PM<sub>10</sub> in Seoul from the Clean Air Policy Support System (CAPSS) inventory (Lee et al., 2011; NIER, [20162018](#)), the national gross domestic product (GDP) growth (IMF, 2017) and the annual consumptions of the final energy, petroleum, and anthracite in Seoul (KEEI, 2016) were used in this study.

### 3 Decomposition of air pollutant time series

#### 3.1 Temporal decomposition of air pollutant time series

Time series of daily air pollutant concentration at a given place comprises mainly three components: trend, seasonality, and white noise (Rao and Zurbenko, 1994). The trend (long-term component) is attributable to the long-term variations in local and regional emissions related to socioeconomic status and policies (Mijling et al., 2013) and the long-term changes in meteorological conditions, which can affect atmospheric chemistry and regional transport patterns (Cai et al, 2017; Zou et al., 2017). The seasonality (seasonal component) arises from the seasonal variations of meteorological conditions (Kim et al., 2018)

and energy consumption pattern (Zhu et al., 2013). The white noise (short-term component) is a day-to-day variation mainly related to synoptic-scale weather changes, which control accumulation/ventilation of local air pollutants and transport of regional air pollutants (Seo et al., 2017) and is partly associated with short-term fluctuations in local emissions (Russell et al., 2010).

- 5 To decompose the air pollutant time series into these three components in different timescales, we used the KZ filter that has been employed in many air pollution time series studies, particularly on O<sub>3</sub> and PM (Wise and Comrie, 2005; Seo et al., 2014; Ma et al., 2016; P. Li et al., 2017). The KZ filter, here denoted as  $KZ_{(m,p)}$ , represents  $p$  times iteration of moving average of time width  $m$  (days) and removes high-frequency component, of which period is smaller than the effective filter width,  $N$  ( $\geq m \times p^{1/2}$ ). The KZ filter method is applicable to the time series with missing data owing to iterative moving average process and provides high accuracy level comparable to that of the wavelet transform method, albeit its simplicity (Eskridge et al., 1997). Here we applied  $KZ_{(15,5)}$  and  $KZ_{(365,3)}$  filters, which can remove variabilities of the periods shorter than 33 days and 1.7 yr, to filter out the short-term component and to leave the long-term component, respectively. The long-term component can be further separated into the emission-related trend and the meteorology-related trend by isolation of the emission-related component using a multiple linear regression model with representative meteorological variables.
- 15 Note that the original concentration ( $\chi$ ) was transformed into its natural logarithm values ( $X = \ln\chi$ ) prior to the decomposition. Because the number distribution of daily air pollutant concentrations is usually lognormal (e.g., Fig. S2), the natural log-transformation of the original concentration data is required for proper temporal decomposition with the KZ filter (Eskridge et al., 1997). The detailed decomposition procedure is described in followings and schematically summarized with PM<sub>10</sub> time series in Seoul in Fig. 2.
- 20 The time series of log-scaled pollutant concentration can be expressed as the sum of short-term ( $X_{ST}$ ), seasonal ( $X_{SN}$ ), and long-term ( $X_{LT}$ ) components.

$$X(t) = X_{ST}(t) + X_{SN}(t) + X_{LT}(t) \quad (1)$$

The sum of seasonal and long-term components is a baseline ( $X_{BL} = X_{SN} + X_{LT}$ ).  $X_{BL}$  and  $X_{ST}$  can be easily decomposed by applying the  $KZ_{(15,5)}$  filter to  $X$ , which filters out the white noise-like  $X_{ST}$ , as follows:

25  $X_{BL}(t) = KZ_{(15,5)}[X(t)] = X(t) - X_{ST}(t)$   
(2)

Now the baseline can be assumed to consist of its repeated climatological seasonal cycle ( $X_{BL}^{clm}$ ) and residuals ( $\varepsilon$ ).

$$X_{BL}(t) = X_{BL}^{clm}(t) + \varepsilon(t) \quad (3)$$

- Although  $X_{BL}^{clm}$  obviously occupies most of the seasonality in  $X_{BL}$ ,  $\varepsilon$  also contains some minor seasonal variability unconsidered in  $X_{BL}^{clm}$  together with the long-term trend. To obtain the long-term component ( $X_{LT}$ ) by filtering out the minor seasonality, the  $KZ_{(365,3)}$  filter is applied to  $\varepsilon$  as follows:
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$$X_{LT}(t) = KZ_{(365,3)}\{\varepsilon(t)\}[\varepsilon(t)] = X_{BL}(t) - X_{SN}(t) \quad (4)$$

Then the seasonal component ( $X_{SN}$ ), which represents the sum of the pure seasonal climatology ( $X_{BL}^{clm}$ ) and the minor seasonality ( $\varepsilon - KZ_{(365,3)}\{\varepsilon\}[\varepsilon]$ ), can be obtained by difference between  $X_{BL}$  and  $X_{LT}$ .

- 5 Note that if we define  $\chi_{BL} = \exp(X_{BL})$  and  $\chi_{ST} = \exp(X_{ST})$  and employ the similar concept to the relationship between the original concentration and its log-transformation ( $\chi = \exp(X)$ ),  $\chi_{BL}$  represents the baseline concentration of the air pollutant, and  $\chi_{ST}$  becomes the ratio of original concentration to baseline concentration ( $\chi/\chi_{BL}$ ). Similarly,  $\exp(X_{SN})$  and  $\exp(X_{LT})$  can be defined as  $\chi_{SN}$  and  $\chi_{LT}$ , respectively. Then  $\chi_{SN}$  represents the seasonal change in concentration without trend, and  $\chi_{LT}$  becomes the ratio of baseline concentration to detrended seasonal concentration ( $\chi_{BL}/\chi_{SN}$ ).
- 10 As shown in the example with  $PM_{10}$  in Seoul, the  $KZ_{(15,5)}$  filter effectively removes  $PM_{10,ST}$ , of which period is smaller than 33 days, and leaves both the seasonality of high  $PM_{10}$  concentrations in winter and spring and the long-term decreasing trend in  $PM_{10,BL}$  (Figs. 2 and S4b).  $PM_{10,SN}$  and  $PM_{10,LT}$  well represent the seasonal variation, of which periods are between 33 days and 1.7 yr with representative periodicities of 0.5 yr and 1 yr, and the long-term variations, of which period are longer than 1.7 yr, respectively (Figs. 2 and S4c–d). The high levels in winter and spring in  $PM_{10,SN}$  in Seoul is attributable to the shallow
- 15 boundary layer that traps local pollutants near the ground and frequent regional transport from China during the cold season (Kim et al., 2018).

### 3.2 Separation of emission- and meteorology-related trends

Since the long-term variability in air pollutant concentrations can be affected not only by changes in local and regional emissions but also by changes in meteorological conditions, the long-term trend is assumed to be consisted of meteorologically-

20 adjusted (emission-related) long-term component ( $X_{LT}^{emis}$ ) and meteorology-related long-term component ( $X_{LT}^{met}$ ). Therefore, the baseline can be represented as follows:

$$X_{BL}(t) = X_{SN}(t) + X_{LT}^{met}(t) + X_{LT}^{emis}(t) \quad (5)$$

To isolate the term  $X_{LT}^{emis}$  in Eq. (5), we built a multiple linear regression model employing the baseline time series of the five representative meteorological variables ( $MET_{BL}$ ), such as temperature ( $T_{BL}$ ), sea level pressure ( $P_{BL}$ ), relative humidity ( $RH_{BL}$ ),

25 wind speed ( $WS_{BL}$ ), and solar irradiance ( $SI_{BL}$ ), which are obtained by the  $KZ_{(15,5)}$  filter, as follows:

$$X_{BL}(t) = a_0 + \sum_i a_i MET_{BL_i}(t) + \varepsilon'(t),$$

$$MET_{BL} = [T_{BL}, P_{BL}, RH_{BL}, WS_{BL}, SI_{BL}] \quad (6)$$

where  $\varepsilon'$  is a sum of the non-meteorological long-term variability ( $X_{LT}^{emis}$ ) and the minor seasonal variability unexplained by the multiple linear regression model ( $\varepsilon' - X_{LT}^{emis}$ ). Note that daily maximum temperature ( $T_{max,BL}$ ) was used for the model of



$O_{3\ 8h_{BL}}$  instead of daily average temperature ( $T_{BL}$ ) considering daytime  $O_3$  as in previous study (Seo et al., 2014). By removing the minor seasonality from  $\varepsilon'$  using the  $KZ_{(365,3)}$  filter,  $X_{LT}^{emis}$  can be isolated as follows:

$$X_{LT}^{emis}(t) = KZ_{(365,3)}\varepsilon'(t)[\varepsilon'(t)] = X_{LT}(t) - X_{LT}^{met}(t) \quad (7)$$

5 Then  $X_{LT}^{met}$  can be simply obtained by difference between  $X_{LT}$  and  $X_{LT}^{emis}$ .

In the example displayed in Fig. 2,  $PM_{10_{LT}}$  in Seoul show a continuous decrease between 2003 and 2012. Such a decreasing trend in  $PM_{10}$  has been recognized as a result of the reduction in diesel vehicle emissions and fugitive dust in Seoul (Ghim et al., 2015; Kim and Lee, 2018). However, a recent modeling study suggested that the long-term increase in wind speed might additionally contribute to the past improvement of  $PM_{10}$  air quality in Seoul (Kim et al., 2017). In fact, both  $PM_{10_{LT}}^{emis}$  and

10  $PM_{10_{LT}}^{met}$  in Fig. 2 show decreasing patterns, and this supports probable influences of both emission controls and meteorology on the  $PM_{10}$  trend in Seoul.

