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2 **Atmospheric CO₂ inversions at the mesoscale using data driven**
3 **prior uncertainties. Part2: the European terrestrial CO₂ fluxes**

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1 **Abstract**

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3 Optimized biogenic carbon fluxes for Europe were estimated from high resolution regional scale
4 inversions, utilizing atmospheric CO₂ measurements at 16 stations for the year 2007. Additional
5 sensitivity tests with different data-driven error structures were performed. As the atmospheric
6 network is rather sparse and consequently contains large spatial gaps, we use a priori biospheric
7 fluxes to further constrain the inversions. The biospheric fluxes were simulated by the
8 Vegetation Photosynthesis and Respiration Model (VPRM) at a resolution of 0.1° and optimized
9 against Eddy covariance data. Overall we estimate an a priori uncertainty of 0.54 GtC y⁻¹ related
10 to the poor spatial representation between the biospheric model and the ecosystem sites. The sink
11 estimated from the atmospheric inversions for the area of Europe (as represented in the model
12 domain) ranges between 0.23 and 0.38 GtC y⁻¹ (0.39 and 0.71 GtC y⁻¹ up-scaled to geographical
13 Europe). This is within the range of posterior flux uncertainty estimates of previous studies using
14 ground based observations.

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1 Introduction

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3 Global and regional atmospheric inversions assimilate atmospheric CO₂ measurements made by
4 a global network for two decades, to infer terrestrial carbon fluxes using surface in situ or flask
5 measurements of CO₂ dry mole fractions (Tans et al., 1989; Enting and Mansbridge, 1989,
6 Conway et al., 1994, Fan et al., 1998; Rödenbeck et al., 2003). The optimization of CO₂
7 biospheric fluxes for the European domain has been of high interest in previous studies either
8 using pseudo or real data (Peters et al., 2010; Carouge et al., 2010a; Carouge et al., 2010b; Rivier
9 et al., 2010; Broquet et al., 2011; Broquet et al., 2013; Peylin et al., 2013). Retrieved fluxes from
10 most of the inversions are obtained from global systems at coarse resolution; hence, the spatial
11 and temporal flux variability at finer scales can not be resolved. Large uncertainties in the flux
12 retrievals are introduced due to the coarse resolution of the transport models used and due to the
13 network sparseness (Peters et al., 2010).

14 Apart from ground based observations, satellite measurements have also been recently used in
15 atmospheric inversions to infer terrestrial fluxes (Basu et al., 2013; Deng et al., 2014; Chevallier
16 et al., 2014). The advantage of using space-borne measurements lies on the high density of the
17 observations providing the opportunity to constrain regions not seen by the ground network.
18 However satellite based inversions significantly differ from ground based inversions, reporting a
19 larger annual uptake for Europe. A characteristic example is the estimated European uptake in
20 the study by Reuter et al. (2014). They calculated an uptake of 1.02 GtCy⁻¹ which triggered an
21 ongoing debate on whether those estimates are data driven or they lack robustness due to
22 deficiencies in the satellite observations and in the inverse modeling (Feng et al., 2016).

23 One of the largest sources of uncertainty in inversions is the atmospheric transport uncertainty.
24 Modeled dry mole fractions are biased particularly due to uncertainties in vertical mixing near the
25 surface (Gurney et al., 2003; Gerbig et al., 2008; Houweling et al., 2010). As a consequence,
26 posterior flux estimates are also biased because biases in concentrations due to transport model
27 errors are translated into biases in fluxes through the optimization procedure. Propagation of
28 uncertainties in winds (Lin and Gerbig, 2005) and in mixing heights (Gerbig et al., 2008) for
29 summer months with active vegetation resulted in uncertainties in simulated dry air mole
30 fractions of 5.9 ppm and 3.5 ppm respectively.

1 The current study is a continuation of the Kountouris et al. (2016) study (hereafter referred to as
2 Ko16) in which the inversion system and its set-up were assessed based on pseudo data. As a
3 next step we apply the modeling framework to real CO₂ atmospheric observations. Our main
4 objectives in the second part of this work are to investigate the potential to infer flux estimates
5 for Europe with reduced uncertainties, and to estimate biospheric fluxes at high spatial resolution
6 and for a full year. We use a spatial flux resolution of 0.25° x 0.25° to couple fluxes with the
7 atmospheric transport model, and the state space allows optimizing 3-hourly NEE corrections to
8 the prior NEE fluxes at a nominal spatial resolution of 0.5° x 0.5°. A data driven error structure is
9 implemented consistent with model-data flux mismatches (Kountouris et al., 2015) as tested in
10 part 1 (Ko16) of this study. Further, different error structures are used and assessed including
11 also a spatial error structure with a hyperbolic correlation shape as suggested by Chevallier et al.
12 (2012). Since spatial autocorrelations have been found to be very short, the annual aggregated
13 uncertainty over the European domain is smaller than traditionally assumed (see also Ko16). The
14 error inflation necessity and implementation was addressed in Ko16 either by inflating the error
15 covariance or, more formally, by introducing a bias term. However the hyperbolic correlation
16 shape suggested by Chevallier et al. (2012) has a stronger impact from larger distances compared
17 to the exponential shape, leading to an aggregated uncertainty which does not require to be
18 inflated. We perform also a number of sensitivity tests to account for misrepresentation of the
19 fossil fuel signal and also for transport uncertainties due to vertical mixing.

20 This paper is structured as follows: Section 2 describes the inversion system, the network and
21 station data which are used and details the assumed error structure. Section 3 shows the results of
22 the goodness of fit, and the retrieved fluxes. The data fitting and the reliability of the posterior
23 fluxes are extensively discussed in section 4.

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26 **2 Methods**

27 **2.1 Two-step inversion**

28

1 Real-data inversions require a nested inversion scheme, since observations contain also
2 contributions from regions outside of the Domain of Interest (DoI). As in part 1 of this study
3 (Ko16), the Jena Inversion System (Rödenbeck 2005) including the two-step nesting scheme
4 (Rödenbeck et al., 2009; Trusilova et al., 2010) was used. This scheme allows for combining
5 regional and global inversions within a consistent system. Here we only provide a brief
6 description as details are given in Rödenbeck et al. (2009) and Trusilova et al. (2010). The
7 atmospheric transport models TM3 (step 1) (Heimann and Körner, 2003) and STILT (step 2)
8 (Lin et al., 2003) were used for transport at the global and regional domain, respectively. For the
9 global runs, TM3 was used at a spatial resolution of 4° latitude x 5° longitude, driven by
10 meteorological fields from the ERA-Interim reanalysis produced by ECMWF (Dee et al., 2011).
11 The transport matrix for the regional inversions was identical to the one used for the synthetic
12 data study in part 1.

13 In the first step, a global inversion is performed using the global transport model. The outcome is
14 an optimized flux field, at coarser scale for the full period (FP) and the global domain. Then two
15 forward runs are performed. The first run uses the global transport model over the FP, computing
16 the modeled mixing ratios Δc_{mod1} . The second run initializes again the global transport model but
17 only within the regional DoI. This can be regarded as a regional simulation, but with coarse
18 resolution, yielding modeled mixing ratios Δc_{mod2} . Then the “remaining mixing ratio” is
19 calculated for all the observing sites inside the DoI:

$$20 \quad \Delta c_{remain} = c_{meas} - (\Delta c_{mod1} - \Delta c_{mod2} + c_{ini})$$

21 (1)

22 were c_{ini} the initial condition which corresponds to a well mixed atmosphere with a given initial
23 tracer mixing ratio.

24 In step two, the high-resolution transport model is used for the regional inversion within the DoI,
25 where all fluxes are represented at fine resolution. For this inversion the vector containing the
26 measured mixing ratios c_{meas} are replaced by the “remaining mixing ratios” Δc_{remain} . The
27 optimized fluxes from this step are the high-resolution fluxes of interest.

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1 **2.2 Atmospheric network and data**

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3 For step 1 we used the same station network as in version s04_v3.6 of the Jena CarboScope CO₂
4 inversion (http://www.bgc-jena.mpg.de/CarboScope/?ID=s04_v3.6), with 64 stations globally. For
5 step 2 (regional inversion) continuous and flask measurements from 16 stations within Europe
6 were used as described in Ko16 (see also Table 1). Of those 16 stations 7 are already included in
7 the step 1 inversion. All provided valid values were used, except those paired flask
8 measurements that differ more than 0.34 ppm which were omitted. Measurements from the
9 continuous stations were aggregated to hourly values where needed. Night and day time
10 observations were selected depending on the type of station (Ko16). As all institutions report
11 mixing ratio values traceable to WMO (World Meteorological Organization) calibration scale,
12 we expect compatibility between the different datasets (also see Rödenbeck et al., 2006).

13 In this study we use the site HEI (Heidelberg) which is traditionally not used for European CO₂
14 flux inversions as being considered too local (Broquet et al., 2013; Rödenbeck et al., 2009;
15 Rivier et al., 2010). The Heidelberg region is considered to be one of the most polluted regions in
16 Germany (Fiedler et al. 2005) and therefore could bias the flux estimates. Moreover the WES
17 (Westerland) site contains long periods with no data. This could potentially affect posterior flux
18 estimates since extended data gaps can lead to jumps in the presence of biases. Thus we evaluate
19 the performance and the sensitivity of the European flux estimates to the network configuration,
20 by performing also an inversion (referred to as nBV14, see Table 2) excluding HEI and WES.

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23 **2.3 A-priori information and uncertainties**

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25 A set of inversion cases differing in the prior information, the error structure and the station
26 configuration was realized (see overview in Table 2). Prior information derived from both
27 biosphere models (VPRM and GBIOME-BGCv1) is used to investigate the impact of the prior

1 fields to the posterior flux estimates. Furthermore an ensemble of inversions using different error
2 structures is used to investigate the impact on the posterior flux estimates and uncertainties.

3 Similarly to the synthetic inversion (Ko16) the model-data mismatch uncertainties are the same
4 as in the Ko16 study (see also fig. 2 therein). Further, we use the base case nBV (No Bias VPRM
5 as prior, B1 in Ko16) which inflates the prior uncertainty by up-scaling the error covariance
6 matrix, and case BVR (Bias VPRM as prior Respiration as shape, S1 in Ko16) which includes a
7 bias term. In the base case the VPRM model provides the prior flux field, and exponentially
8 decaying correlations are assumed. The bias component in the BVR scenario will always have a
9 correction with the same sign for all grid-cells as it just scales a predefined flux field. In the BVR
10 case it follows the shape of the annually averaged respiration flux, in the BVN case that of the a
11 priori Net biogenic flux, and in the BVRT case again that of the annually averaged respiration
12 flux, but with monthly temporal resolution of the bias term to allow for some temporal
13 flexibility. The nBB inversion refers to the scenario where GBIOME-BGCv1 was used as a
14 priori information instead of VPRM, and the error structure does not contain a bias term. With
15 this case we can evaluate how sensitive the posterior flux estimates are with respect to the prior
16 information which has been used. We also examine a spatial error structure based on a
17 hyperbolic (instead of an exponential) spatial correlation shape as suggested in Chevallier et al.
18 (2012) which we will refer to as nBVH scenario.

