

Manuscript

“Contribution of Surface Solar Radiation and Precipitation to Spatiotemporal Patterns of Surface and Air Temperature Warming in China from 1960 to 2003” by Jizeng Du et al.

Response to Reviewer # 1

1. Major

Comment: While the regional trend patterns of T_s , T_a , R_s and P and their relations are interesting, the authors did not give fully explanation of the possible reasons causing the formation of these patterns.

Reply: Following the reviewers comments, we added two paragraphs to explain why the changes of R_s and P caused the pattern of T_s and T_a (Lines 120-144):

“It is well known that the diurnal cycles in T_a and T_s are primarily determined by the surface energy budget. After sunrise, the surface absorbs solar radiation, and the surface net radiation becomes positive and heats the surface first. As a result, the air above the surface becomes unstable. Surface net radiation can be partitioned into three parts: ground heat flux, sensible heat flux, and latent heat flux. Ground heat flux heats the surface and stores energy during the daytime, and this energy may be re-emitted at night. Sensible heat flux directly heats the air above the surface. Latent heat flux is the energy employed to vaporize water during the surface water evaporation and vegetation transpiration processes. How surface net radiation partitions into ground heat flux, sensible heat flux, and latent heat flux is determined by both surface and atmospheric conditions (Wang et al., 2010a, b; Wang and Dickinson, 2012), i.e., surface wetness. Daytime surface net radiation is primarily determined by R_s (Wang and Liang, 2008) and precipitation or surface wetness control partition of surface net radiation into latent and sensible fluxes (Wang and Liang, 2008). Therefore, it is expected that changes in R_s and P play a key role in the variability of T_s and T_a (Wild, 2012; Manara et al., 2015; Hartmann et al., 1986).”

However, quantitative assessments of the impact of R_s on T_s and T_a are still lack due

the shortness of high quality of long-term estimates of R_s . In this study, we used sunshine duration derived R_s (Wang, 2014; Wang et al., 2015) to quantitate the impact of R_s on the spatial pattern of T_a and T_s . To our knowledge, this study presents the first analysis of the relationship between R_s (and P) and T_a (and T_s) based on their spatiotemporal patterns and we further quantified the effect of variations of R_s (and P) on T_a (and T_s) in China for the period 1960-2003.”

Comment: Unless I downloaded the wrong version, I found there are some mislabelling of Figures in the text. e.g., I can’t find Fig S9g-i (line 383).

Reply: The reviewer has an incorrect version of the supplementary material. The supplementary material has been substantially revised during last revision. We double checked and make sure it is correct in this upload.

Comment: I noticed the authors were asked to justify the calculation of adjusted T , but I can’t see the justification in the current version. These parameters (T_a , T_s , R_s and P) interacts with each other, it is not easy to distinguish the casual relations. I am not sure I understand the actual implication of the adjusted T in section 3.3.

Reply: We agree with the reviewer that the regression made in this paper does not provide any information on the casual relation. Following the reviewer’s suggestion, we added two paragraphs to explain why the changes of R_s and P impacted the pattern of T_s and T_a (Lines 120-142). The reason is well known, however, the quantitative assessment on this aspect is still lacking because of the shortness of data. Due to continuous effort, we reconstructed long-term time series of surface solar radiation based on the sunshine duration observations. Furthermore, this study first found that the T_s is most suitable parameter to study the issue, which has been ignored by the other researchers. We revised the equations and justifications based on the reviewers’ suggestion during last round of review.

2. Minor

Comment: Line 54: Please make Figure or Fig. consistent. E g Figure 1, but Fig. S1.

Reply: We have revised all Figure X to Fig. X in new manuscript.

Comment: Line 22: “stronger in northwest China”. Are you sure it is not the northeast China?

Reply: Thanks for your comments. We agree with you that the warming trend is not only stronger in “northwest China” but also in “northeast China”. So we have corrected this sentence in Line 20-21: “the pattern was stronger in northwest and northeast China and weaker or negative in South China and the North China Plain.”

Comment: Line 22: “and weaker in South China”. Is this before adjust or after adjust? Before adjust as shown in Fig. 4, the trend is negative in South China, not “weaker”.

Reply: This is warming trend before adjusting. We agree with your comments and have corrected this sentence in Line 20-21: “the pattern was stronger in northwest and northeast China and weaker or negative in South China and the North China Plain”.

Comment: Line 23: Rs is redefined as SSR in Fig. S4. Make it consistent.

Reply: The reviewer has the incorrect version of the supplementary material. It is correct in the new version.

