

## Author's response:

We thank the Referees for the revision and comments which helped improving the quality of the manuscript. A point-by-point answer (in regular typeset) to the referees' remarks (in the *italic typeset*) follows, while changes to the manuscript are indicated in **blue font**. In the following page and lines references refer to the manuscript version submitted on 14<sup>th</sup> September 2016.

# Anonymous Referee #1

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*I recommend that the authors should add a short bit of text, either to the manuscript or the supplemental that addresses the reviewer 2's comment- that volatility and seasonal trends are linked.*

In the revised manuscript we replaced P22, L21-22 with the following discussion: In general, OOA factors with different seasonal behaviors can be characterized by different volatilities. However in this work the offline-AMS OOA separation is not driven by volatility, given the low correlation between  $\text{NO}_3^-$  and our OOA factors (also reflected by the low  $\text{NO}_3^-$ -related SOA correlation with B-OOA and S-OOA, Table 2). Additionally, the partitioning of semi-volatile OA at low temperatures would lead to a less oxidized OOA fingerprint during winter than in summer; however, this was not the case. We observed a less oxidized OOA factor during summer, whose mass spectral fingerprint closely resembles that of SOA from biogenic precursors. Meanwhile similar to OOA from aging of biomass burning emissions, OOA during the cold season is more oxidized. This has been also reported in an urban environment in central Europe (Zurich) using an online-ACSM (Canonaco et al., 2015). Table 2 was moved below this section.

# Anonymous Referee #2

Received and published: 30 June 2016

## **General Comments:**

*I thank the authors for taking time to revise the manuscript. The authors have addressed the comments adequately. However, I have two minor comments.*

*1. OM/OC ratio. While I agree with the authors that both Aiken and Canagaratna parameterizations are uncertain for this dataset, I want to point out that the OM/OC ratio would affect the recovery ratios determined by Eq. (6). Higher OM/OC ratio from Canagaratna parameterization would lead to lower recovery ratio and hence higher ambient concentration of factors.*

The recovery estimates are independent of the choice of Aiken or Canagaratna's OM:OC parameterizations. Indeed the recovery fitting equation (Eq. 6) explicitly contains the PMF factors OM:OC ratio. However the water-soluble PMF factor concentrations (Eq. 6) implicitly depend on the bulk OM:OC ratio used to determine the bulk WSOM concentration ( $\text{WSOM}_i = \text{WSOC}_i \text{OM/OC}_i$ ) which was used as input for our PMF model. This leads to canceling corrections making the recovery estimates independent of the choice of the Aiken's or Canagaratna's OM:OC parameterizations.

This information was added in the revised SI.

*2. It is important to discuss why the same  $R_z$  is selected for both B-OOA and S-OOA (i.e., response to comment#20) and mention that the  $R_z$  of OOA factors warrants further investigation in the manuscript.*

The factor recoveries determined in this work enabled properly fitting the OC time series according to Eq. (6). The OC fitting residuals were unbiased within our uncertainty in

different seasons (summer and winter) and at the different stations. Therefore there's no reason to consider statistically different recoveries for S-OOA and W-OOA. We also fitted the factor recoveries according to Eq. (6) without any *a-priori* constrain from Daellenbach et al. (2016), and assuming different recoveries for S-OOA and B-OOA. The measured OC vs. fitted OC correlation was not statistically higher (95% confidence interval) than the correlation obtained when constraining the OOAs and BBOA factor recoveries according to Daellenbach et al. (2016). This suggests that the measured OC is equally well explained by the two fits. The completely unconstrained fit returned a wide  $R_{S-OOA}$  range (Fig. D10, only solutions associated to unbiased OC residuals and  $R_k$ s values comprised between 0 and 1 were retained). This occurs despite the considerable contribution of S-OOA, at all sites. This suggests that the least square algorithm fails to independently estimate the recoveries of factors and *a priori* constrains are needed to get unambiguous results. We have assumed  $R_{S-OOA} = R_{B-OOA}$  based on the comparison between offline-AMS and online ACSM, although obtained at another site, especially that this assumption fits our knowledge of OOA water solubility and returned a mathematically equivalent OC reconstruction compared to the completely unconstrained model.

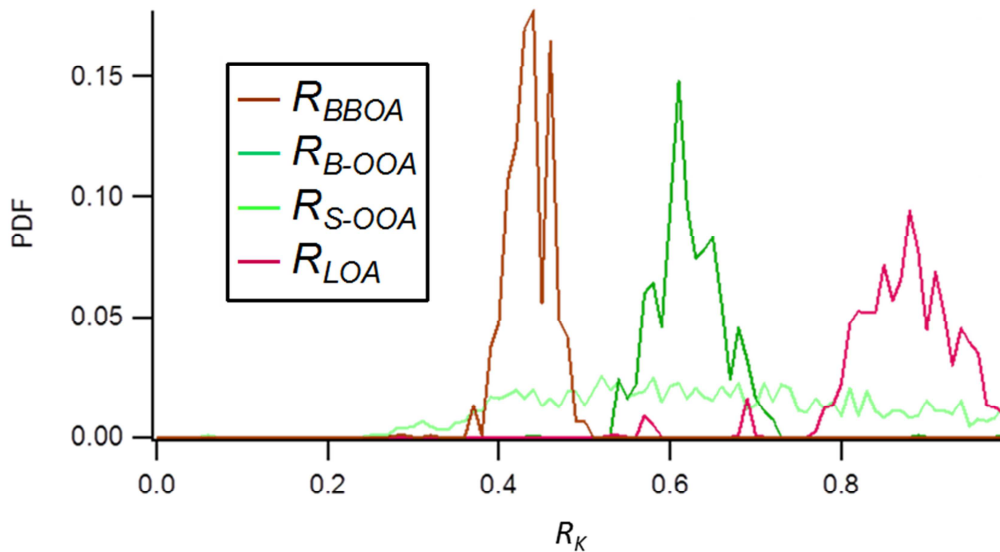


Fig. D10.  $R_k$  probability density functions obtained by fitting Eq. (6) assuming  $R_{S-OOA} \neq R_{B-OOA}$  and without *a-priori*  $R_k$  information.

1 **Argon offline-AMS source apportionment of organic**  
2 **aerosol over yearly cycles for an urban, rural and marine**  
3 **site in Northern Europe**

4  
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18

19 **Abstract**

20 The widespread use of Aerodyne aerosol mass spectrometers (AMS) has greatly improved  
21 real-time organic aerosol (OA) monitoring, providing mass spectra that contain sufficient  
22 information for source apportionment. However, AMS field deployments remain expensive  
23 and demanding, limiting the acquisition of long-term datasets at many sampling sites. The  
24 offline application of aerosol mass spectrometry entailing the analysis of nebulized water  
25 extracted filter samples (offline-AMS) increases the spatial coverage accessible to AMS  
26 measurements, being filters routinely collected at many stations worldwide.

27 PM<sub>1</sub> (particulate matter with an aerodynamic diameter <1 μm) filter samples were collected  
28 during an entire year in Lithuania at three different locations representative of three typical

1 environments of the South-East Baltic region: Vilnius (urban background), Rūgšteliškis (rural  
2 terrestrial), and Preila (rural coastal). Aqueous filter extracts were nebulized in Ar, yielding  
3 the first AMS measurements of water-soluble atmospheric organic aerosol (WSOA) without  
4 interference from air fragments. This enables direct measurement of the  $\text{CO}^+$  fragment  
5 contribution, whose intensity is typically assumed to be equal to that of  $\text{CO}_2^+$ . Offline-AMS  
6 spectra reveal that the water soluble  $\text{CO}_2^+:\text{CO}^+$  ratio not only shows values systematically  $\ll 1$   
7 but is also dependent on season, with lower values in winter than in summer.

8 AMS WSOA spectra were analyzed using positive matrix factorization (PMF), which yielded  
9 4 factors. These factors included biomass burning OA (BBOA), local OA (LOA) contributing  
10 significantly only in Vilnius, and two oxygenated OA (OOA) factors, summer OOA (S-OOA)  
11 and background OOA (B-OOA) distinguished by their seasonal variability. The contribution  
12 of traffic exhaust OA (TEOA) was not resolved by PMF due to both low concentrations and  
13 low water solubility. Therefore, the TEOA concentration was estimated using a chemical  
14 mass balance approach, based on the concentrations of hopanes, specific markers of traffic  
15 emissions. AMS-PMF source apportionment results were consistent with those obtained from  
16 PMF applied to marker concentrations (i.e. major inorganic ions, OC/EC, and organic  
17 markers including polycyclic aromatic hydrocarbons and their derivatives, hopanes, long-  
18 chain alkanes, monosaccharides, anhydrous sugars, and lignin fragmentation products). OA  
19 was the largest fraction of  $\text{PM}_{10}$  and was dominated by BBOA during winter with an average  
20 concentration of  $2 \mu\text{g m}^{-3}$  (53% of OM), while summer-OOA (S-OOA), probably related to  
21 biogenic emissions was the prevalent OA source during summer with an average  
22 concentration of  $1.2 \mu\text{g m}^{-3}$  (45% of OM).

23 PMF ascribed a large part of the  $\text{CO}^+$  explained variability (97%) to the OOA and BBOA  
24 factors. Accordingly we discuss a new  $\text{CO}^+$  parameterization as a function of  $\text{CO}_2^+$ , and  
25  $\text{C}_2\text{H}_4\text{O}_2^+$  fragments, which were selected to describe the variability of the OOA and BBOA  
26 factors.

## 27 **1 Introduction**

28 Atmospheric aerosols affect climate (Lohmann et al., 2004, Schwarze et al., 2006), human  
29 health (Dockery et al., 2005, Laden et al., 2000), and ecosystems on a global scale.  
30 Quantification and characterization of the main aerosol sources are crucial for the  
31 development of effective mitigation strategies. The Aerodyne aerosol mass spectrometer  
32 (AMS, Canagaratna et al., 2007) and aerosol chemical speciation monitor (ACSM, Ng et al.,

1 2011, Fröhlich et al., 2013) have greatly improved air quality monitoring by providing real-  
2 time measurements of the non-refractory (NR) submicron aerosol ( $PM_{10}$ ) components.  
3 Analysis of organic mass spectra using positive matrix factorization (PMF, Paatero, 1997;  
4 Paatero and Tapper, 1994) has enabled the quantitative separation of OA factors, which can  
5 be subsequently related to major aerosol sources and formation processes (e.g. Lanz et al.,  
6 2007; Lanz et al., 2010; Zhang et al., 2011; Ulbrich et al., 2009; Elser et al., 2016 a). Despite  
7 its numerous advantages, AMS field deployment remains expensive and demanding, and  
8 therefore most of the studies are typically restricted to short-time periods and a single (or few)  
9 sampling site(s). The limited amount of long-term datasets suitable for OA source  
10 apportionment severely limits model testing and validation (Aksoyoglu et al., 2011;  
11 Aksoyoglu et al., 2014; Baklanov et al., 2014), as well as for the development of appropriate  
12 pollution mitigation strategies. AMS analysis of aerosol filter samples (Lee et al., 2011; Sun  
13 et al., 2011; Mihara and Mochida, 2011; Daellenbach et al., 2016), which are routinely  
14 collected at many stations worldwide, broadens the temporal and spatial scales available for  
15 AMS measurements.

16 In this study we present the application of the offline-AMS methodology described by  
17 Daellenbach et al. (2016) to yearly cycles of filter samples collected in parallel at three  
18 different locations in Lithuania between September 2013 and August 2014. The methodology  
19 consists of water extraction of filter samples, followed by nebulization of the liquid extracts,  
20 and subsequent measurement of the generated aerosol by high-resolution time-of-flight AMS  
21 (HR-ToF AMS). In this work, organic aerosol water extracts were nebulized in Ar, permitting  
22 direct measurement of the  $CO^+$  ion (Fig. S1), which is typically not directly quantified in  
23 AMS data analysis due to interference with  $N_2^+$ , but is instead estimated as being equal to  
24  $CO_2^+$  (Aiken et al., 2008). Direct measurement of  $CO_2^+$  better captures the variability in the  
25 total OA mass and its elemental composition as well as potentially improving source  
26 apportionment of ambient aerosol. Aerosol elemental ratios and oxidation state are of  
27 particular relevance as they provide important constraints for understanding aerosol sources,  
28 processes, and for the development of predictive aerosol models (Canagaratna et al., 2015).

29 Aerosol composition in the south-east Baltic region has so far received little attention. To our  
30 knowledge the only investigation of OA sources in this area was during a five-day period of  
31 intense land clearing activity occurring in the neighboring Russian enclave of Kaliningrad  
32 (Ulevicius et al., 2016; Dudoitis et al., 2016), in which transported biomass burning emissions

1 dominated the aerosol loading. OA source contributions under less extreme conditions remain  
2 unstudied, with the most relevant measurements performed in Estonia with a mobile lab  
3 during March 2014 at two different locations (Elser et al., 2016b). On-road measurements  
4 revealed large traffic contributions with an increase of 20% from rural to urban environments.  
5 Also, residential biomass burning (BB) and oxygenated OA (OOA) contributions were found  
6 to be substantial.

7 In this study we present a complete source apportionment of the submicron OA fraction  
8 following the methodology described by Daellenbach et al. (2016) in order to quantify and  
9 characterize the main OA sources affecting the Lithuanian air quality. The three sampling  
10 stations were situated in the Vilnius suburb (urban background), Preila (rural coastal  
11 background), and Rūgšteliškis (rural terrestrial background), covering a wide geographical  
12 domain and providing a good overview of the most typical Lithuanian and south-eastern  
13 Baltic air quality conditions and environments. PMF analysis of offline-AMS measurements  
14 are compared with the results reported by Ulevicius et al. (2016) and with PMF analysis of  
15 chemical marker measurements obtained from the same filter samples.

## 16 **2 Sampling and offline measurements**

### 17 **2.1 Site description and sample collection**

18 We collected 24-h integrated PM<sub>1</sub> filter samples at 3 different stations in Lithuania from 30  
19 September 2013 to 2 September 2014 using 3 High-Volume samplers (Digitel DHA80, and  
20 DH-77) operating at 500 L min<sup>-1</sup>. In order to prevent large negative filter artifacts, the high  
21 volume were equipped with temperature control systems maintaining the filter storage  
22 temperature always below 25°C, which is lower or comparable to the maximum daily  
23 temperature during summer. The particulate matter was collected on 150-mm diameter quartz  
24 fiber filters (Pallflex Tissuquartz 2500QAT-UP / pure quartz, no binder) pre-baked at 800°C  
25 for 8 h. Filter samples were wrapped in pre-baked aluminum foils (400°C for 6 h), sealed in  
26 polyethylene bags and stored at -20°C after exposure. Field blanks were collected and stored  
27 following the same procedure.

28 Sampling was conducted at urban (Vilnius), rural terrestrial (Rūgšteliškis) and rural coastal  
29 (Preila) monitoring sites (Fig. 1). The rural terrestrial site of Rūgšteliškis serves as a baseline  
30 against which urban-specific sources in the major population center of Vilnius can be

1 compared. The rural coastal site of Preila provides an opportunity to distinguish terrestrial and  
2 marine sources.