19 Note that in most of the inversions performed, VPRM fluxes were used as prior information.
20 Those fluxes are already optimized using EC measurements, therefore evaluation of the posterior
21 flux estimates against EC data at the local scale could result in posterior fluxes that are limited or
22 even not further constrained (since they are already optimized). In contrast, posterior fluxes
23 produced with BIOME-BGC used as prior are expected to show significantly larger corrections
24 compared to the prior estimations, and are therefore used for evaluation against EC data.
25 Nevertheless in most cases we use VPRM as prior in order to keep our estimates as data-driven
26 as possible through the overall optimization procedure; at local scale by using EC data, and at
27 regional scale using the atmospheric dry mole fractions.

28 As in the synthetic experiment (Ko16) the temporal decorrelation time was set to 31 days. In
29 Kountouris et al. (2015), model-data comparisons representative at site scale (around 1 km)
30 showed spatial correlation lengths of 40 km whilst model-model comparisons representative at

1 50 km resolution identified a correlation scale of 370 km. Considering also that the state space
2 has a resolution of 50 km, the spatial decorrelation length was chosen to be approximately 100
3 km (66 km in meridional, and 130 km in zonal direction). In the prior error covariance, diagonal
4 elements of $2.27 \mu\text{molm}^{-2}\text{s}^{-1}$ were assumed, consistent with the model-data flux mismatches as
5 calculated in Kountouris et al. (2015). Propagating this spatiotemporal error structure yields a
6 domain-integrated uncertainty (E_{st}) of 0.15 GtC y^{-1} . Note that this is substantially smaller than
7 for the synthetic experiment due to the much shorter spatial correlation length scales. A total
8 annual, domain integrated uncertainty E_{tot} of 0.3 GtC y^{-1} was assumed, which corresponds to
9 twice the standard deviation of annual terrestrial flux estimates for 2007 between terrestrial
10 biosphere models taken from the global carbon atlas (<http://www.globalcarbonatlas.org>). This is
11 also consistent with the prior uncertainty (for Europe) assumed for the global inversions
12 performed by the Jena inversion system. For those inversions in which the additional bias term
13 was considered (BVR, BVN, and BVRT scenarios), its error E_{BT} was calculated using

$$14 \quad E_{tot}^2 = E_{ST}^2 + E_{BT}^2 \quad (2)$$

15 For the nBVH scenario using hyperbolic correlations similar to Chevallier et al. (2012) ($\frac{1}{1+d}$), a
16 characteristic value d (lag distance) was used such that the correlation drops after around 60 km
17 to $1/e$ of its initial value, consistent with the hyperbolic fit to the model-data flux residual
18 autocorrelation in Kountouris et al. (2015). For this case no additional bias term was needed, as
19 the spatially and temporally aggregated uncertainty was found to be 0.32 GtC y^{-1} , which is very
20 close to the uncertainty assumed for the inversions (0.3 GtC y^{-1}).

21 Furthermore, we include ocean fluxes from Mikaloff-Fletcher et al. (2007), and anthropogenic
22 emissions from the EDGAR v4.1 inventory scaled at national level for individual years
23 according to the BP (British Petroleum) statistical review of world energy (BP, 2012) following
24 Steinbach et al. (2011). Anthropogenic emissions are considered to be perfectly known (with no
25 prior uncertainty), as one typically assumes that there is more a-priori knowledge regarding the
26 anthropogenic emissions as compared to biogenic fluxes. As the inversion cannot distinguish
27 between biogenic and anthropogenic signals, any errors in the a-priori anthropogenic emissions
28 will be included as corrections to the NEE flux.

1 2.4 Diagnostics and aggregation of fluxes

2

3 Similar to Ko16 we use the χ_c^2 metric to evaluate the goodness of fit for each station (Eq. 3)

$$4 \quad \chi_c^2 = \frac{\sum_t \frac{(\Delta c_t)^2}{\sigma_t^2}}{n} \quad (3)$$

5 where Δc_t is the model-data mismatch in dry mole fractions for a given observation time t , n the
6 number of observations and σ_t the assumed uncertainty. Further we make use also of the reduced
7 χ_r^2 (Eq. 4) where J_{min} is the cost function at its minimum

$$8 \quad \chi_r^2 = 2 \frac{J_{min}}{n} \quad (4)$$

9 For more details about the chi-square metric the reader is referred to Ko16 study.

10 The optimized fluxes are derived at 0.25° spatial and daily temporal resolution from the inversion
11 system. We post-process the fluxes by aggregating them spatially at country/domain-wide scales
12 and temporally at monthly/annual scales.

13 Flux comparisons with other studies require that both fluxes refer to the same geographical
14 region. Typically studies refer to TransCom regions with a European domain that expands more
15 into the Eurasian region. To scale our results to the TransCom EU region, we calculated the area
16 ratio between the TransCom EU region and our European domain. This ratio (about 1.69) was
17 used to scale our posterior estimates and the corresponding uncertainties assuming linearity in
18 the variances (presented in Fig. 8).

19 3 Results

20

21 3.1 Simulated CO₂ and goodness of fit

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23 Figure 1 presents a comparison of observed and modeled daily averages of the nighttime
24 (hours 23, 00, 1, 2, 3, 4 UTC) CO₂ dry air mole fractions for the Schauinsland station (SCH), a

1 mountain station, for the year 2007. The prior estimates (gray line) as derived from a forward
2 model run using VPRM flux fields are systematically lower than the observations (black line)
3 with the most divergent values occurring during the growing season. A similar pattern was found
4 for the other atmospheric stations. Posterior CO₂ timeseries from all the inversions are in much
5 closer agreement with the observations.

6 Table 3 summarizes the statistics between the modeled and the observed CO₂ dry mole fractions
7 for all stations based on daily averages using the respective sampling times (see also Ko16) for
8 mountain (nighttime) and other stations (daytime). Of note is that the real data inversions include
9 errors due to the modeling of transport, which is not the case in the synthetic experiment in Ko16
10 as the same transport model was used for forward and inversion runs. Standard deviations of the
11 posterior residuals (observed – modeled) show an average decrease for all inversion setups and
12 for all stations of 59% compared to the prior residuals. Correlations between prior and observed
13 as well as posterior and observed mole fractions (also Table 3) were likewise increased on
14 average from 0.48 to 0.93. Of note is that nBV and nBB, which use an inflated prior error
15 covariance for the spatiotemporal component, show larger improvement relative to the prior in
16 RMSD and some limited improvement in correlation coefficient, compared to those inversions
17 where a bias component was included (BVR, BVN, BVRT). Figure 2 visually summarizes the
18 goodness of fit in a Taylor diagram for cases nBV and BVR, presenting prior and posterior
19 estimates of the correlation and the normalized standard deviation between modeled and
20 observed CO₂ dry mole fraction time-series. It is obvious that the additional flexibility of nBV
21 in the spatiotemporal flux distribution results in a better reproduction of the concentration
22 variability. The same picture emerges when comparing the nBV and nBB inversions to nBVH
23 (see Table 3). Although all these cases assume no explicit bias term in the error structure, the
24 larger correlations from areas farther away for the nBVH case with a hyperbolic correlation
25 causes a reduced number of effective degrees of freedom, which results in larger residuals in
26 posterior-observed mole fractions (Table 3) comparable to those of the BVR case.

27 Calculating the goodness of fit using the station-specific χ_c^2 values from Eq. (3), most of the sites
28 (Table 3) show values around 1, indicating that the misfits are inside the 1 sigma site specific
29 uncertainty. For the CBW, HEI, JFJ, KAS sites, values above 1 regardless the error structure

1 were found, with the most extreme value of 5.17 for the HEI site in the nBVH inversion. This
2 could suggest that for a polluted site as HEI larger uncertainties should be considered.

3 The reduced χ_r^2 values regarding the overall model performance (Eq. 4) for all inversion set ups
4 is found to be close to 1 with χ^2 values of 1.08 (nBV), 1.16 (nBB), 1.17 (BVR), 1.17 (BVN),
5 1.19 (BVRT), 0.89 (nBV14) and 1.25 (nBVH), suggesting that the assumed prior uncertainty
6 describes well the actual uncertainties.

7

8 **3.2 Posterior flux estimates at different scales**

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10 The annually integrated spatial flux distribution is presented in Fig. 3 for all the different
11 inversion settings. Differences between the results based on the two general error structures (with
12 and without the bias term) were observed mainly in central and Western Europe (longitudes less
13 than 20° E), where the network provides a strong constraint. This difference is characterized by
14 stronger spatial flux variability for the general nBV case, with multiple transitions between
15 carbon sources and sinks at regional scales. The same picture emerges for the western part of
16 Europe. In contrast, all the inversions including a bias component (BVR, BVN, BVRT) yield a
17 more homogeneous flux distribution with somewhat finer structure in the flux retrievals (e.g.
18 France and north-east part of Europe). Comparisons between BVR, BVN, BVRT flux
19 distributions do not show any significant difference. Almost the same picture emerges when
20 comparing nBV and nBV14 cases, indicating that excluding the 2 stations does not have a very
21 strong influence on our annual flux estimates. However spatial differences were observed for the
22 areas close to the two sites. The most important one applies for the area near the HEI station
23 where we observed a transition from source to net carbon sink when excluding the corresponding
24 site. The choice of the prior does only have a small impact on the mean flux as can be seen by
25 comparing posterior fluxes from nBV and nBB despite the significant differences in the flux
26 innovations (Fig. 3). All innovations show that positive fluxes were added mainly in central
27 Europe and more intensively for the cases where no bias term was used. The positive flux
28 corrections is something to be expected since prior fluxes from VPRM show a strong European
29 sink of 0.96 GtC y⁻¹ which is most likely to be unrealistic. Overall the results suggest that the

1 general error structure matters, i.e. whether or not to include a bias term, but how the bias is
2 implemented is of less importance for the retrieved flux patterns. One would expect that the flux
3 distribution from the nBVH case would follow the general flux structure from the inversions
4 without the bias term. Interestingly the distribution is similar to the one obtained from the
5 inversions with the bias term (cases BVR, BVN, and BVRT). This shows that inversions
6 assuming correlations with a strong contribution from the far field have similar characteristics as
7 inversions that assume a flat bias term.