### 3.3 Contributions of local emissions and transport to the long-term trends

$X_{LT}^{emis}$  contains both changes in local emissions ( $X_{LT}^{emis(L)}$ ) and changes in transport of regional emissions ( $X_{LT}^{emis(T)}$ ), and thus  $X_{LT}$  can be represented as Eq. (8).

$$15 \quad X_{LT}(t) = X_{LT}^{emis(T)}(t) + X_{LT}^{emis(L)}(t) + X_{LT}^{met}(t) \quad (8)$$

Then the rate of change of  $X_{LT}$  should satisfy a simple continuity equation as follows:

$$\frac{\partial X_{LT}}{\partial t} = -\vec{V}_{LT} \cdot \nabla X_{LT} + S_{LT} \quad (9)$$

where the advection term,  $-\vec{V}_{LT} \cdot \nabla X_{LT}$ , represents the transport of regional emissions mainly by horizontal winds, and the sources and sinks term,  $S_{LT}$ , can be regarded as the long-term changes by both local emissions and chemical production,

20 accumulation, and dissipation by meteorological factors. Because  $X_{LT}$  and  $X_{LT}^{met}$  were already known, the rates of change of

$X_{LT}^{emis(T)}$  and  $X_{LT}^{emis(L)}$  can be derived as follows: [\(see Appendix S1 in the Supplement\)](#):

$$\frac{\partial}{\partial t} (X_{LT}^{emis(T)}) = -\vec{V}_{LT} \cdot \nabla X_{LT} = -u_{LT} \left( \frac{\partial X_{BL}}{\partial x} - \frac{\partial X_{SN}}{\partial x} \right) - v_{LT} \left( \frac{\partial X_{BL}}{\partial y} - \frac{\partial X_{SN}}{\partial y} \right) \quad (10)$$

$$\frac{\partial}{\partial t} (X_{LT}^{emis(L)}) = \frac{\partial X_{LT}}{\partial t} - \left[ \frac{\partial X_{LT}^{met}}{\partial t} + \frac{\partial}{\partial t} (X_{LT}^{emis(T)}) \right] \quad (11)$$

where  $\vec{V}_{LT} = (u_{LT}, v_{LT})$  is long-term zonal and meridional winds components representative for the target area, and  $\nabla X_{LT}$  is horizontal gradient of long-term component of air pollutant for the larger area, which are identical to difference between  $\nabla X_{BL}$  and  $\nabla X_{SN}$ . If  $\nabla X_{LT} > 0$  at specific time, therefore, the horizontal gradient of baseline ( $\nabla X_{BL}$ ) at that time is steeper than that of

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seasonal climatology ( $\nabla X_{SN}$ ). Note that the above method to evaluate the changes in  $X_{LT}^{emis(T)}$  and  $X_{LT}^{emis(L)}$  is applicable not to data from the individual site but to data from the wide area, because of the requirement of horizontal gradient term in Eq. (10).

In this study, we choose 70 air quality monitoring sites over the SMA, which are spread over the area within a radius of 50 km from the Seoul weather station, based on data availability of more than 75% for 1999–2016 (Fig. 1a). Daily  $X_{BL}$  and  $X_{SN}$  at those sites were utilized to determine  $\nabla X_{BL}$  and  $\nabla X_{SN}$  by linear regressions of  $X_{BL}$  and  $X_{SN}$  on the zonal and meridional distances from the weather station and finally to obtain  $\nabla X_{LT}$  ( $= \nabla X_{BL} - \nabla X_{SN}$ ; Figs. S5 and S6c, e, g, i, and k). Also,  $u_{LT}$  and  $v_{LT}$  in Seoul is calculated by the  $KZ_{(365,3)}$  filter with wind direction and speed at the Seoul weather station (Fig. S6a and b). From  $\vec{V}_{LT}$  and  $\nabla X_{LT}$  of each species, we calculated the long-term transport terms for each air pollutant species in Seoul (Figs. S6d, f, h, j, and l).

- 10 In Eq. (11),  $\frac{\partial X_{LT}}{\partial t}$  and  $\frac{\partial X_{LT}^{met}}{\partial t}$  are three orders of magnitude smaller than  $\frac{\partial}{\partial t}(X_{LT}^{emis(L)})$  and  $\frac{\partial}{\partial t}(X_{LT}^{emis(T)})$ , and thus, approximately  $\frac{\partial}{\partial t}(X_{LT}^{emis(L)}) \approx -\frac{\partial}{\partial t}(X_{LT}^{emis(T)})$ . For example, if we assume that  $\chi \sim 50 \mu\text{g m}^{-3}$ ,  $\frac{\partial \chi}{\partial t} \sim -10 \mu\text{g m}^{-3} \text{ decade}^{-1}$ ,  $\frac{\partial \chi}{\partial x} \sim -10 \mu\text{g m}^{-3} (100 \text{ km})^{-1}$ , and  $u \sim 1 \text{ m s}^{-1}$  at a given place, of which conditions are similar to  $\text{PM}_{10}$  in Seoul and its metropolitan area, the transport term  $-u \frac{\partial \chi}{\partial x}$  ( $= -\frac{u}{\chi} \frac{\partial \chi}{\partial x}$ )  $\sim -2 \times 10^{-6} \text{ s}^{-1}$ , while the tendency term  $\frac{\partial \chi}{\partial t}$  ( $= \frac{1}{\chi} \frac{\partial \chi}{\partial t}$ )  $\sim -2.5 \times 10^{-9} \text{ s}^{-1}$ . If we further assume that the local  $\text{PM}_{10}$  emission in Seoul (area of  $605 \text{ km}^2$ ) has the same order of magnitude as the transport term in concentration
- 15 ( $u \frac{\partial \chi}{\partial x} \sim 1 \times 10^{-4} \mu\text{g m}^{-3} \text{ s}^{-1}$ ) and the boundary layer height over the area is about 1 km, the total amount of yearly  $\text{PM}_{10}$  emission in Seoul will be approximately 1.9 kt, which is consistent with estimation by the CAPSS emission inventory (Fig. 5a; NIER, [20162018](#)). In fact, because air pollutant concentrations are close to the steady state for the long-term period, its tendency is a small residual between two large terms: influx of clean or polluted air (transport) and local emissions/production or dissipation/deposition (sources and sinks). It should be noted that the transport term and the source/sink term, however, are
- 20 not balanced against each other in the short-term timescale because the day-to-day rate of change of the air pollution concentration is comparable to both transport and source/sink terms.

#### 4 Application to air pollutants in Seoul

Variances of each timescale component of  $\text{PM}_{10}$ ,  $\text{SO}_2$ ,  $\text{NO}_2$ , and  $\text{CO}$  in Seoul were generally largest in the short-term fluctuation (~~~50–70~~46–68%), while that of  $\text{O}_3$  were largest in its seasonal cycle (Table 1). This indicates an important role of synoptic-scale weather on the day-to-day variations of the primary air pollutants concentrations in Seoul. The changes in synoptic flow pattern control local air quality by accumulation/ventilation and regional transport of air pollutants (Zheng et al., 2015; Seo et al., 2017). On the other hand,  $\text{O}_3$  is a photochemical product and thus is controlled largely by the annual cycle of solar irradiance in such a  $\text{NO}_x$ - and VOCs-rich urban area, although its short-term variability, which is caused by the changes in synoptic weather and related precursor concentrations and irradiance, is comparable to the seasonal variability.

In terms of the long-term trends, variabilities related to the long-term local and regional emission changes ( $X_{LT}^{emis}$ ) are comparable to those induced by the long-term meteorological effects ( $X_{LT}^{met}$ ) for  $PM_{10}$ ,  $NO_2$ ,  $CO$ , and  $O_3$   $_{8h}$ , while  $X_{LT}^{emis}$  is dominant for total long-term variabilities of  $SO_2$  (Table 1). Since the local-scale meteorology could affect more on the local emission-related long-term trend rather than on the regional background-related long-term trend, the smaller variability of  $X_{LT}^{met}$  imply the less influence of the local emission changes compared to the regional background level changes on  $X_{LT}^{emis}$ .

#### 4.1 Meteorological effects on seasonality in air pollution

In the process of isolating  $X_{LT}^{emis}$  from  $X_{LT}$ , the multiple linear regression model predicting  $X_{BL}$  from meteorological predictors ( $MET_{BL}$ ) were used. The model using the meteorological variables explains well  $O_3$   $_{8hBL}$  (adjusted  $R^2 = 0.86$ ), while the baseline of  $PM_{10}$  ( $PM_{10BL}$ ) is much less explained by the model (adjusted  $R^2 = 0.51$ ; Table 2). This indicates that the long-term non-meteorological (or emission-related) variability in  $PM_{10BL}$  is much larger than that in  $O_3$   $_{8hBL}$ . In the linear regression model,  $PM_{10BL}$  and the baselines of primary gaseous pollutants of  $SO_2$ ,  $NO_2$ , and  $CO$  commonly show strong positive correlation with  $P_{BL}$  but strong negative correlations with  $T_{BL}$  and  $RH_{BL}$ , while  $O_3$   $_{8hBL}$  shows strong positive correlations ( $p < 0.05$ ) with  $SI_{BL}$  and  $T_{maxBL}$  but strong negative correlation ( $p < 0.05$ ) with  $P_{BL}$  (Table 2). Meteorological conditions over the Korean Peninsula in winter is affected by a cold and dry Siberian High that induces both the shallow boundary layer to trap the primary air pollutants near the surface and the northwesterly winds to transport the regional pollutants from the continent (Kim et al., 2018) and thus elevates  $PM_{10BL}$  in winter to early spring (Fig. S3b). On the other hand, enhanced photochemistry by strong irradiance together with temperature effects on  $O_3$  formation (increasing of biogenic hydrocarbons and hydroxyl radicals and enhanced thermal decomposition of peroxyacetyl nitrate in warm condition; e.g. Sillman and Samson, 1995) elevates  $O_3$   $_{8h}$  levels in late spring and summer in Seoul (Fig. S3b).

#### 4.2 Synoptic influences on short-term air pollution events

The short-term air pollution variability is closely related to the day-to-day variations of local meteorological factors. Intercorrelations among  $X_{ST}$  of air pollutant species and meteorological variables in Seoul are summarized in Table 3. Note that using  $X_{ST}$  instead of the original daily concentrations has an advantage to provide short-term features unbiased to seasonal characteristics and background levels because  $\exp(X_{ST})$  is equivalent to the ratio of measured concentration to filtered baseline concentration as aforementioned in Sect. 3.1.

Significant positive intercorrelations ( $p < 0.01$ ) among  $PM_{10ST}$  and primary gaseous pollutants ( $SO_{2ST}$ ,  $NO_{2ST}$ , and  $CO_{ST}$ ) together with their strong negative relationships to  $T_{ST}$  and  $WS_{ST}$  ( $p < 0.01$ ) indicate that the high  $PM_{10}$  episodes in Seoul were occurred in warm and stagnant weather conditions. Such warm and stagnant conditions can be induced by the migratory high-pressure system (e.g., negative relationships of  $WS_{ST}$  to  $T_{ST}$  and  $P_{ST}$  in Table 3) or by the warm front of the extratropical cyclone (e.g., negative correlations of  $T_{ST}$  to  $WS_{ST}$  and  $P_{ST}$  but its positive correlation to  $RH_{ST}$  in Table 3). The lag composites

of the geopotential height and wind anomalies at 850 hPa (about 1.5 km of altitude) relative to each day of  $\exp(\text{PM}_{10\text{ST}})$ , that is equivalent to the ratios of measured concentration to the filtered baseline concentration of  $\text{PM}_{10}$ , exceeding the sum of its mean and one standard deviation ( $> 1.50$ ) indicate that the high  $\text{PM}_{10}$  episodes in Seoul is associated with the former condition (Fig. 3). As shown in Fig. 3a–c, a high-pressure anomaly develops over southern China and moves eastward in three days before the high  $\text{PM}_{10}$  days. This kind of synoptic pattern can induce both slow regional transport of secondary precursors from China along the northern boundary of the high-pressure system (Fig. 3b and c) as well as gradual accumulation of primary and secondary aerosols in local area in the high-pressure system (Fig. 3d; Seo et al., 2017). Strong correlations of  $\text{PM}_{10\text{ST}}$  with  $\text{SO}_{2\text{ST}}$  and  $\text{NO}_{2\text{ST}}$  ( $p < 0.01$ ; Table 3) imply possible large contributions of the secondary species such as sulfate and nitrate aerosols to the haze episodes in Seoul. In fact, previous studies reported that the secondary inorganic aerosols proportion to the fine mode PM ( $\text{PM}_{2.5}$ ) was measured as high as  $\sim 50\%$  in average and even reached to  $\sim 75\%$  during the severe haze event (Seo et al., 2017; Kim et al., 2018).