3 The sampling station in Vilnius is located at the Center for Physical Sciences and Technology  
4 campus (54°38' N, 25°10' E, 165 m a.s.l.) 12 km southwest of the city center (population:  
5 535000) and is classified as an urban background site. The site is relatively far from busy  
6 roads, and surrounded by forests to the north/northeast, and by a residential zone to the  
7 south/east. It is ca. 350 km distant from the Baltic coast, and 98 km from the Rūgšteliškis  
8 station (Fig. 1).

9 The station in Preila (55°55' N, 21°04' E, 5 m a.s.l.) is a representative rural coastal  
10 background site, situated in the Curonian Spit National Park on the isthmus separating the  
11 Baltic Sea from the Curonian Lagoon. The monitoring station is located <100 m from the  
12 Baltic shore. The closest populated area is the village of Preila (population: 200 inhabitants),  
13 located 2 km to the south.

14 The rural terrestrial station of Rūgšteliškis (55°26' N and 26°04' E, 170 m a.s.l.) is located in  
15 the eastern part of Lithuania, about 350 km from the Baltic Sea. The site is surrounded by  
16 forest and borders the Utenas Lake in the southwest. The nearest residential areas are  
17 Tauragnai, Utena (12 km and 26 km west of the station, population: 32000 inhabitants) and  
18 Ignalina (17 km southeast of the station, population: 6000 inhabitants).

## 19 **2.2 Offline-AMS analysis**

20 The term *offline-AMS* will be used herein to refer to the methodology described by  
21 Daellenbach et al. (2016) and summarized below. For each analyzed filter sample, four 16-  
22 mm diameter filter punches were subjected to ultrasonic extraction in 15 mL of ultrapure  
23 water (18.2 MΩ cm at 25°C, total organic carbon (TOC) < 3 ppb) for 20 min at 30°C.

24 The choice of water instead of an organic solvent is motivated by two arguments:

- 25 - Water yields the lowest background and hence the highest signal to noise compared to  
26 other highly pure solvents (including methanol, dichloromethane and ethyl acetate).
- 27 - In contrast to the water extraction, the use of organic solvents precludes the  
28 quantification of the organic content in the extracts (e.g. by using a total OC analyzer),  
29 which in turn prevents a quantitative source apportionment.



1 Liquid extracts were then filtered and atomized in Ar ( $\geq 99,998$  % Vol. abs., Carbagas, CH-  
2 3073 Gümliigen, Switzerland) using an Apex Q nebulizer (Elemental Scientific Inc., Omaha  
3 NE 68131 USA) operating at 60°C. The resulting aerosol was then dried by passing through a  
4 Nafion drier (Perma Pure, Toms River NJ 08755 USA), and subsequently analyzed by a HR-  
5 ToF-AMS. 12 mass spectra per filter sample were collected (AMS V-mode,  $m/z$  12-232, 30 s  
6 collection time per spectrum). A measurement blank was recorded before and after each  
7 sample by nebulizing ultrapure water for 12 minutes. Field blanks were measured following  
8 the same extraction procedure as the collected filter samples, yielding a signal not statistically  
9 different from that of nebulized milliQ water. Finally we registered the AMS fragmentation  
10 spectrum of pure gaseous CO<sub>2</sub> ( $\geq 99,7$  % Vol, Carbagas, CH-3073 Gümliigen, Switzerland), in  
11 order to derive its CO<sub>2</sub><sup>+</sup>:CO<sup>+</sup> ratio.

12 Offline-AMS analysis was performed on 177 filter samples in order to determine the bulk  
13 water-soluble organic matter (WSOM) mass spectral fingerprints. In total, 63 filters from  
14 Rūgšteliškis, 42 from Vilnius, and 71 from Preila were measured in Ar. The reader is referred  
15 to DeCarlo et al. (2006) for a thorough description of the AMS operating principles and  
16 calibration procedures.

17 HR-ToF-AMS analysis software SQUIRREL (SeQUential Igor data RetRiEvaL, D. Sueper,  
18 University of Colorado, Boulder, CO, USA) v.1.53G and PIKA (Peak Integration by Key  
19 Analysis) v.1.11L for IGOR Pro software package (Wavemetrics, Inc., Portland, OR, USA)  
20 were utilized to process and analyze the AMS data. HR analysis of the AMS mass spectra was  
21 performed in the  $m/z$  range 12-115.

## 22 **2.3 Supporting measurements**

23 Additional offline analyses were carried out in order to validate and corroborate the offline-  
24 AMS source apportionment results. This supporting dataset was also used as input for PM<sub>1</sub>  
25 source apportionment as discussed below. The complete list of the measurements performed  
26 can be found in Table 1 and Table S1. Briefly, major ions were measured by ion  
27 chromatography (IC; Jaffrezo et al., 1998); elemental and organic carbon (EC, OC) were  
28 quantified by thermal optical transmittance following the EUSAAR2 protocol (Cavalli et al.,  
29 2010); water-soluble OC (WSOC) was measured by water extraction followed by catalytic  
30 oxidation and non-dispersive infrared detection of CO<sub>2</sub> using a total organic carbon analyzer  
31 (Jaffrezo et al., 2005). Organic markers were determined by gas chromatography-mass

1 spectrometry (GC-MS; Golly et al., 2015); high performance liquid chromatography (HPLC)  
 2 associated with a fluorescence detector (LC 240 Perkin Elmer) and HPLC-pulsed  
 3 amperometric detection (PAD; Waked et al., 2014) for 67 composite samples. Composites  
 4 were created merging two consecutive filter samples, but no measurements are available for  
 5 Vilnius during summer. Measurements included 18 polycyclic aromatic hydrocarbons  
 6 (PAHs), alkanes (C21-C40), 10 hopanes, 13 methoxyphenols, 13 methyl-PAHs (Me-PAHs), 6  
 7 sulfur-containing-PAHs (S-PAHs), 3 monosaccharide anhydrides, and 4 monosaccharides  
 8 (including glucose, mannose, arabitol, and mannitol). In this work ion concentrations always  
 9 refer to the IC measurements.

10 Table 1. Overview of supporting measurements. A complete list of measured compounds can  
 11 be found in table S1.

Analytical Method	Measured compounds	Filters measured
IC (Jaffrezo et al., 1998)	Ions	All
	EC/OC	
Thermal optical transmittance using Sunset Lab Analyzer (Birch and Cary, 1996) using EUSAAR2 protocol (Cavalli et al., 2010)		All
TOC analyzer using persulphate oxidation at 100°C of the OM, followed by CO <sub>2</sub> quantification with a non-dispersive infrared spectrophotometer (Jaffrezo et al., 1998)	WSOC	All
HPLC associated with fluorescence detector (LC 240 Perkin Elmer) (Golly et al., 2015, Besombes et al., 2001)	PAHs (table S1)	67 composite samples
GC-MS (with and without derivatization step) (Golly et al., 2015)	S-PAHs, Me-PAHs, alkanes, hopanes, methoxyphenols, others	67 composite samples

HPLC-PAD, (Waked et al., 2014)	Anhydrous sugars, sugars alcohols, monosaccharides	67 composite samples
Chemiluminescence (Environnement S.A., Model AC31M)	NO <sub>x</sub>	Online (Vilnius only)

1 In the following, subscripts *avg*, and *med* will denote average and median values,  
2 respectively.

### 3 3 Source apportionment

4 Positive matrix factorization (PMF, Paatero and Tapper, 1994) is a bilinear statistical model  
5 used to describe the variability of a multivariate dataset as the linear combination of a set of  
6 constant factor profiles and their corresponding time series, as shown in Eq. (1):

$$7 \quad x_{i,j} = \sum_{z=1}^p (g_{i,z} \cdot f_{z,j}) + e_{i,j} \quad (1)$$

8 Here  $x$ ,  $g$ ,  $f$ , and  $e$  denote elements of data, factor time series, factor profiles and residual  
9 matrices, respectively, while subscripts  $i,j$  and  $z$  are indices for time, measured variables, and  
10 factor number. The value  $p$  represents the total number of factors chosen for the PMF  
11 solution. The PMF algorithm iteratively solves Eq. (1) by minimizing the objective function  
12  $Q$ , defined in Eq. (2) Only non-negative  $g_{i,z}$  and  $f_{z,j}$  values are permitted:

$$13 \quad Q = \sum_i \sum_j \left( \frac{e_{i,j}}{s_{i,j}} \right)^2 \quad (2)$$

14 Here the  $s_{i,j}$  elements represent entries in the input error matrix.

15 In this work the PMF algorithm was run in the robust mode in order to dynamically  
16 downweigh the outliers. The PMF algorithm was solved using the multilinear engine-2 (ME-  
17 2) solver (Paatero, 1999), which enables an efficient exploration of the solution space by *a*  
18 *priori* constraining the  $g_{i,z}$  or  $f_{z,j}$  elements within a certain variability defined by the scalar  $a$   
19 ( $0 \leq a \leq 1$ ) such that the modelled  $g_{i,z}$  and  $f_{z,j}$  satisfy Eq. (3):

$$20 \quad \left| \frac{(1-a)f_{z,n}}{(1+a)f_{z,m}} \leq \frac{f_{z,n}}{f_{z,m}} \leq \frac{(1+a)f_{z,n}}{(1-a)f_{z,m}} \right. \quad (3)$$

1 Here  $n$  and  $m$  are any two arbitrary columns (variables) in the normalized F matrix. The  
2 Source Finder toolkit (SoFi, Canonaco et al., 2013, v.4.9) for Igor Pro software package  
3 (Wavemetrics, Inc., Portland, OR, USA) was used to configure the ME-2 model and for post-  
4 analysis. PMF analysis was applied to two complementary datasets: (1) organic mass spectra  
5 from offline-AMS measurements for the apportionment of OM sources and (2) molecular  
6 markers for the apportionment of the measured PM<sub>1</sub> mass. These two analyses are discussed  
7 separately below.

8

### 9 3.1 Offline-AMS PMF

10 In the following section we describe the offline-AMS source apportionment implementation,  
11 optimization and uncertainty assessment. Briefly, we selected the number of PMF factors  
12 based on residual analyses and solution interpretability; subsequently we explored the  
13 rotational uncertainty of our source apportionment model and discarded suboptimal solutions  
14 providing insufficient correlation of factor time series with external tracers.

15 The offline-AMS source apportionment returns the water soluble PMF factor concentrations.  
16 Daellenbach et al. (2016) determined factor specific recoveries (including PMF factor  
17 extraction efficiencies), by comparing offline-AMS and online-ACSM OA source  
18 apportionments. In that work, filter samples were collected for one year during an online-  
19 ACSM monitoring campaign conducted at the same sampling station. Briefly, the factor  
20 recoveries were determined as the ratio between the water soluble OA PMF-factor  
21 concentrations from offline-AMS source apportionment divided by the OA PMF factor  
22 concentrations obtained from ACSM source apportionment. Factor specific recoveries and  
23 corresponding uncertainties were determined for HOA, BBOA, COA, and OOA. In this work  
24 we applied these the factor recoveries from Daellenbach et al. (2016) enabled to scaling the  
25 water soluble factor concentrations retrieved from offline-AMS PMF to the corresponding  
26 bulk OA concentrations. We conducted Aa sensitivity analysis of on the appliedse recoveries  
27 was reported in (Section 3.1.3), and the corresponding uncertainty was propagated to the  
28 source apportionment results.

29 A second solution selection step was carried out on the rescaled solutions as described in  
30 section 3.1.3.

1 In general, the offline-AMS technique assesses less precisely the contribution of the low  
2 water soluble factors. The higher uncertainty mostly stems from the larger PMF rotational  
3 ambiguity when separating factors characterized by low concentration in the filter extracts  
4 (i.e. low water solubility). Nevertheless, the uncertainty is dataset dependent, as the separation  
5 of such sources can be improved in case of distinct time variability of these sources. The low  
6 aqueous concentration of scarcely water soluble sources in fact can be partially overcome by  
7 the large signal/noise characterizing the offline-AMS technique (170 on average for this  
8 dataset).

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9 The offline-AMS source apportionment results presented in this study represent the average  
10 of the retained rescaled PMF solutions, while their variability represents our best estimate of  
11 the source apportionment uncertainty.

### 12 3.1.1 Inputs

13 The offline-AMS input matrices include in total 177 filter samples (62 filters from  
14 Rūgšteliškis, 42 from Vilnius, and 73 from Preila). Each filter sample was represented on  
15 average by 12 mass spectral repetitions to explore the effect of AMS and nebulizer stability  
16 on PMF outputs. A corresponding measurement blank was subtracted from each mass  
17 spectrum. The input PMF matrices included 269 organic fragments fitted in the mass range  
18 (12-115). The input error  $s_{i,j}$  elements include the blank variability ( $\sigma_{i,j}$ ) and the uncertainty  
19 related to ion counting statistic and ion-to-ion signal variability at the detector ( $\delta_{i,j}$ , Allan et  
20 al., 2003; Ulbrich et al., 2009):

$$21 \quad s_{i,j} = \sqrt{\delta_{i,j}^2 + \sigma_{i,j}^2} \quad (4)$$

22 We applied a minimum error to the  $s_{i,j}$  matrix elements according to Ulbrich et al. (2009), and  
23 a down-weighting factor of 3 to all fragments with an average signal to noise lower than 2  
24 (Ulbrich et al., 2009). Input data and error matrices were rescaled such that the sum of each  
25 row is equal to the estimated WSOM concentration, which is calculated as the product of the  
26 measured WSOC multiplied by the OM:OC<sub>i</sub> ratios determined from the offline-AMS PMF  
27 results.

### 1 3.1.2 Overview of retrieved factors and estimate of traffic exhaust OA (TEOA)

2 We used a 4-factor solution to represent the variability of the input data. The 4 separated OA  
3 factors included the following:

4 1/ a biomass burning OA (BBOA) factor highly correlated with levoglucosan originating from  
5 cellulose pyrolysis;

6 2/ a local OA (LOA) factor explaining a large fraction of N-containing fragments variability  
7 and contributing mostly in Vilnius during summer and spring;

8 3/ a background oxygenated-OA (B-OOA) factor showing relatively stable contributions at all  
9 seasons;

10 4/ a summer-OOA (S-OOA) factor showing increasing concentrations with the average daily  
11 temperature.

12 If the number of factors is decreased to 3, a mixed BBOA/B-OOA factor is retrieved, and  
13 significant structure appears in the residuals during winter (Fig. S2, S3, S4). Increasing the  
14 number of factors to 5 and 6, leads to a splitting of OOA factors that cannot be interpreted in  
15 terms of specific aerosol sources/processes (Fig. S2, S3). The further separated OOA factor in  
16 the 5-factor solution possibly derived from the splitting of B-OOA; in fact the sum of the  
17 newly separated OOA and B-OOA in the 5-factor solution correlated well with the B-OOA  
18 time series from the 4-factor solution ( $R = 0.93$ ). Overall, a clear structure removal in the  
19 residual time-series was observed until a number of factors equal to 4 (Fig. S4, S5).