8 Figure 3 shows the spatially aggregated posterior flux estimates for the full domain with the
9 corresponding uncertainties integrated at monthly and at annual temporal scales. The same prior
10 uncertainty was used for cases nBV and nBB although they differ in prior flux field. Posterior
11 estimates from nBV (blue line/shading) and nBB (green line/shading) inversions do not show
12 any significant difference at monthly and annual scales despite the large difference in prior
13 fluxes. We observe that the maximum uptake occurs slightly earlier for the nBB case. Monthly
14 fluxes from the nBVH inversion also show the same temporal evolution. We do not observe any
15 significant difference in monthly fluxes for the BVR (red line/shading) and BVN (violet
16 line/shading) inversions. Both cases are comparable to the nBV and nBB cases at monthly and
17 annual scales. A slightly different picture emerges from the BVRT inversion, where the bias term
18 allowed for more degrees of freedom for monthly corrections. The resulting seasonal cycle is
19 somewhat smaller, with reduced summer carbon uptake. Inversions that included the bias term
20 yielded smaller posterior uncertainties at both temporal scales, which is expected as the
21 spatiotemporal component of the uncertainty was not inflated as was the case for the nBV
22 scenario. Flux retrievals from the reduced network (sensitivity case nBV14) show a slightly
23 deeper sink, but the differences to the base case nBV are insignificant (i.e. clearly within the
24 posterior uncertainties).

25 All of the inversions suggest Europe to be a carbon sink, with a range of $-0.23 \pm 0.13 \text{ GtC y}^{-1}$ to
26 $-0.38 \pm 0.17 \text{ GtC y}^{-1}$ for the BVRT and nBV inversions respectively. The mean annual posterior
27 flux estimate for Europe averaged over different inversions amounts to -0.32 GtC y^{-1} .

28 Posterior monthly flux estimates at smaller spatial scales (country level) are shown in Fig. 6.
29 Areas that are not well constrained by the current network show some divergence in the posterior
30 flux estimates although not significant considering the uncertainty range. For example Germany,

1 which is better constraint, shows a limited spread of the posterior fluxes with an annually
2 averaged standard deviation between the different posterior flux estimates being 0.9 MtC y^{-1} ,
3 while United Kingdom (which is less well constrained) shows a slightly larger spread of the
4 posterior estimates with an annually averaged standard deviation of 2 MtC y^{-1} . Note that the
5 posterior uncertainties are smaller by about 36% for the BVR case, which is related to the
6 smaller prior uncertainties at monthly time scales (see also section 3.2 in Ko16).

7

8 **3.3 Validation against eddy covariance measurements**

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10 As shown in Broquet et al. (2013) and in Ko16, eddy covariance measurements in principle have
11 the potential for quantitative evaluation of the retrieved fluxes from the inversions. Here we used
12 posterior flux estimates from the nBB inversion for evaluation against eddy covariance
13 measurements, as the prior flux fields in nBB (GBIOME-BGCv1) were not optimized using eddy
14 covariance measurements. Gap-filled data were downloaded from the European Fluxes Database
15 Cluster (<http://www.europe-fluxdata.eu>). A modified flux-site network compared to the one
16 reported in Kountouris et al. (2015) was used. Specifically we omitted sites that they have not
17 been used for the VPRM optimization (*CH-Fru, CH-Lae, CH-Oe1, ES-LMa, FR-Avi, FR-Mau,*
18 *IT-Cas, IT-LMa, IT-Ro2, NL-Dij, NL-Lut, SE-Sk1, SK-Tat*) as well as sites that were not available
19 as gap-filled data (*CH-Dav, ES-Agu, FR-Aur*). Further some more sites were added both for the
20 VPRM optimization and for the flux comparisons (*CZ-wet, DK-Sor, HU-Bug, IT-Non, NL-Cal,*
21 *PL-wet, RU-Fyo, UK-PL3*). Monthly averaged fluxes were extracted, with weights for each
22 vegetation class that compensate for the asymmetry between number of flux towers per
23 vegetation type and the fraction of land area covered by the specific vegetation type, similar to
24 Ko16.

25 The analysis of the monthly prior biospheric fluxes in Fig. 7 reveals significant differences
26 between observed and prior fluxes from the inversion. The GBIOME-BGCv1 model
27 systematically overestimates the observed fluxes throughout the year. The retrieved fluxes from
28 the inversion (dark green line) are closer to the observed fluxes, with a stronger uptake compared
29 to the prior during spring and summer time. The timing of the peak uptake is shifted to one

1 month earlier in comparison to the observations. The mean absolute bias (averaged absolute
2 differences between prior/posterior and observed fluxes) is significantly reduced by 52% from
3 0.84 to 0.40 $\text{gCm}^{-2}\text{day}^{-1}$. The standard deviation of the residuals is reduced by around 24%, from
4 0.68 for the prior to 0.40 $\text{gCm}^{-2}\text{day}^{-1}$ for the posterior residuals. Splitting the sites into two main
5 categories, the first only with crops, and the second with non crop sites, revealed differences on
6 how well those sites can be represented. Clearly best matches were found for the non crop sites
7 with a reduction in the mean absolute bias of 51% whilst for the crop sites it is limited to 38%.

8

9 **4 Discussion**

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11 We performed a series of atmospheric CO_2 inversions based on atmospheric data taken from 16
12 European stations for 2007. Different data-driven error structures in the prior error covariance
13 were assessed, and optimized biospheric fluxes were retrieved and post-processed at various
14 temporal and spatial scales for further evaluation. In this part we discuss the fitting performance
15 of the inversion system, and we detail the comparisons between our flux estimates at grid,
16 national and continental scales against eddy covariance data and reported flux estimates from
17 previous studies. Finally we discuss how sensitive flux retrievals are in the presence of erroneous
18 representation of the fossil fuel fluxes, and the site selection.

19

20 **4.1 Goodness of fit**

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22 Site-specific misfits show a reasonable fit to the atmospheric data. Nevertheless in 4 cases
23 (CBW, HEI, JFJ, and KAS) site-specific χ_c^2 values were found to be larger than 1 (see also Table
24 3), indicating that either the model-data mismatch errors were chosen too small, or the
25 spatiotemporal resolution of the flux model is too coarse compared to the biosphere fluxes and
26 therefore small scale variations are not resolved (Rödenbeck et al., 2003). In fact this seems to be
27 the case for the JFJ and KAS sites as those are high altitude sites with steep cliffs. In such a
28 complex terrain the atmospheric circulation is hard to be simulated from the transport models.

1 Regarding CBW and in particular HEI, those are polluted sites and it would be reasonable to
2 assume larger model-data mismatch uncertainty since the model is too coarse to resolve the fossil
3 fuel emission patterns. One could argue that using higher spatial resolution to couple fossil fuel
4 fluxes with transport models might reduce the model-data mismatch uncertainties, and hence
5 improve posterior fluxes. To investigate that, we performed a forward run at coarser (0.25°) and
6 higher ($1/12^\circ$ lat. X $1/8^\circ$ lon.) spatial resolution using only the fossil fuel emissions. As we use a
7 Lagrangian transport model, fluxes at higher resolution than that of the meteorological fields can
8 be used such that the simulated fossil fuel signals contain more spatially detailed information
9 (Lin et al., 2003). The derived concentration signal was subtracted from the observations and
10 subsequently an atmospheric inversion was performed. We report no significant differences
11 between the retrieved fluxes indicating that simply increasing the spatial resolution to about 10
12 km is not enough to correctly represent the fossil fuel distribution.

13 A common approach in atmospheric inversion studies to evaluate the defined uncertainties is to
14 examine the reduced χ_r^2 values. However, this might not always be a sufficient metric (Michalak
15 et al., 2005; Chevallier, 2007). The reduced χ_r^2 values in our study (between 1.08 and 1.25) are
16 larger than those found by Tolk et al. (2011) where values between 0.34 and 0.78 were found for
17 their pixel based inversion, indicating a more conservative choice for their model-data mismatch
18 errors. Even lower values were reported in the study by Peylin et al. (2005) with values ranging
19 from 0.01 up to 0.6 depending on the assumed correlations. χ^2 values from Zhang et al. (2015)
20 were within a range of 1 to 4, but were modified by inflating the error covariances through an
21 iterative procedure, resulting in χ_r^2 values comparable to ours. Concluding, the χ_r^2 values give
22 confidence that the assumed prior uncertainties are well defined.

23

24 **4.2 Validation against eddy flux measurements**

25

26 At the local scale the inversion shows ability to capture the observed flux variability at monthly
27 scale, as shown for the nBB case (see Fig. 7). The residuals between posterior model and eddy
28 covariance flux-data for monthly and site averaged fluxes show a range of misfits not exceeding
29 $1.04 \text{ gCm}^{-2}\text{day}^{-1}$ which is comparable with Broquet et al. (2013), where misfits up to

1 1.5 $\text{gCm}^{-2}\text{day}^{-1}$ were found using 6 years of data (2002-2007). Of note is that the estimated
2 carbon uptake agrees well with the estimated uptake for 2007 in Broquet et al. (2013) (within the
3 uncertainty range). However, in contrast to the synthetic inversion of Ko16, the real data
4 inversion showed a larger monthly averaged posterior bias equal to $0.40 \text{ gCm}^{-2}\text{day}^{-1}$ compared to
5 the $-0.04 \text{ gCm}^{-2}\text{day}^{-1}$ for the synthetic case. The poorer performance in terms of bias compared to
6 the synthetic case is presumably mainly caused by the representation error. In the synthetic
7 inversion we created a true flux field at the same spatial resolution as the posterior flux
8 estimates, and sampled this true flux distribution at the specific eddy covariance measurement
9 location. This does not include any spatial representation error of the EC measurements
10 (footprint about 1 km) with respect to the spatial resolution of 25 km at which the fluxes are used
11 within the inversion. A further cause for this poorer performance is related to the transport error,
12 as in the synthetic case the same transport was used to create the synthetic observations and to
13 perform the inversion, while in the real data inversions the observed atmospheric mole fraction
14 are a result of real transport which can only be approximated with the transport model used for
15 the inversion.

16 Differences between posterior flux retrievals and observed NEE fluxes at the eddy covariance
17 stations are clearly driven by the crop sites. The good agreement between posterior inverse flux
18 estimates and fluxes measured with the eddy covariance technique at non-crop sites can be
19 attributed to the relatively stable, within the year, land condition. Contrastingly, crop areas are
20 subject to human activities throughout the year. Soil enrichment with organic fertilizers,
21 irrigation and harvesting, can severely influence the carbon balance of the local ecosystem. Thus
22 the poor performance between inverse estimates and eddy covariance flux measurements at crop
23 sites can be linked to the extensive anthropogenic influence on those areas. Further, another
24 difficulty which is common for all the ecosystems, is the fact that atmospheric concentrations
25 implicitly contain more components than just the NEE signal e.g. fire emissions. Such emissions
26 are captured in the atmospheric observations (representative scale of hundreds of km) but might
27 not be captured from the eddy covariance flux measurements which they have a very short
28 representative scale of around 1 km.

29 Posterior fluxes showed a shift by one month earlier (in May), for the maximum carbon uptake
30 (see also fig. 7). An initial hypothesis that this might be driven from sites which are difficult to

1 simulate, such as those located in mountain regions, can not be justified. In specific, mountain
2 sites were excluded in an additional sensitivity analyses, but the temporal shift remains.
3 However, looking into the error of the difference between two months suggests that the flux
4 difference between May and June is not significant. The error of the difference was calculated
5 using a Monte Carlo experiment. Fluxes were averaged over the stations and the monthly
6 differences were calculated. Then we used the standard deviation of the differences over the
7 ensemble members to describe the month-to-month uncertainty.