In terms of short-term variability in  $\text{O}_3$   $8\text{h}$  levels in Seoul,  $\text{O}_3$   $8\text{hST}$  are strongly correlated positively with  $\text{SI}_{\text{ST}}$  and  $T_{\text{maxST}}$  but negatively with  $\text{RH}_{\text{ST}}$  ( $p < 0.01$ ; Table 3). This likely reflects favorable meteorological environment for  $\text{O}_3$  formation: strong irradiance with warm and dry conditions. Interestingly,  $\text{O}_3$   $8\text{hST}$  shows positive correlation with  $\text{WS}_{\text{ST}}$ , although stagnant condition is regarded as conducive to  $\text{O}_3$  formation in general. This is related to the strong negative relationship between  $\text{O}_3$   $8\text{hST}$  and  $\text{NO}_{2\text{ST}}$  and that between  $\text{WS}_{\text{ST}}$  and  $\text{NO}_{2\text{ST}}$ . Since the reaction with NO, which is produced by photolysis of  $\text{NO}_2$ , is a major sink of  $\text{O}_3$ , high  $\text{NO}_2$  concentration reduces  $\text{O}_3$  on high irradiance days, while the low  $\text{NO}_2$  concentration can increase  $\text{O}_3$ . In addition,  $\text{NO}_{2\text{ST}}$  and  $\text{CO}_{\text{ST}}$  are strongly correlated ( $r = +0.86$ ) because both CO and  $\text{NO}_x$  are mainly emitted by transportation exhaust in Seoul (82% of CO and 67% of  $\text{NO}_x$  emissions; NIER, [20162018](#)), and thus  $\text{O}_3$   $8\text{hST}$  has negative relationship with  $\text{CO}_{\text{ST}}$ , as like that with  $\text{NO}_{2\text{ST}}$ . The composites of the 850 hPa geopotential height and wind anomalies for high  $\text{O}_3$   $8\text{h}$  days that  $\exp(\text{O}_3$   $8\text{hST})$  exceeds the sum of its mean and one standard deviation ( $> 1.41$ ) shows that the high  $\text{O}_3$   $8\text{h}$  event frequently occurs at the transition of weather from the cyclonic anomaly to the anticyclonic anomaly (Fig. 4a). As the anomalous low-pressure system moves eastward out of the Korean peninsula, the decrease in cloud cover (Fig. 4b) results in the increase in surface irradiance (Fig. 4c) and thus the  $\text{O}_3$  level. Since this kind of anomalous synoptic pattern enhances the mean westerly flow by the anomalous northwesterly (Figs. 4a and S7a), the wind speed in this region is increased. As shown by the strong positive correlations between  $\text{WS}_{\text{ST}}$  and  $\text{NO}_{2\text{ST}}$  (Table 3),  $\text{NO}_2$  can be reduced by the high winds and thus could contribute to the high  $\text{O}_3$  concentration in short-term timescale.

### 4.3 Long-term trends of air pollution in Seoul

Times series of  $X_{\text{LT}}$  and its two decomposed components,  $X_{\text{LT}}^{\text{emis}}$  and  $X_{\text{LT}}^{\text{met}}$ , for each air pollutant species are shown in Fig. 5, and their linear trends are summarized in Table 4. Note that these time series range from July 2000 to June 2015 because 546 days at the beginning and ending of data were lost due to truncation effect of the  $\text{KZ}_{(365,3)}$  filter. The linear trend of  $X_{\text{LT}}$

represents a fractional change rate (% yr<sup>-1</sup>) of the baseline concentration ( $\chi_{BL}$ ) because  $\frac{\partial X_{LT}}{\partial t}$  is equivalent to  $\frac{1}{\chi_{BL}} \frac{\partial \chi_{BL}}{\partial t}$ . The fractional change rate can be converted into an equivalent concentration change rate by multiplying with time-mean  $\chi_{BL}$  for the analysis period.

In Seoul, there are significant decreasing trends ( $p < 0.05$ ) in  $PM_{10LT}$  (-3.6% yr<sup>-1</sup>) and  $CO_{LT}$  (-2.9% yr<sup>-1</sup>) and an increasing trend ( $p < 0.1$ ) in  $O_{3\ 8hLT}$  (+3.1% yr<sup>-1</sup>) for the recent fifteen years, while  $NO_{2LT}$  (-1.4% yr<sup>-1</sup>) and  $SO_{2LT}$  (+0.8% yr<sup>-1</sup>) doesn't show statistically significant trends (Table 4). The decrease of  $PM_{10}$  concentration and increase of  $O_{3\ 8h}$  level in their long-term linear trends can be clearly identified in temporal variation of  $X_{LT}$  (Fig. 3a and e) and are consistent with the temporal characteristics in their annual average concentrations (Fig. 1b and c). In terms of CO, abrupt decrease in the early 2000s, which was reported to be associated with introduction of natural gas vehicle supply and improvement of fuel quality (Kim and Shon, 2011), affects such a strong linear trend (Fig. 5b). On the other hand,  $SO_2$  concentration in Seoul has already been stabilized at ~5 ppb during the recent decade (Fig. 5c), although it decreased significantly in 1980s and 1990s owing to the government's efforts to control the  $SO_x$  emissions by use of low-sulfur fuel and natural gas for industry and transportation (Ray and Kim, 2014; NIER, 2017).  $NO_2$  concentration has been varied between 30 and 40 ppb in the recent decade and also shows relatively weak trend (Fig. 5d; NIER, 2017). Interestingly,  $NO_2$  level has been stabilized ~~and  $NO_x$  level has been even decreased for the period~~ despite increasing of the number of ~~vehicle vehicles~~ in Seoul from 2.3 million in 1999 to 3.1 million in 2016; probably owing to implementation of natural gas vehicles and low emission diesel engines, ~~and  $NO_x (= NO + NO_2)$  level has been even decreased from ~70 ppb to ~50 ppb for the same period~~ (Shon and Kim, 2011; Kim and Lee, 2018). ~~One explanation for the less decreasing~~ Such an increase of  $NO_2$  ~~compared to NO to  $NO_x$  ratio implies that additional conversion of NO to  $NO_2$  occurs somewhere before the emission (e.g., exhaust line of the vehicle) or in the atmosphere. Although further evidence is the~~ ~~extension~~ required, this can be attributable to expanding of diesel particulate filter (DPF) ~~and diesel oxidation catalyst (DOC)~~ usage for diesel vehicles ~~that could additionally convert NO to  $NO_2$  or increase of atmospheric oxidative potential~~ in the ~~exhaust line~~ SMA (Alvarez et al, 2008; Kim and Lee, 2018). Since Seoul is a  $NO_x$ -saturated regime area (Jin et al., 2012), such a decreasing trend in ambient  $NO_x$  concentration in recent decades may be one of the causes of the long-term increasing of  $O_{3\ 8h}$  in Seoul.

Although the decreasing trends in  $PM_{10LT}$ ,  $CO_{LT}$ , and  $NO_{2LT}$  in the recent decade have been regarded as the result of efforts to reduce the local emissions (Kim and Lee, 2018), their  $X_{LT}^{emis}$  linear trends are less than half of  $X_{LT}$  linear trends (Table 4). This indicates the more important role of the long-term changes in local meteorology in the long-term air pollution trends in Seoul.

In contrast to the comparable linear trends of  $X_{LT}^{emis}$  and  $X_{LT}^{met}$  of  $PM_{10}$ , CO, and  $NO_2$ , the linear trend of  $SO_{2LT}$  is dominantly contributed by that of  $SO_{2LT}^{emis}$  (Table 4). This suggests the larger influence of the regional background  $SO_2$  than that of the local  $SO_x$  emission on  $SO_{2LT}^{emis}$  because the long-term effects of local meteorological conditions on the local emission-related air pollution trend must be limited if the long-term change in local emission is negligible. In fact, the estimated  $SO_x$  emission

intensity of Seoul ( $7.4 \text{ t km}^{-2}$ ) in the reference year 2010 was much lower than that of an industrial port city in the west adjacent to Seoul, Incheon ( $18.1 \text{ t km}^{-2}$ ), or the Chinese eastern coastal region of Jing-Jin-Ji (Beijing-Tianjin-Hebei), Shandong, Jiangsu, and Shanghai ( $14.3 \text{ t km}^{-2}$ ; [NIER, 2016](#); M. Li et al., 2017; [NIER, 2018](#)). In consideration with prevailing westerly in this region (Fig. S7a), the long-term  $\text{SO}_2$  concentration trend in Seoul can be more affected by the long-term change in regional background level rather than that in local emission.

In the reference year 2010, estimated emission intensities of CO and  $\text{NO}_x$  in Seoul in 2010 were estimated as  $215.3 \text{ t km}^{-2}$  and  $117.4 \text{ t km}^{-2}$ , and these are obviously much higher than those in Incheon ( $44.1 \text{ t km}^{-2}$  for CO and  $51.3 \text{ t km}^{-2}$  for  $\text{NO}_x$ ) or the Chinese eastern coastal region ( $98.1 \text{ t km}^{-2}$  for CO and  $17.4 \text{ t km}^{-2}$  for  $\text{NO}_x$ ; [NIER, 2016](#); M. Li et al., 2017; [NIER, 2018](#)). Therefore,  $\text{CO}_{\text{LT}}^{\text{emis}}$  and  $\text{NO}_{2\text{LT}}^{\text{emis}}$  in Seoul are probably much more affected by the long-term changes in local emissions rather than the changes in regional background levels.

#### 4.3.1 Meteorology-related long-term trends

For the most of species in Seoul except  $\text{SO}_2$ , the larger decreasing trends of  $X_{\text{LT}}^{\text{met}}$  than that of  $X_{\text{LT}}^{\text{emis}}$  have been found (Table 4). For example, equivalent linear trend of  $\text{PM}_{10\text{LT}}$  was  $-17.5 \mu\text{g m}^{-3} \text{ decade}^{-1}$ , while that of  $\text{PM}_{10\text{LT}}^{\text{emis}}$  was only  $-8.1 \mu\text{g m}^{-3} \text{ decade}^{-1}$ . Thus, the recent achievement of ambient  $\text{PM}_{10}$  reduction in Seoul would be less successful without the contribution by  $\text{PM}_{10\text{LT}}^{\text{met}}$  trend ( $-9.4 \mu\text{g m}^{-3} \text{ decade}^{-1}$ ; Table 4 and Fig. 5a). In terms of  $\text{O}_3$ , equivalent linear trend of  $\text{O}_{3\text{ 8hLT}}$  ( $+8.8 \text{ ppb decade}^{-1}$ ) was also slightly more contributed by  $\text{O}_{3\text{ 8hLT}}^{\text{met}}$  ( $+4.7 \text{ ppb decade}^{-1}$ ) compared to  $\text{O}_{3\text{ 8hLT}}^{\text{emis}}$  ( $+4.1 \text{ ppb decade}^{-1}$ ; Table 4 and Fig. 5e). Therefore, the long-term changes in meteorology that helps improving PM air quality may have made worse the  $\text{O}_3$  air quality.