20 We also explored a 5-factor solution in which a hydrocarbon-like OA (HOA) profile from  
21 Mohr et al. (2012) was constrained to estimate the TEOA contribution. However, the water-  
22 soluble TEOA (WSTEOA) contribution to WSOM was estimated as  $0.2\%_{\text{avg}}$  (section 3.1.4),  
23 likely too small for PMF to resolve. We performed 100 PMF runs by randomly varying the  
24 HOA  $\alpha$ -value. The obtained results showed a low TEOA correlation with hopanes ( $R_{\text{max}} =$   
25  $0.25$ ,  $R_{\text{min}} = -0.15$ ) with 45% of the PMF runs associated with negative Pearson correlation  
26 coefficients, supporting the hypothesis that this factor has too small contribution in the water  
27 extracts to be resolved. Therefore, we selected the 4-factor solution as our best representation  
28 of the data, while TEOA was instead estimated by a chemical mass balance (CMB) approach  
29 and not based on AMS mass spectral features.

30 TEOA concentrations are estimated using a CMB approach that assumes hopanes, present in  
31 lubricant oils engines, (Subramanian et al., 2006) to be unique tracers for traffic. However,

1 hopanes can also be emitted upon combustion of different types of fossil fuel, in particular by  
2 coal combustion (Rutter et al., 2009), therefore the traffic contribution estimated here,  
3 although very small (as discussed in the result section) should be considered as an upper  
4 estimate. Still, the  $EC_{2:4}/hopanes$  ratio determined in this work ( $900\pm 100$ ) is consistent with  
5  $EC_{2:4}/hopanes$  for TE ( $1400\pm 900$ ; He et al., 2006; He et al., 2008; El Haddad et al., 2009;  
6 Fraser et al., 1998) and not with the coal  $EC_{2:4}/hopanes$  from literature profiles ( $300\pm 200$ ;  
7 Huang et al., 2014; supplementary information (SI)). To assess the traffic exhaust OC  
8 (TEOC) contribution we used the sum of the four most abundant hopanes (17a(H),21b(H)-  
9 norhopane, 17a(H),21b(H)-hopane, 22S,17a(H),21b(H)-homohopane, and  
10 22R,17a(H),21b(H)-homohopane ( $hopanes_{sum}$ )). The TEOC contribution was estimated from  
11 the average  $hopanes_{sum}/TEOC$  ratio ( $0.0012\pm 0.0005$ ) from tunnel measurements reported by  
12 He et al. (2006), He et al. (2008), El Haddad et al. (2009), and Fraser et al. (1998), where the  
13 four aforementioned hopanes were also the most abundant. In order to rescale TEOC to the  
14 total TEOA concentration we assumed an  $OM:OC_{TEOA}$  ratio of  $1.2\pm 0.1$  (Aiken et al., 2008,  
15 Mohr et al., 2012, Docherty et al., 2011, Setyan et al., 2012). The uncertainty of the estimated  
16 TEOA concentration was assessed by propagating the uncertainties relative to the  
17  $OM:OC_{TEOA}$  ratio (8.3%), the  $hopanes_{sum}/TEOC$  ratio (41.7%), the hopane measurement  
18 repeatability (11.5%), and detection limits ( $7 \text{ pg m}^{-3}$ ).

19

### 20 3.1.3. Source apportionment uncertainty

21 A common issue in PMF is the exploration of the rotational ambiguity, here addressed by  
22 performing 100 PMF runs initiated using different input matrices. We adopted a bootstrap  
23 approach (Davison and Hinkley, 1997) to generate the new input data and error matrices  
24 (Brown et al., 2015). Briefly, the bootstrap algorithm generates new input matrices by  
25 randomly resampling mass spectra from the original input matrices. As already mentioned,  
26 the input matrices contained ca. 12 mass spectral repetitions per filter sample; therefore the  
27 bootstrap approach was implemented in order to resample random filter sample mass spectra  
28 together with the corresponding measurement repetitions. Each newly generated PMF input  
29 matrix had a total number of samples equal to the original matrices (177 samples), although  
30 some of the original 177 filter samples are represented several times, while others are not  
31 represented at all. Overall we resampled on average  $63\pm 2\%$  of the filter samples per bootstrap  
32 run. The generated data matrices were finally perturbed by varying each  $x_{i,j}$  element within

1 twice the corresponding uncertainty ( $s_{ij}$ ) assuming a normal distribution of the errors.  
 2 Solutions were selected and retained according to three acceptance criteria based on PMF  
 3 factor correlations with corresponding tracers: BBOA vs. levoglucosan, B-OOA vs.  $\text{NH}_4^+$ ,  
 4 and S-OOA vs. average daily temperature. In order to discard suboptimal PMF runs, we only  
 5 retained solutions associated with positive Pearson correlation coefficients for each criterion,  
 6 for both the individual stations and the entire dataset. In total 95% of the solutions were  
 7 retained following this approach. We note that no solution was discarded based on the first  
 8 two criteria.

9  
 10 The offline-AMS PMF analysis provides the water-soluble contribution of the identified  
 11 aerosol sources. In order to rescale the water-soluble organic carbon concentration of a  
 12 generic factor  $z$  (WSZOC) to its total OC concentration (ZOC) we used the factor recoveries  
 13 ( $R_z$ ) determined by Daellenbach et al. (2016) according to Eq. (5):

$$ZOC_i = \frac{WSZOC_i}{R_z} \quad (5)$$

15 For each PMF factor (BBOA, W-OOA, and S-OOA), the water-soluble organic carbon  
 16 contribution was determined from the OM:OC ratio calculated from the (water-soluble) factor  
 17 mass spectrum (Aiken et al. 2008). For LOA, whose recovery was not previously reported,  
 18  $R_{LOA}$  was estimated from a single parameter fit according to Eq. (6)

$$OC = TEOC + \frac{WSBBOA}{(OM:OC)_{WSBBOA} \cdot R_{BBOA}} + \frac{WSB-OOA}{(OM:OC)_{WSB-OOA} \cdot R_{OOA}} + \frac{WSS-OOA}{(OM:OC)_{WSS-OOA} \cdot R_{OOA}} + \frac{WSLOA}{(OM:OC)_{LOA} \cdot R_{LOA}} \quad (6)$$

20 Here the water-soluble OA factor concentrations were converted to the corresponding water-  
 21 soluble OC concentrations to fit the measured OC concentrations. For each of the 95 retained  
 22 PMF solutions, Eq. (6) was fitted 100 times by randomly selecting a set of 100  $R_{BBOA}$ ,  $R_{OOA}$   
 23 ~~value~~-combinations from those determined by Daellenbach et al. (2016). Each fit was initiated  
 24 by perturbing the input  $OC_i$  and  $TEOC_i$  within their uncertainties, assuming a normal  
 25 distribution of the errors. Additionally we also perturbed the OC and WSOC inputs in order to  
 26 explore the effect of possible bulk extraction efficiency (WSOC:OC) systematic biases on our  
 27  $R_z$  estimates. Specifically, we assumed an estimated accuracy bias of 5% for each of the  
 28 perturbed parameters, which corresponds to the OC and WSOC measurement  
 29 accuracy.~~Additionally, in order to explore the effect of possible bulk extraction efficiency~~  
 30 ~~(WSOC/OC) systematic measurement biases on our  $R_z$  estimates, we also perturbed the OC,~~  
 31 ~~WSOC (Daellenbach et al., 2016) inputs. Specifically, we assumed an estimated accuracy bias~~



1 ~~of 5% for each of the perturbed parameters, which corresponds to the OC and WSOC~~  
2 ~~measurement accuracy.~~ In a similar way, we also perturbed the input  $R_{BBOA}$  and  $R_{OOA}$   
3 assuming an accuracy estimate of 5% deriving from a possible OC measurement bias in  
4 Daellenbach et al. (2016) which could have affected the  $R_Z$  determination. In total  $9.5 \cdot 10^3$  fits  
5 were performed (Eq. 6) and we retained only solutions (and corresponding perturbed  $R_Z$   
6 combinations) associated with average OC residuals not statistically different from 0 within  
7  $1\sigma$  for each station individually and for summer and winter individually (~8% of the  $9.5 \cdot 10^3$   
8 fits, Fig. S6). The OC residuals of the accepted solutions did not manifest a clear correlation  
9 with the LOA concentration (Fig. S7), indicating that the estimated  $R_{LOA}$  was properly fitted,  
10 without compensating for unexplained variability of the PMF model or biases from the other  
11  $R_z$ . Fig. S8 shows the probability density functions (PDF) of the retained perturbed  $R_z$  which  
12 account for all uncertainties and biases mentioned above.  $R_{LOA,med}$  was estimated to be equal  
13 to 0.66 (1<sup>st</sup> quartile 0.61, 3<sup>rd</sup> quartile 0.69, Fig. S8), while the retained  $R_{BBOA}$  and  $R_{OOA}$  values  
14 ( $R_{BBOA,med}$  0.57, 1<sup>st</sup> quartile 0.55, 3<sup>rd</sup> quartile 0.60;  $R_{OOA,med}$  0.84, 1<sup>st</sup> quartile 0.81, 3<sup>rd</sup> quartile  
15 0.88) were systematically lower than those reported by Daellenbach et al. (2016), reflecting  
16 the lower bulk extraction efficiency (bulk EE =  $WSOC_z/OC$ ) measured for this dataset  
17 (median = 0.59, 1<sup>st</sup> quartile = 0.51, 3<sup>rd</sup> quartile = 0.72 vs. median = 0.74, 1<sup>st</sup> quartile = 0.66, 3<sup>rd</sup>  
18 quartile 0.90 in Daellenbach et al. (2016)). All the retained  $R_k$  combinations are available at  
19 DOI: [doi.org/10.5905/ethz-1007-53](https://doi.org/10.5905/ethz-1007-53).

20 Source apportionment uncertainties ( $\sigma_{S.A.}$ ) were estimated for each sample  $i$  and factor  $z$  as the  
21 standard deviation of all the retained PMF solutions (~8% of the  $9.5 \cdot 10^3$  fits). In addition to  
22 the rotational ambiguity of the PMF model (explored by the bootstrap technique) and  $R_Z$   
23 uncertainty, each PMF solution included on average 10 repetitions for each filter sample, and  
24 hence  $\sigma_{S.A.}$  accounted also for measurement repeatability. In this work, the statistical  
25 significance of a factor contribution is calculated based on  $\sigma_{S.A.,zi}$  (Tables S2 and S3).

26 In general the recovery estimates reported in Daellenbach et al. (2016) represent the most  
27 accurate estimates available, being constrained to match the online-ACSM source  
28 apportionment results. The  $R_Z$  combinations reported by Daellenbach et al. (2016)  
29 demonstrated to positively apply to this dataset, enabling properly fitting the measured Bulk  
30 EE ( $WSOC_z/OC$ ) with unbiased residuals and therefore providing a further confidence on  
31 their applicability (we note that in Eq. 6 we fitted OC as function of  $(R_z)^{-1}$  and  $WSOC_{Z,i}$ ,  
32 therefore  $R_Z$  fitted  $WSOC_z/OC = \text{Bulk EE}$ ). In general further  $R_Z$  determinations calculated

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1 comparing offline-AMS and online-AMS source apportionments would be desirable in order  
2 to provide more robust  $R_Z$  estimates. In absence of a-priori  $R_Z$  values for specific factors (e.g.  
3 for LOA in this study) we recommend constraining the  $R_Z$  combinations reported by  
4 Daellenbach et al. (2016) as a-priori information to fit the unknown recoveries, with the  
5 caveat that the  $R_Z$  combinations reported by Daellenbach et al. (2016) were determined for  
6 filter samples extracted with water following a specific procedure; therefore we recommend  
7 adopting these  $R_Z$  combinations for filter samples extracted in the same conditions.  
8 Nevertheless the  $R_Z$  combinations reported by Daellenbach et al. (2016) should be tested also  
9 for filters water extracted in different conditions to verify whether they can properly fit the  
10 Bulk EE. In case the  $R_Z$  combinations reported by Daellenbach et al. (2016) would not apply  
11 for a specific location or extraction procedure (i.e. not enabling a proper fit of Bulk EE) we  
12 recommend a  $R_Z$  redetermination by comparing the offline-AMS source apportionment results  
13 with well-established source apportionment techniques. In absence of data to perform a well-  
14 established source apportionment, we recommend to fit all the  $R_Z$  to match the bulk EE (i.e.  
15 fitting all the recoveries similarly as in Eq. 6 without constraining any a-priori  $R_Z$  value).

16 ~~In general, the offline-AMS technique assesses less precisely the contribution of the lower~~  
17 ~~water soluble factors. The higher uncertainty mostly stems from the larger PMF rotational~~  
18 ~~ambiguity when separating a factor characterized by low concentration in the aqueous filter~~  
19 ~~extracts. Nevertheless, the uncertainty is dataset dependent, as the separation of source~~  
20 ~~components with low water solubility can be improved in case of distinct time variability~~  
21 ~~characterizing those sources in comparison with the other aerosol sources. The low aqueous~~  
22 ~~concentration of scarcely water soluble sources in fact can be partially overcome by the large~~  
23 ~~signal/noise characterizing the offline-AMS technique (170 on average for this dataset).~~

Comment [BA551]: Section moved up

#### 26 3.1.4. Sensitivity of PMF to the un-apportioned TEOA fraction

27 Despite representing only a small fraction, the un-apportioned water-soluble TEOA  
28 (WSTEOA) contribution could in theory affect the apportionment of the other sources in the  
29 PMF model. To assess this, we performed a PMF sensitivity analysis by subtracting the  
30 estimated WSTEOA concentration from the input PMF data matrix, and by propagating the  
31 estimated WSTEOA uncertainty (section 3.1.2) in the input error matrices. To estimate the

1 WSTEOA concentration we assumed  $R_{TEOA}$  of  $0.11 \pm 0.01$  (Daellenbach et al., 2016) and we  
2 used the HOA profile reported by Mohr et al. (2012) as surrogate for the TEOA mass spectral  
3 fingerprint. This approach is equivalent to constraining both the WSTEOA time series and  
4 factor profile. Overall the WSTEOA contribution to WSOM was estimated as  $0.2\%_{avg}$ ,  
5 making a successful retrieval of WSTEOA unlikely (Ulbrich et al., 2009). Consistently, PMF  
6 results obtained from this sensitivity analysis indicated that BBOA and B-OOA were robust,  
7 showing only 1% difference from the average offline-AMS source apportionment results,  
8 with BBOA increased and B-OOA decreased. S-OOA and LOA instead showed larger  
9 deviations from the average source apportionment results (S-OOA increased by 8% and LOA  
10 decreased by 15%), yet within our source apportionment uncertainties. These results highlight  
11 the marginal influence of the un-apportioned WSTEOA fraction on the other factors.

12

### 13 3.2 Marker-PMF: measured $PM_1$ source apportionment

14 In the following section we describe the implementation of source apportionment using  
15 chemical markers (marker-PMF), as well as its optimization and uncertainty assessment. We  
16 discuss the number of factors and the selection of specific constraints to improve the source  
17 separation. Subsequently we discuss the source apportionment rotational uncertainty, and the  
18 sensitivity of our PMF results to the number of source specific markers, and to the assumed  
19 constraints.