Comment: Fig.S6: the caption is copied from the old version; the referenced Figs are not there anymore.

Reply: The reviewer has the incorrect version of the supplementary material. It is correct in the new version.

Comment: Line 28: “were much higher than those in other regions”. It is hard to get this conclusion, since you don’t have enough coverage over other areas.

Reply: Thanks for your comments. We revised this sentence to (Lines 25-27):

“More importantly, the decreasing rates in South China and the North China Plain were stronger than those in other parts of China.”

Comment: Line 32: “North China Plain and the Loess Plateau”. It will be easier to see if you mark these regions on the corresponding map.

Reply: Thanks for your valuable comments. We have marked these regions on the Fig. 9a and introduced them in Line 443-445: “To clearly illustrate these changes, we selected two regions in China for further investigation: R1 primarily included the North China Plain and R2 primarily included the Loess Plateau (see Fig. 9a).”

Comment: Line 92: “sever air pollution “severe air pollution

Reply: We have corrected it in Line 89.

Comment: Line 155: “Our study is the first to use T_s observation as a parameter for identifying regional climate change.” I doubt it.

Reply: We have deleted this sentence in new manuscript.

Comment: Line 196: “reanalyzes”, reanalyses?

Reply: We have revised it in new version.

Comment: Line 241: In Eq (1). Are R_s and P anomalies? Please make it clear.

Reply: Following your comment, we have added a statement that in Line 256-258: “ R_s and P represents the monthly anomalies of surface solar radiation and precipitation respectively”.

Comment: Line 245: “Fig. S3”. Fig. S4?

Reply: The reviewer has the incorrect version of the supplementary material. It is correct for the new version.

Comment: Fig. S4 & 5: Please explain the large difference of the sensitivity in Fig S5a and g.

Reply: We have explained the difference between the sensitivities of T_s and T_a to R_s in Line 339-343: “During the day, T_s is directly determined by the land surface energy

balance, i.e., the incoming energy (including R_s) and atmospheric longwave radiation (Wang and Dickinson, 2013a), and it is partitioned into latent and sensible heat fluxes (Zhou and Wang, 2016). Although T_a is dependent on the land-atmosphere sensible heat flux, it is also affected by local and/or large-scale circulation.”

Comment: Line 342: Fig S7? Please check it is the right Fig.

Reply: The reviewer have the incorrect version of the supplementary material. It is correct for the new version.

Comment: Line 376: Fig. S5? Please check.

Reply: The reviewer has the incorrect version of the supplementary material. It is correct for the new version.

Comment: Line 381-382: Fig S9 doesn't have g-i. Please check.

Reply: The reviewer has the incorrect version of the supplementary material. It is correct for the new version.

Response to Reviewer # 2

1. General Comments

Comment: There is one conceptual issue: the use in L26 and L451 of “caused” for results from the regression analysis between T and R_s . It is not considered acceptable from a regression analysis of A and B to say that A caused B, even though in a simplified surface-boundary layer model, R_s is the primary physical driver of T_{max} . In the real world, other factors (L468) that you have not measured may be involved, and some others such as the outgoing LW are themselves coupled to T_{max} . Your point is that if you use the slope of the partial regression of T_{max} on R_s to remove the spatial variation of R_s , the T_{max} trends became spatially less heterogeneous. Find other wording that allows you to make your point.

Reply: We agree with the reviewer that the regression made in this paper does not

provide any information on the casual relation. Following the reviewer's suggestion, we added two paragraphs to explain why the changes of R_s and P caused the pattern of T_s and T_a (Lines 120-142). The reason is well known, however, the quantitative assessment on this aspect is still lacking because of the shortness of data. Due to the continuous effort, we reconstructed long-term time series of surface solar radiation based on the sunshine duration observations. Furthermore, this study first found that the T_s is most suitable parameter to study the issues, which has been ignored by the other researchers. We revised the equations and justifications based on the reviewers' suggestion during last round of review.

2. Minor

Comment: L92 severe air pollution

Reply: We have corrected it as your comments in new version.

Comment: The notation R_s/P for R_s and P is not conventional as $/$ is normally used for 'divided by'. Also T_{s-max}/T_{a-max}

Reply: We have replaced all " R_s/P " to " R_s and P ", same to " T_{s-max}/T_{a-max} ".

Comment: L278 increased by 0.356 ± 0.0057 °C 0.356 ± 0.057 °C?

Reply: We have revised it in new version.