Since the only meteorological parameter that shows a significant ( $p < 0.1$ ) linear trend in Seoul is wind speed ( $+0.57 \text{ m s}^{-1} \text{ decade}^{-1}$ ; Table 4 and Fig. 6a), the meteorological effects on the long-term decreases of  $\text{PM}_{10}$ , CO, and  $\text{NO}_2$  are probably related to the long-term increase of wind speed that could result in enhancing ventilation of the primary pollutants. The long-term component of wind speed ( $\text{WS}_{\text{LT}}$ ) in Seoul was increased by  $\sim 0.9 \text{ m s}^{-1}$  from 2002 to 2012 (Fig. 6a). Previous statistical research on the dispersion effects of winds reported that the fraction of concentrations reduced by wind speed rising of  $2 \text{ m s}^{-1}$  to  $3 \text{ m s}^{-1}$  was  $\sim 24\%$  for  $\text{PM}_{10}$  and  $\sim 9\%$  for  $\text{NO}_2$  (Lim et al., 2012). These ratios are comparable to the decrease of meteorology-related long-term concentrations ( $\chi_{\text{LT}}^{\text{met}}$ ) in  $\text{PM}_{10}$  and  $\text{NO}_2$  by  $\sim 30\%$  and  $\sim 10\%$  for the period, respectively ( $\Delta\chi_{\text{LT}}^{\text{met}} (= \Delta\chi_{\text{LT}}^{\text{met}} / \chi_{\text{LT}}^{\text{met}})$  is  $-0.3$  for  $\text{PM}_{10}$  and  $-0.1$  for  $\text{NO}_2$ ; Fig. 5a and d). Such a role of wind speed in recent changes in PM concentrations in Seoul was also indicated by regional air quality model simulations (Kim et al., 2017). Interestingly, the recent increasing of wind speed in Seoul is a subregional scale change rather than a synoptic or larger scale climate variability. Linear trend distribution of observed wind speed over South Korea for 1999–2016 shows a rough pattern of increasing in the inland area and decreasing in the coastal area (Fig. 6b). However, the significant long-term increases of wind speed near Seoul seems to be somewhat limited to the Han River basin. In addition, although linear trend pattern of atmospheric circulation at 850 hPa ( $\sim 1.5 \text{ km}$  of altitude) over the western North Pacific/East Asia region shows weakening of the Aleutian low during the period

(Fig. S7a and b), no significant trend can be observed in both 850 hPa wind fields and 10 m wind speeds over and near the Korean Peninsula (Fig. S7b and c).

Since the major increase in wind speed was observed during two periods (2001–2004 and 2010–2011; Fig. 6a), changes in  $PM_{10LT}^{met}$ ,  $CO_{LT}^{met}$ , and  $NO_{2LT}^{met}$  mostly occurred in these periods. Such an increase in wind speed could result in more ventilation of  $NO_x$  and elevation of  $O_3$  levels in the  $NO_x$ -saturated regime. However,  $O_{3\ 8hLT}^{met}$  shows additional variability related to the long-term variation of solar irradiance (e.g., decrease for 2005–2006 and increase for 2007–2009 in both  $O_{3\ 8hLT}^{met}$  and  $SI_{LT}$ , Figs. 5e and 6a).

### 4.3.2 Emission-related long-term trends

Although  $PM_{10LT}$  continuously decreased between 2003 and 2012, a major decrease of  $PM_{10LT}^{emis}$  occurred between 2008 and 2009 (Fig. 5a). Since the decrease of  $X_{LT}^{emis}$  during the period can be found in other primary gaseous pollutants (Fig. 5b–d), it could relate to the reduction of local and/or regional primary emissions. For example, the CAPSS emission inventory shows abrupt decreases in estimated  $SO_x$ ,  $NO_x$ , and  $PM_{10}$  emissions in Seoul from 2007 to 2009 (Fig. 7a). One plausible reason of such reduction of primary emissions during the period can be attributed to enforcement of the Special Act on the Improvement of Air Quality in Seoul Metropolitan Area, which took effect from January 2005. The act aimed to reduce annual average  $PM_{10}$  and  $NO_2$  concentrations to meet  $40\ \mu g\ m^{-3}$  and 22 ppb by the year 2014, by regulations of  $NO_x$ ,  $SO_x$ ,  $PM_{10}$ , and VOCs emissions mainly from industries and transportations (Ghim et al., 2015; Kim and Lee, 2018). Although the act includes the introduction of the emission cap-and-trade system and strengthening VOCs management, most of the budget has been allocated for expanding of DPF/DOC usage for old diesel vehicles, which was considered to be effective for reduction of  $NO_x$  and primary PM emissions (Kim and Lee, 2018). The impact of such policy on the decrease in  $PM_{10}$  and  $NO_x$  emissions (Fig. 7a) is not easily distinguishable from the influence of the decrease in diesel consumption (Fig. 7c). However, the generalized use of DPF/DOC in diesel vehicles may be one reason for relatively stable trends of  $PM_{10}$  and  $NO_x$  emissions despite the rapid increase in diesel consumption mainly by vehicles since 2012.

Socioeconomic status can also be an important factor for emissions and air pollution (Vrekoussis et al., 2013; Tong et al., 2016; Stern and van Dijk, 2017). In terms of local emissions, the rapid reduction of primary emissions in Seoul between 2008 and 2009 coincided with the rapid drops of South Korean GDP growth (Fig. 7b), which were associated with the 2008 global financial crisis triggered by the US subprime mortgage meltdown (Kim et al., 2017). Final energy consumption in Seoul shrank during the recession, primarily by decrease in petroleum products (especially diesel and liquefied petroleum gases) consumption (Fig. 5b) (KEEI, 2016). The decrease in  $PM_{10}$ ,  $NO_x$ , and  $SO_x$  emissions in Seoul in 2007–2009 were attributable to the reductions of diesel consumption and anthracite consumption (Fig. 7a and c). In terms of regional emissions, temporal variation of  $SO_{2LT}^{emis}$ , which shows the large decrease in 2008–2009 and slight recovery in following brief period (Fig. 5c), corresponds to interannual variation of satellite  $SO_2$  observation over China (van der A et al., 2017). The large decrease in

2008–2009 are attributable to slowdown of the Chinese economy, as well as desulfurization program of the Chinese authority (van der A et al., 2017; C. Li et al., 2017).

Temporal variation of  $O_{3\ 8hLT}^{emis}$  shows a general increasing trend through the analysis period (Fig. 5e). Considering  $NO_x$ -rich condition in Seoul, this may results from the long-term decrease of  $NO_{2LT}^{emis}$  (Fig. 5d) and relatively more reduction of  $NO_x$  emission compared to that of VOCs emission in the recent decade (Fig. 7a). However, the factors that caused its steep increase in 2005, slight decrease in 2010, and following recovery between 2011 and 2012 are somewhat vague, without further information of concentration and temporal variation of VOCs in Seoul.

### 4.3.3 Long-term effects of local emissions vs. transport

The change rates of  $X_{LT}^{emis(L)}$  and  $X_{LT}^{emis(T)}$  have nearly same magnitude but opposite signs. Since magnitudes of the change rates of  $X_{LT}$ ,  $X_{LT}^{met}$ , and  $X_{LT}^{emis}$  are much smaller compared to those of  $X_{LT}^{emis(L)}$  and  $X_{LT}^{emis(T)}$  in the long-term timescale, characteristics of their temporal variation is not easy to be explained by using  $X_{LT}^{emis(L)}$  and  $X_{LT}^{emis(T)}$ . However, comparing these two terms could provide some physical and quantitative insights into temporal changes in each contribution of local emissions and transport of regional background emissions to the local air quality, albeit without considering atmospheric chemistry. For example,  $\frac{\partial}{\partial t}(X_{LT}^{emis(L)}) \gg 0$  (or  $\frac{\partial}{\partial t}(X_{LT}^{emis(T)}) < 0$ ) has implications in the long-term increasing contributions of local emissions and, at the same time, flushing out of air pollutants. On the other hand,  $\frac{\partial}{\partial t}(X_{LT}^{emis(L)}) \ll 0 < 0$  (or  $\frac{\partial}{\partial t}(X_{LT}^{emis(T)}) > 0$ ) shows gradual reduction of local emissions and compensation by increase in transport of regional background emissions.

In Seoul,  $\frac{\partial}{\partial t}(X_{LT}^{emis(L)}) > 0$  for  $PM_{10}$  and primary gaseous pollutants (CO,  $SO_2$ , and  $NO_2$ ) in the early 2000s, but the change rates got close to zero owing to the efforts to reduce the primary emissions (Fig. 5). Unlike CO and  $NO_2$ , of which emission intensities are large in Seoul,  $SO_2$  shows obvious positive  $\frac{\partial}{\partial t}(SO_{2LT}^{emis(T)})$  in recent years (Fig. 5c), and this indicates the large effect of regional transport on the variability of  $SO_2$  concentrations in Seoul. Similar to  $SO_2$ ,  $PM_{10}$  also show recent increase in  $PM_{10LT}^{emis(T)}$  with decreasing of  $PM_{10LT}^{emis(L)}$  (Fig. 5a). In terms of  $O_3$ , positive  $\frac{\partial}{\partial t}(O_{3\ 8hLT}^{emis(T)})$  during the 2000s (Fig. 5e) implies that the transport of regional background  $O_3$  played an important role in the meteorologically-adjusted long-term  $O_{3\ 8h}$  trend ( $O_{3\ 8hLT}^{emis}$ ) during the period. This result is consistent with the previous study on  $O_3$  in Seoul for 2002–2006 using the OZone Isopleth Plotting Package for Research (OZIPR) and the KZ filter (Shin et al., 2012). However, because  $O_{3\ 8hLT}^{emis(L)}$  has been gradually changed from the decreasing phase to the increasing phase, and this implies that the over the analysis period (Fig. 5e), the recent increase of  $O_3$  concentration in Seoul changes in  $O_{3\ 8hLT}^{emis}$  in the 2010s are probably more related to the local secondary production by local precursors, rather than the transport of regional background  $O_3$  transport.