8

9 **4.3 Reliability of European flux estimates**

10 **4.3.1 Mismatch in bottom-up and top-down methods**

11

12 Of note is the strong flux correction when using a-priori fluxes from VPRM with an uptake of
13 0.96 GtC y^{-1} compared to the 0.3 GtC y^{-1} after the inversion. The large correction of about 0.66
14 GtC y^{-1} corresponds to roughly twice the prior uncertainty. We note that VPRM is a diagnostic
15 model which uses simple light use efficiency and respiration equations and MODIS indices, with
16 parameters optimized to match hourly observations of NEE fluxes (Mahadevan et al., 2008). It
17 does not account for land management and land use changes (i.e. crop harvest, deforestation),
18 thus it will estimate a strong sink even for lands that have been harvested, with the respiration
19 fluxes resulting from the use of the harvest (e.g. as food) not included. Those so-called lateral
20 carbon fluxes, that are seen by the atmospheric inversion, account for approximately 0.165 GtC
21 y^{-1} of the prior-posterior flux difference (Ciais et al., 2008). The rest of the difference of about
22 0.5 GtC y^{-1} might be related to local characteristics of eddy covariance sites, which VPRM is not
23 able to represent. Spatial variations of NEE from VPRM are driven by those of EVI (Enhanced
24 Vegetation Index), which is used at a spatial resolution of 1 km. For example, a crop field with
25 typical dimensions of 100 m – 200 m surrounded by other fields with different crop rotation (and
26 differing phenology) are hard to represent with 1 km resolution EVI (even with the highest
27 possible resolution of 250 m for MODIS reflectances). To quantitatively assess the impact of this
28 representation error in combination with the selection of sites used for the VPRM optimization,
29 the annual domain wide C-budget from VPRM was recalculated after omitting one site per

1 vegetation type at a time and optimizing the VPRM parameters (Jackknife delete-1 method).
2 Detailed results are shown in Table 4. The derived Jackknife standard error amounted to 0.54
3 GtC y^{-1} , with a dominant contribution from the cropland vegetation class (0.50 GtC y^{-1}). This
4 uncertainty can fully explain the mismatch between the a priori and the posterior fluxes, and it
5 emphasizes the importance of site selection and site representativeness in up-scaling local eddy
6 covariance measurements to larger regions.

7 The estimated uncertainty for VPRM fluxes based on jackknifing is larger than the prior
8 uncertainties assumed for the atmospheric inversions. Hence, one could argue that the prior
9 fluxes using VPRM (which indicate a too strong sink) combined with a too small prior
10 uncertainty in the inversion leads to erroneous posterior flux estimates. However the optimized
11 biogenic fluxes from all inversions converge at the annual and domain-integrated scale. A
12 particular example is that of the nBB inversion. Even though the GBIOME-BGCv1 fluxes differ
13 greatly from those produced by VPRM, this inversion is fully in line with the results from the
14 rest of the inversions, indicating that the optimized flux estimates are not biased by the a priori
15 flux fields but instead are driven by the atmospheric data.

16 **4.3.2 Sensitivity to anthropogenic emissions**

17

18 Another source of biospheric flux misrepresentation is the fossil fuel inventories. As mentioned
19 in section 2.3 we do not allow for corrections in anthropogenic emissions, as they are assumed to
20 be better known than the terrestrial fluxes. An overestimation/underestimation in anthropogenic
21 emissions will thus lead to a stronger/weaker biospheric sink in atmospheric inversions. The
22 anthropogenic emissions we use are 0.32 GtCy^{-1} (27%) lower for the EU-12 countries compared
23 to those used by Rivier et al. (2010) (1.2 GtCy^{-1}). Peylin et al. (2011) estimates the difference
24 between national totals for the different emission inventories to be around 10%. In a study by
25 Ciais et al. (2009) uncertainties of total fossil-fuel CO_2 emissions in the European Union 25
26 member states were estimated to 19%, based on four different emission inventories. For the EU-
27 25 countries, EDGAR emissions were found to be 12% larger than the mean of the GAINS
28 (Greenhouse Gas and Air Pollution Interactions and Synergies), UNFCC (United Nations
29 Framework Convention on Climate Change) and CDIAC (Carbon Dioxide Information Analysis
30 Center) inventories (Ciais et al. 2009, table 2). Sensitivity tests with increased prior fossil fuel

1 emissions showed that the added fossil fuel increases the estimated uptake by almost 50%
2 relative to the added anthropogenic emissions. Taking an extreme scenario where the fossil fuel
3 emissions are increased by 17% or 0.3 GtC y^{-1} (resulting in 1.77 GtC y^{-1} compared to 1.47 GtC
4 y^{-1} total emissions for EU-domain), we estimate a European carbon sink for the nBV set up of -
5 $0.51 \pm 0.17 \text{ GtCy}^{-1}$ compared to $-0.38 \pm 0.17 \text{ GtC y}^{-1}$ for the standard nBV case. Thus the
6 additional assumed fossil fuel emissions increased the estimated uptake by 0.13 GtCy^{-1} , which is
7 about 44% of the added anthropogenic emissions. The fact that the resulting increase in the
8 biospheric sink does not fully correspond to the increase in assumed emissions is likely a result
9 of the sparse network, where emissions from regions further away from the measurement sites
10 are not fully registered in the simulated mole fractions.

11 In this study we assumed that anthropogenic emissions are perfectly known (which is a
12 traditional assumption in atmospheric inversions), although this is not the case. As a result of not
13 allowing for a correction in the fossil fuel component, this correction will be added to the
14 correction of the biogenic signal. In this paragraph we already discussed how uncertain fossil
15 fuel emissions may be. Further, we estimated how the uncertainty in the fossil fuel component
16 impacts, the carbon flux estimates; the magnitude but also spatial and temporal flux distributions
17 may be significantly erroneous. For better future carbon flux estimations, fossil fuel optimization
18 seems to be necessary. However, that would require ^{14}C tracer measurements which are
19 currently not available.

20

21 **4.3.3 Sensitivity to site selection**

22

23 Uncertainties in vertical mixing and especially in the nocturnal boundary layer (Gerbig et al.,
24 2008) should be carefully addressed as they might lead to erroneous estimations of the carbon
25 uptake. Typically, in atmospheric inversions the model-data mismatch error (measurement error
26 covariance) accounts also for uncertainties due to the transport (i.e. wrong representation of the
27 nocturnal boundary layer). The set of network stations includes 7 mountain stations, for which
28 we use night-time observations (day-time for non mountain stations) as these measurements are
29 considered to be representative for the free troposphere. Errors can be introduced if the
30 measurement height assumed in the transport model is within the modeled nocturnal stable

1 boundary layer while in the real world it is not, which would lead to an overestimation in the
2 simulated CO₂ signals from respiration or vice versa. In the inversion this would be compensated
3 by introducing stronger uptake fluxes to match the observed CO₂ time series. In order to
4 investigate whether our results are influenced by the use of mountain stations, we performed an
5 additional inversion using the nBV error structure, but excluding all these stations. The resulting
6 sink in Europe was found to be $-0.41 \pm 0.17 \text{ GtCy}^{-1}$ which is fully in line with nBV inversion
7 using all sites, suggesting that our estimates are not biased due to misrepresentation of the
8 mountain stations at least at annual and domain wide aggregation scales.

9 However, the spatial flux distribution seems to depend on the site selection and in particular on
10 the mountain sites used in a given inversion. Ambiguous carbon fluxes e.g. carbon sinks over
11 non productive areas like Alps, England, and west Czech Republic, as well as carbon sources
12 over cultivated lands like western France, Poland and Ukraine were derived from the inversions
13 (fig. 3). Figure 4 presents the annual spatial flux distribution by using a network of stations with
14 no mountain sites (MS0 case) and using an error structure which does not contain a bias term.
15 This sensitivity test is equivalent to the nBB case where we used also the GBIOME-BGC model
16 as prior. Subsequently we plot the flux distribution by adding one mountain site at a time (cases
17 1:7 where the number denotes how many mountain sites are being used). The add-one mountain
18 site sequence is as follows: CMN, OXK, PTR, JFJ, KAS, SIL, PUY. For the MS0 case, we
19 observe that in the region around the Alps, and the neighboring countries, the sink is smaller
20 compared to the rest inversions. The OXK and the KAS sites seem to be responsible for the sink
21 over the Czech Republic. The KAS site seems also to be the driver for the high carbon flux
22 sources around Poland, Ukraine and the Black Sea coasts.

23 Using an error structure which allows for a bias term as the one in BVR case, seems to moderate
24 the spatial flux misrepresentation. Comparing in fig. 3 the subplots nBV: without bias term,
25 BVR: with bias term, we see that the abovementioned highly productive regions (according to
26 the simulation), show somewhat weaker sinks for the BVR case compared to the nBV (indicated
27 by the less bluish contours). Subsequently, regions that appear to be strong carbon sources (in
28 nBV case), show weaker flux signal when the bias term is used (BVR).

1 Although this study uses as much information as possible, in terms of the available atmospheric
2 observations still, large areas are poorly or not constrained at all from the atmospheric network
3 e.g. West France, the whole East European part. Hence, the spatial flux distribution at those
4 areas, is prone to large uncertainties.

5

6

7 **4.3.4 Retrieved fluxes and comparison to previous inverse estimates**

8

9 The retrieved spatially resolved fluxes showed a sensitivity in their spatial patterns to the a priori
10 error structure, specifically to the inclusion of a bias component, as indicated by differences
11 between the nBV and BVR cases. Such differences were not identified in the synthetic
12 experiment in Ko16, however there a much larger spatial correlation length scale was assumed.
13 In the synthetic inversions the long correlation length (766 km at the zonal and 411 km at the
14 meridional direction) drastically reduces the effective number of degrees of freedom, forcing the
15 fluxes to be smoothly corrected, regardless of the use of the bias component. In the real data
16 inversions the shorter correlation length (around 100 km), combined with the required larger
17 error inflation (compared to the synthetic inversions) for the nBV and nBB cases, increases the
18 effective number of degrees of freedom. By using a bias component (BVR, BVN, BVRT cases)
19 or by using the hyperbolic correlation shape (nBVH) with stronger large-scale correlation,
20 instead of inflating the spatiotemporal error component, fluxes remain less flexible at gridscale.