#### 20 3.2.1 Inputs

21 The marker-PMF yields a source apportionment of the entire measured  $PM_1$  fraction (organic  
22 and inorganic). Measured  $PM_1$  is defined here as the sum of EC, ions measured via IC, and  
23 OM estimated from OC measurements multiplied by the  $(OM:OC)_i$  ratio determined from the  
24 offline-AMS PMF results by summing the factor profiles OM:OC ratios weighted by the time  
25 dependent factor relative contributions (rescaled by the recoveries). PMF was used to analyze  
26 a data matrix consisting of selected organic molecular markers, ions measured by IC, EC, and  
27 the remaining OM fraction ( $OM_{res}$ ) calculated as the difference between OM and the sum of  
28 the organic markers already included in the input matrix.  $OM_{res}$  represented on average  
29  $95 \pm 2\%$  of total OM. The marker-PMF analysis is limited by the lack of elemental  
30 measurements (e.g. metals and other trace elements) typically used to identify mineral dust  
31 and certain anthropogenic sources. All markers showing concentrations above the detection

1 limits for more than 25% of the samples were selected as input variables (72 in total). The  
2 PMF input matrices contain 67 composite samples (31 for Rūgštelīškis, 29 for Preila, and 7  
3 for Vilnius). The errors ( $s_{ij}$ ) were estimated by propagating for each  $j$  variable the detection  
4 limits (DL) and the relative repeatability ( $RR$ ) multiplied by the  $x_{ij}$  concentration according to  
5 Eq. (7) (Rocke and Lorenzato, 1995):

$$6 \quad s_{ij} = \sqrt{(DL_j)^2 + (x_{ij} \cdot RR_{i,j})^2} \quad (7)$$

### 7 3.2.2 Number of factors and constraints

8 We selected a 7-factor solution to explain the variability of the measured PM<sub>1</sub> components.  
9 The retrieved factors were biomass burning (BB), traffic exhaust (TE), primary biological  
10 organic aerosol (PBOA), SO<sub>4</sub><sup>2-</sup>-related secondary aerosol (SA), NO<sub>3</sub><sup>-</sup>-related SA, methane  
11 sulfonic acid (MSA)-related SA, and a Na<sup>+</sup>-rich factor explaining the variability of inorganic  
12 components typically related to resuspension of mineral dust, sea salt, and road salt.

13 We first tested an unconstrained source apportionment. This led to a suboptimal separation of  
14 the aerosol sources, with large mixings of PMF factors associated with contributions of  
15 markers originating from different sources. In particular we observed mixing of BB markers  
16 (e.g. levoglucosan) with fossil fuel combustion markers such as hopanes, as well as with  
17 inorganic ions such as NO<sub>3</sub><sup>-</sup> and Ca<sup>2+</sup>. All these markers, although related to different  
18 emission/formation processes, are characterized by similar seasonal trends, i.e. higher  
19 concentrations during winter than in summer. Specifically, the BB tracers increase during  
20 winter because of domestic heating activity, hopanes presumably because of the accumulation  
21 in a shallower boundary layer and lower photochemical degradation, NO<sub>3</sub><sup>-</sup> because of the  
22 partitioning into the particle phase at low temperatures, and Ca<sup>2+</sup> because winter was the  
23 windiest season and therefore was associated with the most intense resuspension.

24 We subsequently exploited the markers' source-specificity to set constraints for the profiles  
25 output by our model: for each individual source, we treated the contribution of the unrelated  
26 source-specific markers as negligible (e.g. we assumed that TE, SA, Na-rich factor and PBOA  
27 do not contribute to levoglucosan). In contrast, the non-source specific variables (EC, OM<sub>res</sub>,  
28 (Me-)PAHs, S-PAHs, inorganic ions, oxalate, alkanes) were freely apportioned by the PMF  
29 algorithm. In a similar way we set constraints for primary markers (e.g. K<sup>+</sup> and Ca<sup>2+</sup>) and  
30 combustion related markers (e.g. PAHs), which are not source-specific but the contribution of

1 which can be considered as negligible in the SA factors. In this case the algorithm can freely  
2 apportion these markers to all the primary factors and combustion-related factors,  
3 respectively.

4 In details, EC, PAHs, and methyl-PAHs were constrained to zero in non-combustion sources,  
5 i.e. all profiles but TE and BB. While EC could partially derive from dust resuspension,  
6 literature profiles for this source suggest an EC contribution below 1% (Chow et al., 2003).  
7 This is expected to be also the case here given the distance of the three stations from  
8 residential areas and busy roads. Methoxyphenols and sugar anhydrides, considered to be  
9 unique BB markers, were constrained to zero in all sources but BB. Similarly, hopanes were  
10 constrained to zero in all factors but TE. We also assumed no contribution from glucose,  
11 arabitol, mannitol, and sorbitol to all secondary factors, and traffic exhaust. The  $\text{SO}_4^{2-}$   
12 contribution from primary traffic emissions was estimated to be negligible, given the use of  
13 desulfurized fuel for vehicles in Lithuania. Likewise, alkane contributions were assumed to be  
14 zero in the SA factors, similar to the contribution of  $\text{Ca}^{2+}$ ,  $\text{Na}^+$ ,  $\text{K}^+$  and  $\text{Mg}^{2+}$  in the SA factors  
15 and TE.

16 The number of factors was increased until no mixing between source-specific markers for  
17 different aerosol sources/processes was observed any more. Secondary sources instead were  
18 explained by three factors because of the distinct seasonal and site-to-site variability of MSA,  
19  $\text{NO}_3^-$  and  $\text{SO}_4^{2-}$ . Oxalate correlated well with  $\text{NH}_4^+$  ( $R=0.62$ ) and the latter well with the sum  
20 of  $\text{SO}_4^{2-}$  and  $\text{NO}_3^-$  equivalents ( $R=0.98$ ). Note that the aforementioned secondary tracers were  
21 not constrained in any factor with the exception of  $\text{SO}_4^{2-}$  contributions which were assumed to  
22 be negligible in the TE factor. Moreover the 7-factor solution showed unbiased residuals  
23 (residual distribution centered at 0 within  $1\sigma$ ) for all the stations together and for each station  
24 individually, while lower order solutions showed biased residuals for at least one station or all  
25 the stations together.

26 PMF results obtained assuming only the aforementioned constraints returned suboptimal  
27 apportionments of  $\text{OM}_{\text{res}}$  and  $\text{Na}^+$  between the BB and the  $\text{Na}^+$ -rich factor, with unusually  
28 high  $\text{OM}_{\text{res}}$  fractional contributions in the  $\text{Na}^+$ -rich factor and unusually high  $\text{Na}^+$   
29 contributions in the BB profile in comparison with literature profiles (Chow et al., 2003;  
30 Huang et al., 2014 and references therein; Schauer et al., 2001). Similarly the  $\text{EC}:\text{OM}_{\text{res}}$   
31 value for TE was substantially lower than literature profiles (El Haddad et al., 2013 and  
32 references therein). Other constraints were therefore introduced to improve the separation of

1 these three variables. Specifically, EC and OM<sub>res</sub> were constrained in the traffic profile to be  
2 equal to 0.45 and 0.27 (*a*-value = 0.5) according to El Haddad et al. (2013), while EC:BB  
3 ratio was constrained to 0.1 (*a*-value = 1) according to Huang et al. (2014) and references  
4 therein. Na<sup>+</sup> was constrained to 0.2% (*a*-value = 1) in BB according to Schauer et al. (2001),  
5 while OM<sub>res</sub> was constrained to zero in the Na<sup>+</sup>-rich factor to avoid mixing with BB. Although  
6 this represents a strict constraint, we preferred avoiding constraining OM<sub>res</sub> to a specific value  
7 for the Na<sup>+</sup>-rich factor which could not be linked to a unique source but possibly represents  
8 different resuspension-related sources (e.g. sea salt, mineral dust and road dust). However, we  
9 expect none of the aforementioned sources to explain a large fraction of the submicron OM<sub>res</sub>  
10 (the OC:dust ratio for dust profiles is 1-15% according to Chow et al., 2003). The sensitivity  
11 of our source apportionment to the constraints listed in this section is discussed in the next  
12 section.

13

### 14 3.2.3. Source apportionment uncertainty and sensitivity analyses

15 We explored the model rotational uncertainty by performing 20 bootstrap PMF runs, and by  
16 perturbing each input  $x_{i,j}$  element within  $2 \cdot s_{i,j}$  assuming a normal distribution of the errors.  
17 Results and uncertainties of the PMF model reported in this paper represent the average and  
18 the standard deviation of the bootstrap runs.

19 As discussed in section 3.2.2, we assumed the contribution of specific markers to be 0 in  
20 different factor profiles. Such assumptions preclude the PMF model to vary the contributions  
21 of these variables from 0 (Eq. 3). In order to explore the effect of such assumptions on our  
22 PMF results we loosened all these constraints assuming variable contributions equal to 50%,  
23 37.5%, 25%, and 12.5% of their average relative contribution to measured PM<sub>1</sub>. In all cases  
24 the *a*-value was set to 1. The average factor concentrations for the 12.5% case and the fully  
25 constrained average bootstrap PMF solutions were not statistically different (confidence  
26 interval of 95%, Fig. S9). Statistically significant differences arose for the of the SO<sub>4</sub><sup>2-</sup>-related  
27 SA in the 50% and 37.5% cases, and the Na<sup>+</sup>-rich factor in the 25% and 37.5% cases,  
28 indicating that loosening the constraints allowed additional rotational uncertainty in  
29 comparison to the uncertainty explored by the bootstrap approach. By contrast, the factors  
30 associated with large relative uncertainties from the marker source apportionment (TE and  
31 PBOA, Table S3) showed the best agreement in terms of concentrations (Fig. S9) with the  
32 fully constrained solution, suggesting that the variability introduced by loosening the

1 constraints did not exceed that already accounted for by the bootstrap approach. As previously  
2 mentioned, the largest contribution discrepancies were observed for the  $\text{SO}_4^{2-}$ -related SA and  
3  $\text{Na}^+$ -rich factor. Looser constraints increased the explained variability of primary components  
4 such as EC, arabitol, sorbitol,  $\text{K}^+$ ,  $\text{Mg}^{2+}$ , and  $\text{Ca}^{2+}$  by the (secondary)  $\text{SO}_4^{2-}$ -related SA factor.  
5 The  $\text{Na}^+$ -rich factor showed increasing contributions from  $\text{OM}_{\text{res}}$  and from BB components  
6 such as methoxyphenols, and anhydrous sugars, which exhibited similar seasonal trends as the  
7  $\text{Na}^+$ -rich factor. None of the marker-PMF factors showed statistically different average  
8 contributions (confidence interval of 95%) when tolerating a variability of the constrained  
9 variables within 12.5% of their relative contribution to  $\text{PM}_{10}$ . Note that with this degree of  
10 tolerance the contribution of OM to the  $\text{Na}^+$ -rich was 28%, which is unrealistically high  
11 compared to typically reported values for OM:dust ratios (<15% Chow et al., 2003).  
12 Therefore, we consider the fully constrained PMF solution to represent best the average  
13 composition of the contributing sources.

14 The marker-PMF source apportionment depends strongly on the input variables (i.e. measured  
15 markers), as these are assumed to be highly source specific. That is, minor sources, such as  
16 MSA-related SA and PBOA, are separated because source-specific markers were used as  
17 model inputs. Meanwhile, more variables were used as tracers for TE and BB  
18 (methoxyphenols (5 variables), sugar anhydrides (3 variables), and hopanes (5 variables)),  
19 which gives more weight to these specific sources. We explored the sensitivity of the PMF  
20 results to the number and the choice of traffic and wood burning markers, by replacing them  
21 with randomly selected input variables. In total 20 runs were performed and the average  
22 contribution of the different sources to  $\text{OM}_{\text{res}}$  was compared with the marker source  
23 apportionment average results, where bootstrap was applied to resample time points. Results  
24 displayed in Fig. S10 are in agreement the apportionment of  $\text{OM}_{\text{res}}$  from BB within 11%<sub>avg</sub>,  
25 highlighting its robustness. The agreement for TE was lower, which is not surprising given  
26 the lower contribution of this source and the smaller number of specific markers (hopanes).  
27 However, these uncertainties were within the marker source apportionment uncertainty (Fig.  
28 S10), implying that the results were not significantly sensitive to the number and the choice of  
29 input markers for BB and traffic exhaust.

30

## 1 4 Results and Discussion

### 2 4.1 PM<sub>1</sub> composition

3 An overview of the measured PM<sub>1</sub> composition can be found in Fig. 1. Measured PM<sub>1</sub>  
4 average concentrations were in general low, with lower values detected at the rural terrestrial  
5 site of Rūgšteliškis ( $5.4 \mu\text{g m}^{-3}_{\text{avg}}$ ) than in Vilnius ( $6.7 \mu\text{g m}^{-3}_{\text{avg}}$ ) and Preila ( $7.0 \mu\text{g m}^{-3}_{\text{avg}}$ ).  
6 OM represented the major fraction of measured PM<sub>1</sub> for all seasons and stations, with  $57\%_{\text{avg}}$   
7 of the mass. The average OM concentrations were higher during winter ( $4.2 \mu\text{g m}^{-3}$ ) than in  
8 summer ( $3.0 \mu\text{g m}^{-3}$ ) at all sites probably to a combination of domestic wood burning activity  
9 and accumulation of the emissions in a shallower boundary layer. For similar reasons, EC  
10 average concentrations showed higher values during winter ( $0.42 \mu\text{g m}^{-3}$ ) than in summer  
11 ( $0.25 \mu\text{g m}^{-3}$ ). During summer, the average EC concentration was ~5 times higher in Vilnius  
12 ( $0.54 \mu\text{g m}^{-3}$ ) than in Preila and Rūgšteliškis ( $0.12$  and  $0.11 \mu\text{g m}^{-3}$ , respectively), indicating  
13 an enhanced contribution from combustion emissions. In the absence of domestic heating  
14 during this period, a great part of these emissions may be related to traffic. During winter, EC  
15 concentrations were comparable at all sites (only 25% higher in Vilnius than in Preila and  
16 Rūgšteliškis). This suggests that a great share of wintertime EC may be related to BB, the  
17 average contribution of which is significant at all stations within  $3\sigma$  (table S2). It should be  
18 noted that the highest measured PM<sub>1</sub> concentrations were detected at the remote rural coastal  
19 site of Preila during three different pollution episodes. In particular, the early March episode  
20 corresponded to the period analyzed by Ulevicius et al. (2016) and Dudoitis et al. (2016), and  
21 was attributed to regional transport of polluted air masses associated to an intense land  
22 clearing activity characterized by large scale grass burning in the neighboring Kaliningrad  
23 region.  $\text{SO}_4^{2-}$  represented the second major component of measured PM<sub>1</sub> ( $20\%_{\text{med}}$ ) at all sites  
24 and seasons. Its average concentration remained rather constant with only slightly higher  
25 concentrations in summer than in winter ( $1.2 \pm 0.7 \mu\text{g m}^{-3}$ , and  $1.1 \pm 0.6 \mu\text{g m}^{-3}$  respectively).  
26 Overall  $\text{SO}_4^{2-}$  concentrations did not show large differences from site-to-site, suggestive of  
27 regional sources. By contrast  $\text{NO}_3^-$  showed a clear seasonality with larger contributions in  
28 winter (average  $0.9 \pm 0.8 \mu\text{g m}^{-3}$  equivalent to 12% of measured PM<sub>1</sub>) than in summer  
29 ( $0.03 \pm 0.03 \mu\text{g m}^{-3}$ ), as expected from its semi-volatile nature.