Interestingly, the change rates of  $X_{LT}^{emis(L)}$  reflect some changes in socioeconomic factors. For example,  $PM_{10LT}^{emis(L)}$  and  $SO_{2LT}^{emis(L)}$  turned into the decreasing phase since 2008 (Fig. 5a and c), corresponding to the aforementioned descriptions in the previous subsection. In addition, the variation of  $\frac{\partial}{\partial t}(NO_{2LT}^{emis(L)})$  reflects well the changes in diesel consumption, which were not clearly shown in the CAPSS inventory. The change rate of  $NO_{2LT}^{emis(L)}$  was negative between 2006 and 2011 but turned into positive since 2012 (Fig. 5d), and these changes correspond to the diesel consumption in Seoul, which was decreased from 2007 to 2012 but has increased since then due to policy support for the diesel vehicle market to reduce greenhouse gas emissions (Fig. 7c).

## 5 Conclusions

Based on the statistical approach to the long-term air pollutant measurement data from the urban air quality monitoring sites in Seoul, the present study revealed the important role of synoptic weather conditions on the episodic air pollution events and the meteorological effects on the long-term air pollution trends. Temporal decomposition using the KZ filter technique and the multiple linear regression with meteorological factors allowed separation of the trend-free short-term variability and isolation of the meteorology- and emission-related long-term trends in the daily air pollution data. Simplified continuity equation using the surface wind data in Seoul and air pollutant measurement data from the wider area around Seoul further provided the approximate estimation of the long-term changes in contributions of local emissions and transport to the air pollution in Seoul. In terms of the short-term variabilities, which occupies the largest portion of the total air pollution variabilities, the high  $PM_{10}$  and primary gaseous pollutants concentrations are related to the influence of migratory high-pressure systems, which induces both regional transport and local accumulation of the pollutants in the warm and stagnant conditions. On the other hand, the high  $O_3$  episodes are related to the weather transitions at the rear of cyclones, which accompany the  $NO_2$  reduction by high winds and the strong solar irradiance.

In terms of the long-term trends, Seoul experienced a decreasing of  $PM_{10}$  concentrations ( $-17.5 \mu g m^{-3} decade^{-1}$ ) and an increase in  $O_{3\ sh}$  levels ( $+8.8 ppb decade^{-1}$ ) over the period of 1999–2016. Such long-term changes are largely contributed by the meteorology-related trends (e.g.,  $-9.4 \mu g m^{-3} decade^{-1}$  for  $PM_{10}$  and  $+4.7 ppb decade^{-1}$  for  $O_{3\ sh}$ ) related to the decadal increase in surface wind speeds ( $+0.57 m s^{-1} decade^{-1}$ ). Therefore, the recent improvement of particulate air quality in Seoul was not achieved solely by the emission control policies but was also induced by changes in local meteorological conditions, especially, wind speeds. Although no evidence of the influence of synoptic or larger scale climate variability can be found, the long-term wind speed increase was also observed in the Han River basin around Seoul and probably enhanced the ventilation of local primary pollutants and secondary precursors. Since Seoul is the  $NO_x$ -saturated regime, the long-term decrease in  $NO_2$  by the enhanced winds could be a cause for the long-term increase in  $O_{3\ sh}$  levels. Unlike other species,  $SO_2$  doesn't show a significant meteorology-related trend because of the small influence of local meteorology on its small local emissions.

The isolated emission-related air pollution trends of PM<sub>10</sub> and other primary gaseous pollutants commonly showed a major decrease between 2008 and 2009. Although the enforcement of the Special Act on the Improvement of Air Quality in Seoul Metropolitan Area in 2005 might affect the reduction of local emissions, socioeconomic indices like GDP growth and energy consumptions also indicate probable influence of the 2008 global economic recession on the changes in the emission-related air pollution trends. Contributions of the local emissions and the transport of the regional background air pollutants to the emission-related long-term trends are almost balanced each other because the change rate of air pollution trend is negligibly small compared to those of the local emissions and the transport in long-term. The changes in the contributions of local emissions to PM<sub>10</sub> and primary gaseous pollutants turned into decreasing phase since the mid-2000s in Seoul, and those species, especially PM<sub>10</sub> and SO<sub>2</sub>, became to be more affected by the regional background levels in recent years. The contribution of local emissions to NO<sub>2</sub> has turned into increasing phase in 2010s due to the recent policy support for the diesel vehicle. Also, the recent increasing phase of local contributions to O<sub>3</sub> implies an important role of local secondary production in the increasing trend of O<sub>3</sub> in Seoul.

Although the results from the emission-related long-term air pollution trends showed general agreement with estimated emissions and socioeconomic factors, the statistical technique employed in this study overlooked a role of long-term changes in the chemical conditions and reactions on the urban aerosols and O<sub>3</sub> (e.g., atmospheric oxidants, aerosol acidity and hygroscopicity, and secondary organic aerosol formation) and its relationship with the long-term changes in relevant meteorological factors (e.g., atmospheric water, irradiance, and temperature). Nevertheless, this simple concept of considering the emission-related air pollution trends separately from the meteorology-related trends could give insights into the assessment of past regulations on emissions and help the improvement of current environmental policies on urban air quality. As a complementary way to the chemical transport modeling, our analysis will provide a scientific background for effective air quality management strategy in the SMA.

### **Data availability**

The hourly data of PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3</sub> concentrations over South Korea for 1999–2016 are available in the website managed by the Korea Environment Corporation (<https://www.airkorea.or.kr>). The hourly data of temperature, sea level pressure, relative humidity, wind speed, and solar irradiance at the Seoul weather station for the same period can be found in the website of the KMA (<https://data.kma.go.kr>). The ERA-Interim data can be accessed via the European Centre for Medium-Range Weather Forecasts (ECMWF) data server (<http://apps.ecmwf.int/datasets/>).

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**Table 1: Explained Total variances (%) ( $Var(X)$ ) of short-term components ( $X_{ST}$ ), seasonal components ( $X_{SN}$ ), and long-term components ( $X_{LT}$ ) log-scale times series for Seoul average concentrations of PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3 8h</sub> in Seoul, and relative contributions (%) of variances of and covariances (Cov) among each component to Var(X). Daily data for the period between of July 2000 to JuneJun 2015 that X<sub>LT</sub> data is available were used.**

Components	PM <sub>10</sub>	SO <sub>2</sub>	NO <sub>2</sub>	CO	O <sub>3 8h</sub>	Notes
$X_{ST}Var(X)^a$	<u>69.9%0.2</u> <u>998</u>	<u>54.1%0.1</u> <u>345</u>	<u>72.4%0.1400</u>	<u>50.3%0.1</u> <u>540</u>	<u>47.4%0.406</u> <u>8</u>	Short-term components
$X_{SN}Var(X_{ST})$	<u>27.663.98</u> %	<u>48.92%</u>	<u>67.94%</u>	<u>46.819%</u>	<u>28.742.96</u> %	39.7 % 52.2 % Seasonal components
$Var(X_{SN})$	<u>22.07%</u>	<u>40.58%</u>	<u>23.86%</u>	<u>34.66%</u>		<u>47.06%</u>
$X_{LT}Var(X_{LT})$	<u>8.174%</u>	<u>5.54.50%</u>	<u>3.864%</u>	<u>15.214.32</u> %	<u>5.700%</u>	Long-term components
$Cov(X_{ST}, X_{SN})$	<u>2.91%</u>	<u>2.53%</u>	<u>2.20%</u>	<u>2.00%</u>		<u>2.17%</u>
$Cov(X_{SN}, X_{LT})$	<u>-0.29%</u>	<u>0.47%</u>	<u>0.09%</u>	<u>0.42%</u>		<u>0.32%</u>
$Cov(X_{ST}, X_{LT})$	<u>-0.01%</u>	<u>0.00%</u>	<u>-0.01%</u>	<u>0.00%</u>		<u>0.01%</u>
$X_{LT}^{emis}Var(X_{LT}^{emis})$	<u>6.42.49%</u>	<u>5.4.88%</u>	<u>3.01.8</u> <u>0%</u>	<u>4.397%</u>		Emission-related $X_{LT}$ 1.90%
$Var(X_{LT}^{met})$	<u>2.85%</u>	<u>0.09%</u>	<u>1.08%</u>	<u>4.45%</u>		<u>1.52%</u>
$X_{LT}^{met}Cov(X_{LT}^{emis}, X_{LT}^{met})$	<u>6.91.70%</u>	<u>-0.323%</u>	<u>1.90.38%</u>	<u>10.72.45</u> %	<u>4.0.79%</u>	Meteorology related $X_{LT}$

<sup>a</sup> Values of variances of each pollutant time series

$$Var(X) = Var(X_{ST}) + Var(X_{SN}) + Var(X_{LT}) + 2[Cov(X_{ST}, X_{SN}) + Cov(X_{SN}, X_{LT}) + Cov(X_{ST}, X_{LT})]$$



**Table 2: Correlation coefficients ( $r$ ) between baseline time series of pollutants (PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO, and O<sub>3 8h</sub>) and meteorological variables (T, T<sub>max</sub>, P, RH, WS, and SI), and adjusted  $R^2$  between the baseline time series of pollutants ( $X_{BL}$ ) and multiple linear regression model ( $a_0 + \sum_i a_i \text{MET}_{BLi}$ ).**

		PM <sub>10BL</sub>	SO <sub>2BL</sub>	NO <sub>2BL</sub>	CO <sub>BL</sub>	O <sub>3 8hBL</sub>
Correlation coefficients ( $r$ ) between $X_{BL}$ and $\text{MET}_{BL}$	T <sub>BL</sub>	-0.481 <sup>a</sup>	-0.785 <sup>a</sup>	-0.704 <sup>a</sup>	-0.689 <sup>a</sup>	-
	T <sub>maxBL</sub>	-	-	-	-	+0.726 <sup>a</sup>
	P <sub>BL</sub>	+0.338 <sup>a</sup>	+0.682 <sup>a</sup>	+0.666 <sup>a</sup>	+0.664 <sup>a</sup>	-0.760 <sup>a</sup>
	RH <sub>BL</sub>	-0.480 <sup>a</sup>	-0.627 <sup>a</sup>	-0.635 <sup>a</sup>	-0.383 <sup>a</sup>	+0.169
	WS <sub>BL</sub>	-0.031	+0.287	-0.013	-0.203	+0.217
	SI <sub>BL</sub>	-0.001	-0.366 <sup>a</sup>	-0.241	-0.540 <sup>a</sup>	+0.894 <sup>a</sup>
Adjusted $R^2$ for $a_0 + \sum_i a_i \text{MET}_{BLi}$		0.508	0.594	0.637	0.597	0.863

<sup>a</sup> The correlation is statistically significant at the 95% level or higher ( $p < 0.05$ ).