21 Our knowledge regarding annual CO₂ flux estimates for Europe is still highly uncertain, in part
22 due to the limited number of regional inversions focusing on this domain. Flux estimates from
23 previous studies, mainly global inversions, show a wide range (Fig. 8). We estimated an annual
24 European carbon sink (ranging between -0.23 ± 0.13 and -0.38 ± 0.17 GtC y⁻¹ for the different
25 inversion scenarios, Fig. 5 d)), which is however representative for a smaller European region
26 compared to the TransCom European region typically used in other studies. The up-scaled flux
27 estimates (see also section 2.4) for the TransCom EU region have a range of -0.39 to -0.71 GtC
28 y⁻¹. Ciais et al. (2000) estimated a European sink of -0.3 ± 0.8 GtC y⁻¹ for the target period 1985-
29 1995, however in contrast to our study they used a global system and a gap filling algorithm

1 since 42% of the observational data were missing. A recent study from Peylin et al. (2013)
2 computed the mean European sink for the period 1998-2001 to be $-0.44 \pm 0.45 \text{ GtC y}^{-1}$ by
3 utilizing eleven different global inversion systems. Gurney et al. (2004) performed also global
4 inversions and found the mean European annual fluxes for 1992 – 1996 period to be -0.98 ± 0.4
5 GtC y^{-1} which is larger compared to our estimations. Moreover, our results for the mean net
6 monthly fluxes over Europe agreed very well with Rivier et al. (2010) who estimated for the
7 1998-2001 time frame using five different transport models in their inversion that the maximum
8 seasonal uptake occurs in July and lies between -10 and $-80 \text{ gCm}^{-2}\text{month}^{-1}$, while our results
9 show maximum uptake in June with a range of -33 to $-37 \text{ gCm}^{-2}\text{month}^{-1}$ for the different
10 inversion cases. We note that the annual flux differences between our flux estimates and those
11 from other studies may be also caused due to the interannual flux variability. Nevertheless this
12 should not be expected to critically drive those differences since posterior uncertainties found to
13 be larger than interannual variations (Broquet et al., 2013) making the significance of the
14 variations questionable.

15 A recent study from Reuter et al. (2014) based on inversions using satellite observations
16 estimated the carbon budget for the TransCom European region. For the year 2007 the sink was
17 found to be $-1.1 \pm 0.30 \text{ GtC y}^{-1}$, much larger compared to most of other inversion estimates using
18 ground observations. However Feng et al. (2016) tried to investigate why atmospheric inversions
19 using satellite observations, show an elevated European uptake, through a series of sensitivity
20 tests. They linked the increased uptake when using satellite measurements to potential
21 observation biases and to the emission spatial patterns. Further Feng et al. (2016) highlighted that
22 the large European uptake is related up to 60-90 % from systematically higher modeled CO_2
23 fluxes transported into Europe from regions outside of the domain. As this looks to be a problem
24 related with column measurements this is not the case in our study since ground observations
25 were used. In addition we use the two step inversion scheme which limits the influence from the
26 far field as we calculate the concentration signal from outside the domain and subtract that from
27 the observations. Whilst the flux uncertainties outside the domain are not propagated, still they
28 can be expressed as uncertainties in the observation space. However if biases introduced from
29 the global inversion to the fluxes outside of the domain, then regional flux estimations may
30 differ.

1 At national scale we can compare our results to those obtained by Meesters et al. (2012) for the
2 Netherlands, who estimated the annual national carbon sink to about $-0.017 \pm 0.004 \text{ GtCy}^{-1}$. Our
3 estimations are very close, with a range of $-0.012 \pm 0.004 \text{ GtCy}^{-1}$ (BVR inversion) to $-0.014 \pm$
4 0.005 for the nBB inversion. Of note is that the carbon budget estimates for Netherlands agree
5 remarkably well despite the substantial differences between the two studies: Meesters et al.
6 (2012) used an inversion scheme that solves for scaling factors of the gross prior fluxes. Spatial
7 correlations of 100 km were assumed but only for photosynthetic fluxes within the same land use
8 class. In addition the domain of interest (Netherlands) has a stronger constraint as four stations
9 located within the domain were used, while our inversion only uses one station (CBW), with the
10 rest of the stations being at least 360 km away (WES). Both studies assume approximately the
11 same fossil fuel emissions (0.051 GtC y^{-1} vs. 0.053 GtC y^{-1} in Meesters et al. (2012)).

12 **5 Conclusions**

13
14 This study is a follow up work from Kountouris et al. (2016). In this second part, the inverse
15 modeling framework was deployed using real atmospheric data from 16 stations in Europe, to
16 infer biospheric carbon fluxes. Different prior error structures were assumed to investigate how
17 sensitive posterior fluxes are. The results are validated and compared at different temporal and
18 spatial scales. Satisfactory agreement was found when posterior inverse flux estimates were
19 compared against eddy covariance observations at local scale, as well as against previous studies
20 at national and continental scales, which gives us confidence for our carbon flux estimations. We
21 calculated a sink for the European continent which amounts of $-0.23 \pm 0.13 \text{ GtC y}^{-1}$ to $-0.38 \pm$
22 0.17 GtC y^{-1} depending on the assumed prior error structure.

23 A special effort was also made to avoid potential biased flux estimations due to site selection (i.e.
24 heavily polluted sites, or sites that are within the nocturnal boundary layer e.g. mountain
25 stations) by performing inversions using different network configurations. We did not observe
26 any significant impact for domain-wide aggregated fluxes at least for monthly and annual scales.
27 However changes in spatial flux patterns at the pixel scale should be expected, when then
28 network configuration is changed. Further we studied also how sensitive biospheric carbon
29 fluxes are, when wrong fossil fuel emissions are assumed. We found that due to the network

1 sparseness the fossil fuel emissions are not fully captured in the simulated mole fractions which
2 may bias the flux estimates.

3 What do we learn or should we expect then from the top down approach? The current analysis
4 part one and two, suggests that aggregated fluxes at monthly (temporally) and country (spatially)
5 scales can be successfully retrieved from the inversion system. However, retrieving spatially
6 resolved fluxes at finer scales is still rather challenging. Lack of observations for extended
7 European regions, complexity of the terrain especially in mountainous regions as well as the
8 absence of fossil fuel measurements which would otherwise, allow the separation of fossil fuel
9 signals from biospheric signals in observed CO₂ time-series, complete the mosaic of the current
10 problems that regional inversions are facing. Whilst ICOS (Integrated Carbon Observing
11 System) will introduce more stations in the European continent still, inversions should use all the
12 available information; that could be achieved by assimilating multiple data streams like
13 continuous and flask measurements in combination with satellite derived information, aiming to
14 constrain as tight as possible the European continent. Further, new stations should also aim in
15 measuring combustion tracers. That would be of a great help in future inversion systems to be
16 able to update the anthropogenic emission maps and subsequently to compute more accurately
17 the biogenic signal.

18

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1 Table 1. Information on the stations used for the regional inversions. Same network applied for
 2 the synthetic, and the real data inversions in Kountouris et al. (2016). In first column the term
 3 “type” stands for continuous (C) or flask (F) data. Under “Data origin” WDCGG means “World
 4 Data Centre for Greenhouse Gases”.

5

Site Code / type	Name	Latitude (°)	Longitude (°)	Height (m.a.s.l.) (m)	Measurement height (above ground) (m)	Model	Data provider	Data origin	Citation
BAL/F	Baltic Sea, Poland	55.50	16.67	8	57	28	NOAA	Direct contact	Dlugokency et al. 2015
BIK/C	Bialystok, Poland	53.23	23.03	183	90	90	MPI-BGC	Direct access	Popa et al. (2010)
CBW/C	Cabauw, Netherlands	51.58	4.55	-2	200	200	ECN	Direct contact	Vermeulen et al. (2011)
CMN/C	Monte Cimone, Italy	44.18	10.7	2165	12	670	IAFMS	WDCGG	Alemanno et al. (2014)
HEI/C	Heidelberg, Germany	49.42	8.67	116	30	30	University of Heidelberg	CarboEurope	Hammer et al. (2008)
HPB/F	Hohenpeissenberg, Germany	47.80	11.01	934	50	10	NOAA	Direct contact	-
HUN/C	Hegyhatsal, Hungary	46.95	16.65	248	115	96	HMS	WDCGG	Haszpra et al. (2001)
JFJ/C	Jungfrauoch, Switzerland	46.55	7.98	3572	10	720	University of Bern	CarboEurope	-
KAS/	Kasprowy	49.23	19.93	1987	5	480	UKRAK	CarboEur	Necki et

C	Wierch							, AGH	ope	al. (2013)
LMU/ C	La Muela, Spain	41.36	-1.6	570	79	80	Universi ty of Barcelon a	CarboEur ope	-	
MHD/ C	Mace Head, Ireland	53.33	-9.90	25	10	15	LSCE	WDCGG	Ramonet et al. (2010)	
OXX/ C	Ochsenkopf, Germany	50.03	11.81	1022	163	163	MPI- BGC	CarboEur ope	Thompson et al. (2009)	
PRS/ C	Plateau Rosa, Italy	45.93	7.71	3480	-	500	RSE	WDCGG	Ferrarese et al. (2015)	
PUY/ C	Puy De Dome, France	45.77	2.97	1465	10	400	LSCE	CarboEur ope	Lopez et al. (2015)	
SCH/ C	Schauinsland, Germany	47.92	7.92	1205	-	230	UBA	WDCGG	-	
WES/ C	Westerland, Germany	54.93	8.32	12	-	15	UBA	WDCGG	-	

1 Glossary for the data providers: AGH: University of science and Technology Polland, ECN: Energy research Centre
2 of the Netherlands, HMS: Hungarian Meteorological Service, IAFMS: Italian Air Force Meteorological Service,
3 LSCE: Le Laboratoire des Sciences du Climat et de l'Environnement, MPI-BGC: Max Planck Institute for
4 BioGeoChemistry, NOAA: National Oceanic and Atmospheric Administration, RSE: Ricerca sul Sistema
5 Energetico, UBA: Umweltbundesamt, UKRAK: Department of Environmental Physics Polland

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Table 2. Overview of the inversion scenarios. “Shape” describes the internal structure of the bias component (proportional to respiration R or to Net Ecosystem Exchange NEE), and “Time vary” indicates whether the bias component also has temporal variations or not. The fifth column “Prior” represents the terrestrial model used as prior, and “Correlation shape” describes the functional form used for the spatial prior uncertainty correlation, either exponential (E) or hyperbolic (H). The last column indicates whether the full or the reduced station network was assumed.

Inversion code	Bias component	Shape	Time vary	Prior	Correlation shape	No. of Stations
nBV	-	-	-	VPRM	E	16
nBB	-	-	-	GBIOME	E	16
BVR	Yes	R	Flat	VPRM	E	16
BVN	Yes	NEE	Flat	VPRM	E	16
BVRT	Yes	R	Vary	VPRM	E	16
nBV14	-	-	-	VPRM	E	14
nBVH	-	-	-	VPRM	H	16

Table 3. RMSD (first column in ppm) and correlation coefficients (second column) between observations and prior/posterior CO₂ dry mole fractions for daily “daytime” or “nighttime” averaged values and for each station. The third column shows χ_c^2 , the normalized dry mole fraction mismatch per degree of freedom for 7-day averaged residuals, as a measure of how well the data were fitted. The format for each station is as follows: RMSD | r² | χ^2 .