## 1 4.2 OM source apportionment (Offline-AMS PMF)

2 The apportioned PMF factors were associated to aerosol sources/processes according to their  
3 mass spectral features, seasonal contributions and correlations with tracers. The four  
4 identified factors were BBOA, LOA, B-OOA, and S-OOA, which are thoroughly discussed  
5 below. The TEOA contributions instead were determined using a CMB approach.

6 BBOA was identified by its mass spectral features, with high contributions of  $C_2H_4O_2^+$ , and  
7  $C_3H_5O_2^+$  (Fig. 2), typically associated with levoglucosan fragmentation from cellulose  
8 pyrolysis (Alfarra et al., 2007), accordingly the BBOA factor time series correlated well with  
9 levoglucosan (Pearson correlation coefficient:  $R=0.90$ , Fig. S11). BBOA contributions were  
10 higher during winter and lower during summer (Fig. 3a). We determined the biomass burning  
11 organic carbon (BBOC) concentration from the BBOA time series divided by the  
12 OM:OC<sub>BBOA</sub> ratio determined from the corresponding HR spectrum. The winter  
13 levoglucosan:BBOC ratio was 0.16<sub>med</sub>, consistent with values reported in continental Europe  
14 for ambient BBOC profiles (levoglucosan:BBOC range: 0.10-0.21, Zotter et al., 2014;  
15 Minguillón et al., 2011; Herich et al., 2014).

16 The second factor was defined as LOA because of its statistically significant contribution  
17 (within  $3\sigma$ ) only in Vilnius during summer (table S2), in contrast to other potentially local  
18 primary (e.g. BBOA) and secondary (S-OOA) sources which contributed at all sites. The  
19 LOA mass spectrum was characterized by a high contribution of N-containing fragments  
20 (especially  $C_5H_{12}N^+$ , and  $C_3H_8N^+$ ), with the highest N:C ratio (0.049) among the apportioned  
21 PMF factors (0.029 for BBOA, 0.013 for S-OOA, 0.023 for B-OOA). A similar factor was  
22 also observed by Byčenkienė et al. (2016) using an ACSM at the same station. In that work,  
23 high LOA concentrations were associated with wind directions from N-NW, and the authors  
24 suggested the sludge utilization system of Vilnius (UAB Vilniausvandenys) situated 3.9 km  
25 NW from the sampling station as a probable source.

26 Two different OOA sources (S-OOA and B-OOA) were resolved and exhibited different  
27 seasonal trends. Separation and classification of OOA sources from offline-AMS is typically  
28 different from that of online AMS and ACSM measurements, mainly due to the different time  
29 resolution.

30 Few online-AMS studies reported the separation of isoprene-related OA factor  
31 (Budisulistiorini et al., 2013; Hu et al., 2015, Xu et al., 2015) mostly driven by isoprene

1 epoxides chemistry. Xu et al. (2015) showed that nighttime monoterpene oxidation by nitrate  
 2 radical contributes to less-oxidized OOA. However, the large majority of online-AMS OOA  
 3 factors are commonly classified based on their volatility (semi-volatile OOA and low-  
 4 volatility OOA) rather than on their sources and formation mechanisms. This differentiation is  
 5 typically achieved only for summer datasets when the temperature gradient between day and  
 6 night is sufficiently high, yielding a detectable daily partitioning cycle of the semi-volatile  
 7 organic compounds and  $\text{NO}_3^-$  between the gas and the particle phases. Online AMS datasets  
 8 have higher time resolution than filter sampling, but sampling periods typically cover only a  
 9 few weeks. Therefore the apportionment is driven by daily variability rather than seasonal  
 10 differences. By contrast, in the offline-AMS source apportionment, given the 24-h time  
 11 resolution of the filter sampling and the yearly cycle time coverage, the separation of the  
 12 factors is driven by the seasonal variability of the sources and by the site-to-site differences.

13 In general, OOA factors with different seasonal behaviors can be characterized by different  
 14 volatilities. However in this work the offline-AMS OOA separation is not driven by volatility,  
 15 given the low correlation between  $\text{NO}_3^-$  and our OOA factors (also reflected by the low  $\text{NO}_3^-$ -  
 16 related SOA correlation with B-OOA and S-OOA, Table 2). Additionally, the partitioning of  
 17 semi-volatile OA at low temperatures would lead to a less oxidized OOA fingerprint during  
 18 winter than in summer; however, this was not the case. We observed a less oxidized OOA  
 19 factor during summer, whose mass spectral fingerprint closely resembles that of SOA from  
 20 biogenic precursors. Meanwhile similar to OOA from aging of biomass burning emissions,  
 21 OOA during the cold season is more oxidized. This has been also reported in an urban  
 22 environment in central Europe (Zurich) using an online-ACSM (Canonaco et al.,  
 23 2015). Therefore, the offline-AMS source apportionment separates factors by seasonal trends  
 24 rather than volatility.

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25  
 26 Table 2: Pearson correlation coefficients between non-combustion factors (Other-OA  
 27 components) from offline-AMS and marker-source apportionment.

		<u>Other-OA<sub>marker</sub></u>			
		<u>SO<sub>4</sub><sup>2-</sup>-related SOA</u>	<u>MSA-related SOA</u>	<u>NO<sub>3</sub><sup>-</sup>-related SOA</u>	<u>PBOA</u>
<u>Other-</u>	<u>LOA</u>	<u>0.33</u>	<u>0.16</u>	<u>-0.08</u>	<u>0.10</u>
<u>OA<sub>offline-</sub></u>	<u>B-OOA</u>	<u>0.70</u>	<u>0.22</u>	<u>0.21</u>	<u>0.47</u>
<u>AMS</u>	<u>S-OOA</u>	<u>0.60</u>	<u>0.45</u>	<u>-0.47</u>	<u>0.05</u>

28

1 The resolved B-OOA factor explained a higher fraction than S-OOA. It was associated with  
2 background oxygenated aerosols as no systematic seasonal pattern was observed. However,  
3 B-OOA correlated well with  $\text{NH}_4^+$  ( $R=0.69$ , Fig. S11), and had the highest OM:OC ratio  
4 among the apportioned PMF factors (2.21).

5 Unlike B-OOA, S-OOA showed a clear seasonality with higher contributions during summer,  
6 increasing exponentially with the average daily temperature (Fig. S12a). During summer the  
7 site-to-site S-OOA concentrations were not statistically different within a confidence interval  
8 of 95%, while during winter the site-to-site agreement was lower, possibly due to the larger  
9 model uncertainty associated with the low S-OOA concentrations. A similar S-OOA vs.  
10 temperature relationship was reported by Leaitch et al. (2011) for a terpene dominated  
11 Canadian forest using an ACSM and by Daellenbach et al. (2016) and Bozzetti et al. (2016)  
12 for the case of Switzerland (Fig. S12b), using a similar source apportionment model. This  
13 increase in S-OOA concentration with temperature is consistent with the exponential increase  
14 in biogenic SOA precursors (Guenther et al., 2006). Therefore, even though the behavior of S-  
15 OOA at different sites might be driven by several parameters, including vegetation coverage,  
16 available OA mass, air masses photochemical age and ambient oxidation conditions (e.g.  $\text{NO}_x$   
17 concentration), temperature seems to be the main driver of S-OOA concentrations. Overall  
18 more field observations at other European locations are needed to validate this relation. While  
19 the results indicate a probable secondary biogenic origin of the S-OOA factor, the precursors  
20 of the B-OOA factor are not identified. In section 4.4.2 more insights into the OOA sources  
21 will be discussed.

22 The S-OOA profile showed a  $\text{CO}_2^+/\text{C}_2\text{H}_3\text{O}^+$  ratio of  $0.61_{\text{avg}}$ , placing it in the region of semi-  
23 volatile SOA from biogenic emissions in the  $f_{44}/f_{43}$  space (Ng et al., 2011), as attributed by  
24 Canonaco et al. (2015). Despite the higher summer photochemical activity, the water-soluble  
25 bulk OA showed more oxidized mass spectral fingerprints during winter ( $\text{O:C}=0.61_{\text{avg}}$ ) than  
26 in summer ( $\text{O:C}=0.55_{\text{avg}}$ ), similar to the results presented by Canonaco et al. (2015) for  
27 Zurich. Accordingly, the S-OOA profile also showed a less oxidized water-soluble mass  
28 spectral fingerprint than B-OOA, with an O:C ratio of  $0.40_{\text{avg}}$ , in comparison with  $0.80_{\text{avg}}$  for  
29 B-OOA. Considering the sum of B-OOA and S-OOA, the median OOA: $\text{NH}_4^+$  ratios for  
30 Rūgšteliškis, Preila, and Vilnius were 3.2, 2.4, and 2.5 respectively, higher than the average  
31 but within the range of the values reported by Crippa et al. (2014) for 25 different European  
32 rural sites ( $2.0_{\text{avg}}$ ; minimum value 0.3; maximum 7.3).

1

### 2 **4.3 PM<sub>1</sub> source apportionment (marker-PMF)**

3 The PMF factors in this analysis were associated with specific aerosol sources/processes  
4 according to their profiles, seasonal trends and relative contributions to the key variables. Fig.  
5 4 displays factor profiles, and the relative contribution of each factor to each variable. The  
6 Na<sup>+</sup>-rich factor explained a large part of the variability of Ca<sup>2+</sup>, Mg<sup>2+</sup>, and Na<sup>+</sup> (Fig. 4) and  
7 showed higher contributions during winter than in summer (Fig. 5), suggesting a possible  
8 resuspension of sand and salt typically used during winter in Lithuania for road de-icing. This  
9 seasonal trend is also consistent with wind speed, which showed the highest monthly values  
10 during December 2013 and January 2014. We cannot exclude the possibility that this factor  
11 may include contributions from sea salt, although Na<sup>+</sup> and Cl<sup>-</sup> were not enhanced at the  
12 marine station in comparison with the other stations. The overall contribution of this Na<sup>+</sup>-rich  
13 factor to measured PM<sub>1</sub> was relatively small (1%<sub>avg</sub>), but may be larger in the coarse fraction.

14 The BB factor showed a well-defined seasonality, with high contributions during winter. This  
15 factor explained a large part of the variability of typical wood combustion tracers such as  
16 methoxyphenols, sugar anhydrides (including levoglucosan, mannosan, and galactosan), K<sup>+</sup>,  
17 Cl<sup>-</sup>, EC, PAHs, and methyl-PAHs (Fig. 4). Using the OM:OC<sub>BBOA</sub> ratio (1.88) calculated  
18 from offline-AMS, we estimated the levoglucosan:BBOC ratio to be 0.18<sub>avg</sub>, which is within  
19 the range of previous studies (Ulevicius et al., 2016 and references therein). Note that this  
20 factor explained also large fractions of variables typically associated with non-vehicular fossil  
21 fuel combustion, such as benzo(b)naphtho(2,1-d)thiophene (BNT[2,1]) and 6,10,14-trimethyl-  
22 2-pentadecanone (DMPT, Fig. 4, Manish et al., 2007; Subramanian et al., 2007), indicating a  
23 potential mixing of BB with fossil fuel combustion sources. However, the fossil fuel  
24 combustion contribution to BB is unlikely to be large, considering the low concentrations of  
25 fossil fuel tracers such as hopanes (66% of the samples below quantification limit (<QL)),  
26 BNT[2,1] (64%<QL), and DMPT (55%<QL). Moreover, the above mentioned agreement of  
27 the levoglucosan:BBOC ratio with previous studies corroborates the BB estimate from the  
28 marker-PMF.

29 The traffic exhaust factor explained a significant fraction of the alkane variability, with a  
30 preferential contribution from light alkanes (Fig. 4). Its contribution was statistically  
31 significant within 3σ. However on average the concentration was higher in Vilnius than at the  
32 other stations and in general higher in winter than in summer.

1 The PBOA factor explained the variability of the primary biological components, such as  
2 glucose, mannitol, sorbitol, arabitol, and alkanes with an odd number of carbon atoms  
3 (consistent with Bozzetti et al., 2016 and references therein). Highest PBOA concentrations  
4 were observed during spring, especially at the rural site of Rūgšteliškis. Overall the  
5 contribution of this factor was uncertain with an average relative model error of 160%  
6 probably due to the small PBOA contributions ( $0.6\%_{\text{avg}}$  of the total OM), which hampers a  
7 more precise determination by the model. In particular  $\text{OM}_{\text{res}}$  was the variable showing the  
8 highest mass contribution to the PBOA factor. However, the large contribution and the large  
9 uncertainty of  $\text{OM}_{\text{res}}$  to this factor ( $0.3\pm 0.4$ ) resulted in a large uncertainty in the PBOA  
10 estimated concentration.

11 The last three factors were related to SA, as indicated by the large contributions of secondary  
12 species such as oxalate,  $\text{SO}_4^{2-}$ , MSA, and  $\text{NO}_3^-$  to the factor profiles (Fig. 4). The three factors  
13 showed different spatial and temporal contributions.

14 The  $\text{NO}_3^-$ -related SA exhibited highest contributions during winter, suggesting temperature-  
15 driven partitioning of secondary aerosol components. Moreover the  $\text{NO}_3^-$ -related SA,  
16 similarly to BB and TE, showed the highest concentrations in Vilnius, and the lowest in  
17 Rūgšteliškis suggesting its possible relation with anthropogenic gaseous precursors (e.g.  
18  $\text{NO}_x$ ), as already reported in other studies (e.g. Xu et al., 2016; McMeeking et al., 2012).

19 The MSA-related SA factor manifested the highest concentrations at the marine site of Preila  
20 during summer, and in general larger contributions during summer than winter, suggesting its  
21 relation with marine secondary aerosol. MSA has been reported to be related to marine  
22 secondary biogenic emissions deriving from the photo-oxidation of dimethyl sulfide (DMS)  
23 emitted by the phytoplankton bloom occurring during the warm season (Li et al., 1993,  
24 Crippa et al., 2013 and references therein).

25 The last factor ( $\text{SO}_4^{2-}$ -related SA) showed higher contributions during summer than in winter  
26 without clear site-to-site variability, following the seasonal behavior of  $\text{SO}_4^{2-}$  showing slightly  
27 higher concentrations during summer than in winter, which is probably driven by the  
28 secondary formation from gaseous photochemical reactions and aqueous phase oxidation.  
29 This factor explained the largest part of the oxalate and  $\text{SO}_4^{2-}$  variability and represented  
30  $48\%_{\text{avg}}$  of the measured  $\text{PM}_1$  by mass.