**Table 3: Correlation coefficients ( $r$ ) among short-term components of each pollutant and meteorological variable in Seoul for the period between February 1999 and November 2016.**

$X_{ST}$	$PM_{10ST}$	$SO_{2ST}$	$NO_{2ST}$	$CO_{ST}$	$O_3\ 8h_{ST}$	$T_{ST}$	$T_{maxST}$	$P_{ST}$	$RH_{ST}$	$WS_{ST}$
$SI_{ST}$	+0.072	+0.110 <sup>a</sup>	-0.075	-0.101 <sup>a</sup>	+0.537 <sup>a</sup>	+0.032	+0.283 <sup>a</sup>	+0.336 <sup>a</sup>	-0.674 <sup>a</sup>	-0.015
$WS_{ST}$	-0.342 <sup>a</sup>	-0.372 <sup>a</sup>	-0.687 <sup>a</sup>	-0.578 <sup>a</sup>	+0.213 <sup>a</sup>	-0.258 <sup>a</sup>	-0.315 <sup>a</sup>	-0.296 <sup>a</sup>	+0.015	
$RH_{ST}$	+0.038	-0.048	+0.091 <sup>a</sup>	+0.219 <sup>a</sup>	-0.385 <sup>a</sup>	+0.122 <sup>a</sup>	-0.081 <sup>a</sup>	-0.446 <sup>a</sup>		
$P_{ST}$	+0.048	+0.106 <sup>a</sup>	+0.142 <sup>a</sup>	+0.031	+0.048	-0.194 <sup>a</sup>	-0.053			
$T_{maxST}$	+0.410 <sup>a</sup>	+0.448 <sup>a</sup>	+0.508 <sup>a</sup>	+0.486 <sup>a</sup>	+0.118 <sup>a</sup>	+0.924 <sup>a</sup>				
$T_{ST}$	+0.373 <sup>a</sup>	+0.391 <sup>a</sup>	+0.471 <sup>a</sup>	+0.465 <sup>a</sup>	-0.012					
$O_3\ 8h_{ST}$	+0.072	-0.017	-0.223 <sup>a</sup>	-0.223 <sup>a</sup>						
$CO_{ST}$	+0.744 <sup>a</sup>	+0.718 <sup>a</sup>	+0.859 <sup>a</sup>							
$NO_{2ST}$	+0.627 <sup>a</sup>	+0.687 <sup>a</sup>								
$SO_{2ST}$	+0.720 <sup>a</sup>									

<sup>a</sup> The correlation is statistically significant at the 99% level or higher ( $p < 0.01$ ).

**Table 4: Linear trends of long-term components of air pollutants and meteorological variables in Seoul for the period between July 2000 and June 2015.**

Components	Trends		<i>p</i> -values	Components	Trends		<i>p</i> -values
	(% yr <sup>-1</sup> )	(per decade)			(% yr <sup>-1</sup> )	(per decade)	
PM <sub>10LT</sub>	-3.63 <sup>a</sup>	-17.54 <sup>a</sup> (μg m <sup>-3</sup> )	0.027	O <sub>3 8hLT</sub>	+3.09 <sup>a</sup>	+8.77 <sup>a</sup> (ppb)	0.062
PM <sub>10LT</sub> <sup>emis</sup>	-1.69	-8.15	0.153	O <sub>3 8hLT</sub> <sup>emis</sup>	+1.44	+4.10	0.179
PM <sub>10LT</sub> <sup>met</sup>	-1.95 <sup>a</sup>	-9.39 <sup>a</sup>	0.032	O <sub>3 8hLT</sub> <sup>met</sup>	+1.65 <sup>a</sup>	+4.67 <sup>a</sup>	0.093
SO <sub>2LT</sub>	+0.80	+0.42 (ppb)	0.320	Meteorological variables			
SO <sub>2LT</sub> <sup>emis</sup>	+0.71	+0.37	0.352	T <sub>LT</sub>		-0.10 (°C)	0.778
SO <sub>2LT</sub> <sup>met</sup>	+0.08	+0.04	0.423	T <sub>maxLT</sub>		+0.17	0.710
NO <sub>2LT</sub>	-1.36	-4.53 (ppb)	0.173	P <sub>LT</sub>		+0.02 (hPa)	0.822
NO <sub>2LT</sub> <sup>emis</sup>	-0.56	-1.88	0.326	RH <sub>LT</sub>		-1.82 (%)	0.135
NO <sub>2LT</sub> <sup>met</sup>	-0.80 <sup>a</sup>	-2.65 <sup>a</sup>	0.044	WS <sub>LT</sub>		+0.57 <sup>a</sup> (m s <sup>-1</sup> )	<del>0.06606</del> 5
CO <sub>LT</sub>	-2.91 <sup>a</sup>	-1.73 <sup>a</sup> (ppb)	0.032	SI <sub>LT</sub>		+4.6667 (W m <sup>-2</sup> )	0.294
CO <sub>LT</sub> <sup>emis</sup>	-1.09	-0.65	0.106				
CO <sub>LT</sub> <sup>met</sup>	-1.82 <sup>a</sup>	-1.08 <sup>a</sup>	0.046				

<sup>a</sup> The slope of linear trend is statistically significant at the 90% level or higher (*p* < 0.1).

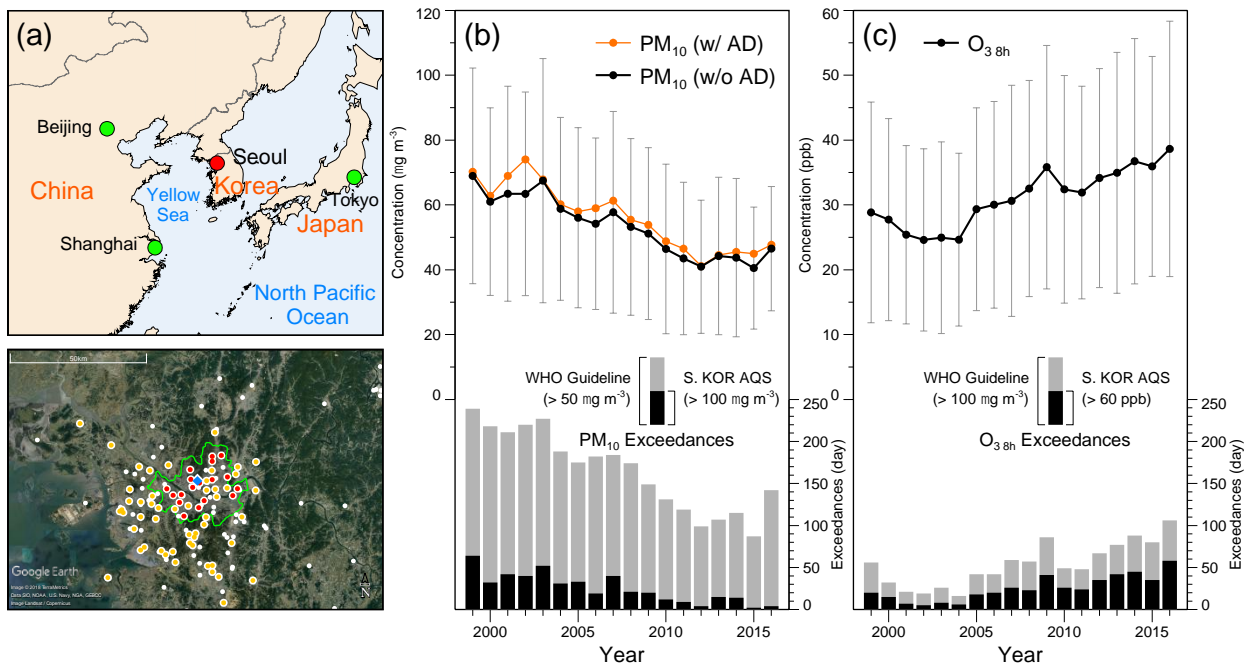


Figure 1: (a) Geographical locations of air quality monitoring sites in the SMA (circles) and the Seoul weather station (a blue solid diamond). Red solid circles denote 18 air quality monitoring sites utilized to obtain average daily air quality data in Seoul. Yellow solid circles show 52 air quality monitoring sites, which together with 18 sites (red solid circles) were used to calculate spatial gradients of the long-term components of air pollutants. Boundary of Seoul is marked with green line, and the satellite image is a courtesy of Google Earth Pro. Annual average concentrations and exceedances of the World Health Organization (WHO) guidelines and the South Korean AQS for (b)  $PM_{10}$  (excluding AD days) and (c)  $O_3$  in Seoul. Annual average  $PM_{10}$  concentrations including AD days are additionally represented in orange line in (b). The vertical bars on the annual average concentrations indicate annual standard deviations.