	Prior	nBV	nBB	BVR	BVN	BVRT	nBV14	nBVH
BAL	7.12 0.20 69.35	1.48 0.97 0.89	1.53 0.97 0.93	2.26 0.93 2.04	2.26 0.93 2.03	2.25 0.93 2.02	1.41 0.97 0.83	2.37 0.92 2.07
BIK	8.20 0.52 60.10	2.93 0.93 0.88	3.17 0.92 0.99	3.52 0.90 1.51	3.52 0.90 1.53	3.51 0.90 1.53	2.93 0.93 0.88	3.78 0.88 1.70
CBW	8.71 0.23 83.98	3.43 0.88 2.05	3.49 0.88 2.18	4.09 0.83 2.47	4.09 0.83 2.48	4.09 0.83 2.49	3.42 0.88 1.99	4.33 0.81 2.61
CMN	4.20 0.40 31.73	1.26 0.96 0.16	1.35 0.95 0.19	1.45 0.94 0.19	1.44 0.95 0.19	1.46 0.94 0.21	1.25 0.92 0.15	1.57 0.94 0.26
HEI	14.04 0.37 31.28	6.93 0.84 3.05	7.07 0.83 3.07	7.92 0.79 4.22	7.91 0.79 4.23	7.92 0.79 4.23	-	8.34 0.77 5.17
HPB	5.06 0.43 15.61	1.41 0.91 0.34	1.70 0.94 0.50	2.00 0.96 0.65	2.01 0.91 0.66	2.00 0.91 0.65	1.41 0.96 0.33	2.03 0.91 0.67
HUN	7.44 0.55 66.36	2.58 0.94 0.84	2.74 0.93 0.88	3.07 0.92 1.32	3.08 0.92 1.34	3.08 0.92 1.33	2.58 0.94 0.87	3.43 0.90 1.98
JFJ	4.52 0.03 21.39	1.96 0.77 1.59	2.23 0.72 1.53	2.07 0.75 1.83	2.07 0.75 1.82	2.07 0.75 1.84	1.95 0.78 1.58	2.10 0.74 1.98
KAS	6.35 0.39 52.58	3.41 0.87 2.90	3.43 0.87 2.89	3.88 0.82 3.96	3.88 0.82 3.99	3.87 0.83 3.93	3.29 0.77 2.77	4.01 0.81 4.67
LMU	6.01 0.05 29.00	1.45 0.94 0.29	1.51 0.94 0.28	1.74 0.92 0.59	1.74 0.92 0.58	1.76 0.92 0.60	1.44 0.95 0.29	1.84 0.91 0.68
MHD	4.50 0.21 22.24	1.23 0.94 0.24	1.20 0.94 0.21	1.29 0.92 0.31	1.74 0.93 0.31	1.76 0.94 0.31	1.23 0.94	1.26 0.94 0.27

0.24

OXK	5.39 0.28 38.95	2.45 0.85 0.79	2.52 0.84 0.85	2.78 0.81 1.19	2.78 0.81 1.20	2.79 0.81 1.20	2.41 0.86 0.70	2.98 0.78 1.59
PRS	2.98 0.07 20.75	1.06 0.89 0.46	1.10 0.88 0.49	1.16 0.87 0.52	1.16 0.87 0.52	1.17 0.87 0.52	1.07 0.89 0.45	1.22 0.86 0.53
PUY	4.86 0.29 39.48	2.05 0.87 0.67	2.16 0.86 0.75	2.40 0.82 0.97	2.40 0.82 0.97	2.40 0.82 0.95	2.02 0.88 0.71	2.48 0.81 1.27
SCH	5.18 0.24 41.77	1.90 0.89 0.27	2.00 0.88 0.28	2.23 0.85 0.51	2.23 0.85 0.51	2.23 0.85 0.51	1.84 0.90 0.24	2.38 0.84 0.70
WES	8.06 0.23 41.77	2.21 0.94 0.27	2.00 0.94 0.28	2.23 0.91 0.51	2.23 0.91 0.51	2.23 0.91 0.51	-	2.38 0.90 0.70

Table 4. Results from Jackknife delete-1 statistics for VPRM estimated domain-wide NEE for different vegetation classes and for all of the land area. The uncertainty in NEE from all land area was derived assuming independence in the vegetation class specific uncertainties. Note the strong asymmetry between the fraction of land area covered by the different vegetation classes and the number of eddy covariance sites used, indicating over/under representation: for example 8 crop sites represent 51% of the land area, while 15 grassland sites represent 5.6% of the land area of Europe.

	NEE [GtC/y]	NEE uncertainty [GtC/y]	Number of sites	Fraction of land area [%]
Evergreen forest	-0.165	0.039	16	16.5
Deciduous forest	-0.174	0.020	5	4.4
Mixed forest	-0.025	0.176	2	8.4
Open shrub ^a	-0.201	-	1	13.8
Savanna ^a	-0.012	-	0	0.3
Crop	-0.443	0.502	8	51.0
Grass	0.059	0.026	15	5.6
Total	0.960	0.536	47	100

^aUncertainties for open shrubland and savanna could not be derived due to the lack of representative eddy covariance sites

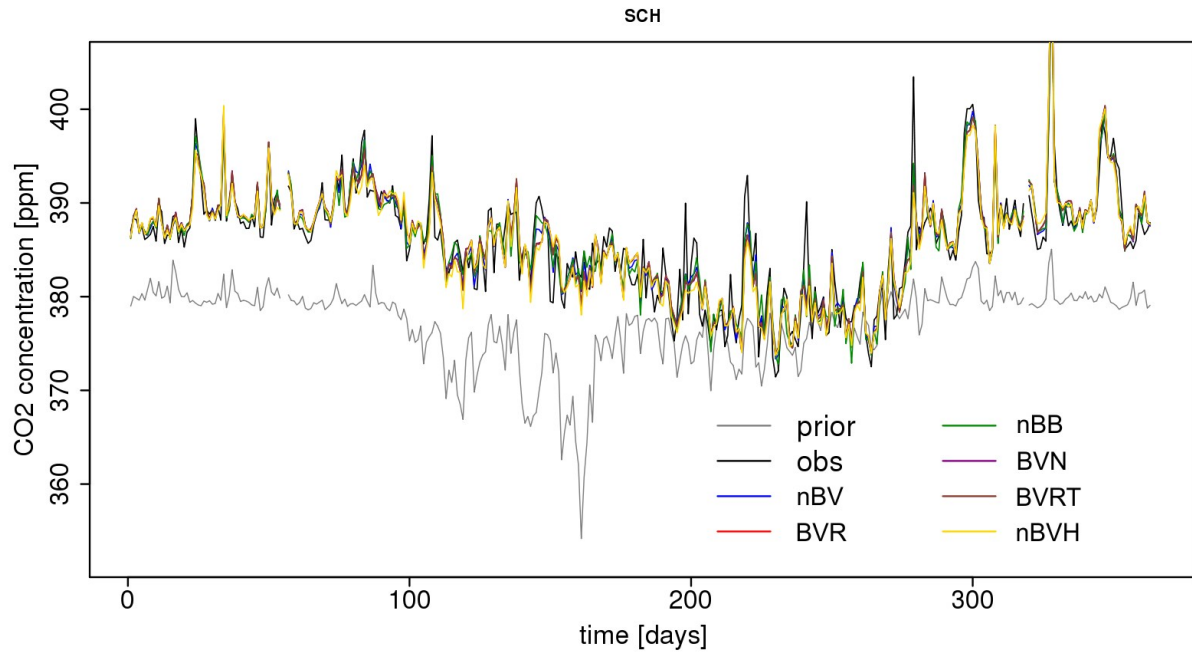


Figure 1. Daily nighttime (23:00-4:00 UTC) averages for prior, true, and posterior CO₂ dry mole fraction time series for the Schauinsland site for the real data inversion. Time starts at 1st January 2007.

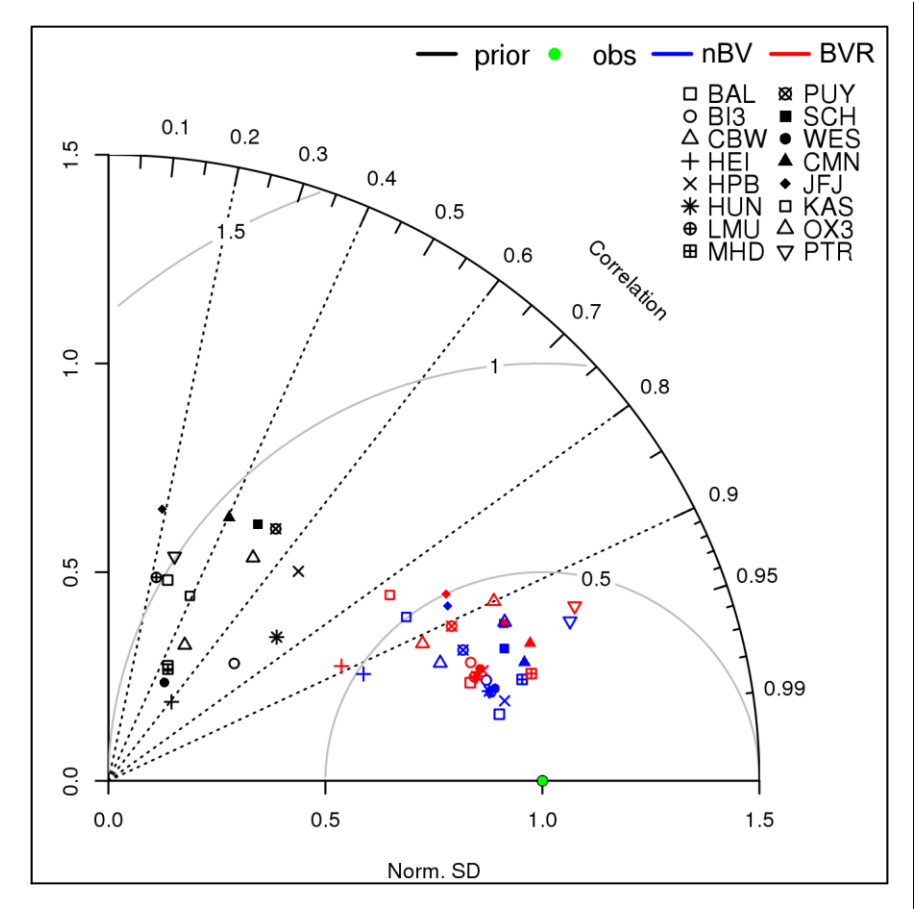


Figure 2 Taylor diagram for modeled and observed time-series of CO₂ dry mole fractions. Prior (black), observed (green, the perfect match of modeled and observed time-series) and the different inversion cases (nBV blue; BVR red) are displayed. Different symbols denote different atmospheric stations. The normalized SD was calculated as the ration of the SD of the modeled time-series to the SD of observations. Gray semi-circles show contours of the standard deviation of the model error.