#### 1 4.4 Comparison of the source apportionment methods

2 In this section we compare the offline-AMS PMF and marker-PMF results. We begin with  
3 BBOA and TE emissions which were resolved by marker-PMF and offline-AMS (TEOA was  
4 not resolved by offline-AMS but determined through a CMB approach). The remaining OM  
5 fraction ( $\text{Other-OA} = \text{OA} - \text{BBOA} - \text{TEOA}$ ) was apportioned by the offline-AMS source  
6 apportionment to B-OOA, S-OOA and LOA ( $\text{Other-OA}_{\text{offline-AMS}}$ ). However, the LOA  
7 contribution was statistically significant (within  $3\sigma$ ) only in Vilnius during summer (Table  
8 S2), while no data were available for these periods from the marker source apportionment.  
9 The marker source apportionment instead attributed the Other-OA mass fraction to 4 factors  
10 ( $\text{Other-OA}_{\text{marker}}$ ): PBOA, as well as to  $\text{SO}_4^{2-}$ ,  $\text{NO}_3^-$ , and MSA-related secondary organic  
11 aerosols (SOA, Fig. S13). The OA concentrations of the factors retrieved from the  $\text{PM}_{10}$   
12 markers source apportionment were obtained by multiplying the factor time series by the sum  
13 of the organic markers and  $\text{OM}_{\text{res}}$  contributions to the normalized factor profiles. The PM  
14 concentrations from the marker PMF factors are displayed in Fig. 5.

##### 15 4.4.1 Primary OA sources

16 Offline-AMS and marker source apportionments provided comparable BBOA estimates, with  
17 concentrations agreeing within a 95% confidence interval (Fig. 6). Results revealed that  
18 BBOA contributed the largest fraction to the total OM during winter in Preila and Vilnius,  
19 while in Rūgšteliškis the largest OA source derived from B-OOA. The average winter BBOA  
20 concentration was  $1.1 \pm 0.8 \mu\text{g m}^{-3}$  in Rūgšteliškis and  $2 \pm 1 \mu\text{g m}^{-3}$  in Vilnius (errors in this  
21 section represent the standard deviation of the temporal variability). Overall the average  
22 BBOA concentrations were higher at the urban background site of Vilnius and lower at the  
23 rural terrestrial site of Rūgšteliškis. Preila showed higher values ( $3 \pm 3 \mu\text{g m}^{-3}$ ) driven by the  
24 grass burning episode occurred at the beginning of March (Ulevicius et al., 2016). Excluding  
25 this episode, the BBOA winter concentration was lower than in Vilnius ( $1.8 \mu\text{g m}^{-3}$ ). During  
26 winter, considering only the samples concomitantly collected, Preila and Vilnius showed well  
27 correlated BBOA time series ( $R = 0.91$ ) and significantly positive correlations were observed  
28 for also for Preila and Rūgšteliškis ( $R = 0.72$ ) and for Vilnius and Rūgšteliškis ( $R = 0.66$ )  
29 (offline-AMS BBOA time series). These results highlight the effect of regional  
30 meteorological conditions on the BBOA daily variability in the south east Baltic region.

1 By contrast, during summer BBOA concentrations were much lower, with 40% of the points  
2 showing statistically not significant contributions within  $3\sigma$  for the offline-AMS source  
3 apportionment and 100% for the marker source apportionment. Between late autumn and  
4 early March the offline-AMS source apportionment revealed three simultaneous episodes  
5 with high BBOA concentrations at the three stations, while the marker source apportionment  
6 which is characterized by lower time resolution did not capture some of these episodes. The  
7 first episode occurred between 19 and 25 December 2013 during a cold period with an  
8 average daily temperature drop to  $-9.7\text{ }^{\circ}\text{C}$  as measured at the Rūgštelėškis station (no  
9 temperature data were available for the other stations). The third episode occurred between 5  
10 and 10 March 2014 and was associated with an intense grass burning episode localized mostly  
11 in the Kaliningrad region (Ulevicius et al., 2016, Dudoitis et al., 2016, Mordas et al., 2016).  
12 The episode was not associated with a clear temperature drop, with the highest concentration  
13 ( $14\text{ }\mu\text{g m}^{-3}$ ) found at Preila on 10 March 2014, the closest station to the Kaliningrad region.  
14 Similarly, at the beginning of February high BBOA concentrations were registered at the  
15 three stations, without a clear temperature decrease. Other intense BBOA events were  
16 detected but only on a local scale, with intensities comparable to the regional scale episodes.  
17 Using the OM:OC<sub>BBOA</sub> ratio calculated from the HR water-soluble BBOA spectrum (1.88),  
18 we estimated the BBOC<sub>avg</sub> concentrations during the grass burning episode (5-10 March  
19 2014) to span between  $0.8$  and  $7.2\text{ }\mu\text{g m}^{-3}$ . On a daily basis our BBOC concentrations are  
20 consistent with the estimated ranges reported by Ulevicius et al. (2016) for non-fossil primary  
21 organic carbon ( $0.6$ - $6.9\text{ }\mu\text{g m}^{-3}$  during the period under consideration), showing also a high  
22 correlation ( $R=0.98$ ).

23 TEOA estimates obtained by CMB and marker-PMF always agreed with each other within  $3\sigma$   
24 (Fig. 6). The two approaches confirm that TEOA is a minor source (Fig. 6).  
25 Consistently Hopane concentrations (used in this work as TEOA tracers), ~~concentrations~~  
26 were below detection limits ( $7\text{ pg m}^{-3}$ ) for 66% of the collected samples. Similarly to  $\text{NO}_x$ ,  
27 hopanes, showed a clear spatial and seasonal variability with higher concentrations in Vilnius  
28 during winter, suggesting an accumulation of traffic emissions in a shallower boundary layer  
29 (Fig. 3b,  $\text{NO}_x$  data available only for Vilnius). During the grass burning event, we observed a  
30 peak in the total hopane concentration, and therefore also a peak of the estimated TEOA ( $2.4$   
31  $\mu\text{g m}^{-3}$  maximum value). This relatively high concentration is most probably not due to a local  
32 increase of TE, but rather due to a regional transport of polluted air masses from neighboring

1 countries (Poland and the Russian Kaliningrad enclave). By assuming an OM:OC<sub>TEOA</sub> ratio of  
2 1.2±0.1 (Aiken et al., 2008, Mohr et al., 2008, Docherty et al., 2011, Setyan et al., 2012), we  
3 determined the corresponding organic carbon content (TEOC). Our TEOC concentration was  
4 consistent within 3σ with the average fossil primary OC over the whole episode estimated by  
5 Ulevicius et al. (2016), (0.4-2.1 μg m<sup>-3</sup>) although on a daily basis the agreement was relatively  
6 poor.

7 Overall, offline-AMS source apportionment and marker-PMF returned comparable results for  
8 BBOA. Similarly the TEOA estimate by markers-PMF and CMB were comparable, therefore  
9 not surprisingly the two approaches yielded OA concentrations also for the Other-OA  
10 fractions which agreed within 3σ.

#### 11 4.4.2 Other-OA sources: offline-AMS and marker-source apportionment 12 comparison

13 The marker-source apportionment, in comparison to the offline-AMS source apportionment  
14 enables resolving well-correlated sources (e.g. BBOA and NO<sub>3</sub><sup>-</sup>-related SOA) as well as  
15 minor sources (e.g. MSA-related SOA and PBOA) because source-specific markers were  
16 used as model inputs. By contrast, the offline-AMS source apportionment is capable of  
17 resolving OA sources for which no specific markers were available such as LOA, which was  
18 separated due to the distinct spatial and temporal trends of some N-containing AMS  
19 fragments. We first briefly summarize the Other-OA factor concentrations and their site-to-  
20 site differences retrieved by the two techniques; subsequently we compare the two source  
21 apportionment results.

22 The Other-OA<sub>offline-AMS</sub> factor time series are displayed in Fig. S13. The B-OOA factor  
23 showed relatively stable concentrations throughout the year with 0.9±0.8<sub>avg</sub> μg m<sup>-3</sup> during  
24 summer and 1.1±0.9<sub>avg</sub> μg m<sup>-3</sup> during winter. Although B-OOA concentrations were relatively  
25 stable throughout the year, higher contributions were observed in Preila and Rūgšteliškis  
26 compared to Vilnius. The extreme average seasonal concentrations were between 0.8 and 1.3  
27 μg m<sup>-3</sup> at Rūgšteliškis during fall and winter, between 0.9 and 1.1 μg m<sup>-3</sup> at Preila during  
28 spring and winter, and between 0.4 and 0.6 μg m<sup>-3</sup> in Vilnius during summer and winter.  
29 These values do not evidence clear seasonal trends, but highlight a site-to-site variability  
30 which will be further discussed in the following. S-OOA instead was the largest contributor to  
31 total OM during summer with an average concentration of 1.2±0.8 μg m<sup>-3</sup>, always agreeing



1 between sites within a confidence interval of 95% (2 tails t-test). By contrast, during winter  
2 the S-OOA concentration dropped to an average value of  $0.3 \pm 0.2 \mu\text{g m}^{-3}$ , with 81% of the  
3 points not statistically different from  $0 \mu\text{g m}^{-3}$  within  $3\sigma$ . Finally, the LOA factor showed  
4 statistically significant contributions within  $3\sigma$  only during summer and late spring in Vilnius.  
5 Despite its considerable day-to-day variability this factor contributed  $1.0 \pm 0.8 \mu\text{g m}^{-3}_{\text{avg}}$  in  
6 Vilnius during summer.

7 The markers source apportionment instead attributed  $85\%_{\text{avg}}$  of the Other- $\text{OA}_{\text{marker}}$  mass to the  
8  $\text{SO}_4^{2-}$ -related SOA, while  $\text{NO}_3^-$ -related SOA, MSA-related SOA, and PBOA explained  
9 respectively  $9\%_{\text{avg}}$ ,  $5\%_{\text{avg}}$  and  $1\%_{\text{avg}}$  of the Other- $\text{OA}_{\text{marker}}$  mass (Fig. S13). The  $\text{SO}_4^{2-}$ -related  
10 SOA average concentration was  $2.4 \mu\text{g m}^{-3}$  during summer and  $1.7 \mu\text{g m}^{-3}$  during winter with  
11 no significant differences from station to station, suggesting a regional origin of the factor.  
12 The  $\text{NO}_3^-$ -related SOA concentration was  $0.4 \mu\text{g m}^{-3}_{\text{avg}}$  during winter, only  $0.03_{\text{avg}} \mu\text{g m}^{-3}$ ,  
13 during summer, corresponding to  $10\%_{\text{avg}}$  and  $1\%$  of the OA, respectively. Moreover, the  $\text{NO}_3^-$   
14 -related SOA during winter showed the highest average concentrations in Vilnius with  $0.5 \mu\text{g}$   
15  $\text{m}^{-3}$  and the lowest in Rūgšteliškis with  $0.3 \mu\text{g m}^{-3}_{\text{avg}}$ . The MSA-related SOA instead  
16 manifested the highest concentrations during summer with an average of  $0.12 \mu\text{g m}^{-3}_{\text{avg}}$ .  
17 Higher values were observed during summer at the rural coastal site of Preila where the  
18 average concentration was  $0.28 \mu\text{g m}^{-3}_{\text{avg}}$  corresponding to  $10\%_{\text{avg}}$  of the OM. Finally, the  
19 PBOA factor exhibited the largest seasonal concentrations during spring at the rural terrestrial  
20 site of Rūgšteliškis with an average of  $0.05 \mu\text{g m}^{-3}_{\text{avg}}$ , while the summer average  
21 concentration was  $0.02 \mu\text{g m}^{-3}$  consistent with the low PBOA estimates reported in Bozzetti et  
22 al. (2016) for the submicron fraction during summer.

23 Many previous studies reported a source apportionment of organic and inorganic markers  
24 concentrations (Viana et al., 2008 and references therein). In these studies  $\text{SO}_4^{2-}$ ,  $\text{NO}_3^-$ , and  
25  $\text{NH}_4^+$  were typically used as tracers for secondary aerosol factors commonly associated with  
26 regional background and long-range transport; here we compare the apportionment of the  
27 SOA factors obtained from the marker source apportionment and the OOA factors separated  
28 by the offline-AMS source apportionment. Moreover, contrasting the two source  
29 apportionments may provide insight into the origin of the OOA factors retrieved from the  
30 offline-AMS source apportionment, and into the origin of the SOA factors resolved by the  
31 offline-AMS source apportionment. To our knowledge an explicit comparison has not yet  
32 been reported in the literature.

Table 2: Pearson correlation coefficients between Other-OA components from offline-AMS and marker source apportionment.

		Other-OA <sub>marker</sub>			
		SO <sub>4</sub> <sup>2-</sup> -related SOA	MSA-related SOA	NO <sub>3</sub> <sup>-</sup> -related SOA	PBOA
Other-OA <sub>offline-AMS</sub>	LOA	0.33	0.16	-0.08	0.10
	B-OOA	0.70	0.22	0.21	0.47
	S-OOA	0.60	0.45	-0.47	0.05

Table 2 reports the correlations between the time series of the Other-OA<sub>marker</sub> factors and the Other-OA<sub>offline-AMS</sub> factors (Figs. 6 and S13). These correlations are mostly driven by seasonal trends as none of these sources shows clear spikes except for LOA during summer in Vilnius. Using the correlations coefficients we can identify the mostly related factors from the two source apportionments.

The SO<sub>4</sub><sup>2-</sup>-related SOA explained the largest fraction of the Other-OA<sub>marker</sub> mass (85%<sub>avg</sub>), and it was the only Other-OA<sub>marker</sub> factor always exceeding the individual concentrations of B-OOA and S-OOA, indicating that the variability explained by the SO<sub>4</sub><sup>2-</sup>-related SOA in the marker-source apportionment is explained by both OOA factors in the offline-AMS source apportionment. Moreover, the SO<sub>4</sub><sup>2-</sup>-related SOA seasonality seems consistent with the sum of S-OOA and B-OOA with higher concentrations in summer than in winter. This observation suggests that the OOA factors resolved by offline-AMS are mostly of secondary origin and the SO<sub>4</sub><sup>2-</sup>-related SOA, typically resolved by the markersource apportionment, explains the largest fraction of the OOA factors apportioned by offline-AMS which includes both biogenic SOA and aged background OA.

The NO<sub>3</sub><sup>-</sup>-related SOA and the PBOA factors were mostly related to B-OOA as they showed higher correlations with B-OOA than with S-OOA (Table 2). The B-OOA factor therefore may explain a small fraction of primary sources (PBOA), which however represents only 0.6%<sub>avg</sub> of the total OA. In detail, the NO<sub>3</sub>-related SOA correlation with B-OOA was poor ( $R = 0.21$ ), ~~but however~~ the correlation with LOA and S-OOA was negative (Table 2), suggesting that the mass attributed by the markers source apportionment to NO<sub>3</sub>-related SOA ~~by the markers source apportionment~~ was fully attributed to the B-OOA factor in the offline-AMS source apportionment. This is also confirmed by the fact that the sum of LOA and S-OOA concentrations during winter (when the NO<sub>3</sub>-related SOA substantially contributed) was

1 | ~~much smaller than can't explain~~ the  $\text{NO}_3^-$ -related SOA concentration, which therefore ~~was has~~  
2 | ~~to be~~ attributed to B-OOA.