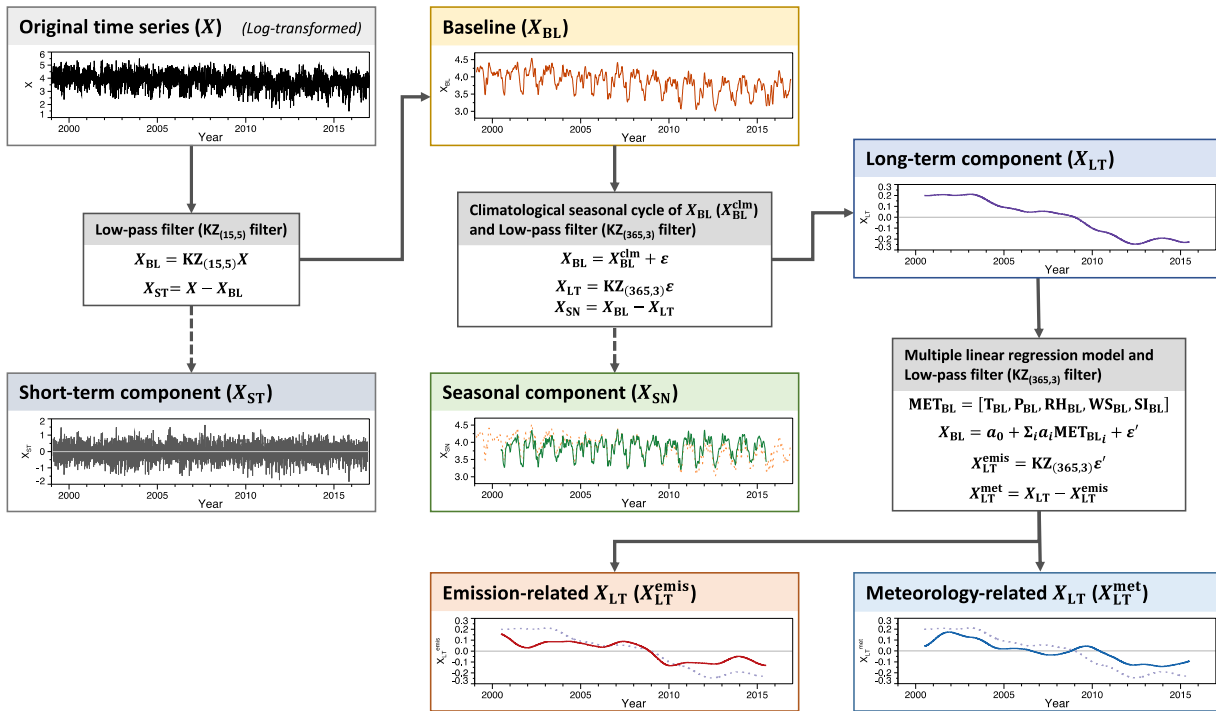
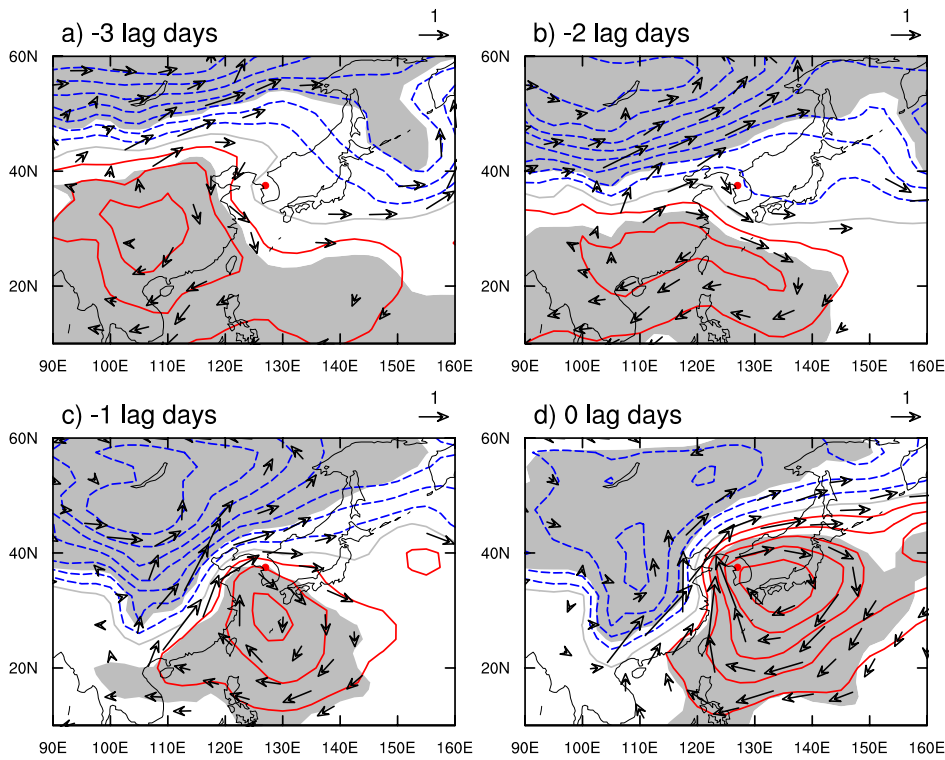
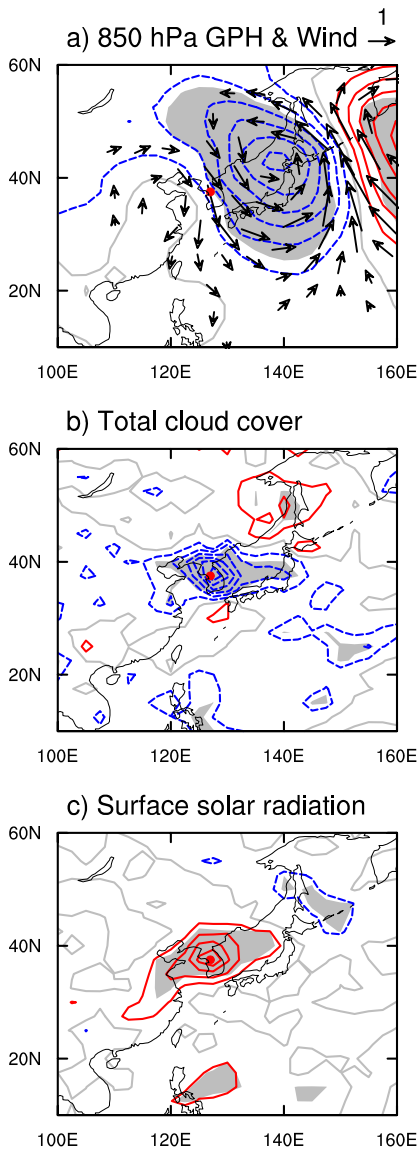


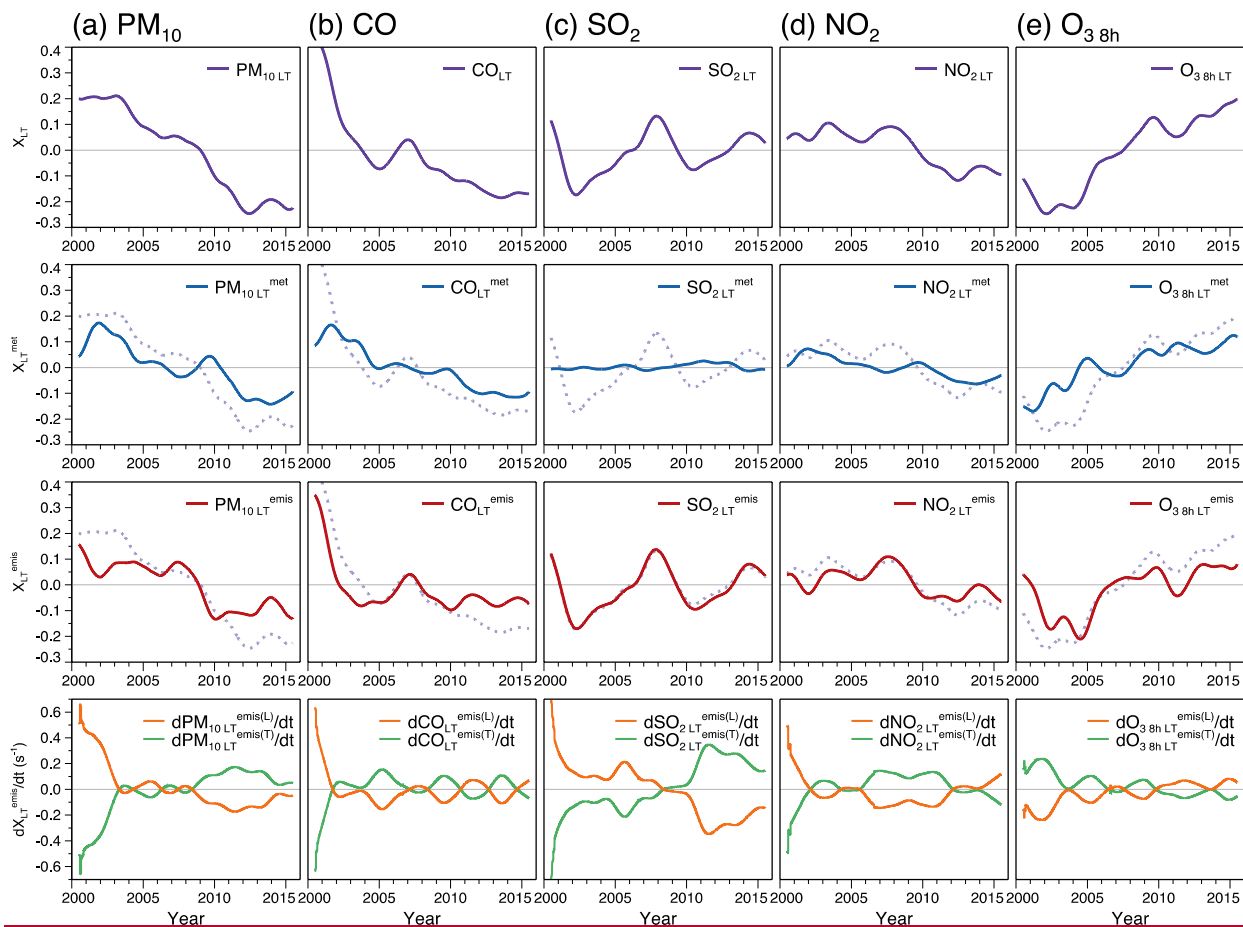
Figure 2: Schematic flowchart of temporal decomposition of air pollution time series (Seoul average daily  $PM_{10}$  in Seoul concentration) into short-term, seasonal, and emission-related and meteorology-related long-term components.



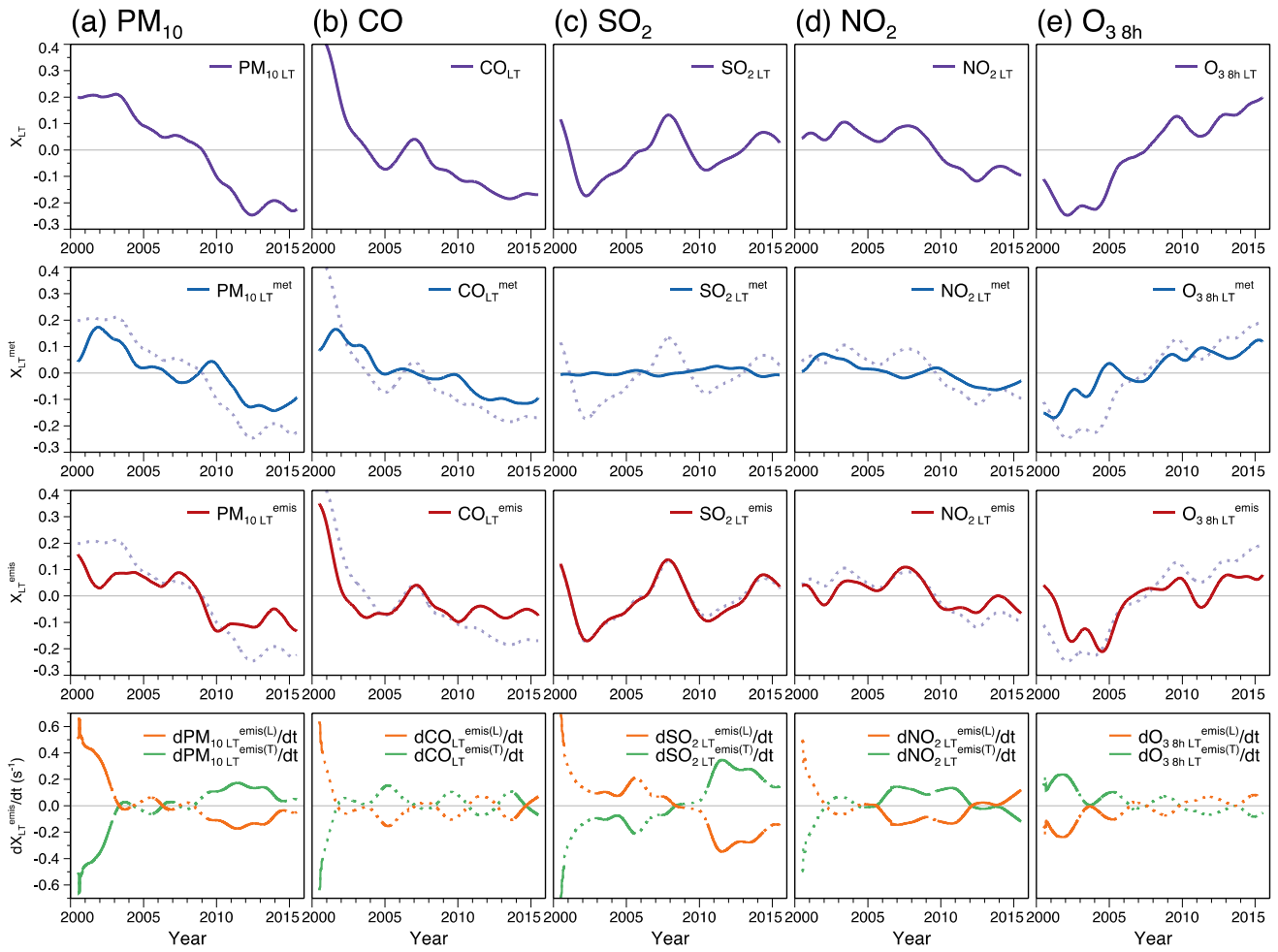
5 **Figure 3: Lag composites of seasonal anomalies of geopotential height (contours with interval of 2.5 gpm) and wind (arrows with reference scale of  $1 \text{ m s}^{-1}$ ) at 850 hPa relative to each first day that  $\exp(\text{PM}_{10\text{ST}})$  exceeds the sum of its mean and one standard deviation. Total number of events, mean, and standard deviation are 283, 1, and 0.50, respectively. Geopotential height and wind statistically significant at 95% confidence level are represented as light gray shading and arrows. Location of Seoul is marked by solid red circles.**