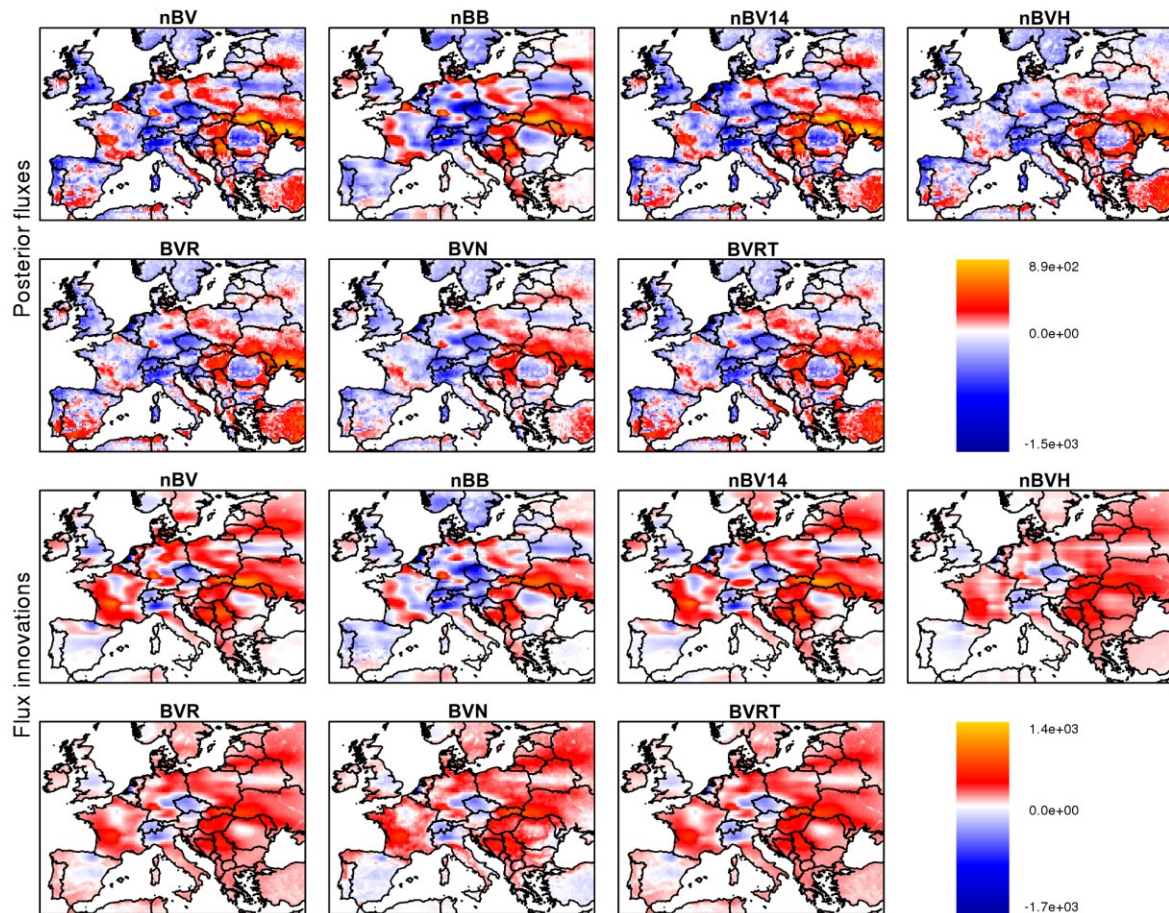


Figure 3. Annual biogenic flux spatial distribution (top two rows) and flux innovations (posterior - prior) (bottom two rows) as estimated from the different inversions for the real data case. Units are in $\text{gCy}^{-1}\text{m}^{-2}$.

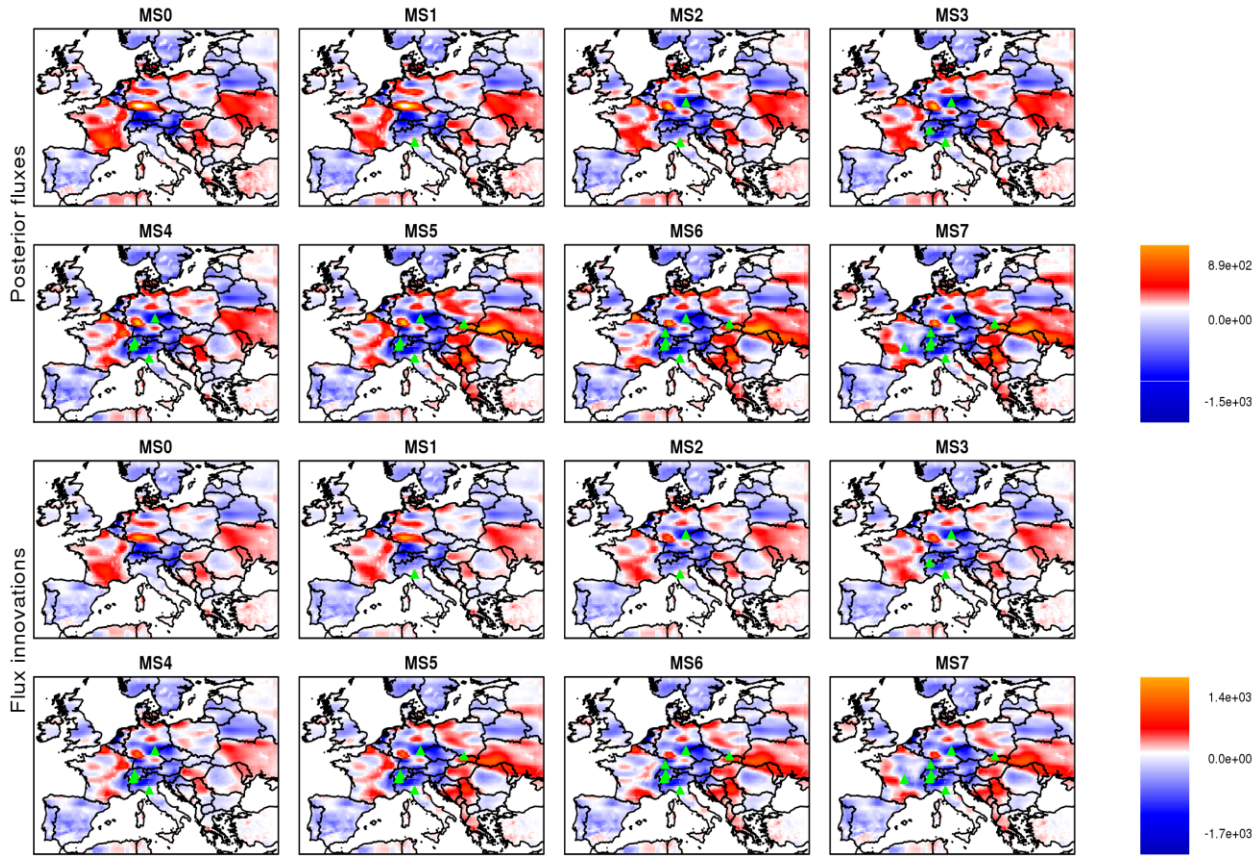


Figure 4. As figure 3, but only for the nBB inversion case. The numbers denote the number of mountain sites used in the inversions e.g. MS0: no mountain site. Units are in $\text{gCy}^{-1}\text{m}^{-2}$.

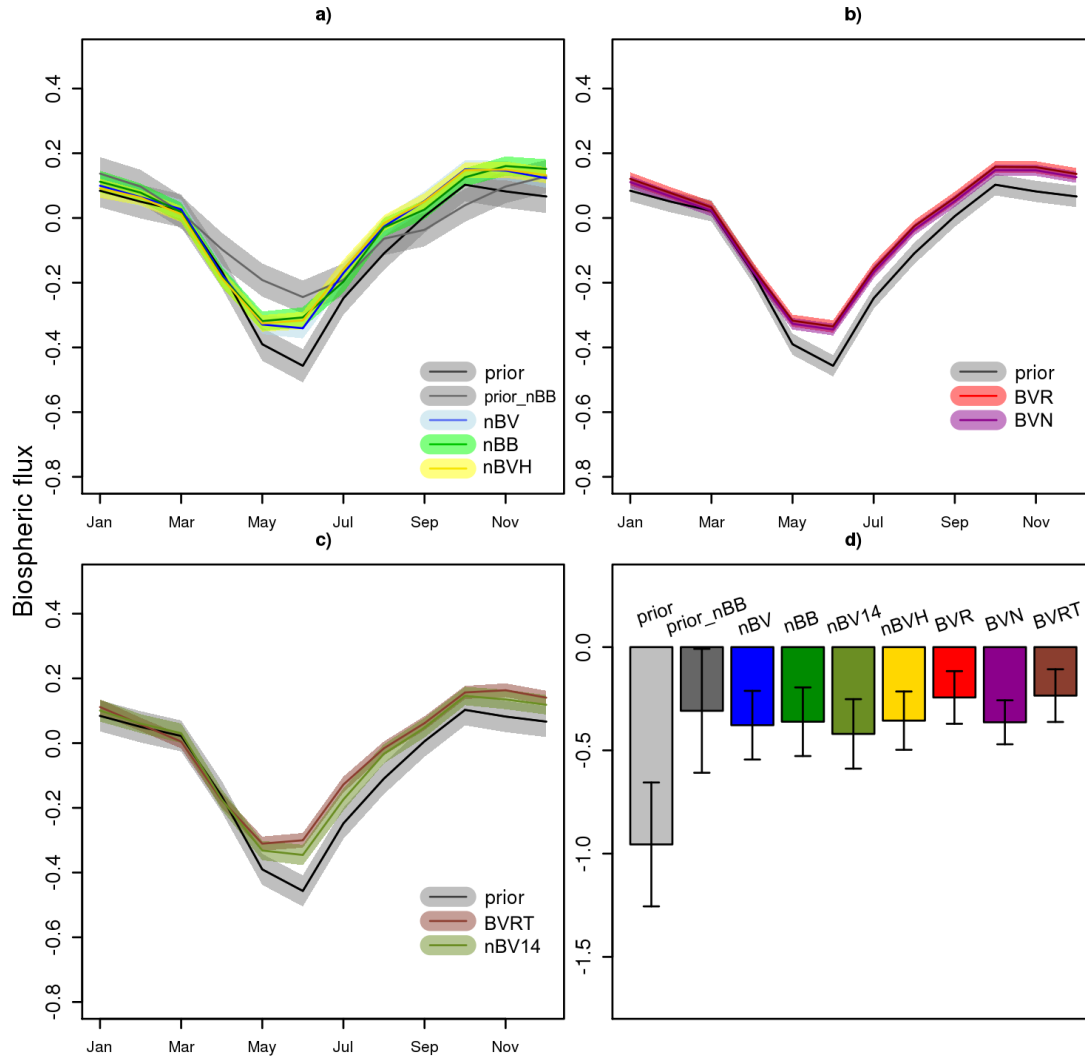


Figure 5. Monthly and annual (panel d) biosphere fluxes integrated over the domain. Panel a) shows nBV, nBB and nBVH cases, b) BVR and BVN and the c) panel shows BVRT and nBV14 cases. Note that all inversions share the same annual prior uncertainty but monthly prior uncertainties differ. Units are in GtC month⁻¹ and GtC y⁻¹ for monthly and annual fluxes, respectively