3 | The MSA-related SOA showed the highest correlation with the S-OOA factor, as the two  
4 | sources exhibited the highest concentrations during summer, although the MSA-related SOA  
5 | preferentially contributed at the rural coastal site of Preila. While we already discussed the  
6 | probable secondary biogenic origin of S-OOA, the correlation with the MSA-related SOA  
7 | suggests that the S-OOA factor, especially at the rural coastal site of Preila, explains also a  
8 | large fraction of the marine biogenic SOA. The correlation between the two factors is  
9 | therefore not surprising as the precursor emissions (dimethyl sulfide, isoprene and terpenes)  
10 | are strongly related to the temperature leading to higher summer MSA-related SOA and S-  
11 | OOA concentrations. Assuming all the MSA-related SOA to be explained by the S-OOA  
12 | factor, we estimate a marine biogenic SOA contribution to S-OOA of 27%<sub>avg</sub> during summer  
13 | at Preila, while this contribution is lower at the other stations (12%<sub>avg</sub> in Rūgšteliškis during  
14 | summer, 7% in Vilnius during spring, no summer data for Vilnius Fig. S13). As already  
15 | mentioned, here we assume all the MSA-related SOA to be related to marine secondary  
16 | biogenic emissions, however other studies also report MSA from terrestrial biogenic  
17 | emissions (Jardine et al., 2015), moreover a certain fraction of the MSA-related SOA can also  
18 | be explained by the B-OOA factor. Overall these findings indicate that the terrestrial sources  
19 | dominate the S-OOA composition, nevertheless the marine SOA sources may represent a  
20 | non-negligible fraction, especially at the marine site.

21 | Another advantage obtained in coupling the two source apportionment results is the  
22 | possibility to study the robustness of the factor analyses by evaluating the consistency of the  
23 | two approaches as we already discussed for the primary OA and Other-OA fractions. Figure  
24 | S14~~b~~~~a~~ displays the ratio between PMF modelled WSOC and measured WSOC for the offline-  
25 | AMS case. A clear bias between Vilnius and the rural ~~sites, can sites, can~~ be observed, with a  
26 | WSOC overestimate of ~5% in Preila and Rūgšteliškis. While this overestimate is negligible  
27 | for the WSOC mass, it might have significant consequences on single factor concentrations.

28 | By contrast, for the markers source apportionment (Fig. S14a), OM residuals are more  
29 | homogeneous ~~for the case of markers PMF (Fig. S14b)~~. As we show in Fig. S6, these  
30 | residuals marginally affect the apportionment of combustion sources, as suggested by the well  
31 | comparing estimates of BBOA and TEOA using the two methods. Therefore, these residuals  
32 | are more likely affecting non-combustion sources (LOA, S-OOA and B-OOA). For the

1 common days, the S-OOA concentration is not statistically different at the different stations  
2 during summer (confidence interval of 95%), indicating that the residuals are more likely  
3 affecting LOA and B-OOA, which instead show site-to-site differences. Now, the PMF  
4 WSOC residuals appear at all seasons, also during periods without significant LOA  
5 contribution in Vilnius. Therefore, we conclude that B-OOA is the factor most significantly  
6 affected by the difference in the WSOC residuals. We could best assess the residual effects by  
7 comparing the B-OOA<sub>offline-AMS</sub> with that estimated using the other technique that seem to  
8 yield more homogeneous residuals: B-OOA<sub>marker</sub>. Here B-OOA<sub>marker</sub> is estimated as Other-  
9 OA<sub>markers</sub> - LOA - S-OOA. While B-OOA<sub>offline-AMS</sub> shows site-to-site differences, B-  
10 OOA<sub>markers</sub> did not show statistically different concentrations at all stations within a  
11 confidence interval of 95%. Based on these observations, we conclude that observed site-to-  
12 site differences in B-OOA concentrations are likely to be related to model uncertainties.

#### 13 **4.5 $f\text{CO}^+$ vs. $f\text{CO}_2^+$**

14 Figure 7 displays the water-soluble  $f\text{CO}^+$  vs.  $f\text{CO}_2^+$  scatter plot. A certain correlation ( $R=0.63$ )  
15 is seen, with  $f\text{CO}^+$  values being systematically lower than  $f\text{CO}_2^+$  ( $\text{CO}_2^+:\text{CO}^+$ : 1<sup>st</sup> quartile 1.50,  
16 median 1.75, 3<sup>rd</sup> quartile 2.01), whereas a 1:1  $\text{CO}_2^+:\text{CO}^+$  ratio is assumed in standard  
17 AMS/ACSM analyses (Aiken et al., 2008; Canagaratna et al., 2007). Comparing the measured  
18  $\text{CO}_2^+:\text{CO}^+$  values for the bulk WSOM and for pure gaseous  $\text{CO}_2$  might provide insight into  
19 the origin of the  $\text{CO}^+$  fragment in the AMS. The fragmentation of pure gaseous  $\text{CO}_2$  returned  
20 a  $\text{CO}_2^+:\text{CO}^+$  ratio of 8.21<sub>avg</sub> which is significantly higher than our findings for the water-  
21 soluble bulk OA (1.75<sub>med</sub>). Assuming thermal decarboxylation of organic acids as the only  
22 source of  $\text{CO}_2^+$  does not explain the observed  $\text{CO}_2^+:\text{CO}^+$  ratio of 1.75<sub>med</sub> and another large  
23 source of  $\text{CO}^+$  has to be assumed. Therefore, the carboxylic acid decarboxylation ~~into  $\text{CO}_2$~~   
24 can be considered as a minor source of  $\text{CO}^+$ .

25 Figure ~~7b-7a~~ and Fig. 8 show that not only does the water-soluble (WS)  $\text{CO}_2^+:\text{CO}^+$  ratio  
26 systematically differ from 1, but it also varies throughout the year with higher  $\text{CO}_2^+:\text{CO}^+$   
27 values associated with warmer temperatures (Fig. ~~7b7c~~). The lower  $\text{CO}_2^+:\text{CO}^+$  ratios in winter  
28 are primarily due to BB, as the WSBBOA factor profile showed the lowest  $\text{CO}_2^+:\text{CO}^+$  ratio  
29 (1.20<sub>avg</sub>) among all the apportioned WS factors (2.00<sub>avg</sub> for B-OOA, 2.70<sub>avg</sub> for S-OOA, and  
30 2.70<sub>avg</sub> for LOA). We observed a seasonal variation of the  $\text{CO}_2^+:\text{CO}^+$  ratio also for the water-  
31 soluble OOA (S-OOA + B-OOA) mass spectral fingerprint. The  $\text{CO}_2^+:\text{CO}^+$  ratio was slightly  
32 lower for B-OOA than for S-OOA (2.00<sub>avg</sub> for B-OOA, 2.70 for S-OOA). Nevertheless, given

1 the low S-OOA relative contribution during winter (Fig. 3), we note that the total OOA  
2 showed a slightly lower  $\text{CO}_2^+:\text{CO}^+$  ratio during winter than in summer (Fig. S15), indicating  
3 that the OOA mass spectral fingerprint evolves over the year, possibly because of different  
4 precursor concentrations, and different photochemical activity.

5 Fig. 7a shows that most of the measured  $\{f\text{CO}^+;f\text{CO}_2^+\}$  combinations lies within the triangle  
6 defined by the BBOA, S-OOA and B-OOA  $\{f\text{CO}^+;f\text{CO}_2^+\}$  combinations. The LOA factor  
7  $\{f\text{CO}^+;f\text{CO}_2^+\}$  combination lies within the triangle as well, but is anyways a minor source and  
8 thus unlikely to contribute to the  $\text{CO}_2^+:\text{CO}^+$  variability. We parameterized the  $\text{CO}^+$  variability  
9 as a function of the  $\text{CO}_2^+$ , and  $\text{C}_2\text{H}_4\text{O}_2^+$  fragment variabilities using a multi-parameter fit  
10 according to Eq. (8).  $\text{CO}_2^+$  and  $\text{C}_2\text{H}_4\text{O}_2^+$  were chosen as B-OOA and BBOA tracers,  
11 respectively, with B-OOA and BBOA being the factors that explained the largest fraction of  
12 the  $f\text{CO}^+$  variability (85% together).

$$13 \quad \text{CO}^+_i = a \cdot \text{CO}_2^+_i + b \cdot \text{C}_2\text{H}_4\text{O}_2^+_i \quad (8)$$

14 Although this parameterization is derived from the WSOM fraction  $\text{CO}_2^+$ ,  $\text{C}_2\text{H}_4\text{O}_2^+$ , and  $\text{CO}^+$   
15 originate from the fragmentation of oxygenated, i.e. mostly water-soluble compounds.  
16 Accordingly, this parameterization might also well represent the total bulk OA (as the offline-  
17 AMS recoveries of these oxygenated fragments are relatively similar:  $R_{\text{CO}_2^+}=0.74$ ,  
18  $R_{\text{C}_2\text{H}_4\text{O}_2^+}=0.61$ , Daellenbach et al., 2016). Note that this parameterization may represent very  
19 well the variation of  $\text{CO}^+$  in an environment impacted by BBOA and OOA, but should be  
20 used with caution when other sources (such as COA) may contribute to  $\text{CO}^+$ ,  $\text{CO}_2^+$  and  
21  $\text{C}_2\text{H}_4\text{O}_2^+$ . In order to check the applicability of this parameterization to a PMF output, we  
22 recommend monitoring the  $\text{CO}_2^+$  and  $\text{C}_2\text{H}_4\text{O}_2^+$  variability explained by the OOA and BBOA  
23 factors. In case a large part of the  $\text{CO}_2^+$  and  $\text{C}_2\text{H}_4\text{O}_2^+$  variability is explained by OOA and  
24 BBOA, the parameterization should return accurate  $\text{CO}^+$  values. The coefficients  $a$  and  $b$  of  
25 Eq. (8) were determined as 0.52 and 1.39 respectively, while the average fit residuals were  
26 estimated to be equal to 10% (Fig. S16). In contrast, parameterizing  $\text{CO}^+$  as proportional to  
27  $\text{CO}_2^+$  only (as done in the standard AMS analysis scheme with coefficients updated to the  
28 linear fit between  $\text{CO}^+$  and  $\text{CO}_2^+$  (1.75)) yielded 20%<sub>avg</sub> residuals, indicating that such a  
29 univariate function describes the  $\text{CO}^+$  variation less precisely.

30 An alternative parameterization is presented in the SI in which the contribution of moderately  
31 oxygenated species (such as S-OOA) to  $\text{CO}^+$  was also considered by using  $\text{C}_2\text{H}_3\text{O}^+$  as an  
32 independent variable. We show that the dependence of  $\text{CO}^+$  on  $\text{C}_2\text{H}_3\text{O}^+$  is statistically

1 significant (Fig. 7b) as also suggested by the PMF results (S-OOA contributes 12% to the  
2  $\text{CO}^+$  variability). However, the parameter relating  $\text{CO}^+$  to  $\text{C}_2\text{H}_3\text{O}^+$  is negative, because the  
3  $\text{CO}^+:\text{CO}_2^+$  and  $\text{CO}^+:\text{C}_2\text{H}_4\text{O}_2^+$  ratios are lower in moderately oxygenated species compared to  
4 species present in BBOA and B-OOA. While this parameterization captures the variability of  
5  $\text{CO}^+$  across the seasons better compared to a 2-parameter fit for the present dataset, it may be  
6 more prone to biases in other environments due to the known contributions of other factors to  
7  $\text{C}_2\text{H}_3\text{O}^+$ . For example, cooking-influenced organic aerosol (COA) often accounts for a  
8 significant fraction of  $\text{C}_2\text{H}_3\text{O}^+$ . For ambient datasets we propose the use of  $\text{CO}_2^+$  and  $\text{C}_2\text{H}_4\text{O}_2^+$   
9 only, which may capture less variation but is also less prone to biases. Although our results  
10 suggest that the available  $\text{CO}^+$  and O:C estimates (Aiken et al., 2008; Canagaratna et al.,  
11 2015) may not well capture the  $\text{CO}^+$  variability, our  $\text{CO}^+$  parameterization should not be  
12 applied to calculate the O:C ratios or recalculate the OA mass from AMS datasets, as those  
13 are calibrated assuming a standard fragmentation table (i.e.  $\text{CO}_2^+ = \text{CO}^+$ ).

14 In a recent work, Canagaratna et al. (2015) reported the Ar nebulization of water soluble  
15 single compounds to study the HR-AMS mass spectral fingerprints in order to improve the  
16 calculation of O:C and OM:OC ratios. Following the same procedure, we nebulized a subset  
17 of the same standard compounds including malic acid, azalaic acid, citric acid, tartaric acid,  
18 cis-pinonic acid, and D(+)-mannose. We obtained comparable  $\text{CO}_2^+:\text{CO}^+$  ratios (within 10%)  
19 to those of Canagaratna et al. (2015) for all the analyzed compounds, highlighting the  
20 comparability of results across different instruments. With the exception of some  
21 multifunctional compounds (citric acid, malic acid tartaric acid, ketobutyric acid, hydroxyl  
22 methylglutaric acid, pyruvic acid, oxaloacetic acid, tartaric acid, oxalic acid and malonic  
23 acid), the water-soluble single compounds analyzed by Canagaratna et al. (2015) mostly  
24 showed  $\text{CO}_2^+:\text{CO}^+$  ratios  $<1$ , systematically lower than the  $\text{CO}_2^+:\text{CO}^+$  ratios measured for the  
25 bulk WSOM in Lithuania (1<sup>st</sup> quartile 1.50, median 1.75, 3<sup>rd</sup> quartile 2.01), which represents a  
26 large fraction of the total OM (bulk EE: median = 0.59, 1<sup>st</sup> quartile = 0.51, 3<sup>rd</sup> quartile = 0.72).  
27 Considering the relatively high bulk EE, and considering that the  $\text{CO}^+$  and  $\text{CO}_2^+$   
28 fragmentation precursors tend to be more water soluble than the bulk OA, the aforementioned  
29 compounds could be representative of a large part of the  $\text{CO}^+$  and  $\text{CO}_2^+$  fragmentation  
30 precursors. This indicates that the selection of appropriate reference compounds for ambient  
31 OA is non-trivial, and the investigation of multifunctional compounds is of high importance.

32

## 1 5 Conclusions

2 PM<sub>1</sub> filter samples were collected over an entire year (November 2013 to October 2014) at  
3 three different stations in Lithuania. Filters were analyzed by water extraction followed by  
4 nebulization of the liquid extracts and subsequent measurement of the generated aerosol with  
5 an HR-ToF-AMS (Daellenbach et al., 2016). For the first time, the nebulization step was  
6 conducted in Ar, enabling direct measurement of the CO<sup>+</sup> ion, which is typically masked by  
7 N<sub>2</sub><sup>+</sup> in ambient air and assumed to be equal to CO<sub>2</sub><sup>+</sup> (Aiken et al., 2008). CO<sub>2</sub><sup>+</sup>:CO<sup>+</sup> values >1  
8 were systematically observed, with a mean ratio of 1.7±0.3. This is likely an upper limit for  
9 ambient aerosol, as only the water-soluble OM fraction is measured by the offline-AMS  
10 technique. CO<sup>+</sup> concentrations were parameterized as a function of CO<sub>2</sub><sup>+</sup>, and C<sub>2</sub>H<sub>4</sub>O<sub>2</sub><sup>+</sup>, and  
11 this two-variable parameterization showed a superior performance to a parameterization based  
12 on CO<sub>2</sub><sup>+</sup> alone, because CO<sup>+</sup> and CO<sub>2</sub><sup>+</sup> show different seasonal trends.