**Figure 4: Composites of seasonal anomalies of (a) geopotential height (contours with interval of 2.5 gpm) and wind (arrows with reference scale of  $1 \text{ m s}^{-1}$ ) at 850 hPa, (b) total cloud cover fraction (contours with interval of 2%), (c) downward solar radiation anomaly at surface (contours with interval of  $5 \text{ W m}^{-2}$ ) for each first day that  $\exp(\text{O}_3_{8\text{h}_{\text{ST}}})$  exceeds sum of the mean and one standard deviation. Total number of events, mean, and standard deviation are 346, 1, and 0.41, respectively. Geopotential height and wind, total cloud cover, and downward solar radiation statistically significant at 95% confidence level are represented by light gray shading and arrows. Location of Seoul is marked by a solid red circle.**







**Figure 5: Long-term components those are unadjusted for the meteorological variables ( $X_{LT}$ ; violet lines), meteorology-related ( $X_{LT}^{met}$ ; blue lines) and emission-related ( $X_{LT}^{emis}$ ; red lines) long-term components, and contributions of local emissions ( $\frac{\partial}{\partial t} X_{LT}^{emis(L)}$ ; orange lines) and transport of regional emissions ( $\frac{\partial}{\partial t} X_{LT}^{emis(T)}$ ; green lines) to the long-term trends of (a)  $PM_{10}$ , (b)  $CO$ , (c)  $SO_2$ , (d)  $NO_2$ , and**

- 5 **(e)  $O_3$  8h in Seoul. Solid lines in  $\frac{\partial}{\partial t} X_{LT}^{emis(L)}$  and  $\frac{\partial}{\partial t} X_{LT}^{emis(T)}$  show that the horizontal gradients of  $X_{LT}(-\vec{V}_{LT} \cdot \nabla X_{LT})$  obtained by linear regression are statistically significant at the 95% level or higher ( $p < 0.05$ ).**

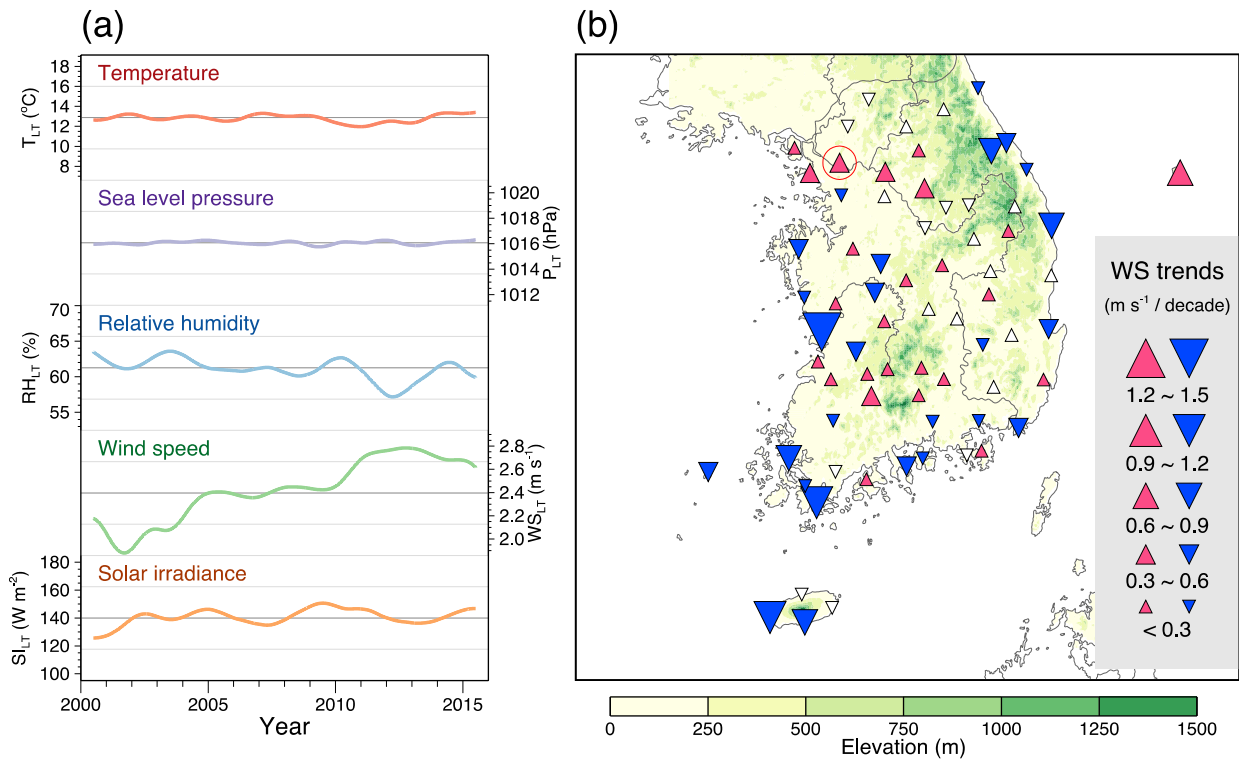
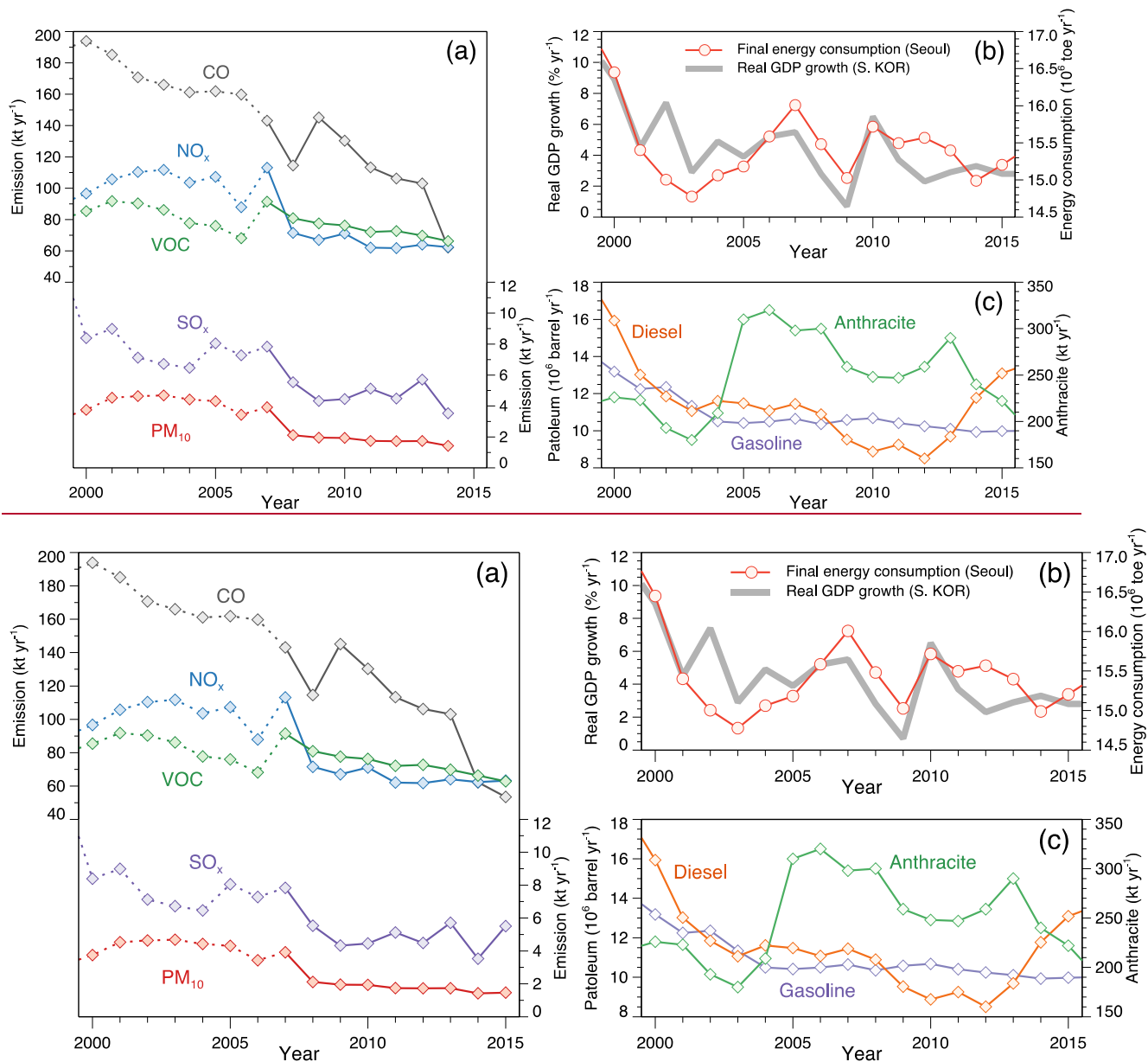


Figure 6: (a) Long-term components of daily average temperature, sea level pressure, relative humidity, wind speed, and solar irradiance. Vertical scales represented by grey horizontal lines are equivalent to 0.3 standard deviations of each original time series of the meteorological variables. (b) Linear trends of wind speed observed at 75 weather stations over South Korea for the period of 1999–2016. Upward and downward triangles with colors indicate the increasing and decreasing trends, which are statistically significant at 95% or more ( $p < 0.05$ ), respectively. Location of Seoul is marked by a red open circle.

5



5 **Figure 7: (a) Annual emissions of SO<sub>x</sub>, NO<sub>x</sub>, CO, VOCs, and PM<sub>10</sub> emissions in Seoul. Note that estimation method for the CAPSS inventory has been continuously updated, but the significant update was made in the year 2007. Emissions before and after 2007 were distinguished by dotted and solid lines. (b) Final energy consumption in Seoul and real GDP growth of South Korea. (c) Diesel, gasoline, and anthracite consumptions in Seoul.**