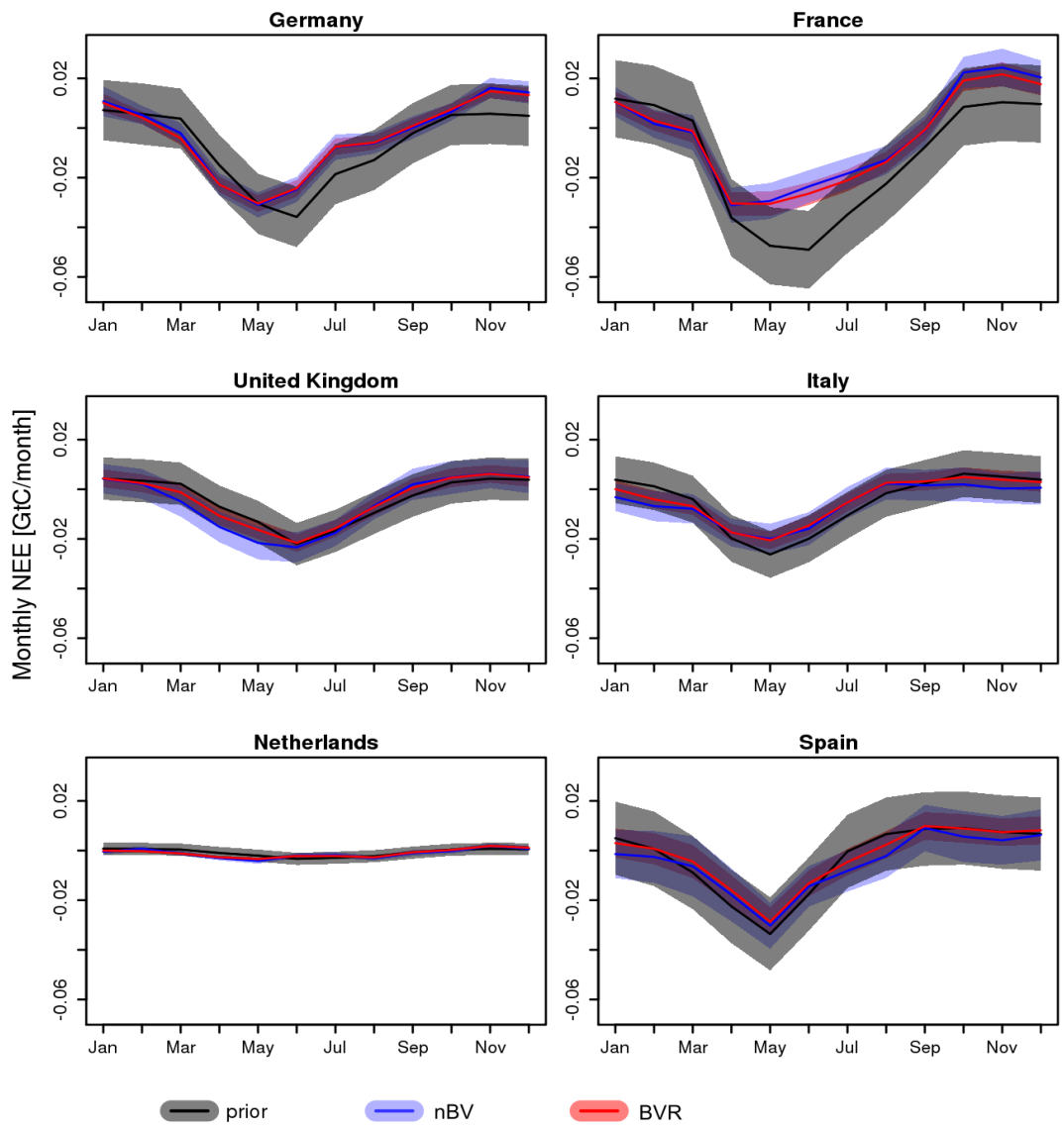


Figure 6. Temporal evolution of prior and posterior monthly NEE for selected European countries.

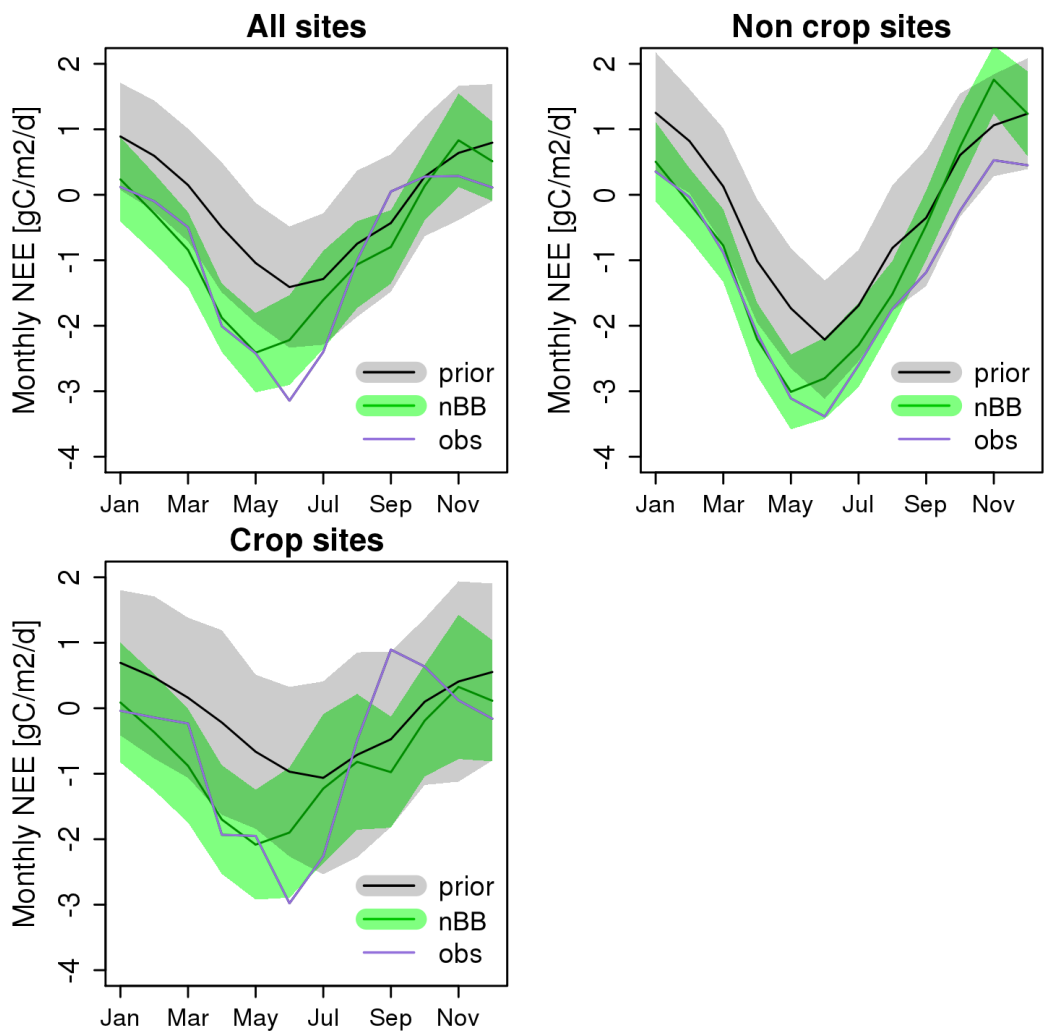


Figure 7. Temporal evolution of monthly NEE ($\text{gCm}^{-2}\text{day}^{-1}$) averaged over all EC sites (top left), excluding crop (top right), and using only crop sites (bottom). Uncertainties (error of the mean monthly NEE) are indicated by the shaded areas.

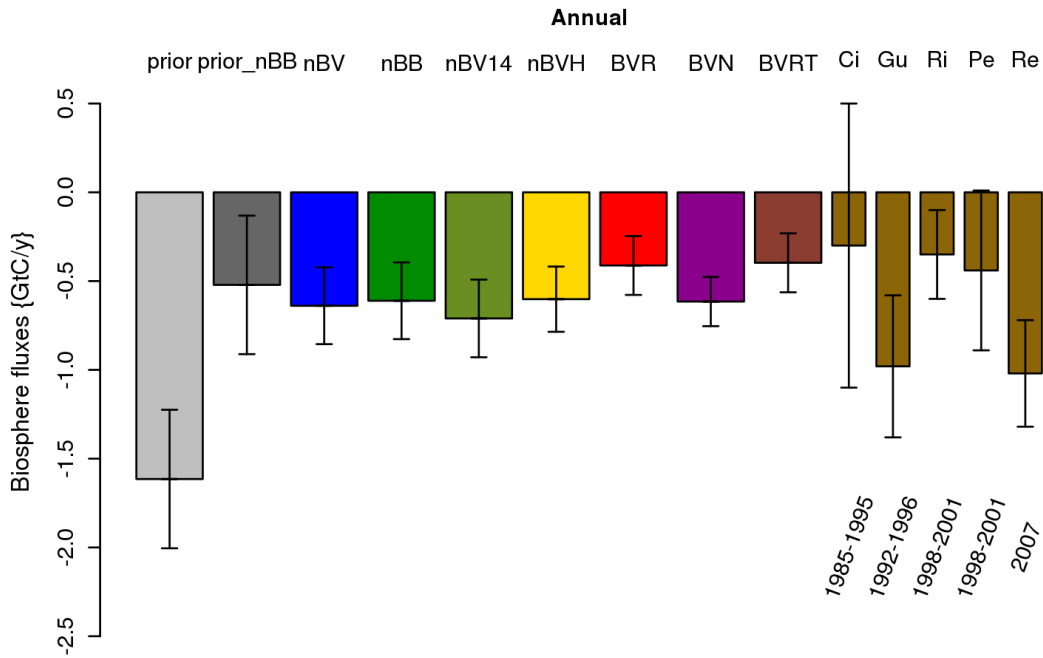


Figure 8. Annual European biogenic CO₂ fluxes in GtCy⁻¹ for the different inversions and comparison to previous studies. Fluxes are upscaled to the TransCom EU domain. Labels of the references are as follows: Ci : Ciais et al. (2000); Gu : Gurney et al. (2004); Ri : Rivier et al. (2010); Pe : Peylin et al. (2013); Re : Reuter et al. (2014). Periods for the inverted fluxes are given below the flux estimates.