13 PMF analysis was conducted on both the offline-AMS data described above and a set of  
14 molecular markers together with total OM. Biomass burning was found to be the largest OM  
15 source in winter, while secondary OA was largest in summer. However, higher concentrations  
16 of primary anthropogenic sources (biomass burning and hopanes here used as traffic markers)  
17 were found at the urban background station of Vilnius. The offline-AMS and marker-based  
18 analyses also identified local emissions and primary biological particles, respectively, as  
19 factors with low overall but episodically important contributions to PM. Both methods  
20 showed traffic exhaust emissions to be only minor contributors to the total OM; which is not  
21 surprising given the distance of the three sampling stations from busy roads.

22 The two PMF analyses apportioned SOA to sources in different ways. The offline-AMS data  
23 yielded factors related to regional background (B-OOA) and temperature-driven (likely  
24 biogenic-influenced) emissions (S-OOA), while the marker-PMF yielded factors related to  
25 nitrate, sulfate, and MSA. For the offline-AMS PMF, S-OOA was the dominant factor in  
26 summer and showed a positive exponential correlation with the average daily temperature,  
27 similar to the behavior observed by Leaitch et al. (2011) in a Canadian boreal forest.  
28 Combining the two source apportionment techniques suggests that the S-OOA factor includes  
29 contributions from both terrestrial and marine secondary biogenic sources, while only small  
30 PBOA contributions to submicron OOA factors are possible. The analysis highlights the  
31 importance of regional meteorological conditions on air pollution in the southeastern Baltic  
32 region, as evidenced by simultaneous high BBOA levels at the three stations during three

1 different episodes in winter and by statistically similar S-OOA concentrations across the three  
2 stations during summer.

3

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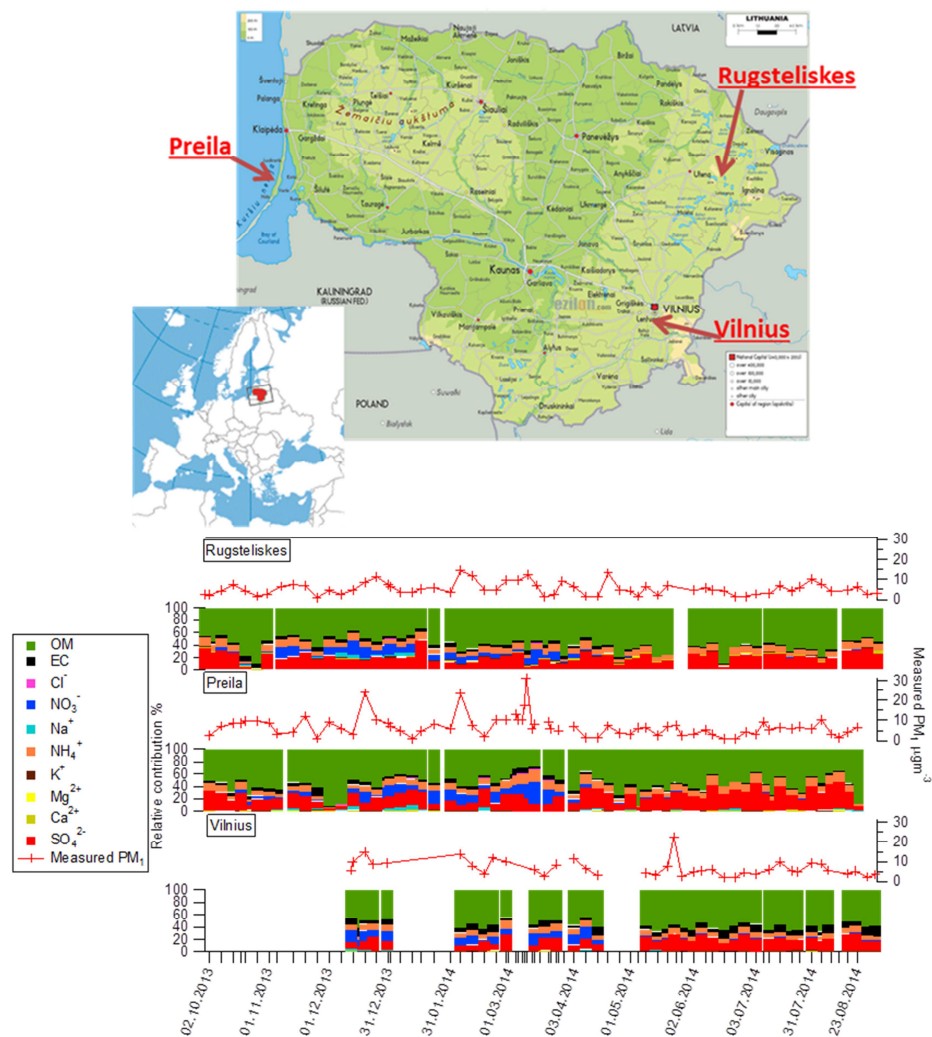
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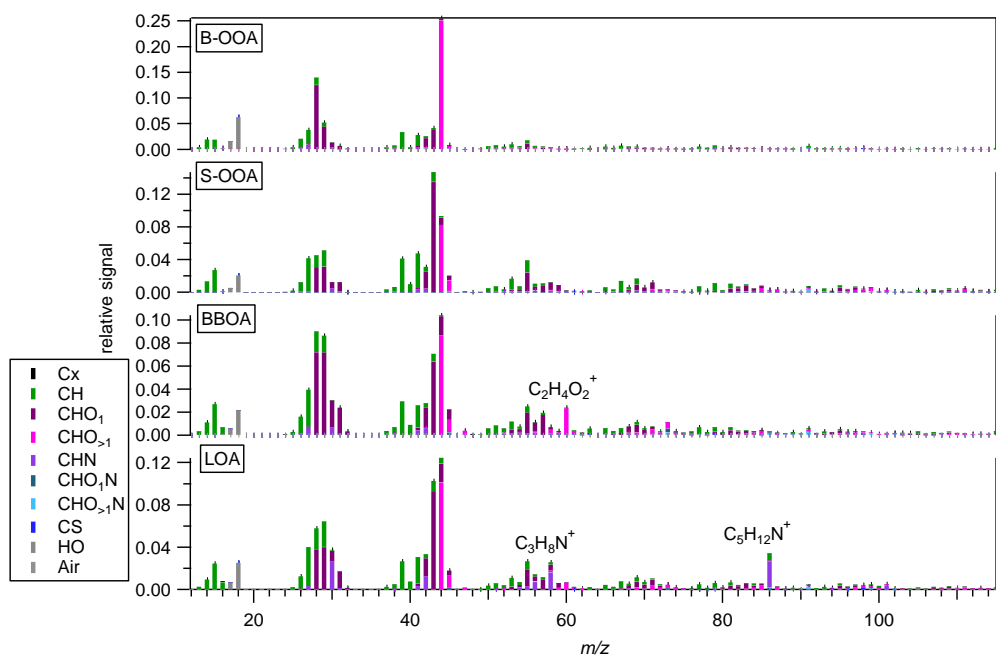
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6 **Figures main text**

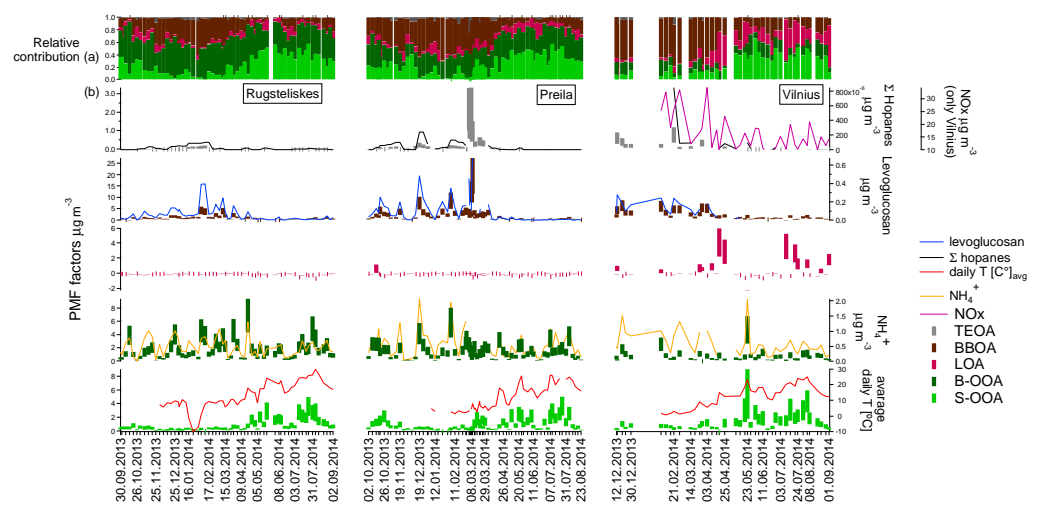


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8 Figure 1. Sampling locations, and measured PM<sub>1</sub> composition.

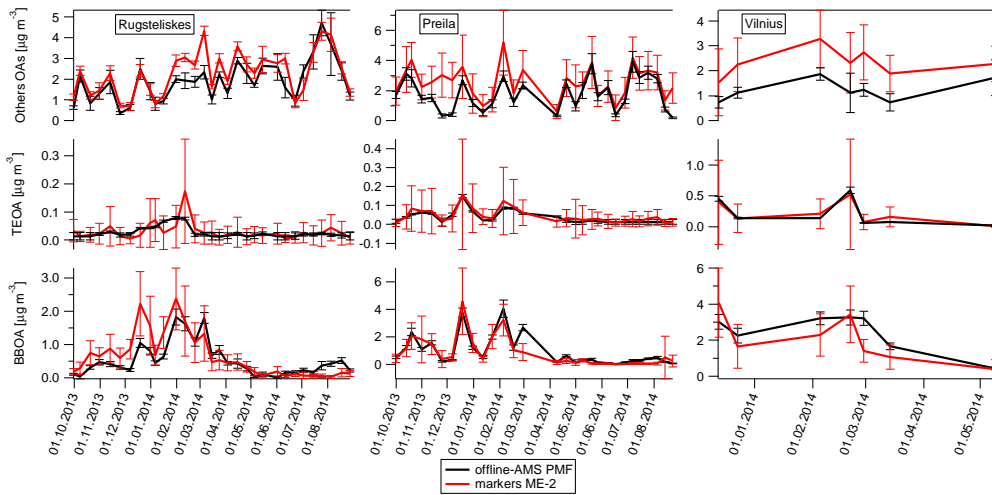


1  
 2 Figure 2. Offline-AMS PMF factor profiles: background oxygenated OA (B-OOA), summer  
 3 oxygenated OA (S-OOA), biomass burning OA (BBOA), local OA (LOA).

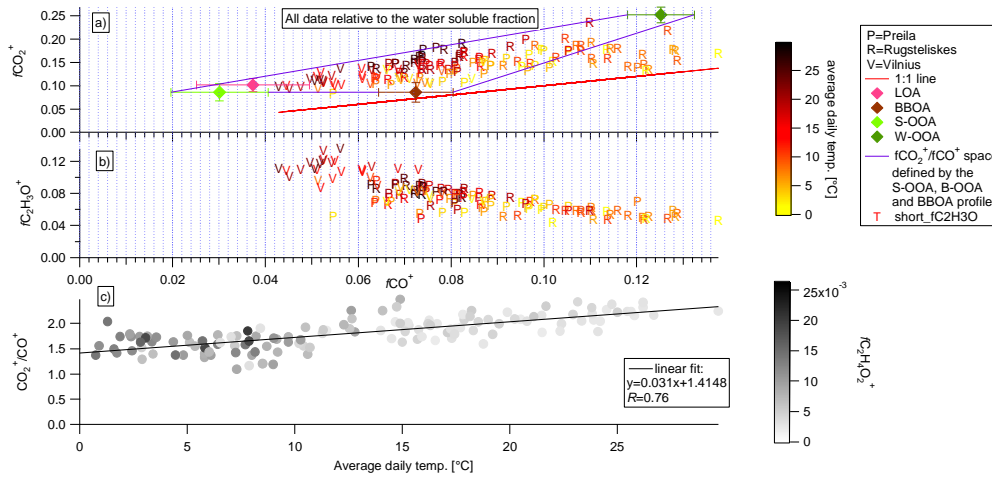


5  
 6 Figure 3. a) Temporal evolutions of relative contributions to the OA factors; b) OA sources  
 7 and corresponding tracers: concentrations and uncertainties (shaded areas).

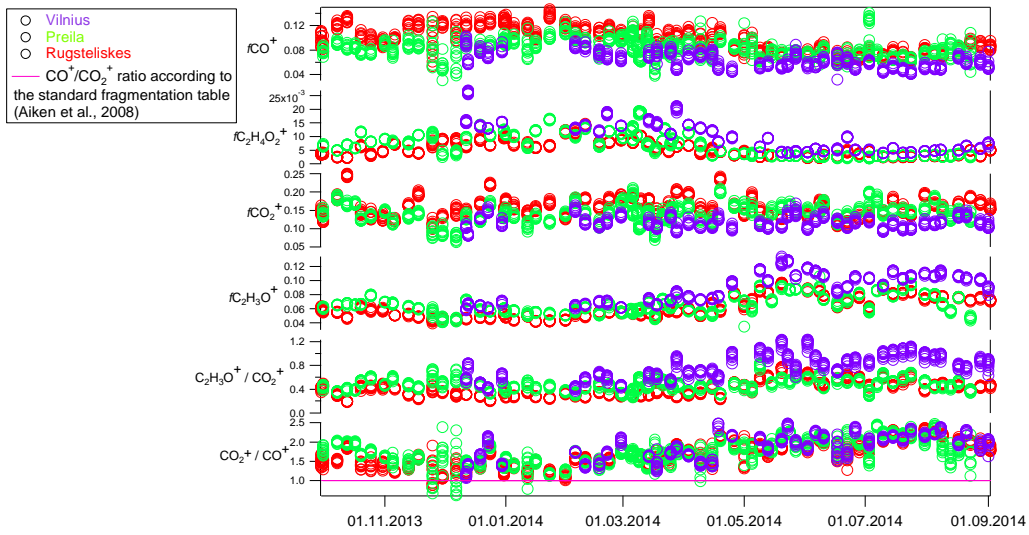




1  
2 Figure 6. Marker-PMF and offline-AMS OM source apportionment comparison.



3  
4 Figure 7. a) Water-soluble  $f\text{CO}_2^+$  vs  $f\text{CO}^+$  scatter plot. Color code denotes the average daily  
5 temperature [°C], diamonds indicate the  $f\text{CO}_2^+ : f\text{CO}^+$  ratio for different PMF factor profiles.  
6 The 1:1 line is displayed in red. Few points from Rūgšteliškis lie outside the triangle,  
7 suggesting they are not well explained by our PMF model. However, Fig. S5 displays flat  
8 residuals for Rūgšteliškis, indicating an overall good WSOM explained variability by the  
9 model. b) Water-soluble  $f\text{C}_2\text{H}_3\text{O}^+$  vs  $f\text{CO}^+$  scatter plot. Color code denotes the average daily  
10 temperature [°C] c) Scatter plot of the water-soluble  $\text{CO}_2^+ : \text{CO}^+$  ratio vs. average daily  
11 temperature. Grey code denotes  $f\text{C}_2\text{H}_4\text{O}_2^+$ .  
12



1

2 Figure 8. Time-dependent fractional contributions ( $f$ ) of typical AMS tracers.