Comment: L307 This dependence has been successfully attributed to amplified dynamics. This comment needs clarification

Reply: Thanks for your comments. We have changed the expression and make it clearer. "Related studies suggested that this dependence was strongly associated with the mode variability in large-scale circulation, such as a negative trend in the North Atlantic Oscillation during this period (Wallace et al., 2012; Ding et al., 2014)." (Line 324-327).

Comment: L789 which did not include the effect of the R_s variations. Are you sure you

mean this? Or ‘after removing the effect of the R_s variations’ (and P ?)

Reply: Thanks for your comments. We have corrected those sentences: “(c) Annual, warm, and cold seasonal scale trends calculated based on the data before adjusting the effect of R_s and P . (d) Annual, warm, and cold seasonal scale trends calculated based on the data after adjusting the effect of R_s and P .” (Line 815-817).

Comment: L451 ‘primary cause of’ (and L26 of abstract). Find other wording since regression shows correlation and association, but not causes. So it is incorrect to say ‘primary cause’. See also L468-469 where other factors, not measured are mentioned. Reword L326 as well.

Reply: Please see also our response to your general comments. We have revised the wording as your suggestion. “During the study period, R_s decreased by $-1.50 \pm 0.42 \text{ W m}^{-2} \text{ 10yr}^{-1}$ in China, which contributed the trends of T_{s-max} and T_{a-max} decreased by about 0.139 and 0.053 $^{\circ}\text{C 10yr}^{-1}$, respectively.” (Line 23-25). “we found that R_s plays a distinctly important role in the spatial warming patterns of T_{s-max} and T_{a-max} ”. (Line 472- 473).

**Contributions of Surface Solar Radiation and Precipitation to the Spatiotemporal
Patterns of Surface and Air-Temperature Warming in China from 1960 to 2003**

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15 **Abstract**

16 Although global warming has been attributed to increases in atmospheric greenhouses
17 gases, the mechanisms underlying spatiotemporal patterns of warming trends remain
18 under debate. Herein, we ~~analysed~~ analyzed surface and air warming observations
19 recorded at 1,977 stations in China from 1960 to 2003. Our results showed a significant
20 spatial pattern for the warming of the daily maximum surface (T_{s-max}) and air (T_{a-max})
21 temperatures, and the pattern was stronger in northwest and northeast China and weaker
22 or negative in South China and the North China Plain. These warming spatial patterns
23 were attributed to surface shortwave solar radiation (R_s) and precipitation (P), which
24 ~~represent the key parameters~~ play a key role of in the surface energy budget. During the
25 study period, R_s decreased by $-1.50 \pm 0.42 \text{ W m}^{-2} \text{ 10yr}^{-1}$ in China, which ~~contributed~~
26 ~~reduced~~ the trends of T_{s-max} and T_{a-max} ~~decreased~~ by about 0.139 and 0.053 °C 10yr⁻¹,
27 respectively. More importantly, the decreasing rates in South China and the North China
28 Plain were ~~stronger~~ much higher than those in other ~~study regions~~ parts of China. The
29 spatial contrasts in the trends of T_{s-max} and T_{a-max} in China were significantly reduced
30 after adjusting for the effect of R_s and P . For example, after adjusting for the effect of
31 R_s and P , the difference in the T_{s-max} and T_{a-max} values between the North China Plain
32 and the Loess Plateau was reduced by 97.8% and 68.3%, respectively, the seasonal
33 contrast in T_{s-max} and T_{a-max} decreased by 45.0% and 17.2%, respectively, and the daily

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34 contrast in the warming rates of the surface and air temperature decreased by 33.0%
35 and 29.1%, respectively. This study showed that the land energy budget plays an
36 essential role in the identification of regional warming patterns.

37

38 **1. Introduction**

39 Increases in observational data and rapid developments in simulation capacity of
40 climate models have provided evidence for the phenomenon of global warming
41 (Hartmann et al., 2013), and the increases in anthropogenic greenhouse gases and other
42 anthropogenic effects are considered as the primary causes. However, significant spatial
43 and temporal heterogeneities in climate warming have been observed. For example,
44 faster warming rates occur in semiarid regions and a “warming hole” has been identified
45 in the central United States (Boyles and Raman, 2003; Huang et al., 2012). These
46 spatiotemporal heterogeneities represent a major barrier to the reliable detection and
47 attribution of global warming (Tebaldi et al., 2005; Mahlstein and Knutti, 2010).
48 Furthermore, uncertainties in model simulations generally increase from global to
49 regional scales because of uncertainty in regional climatic responses to global change
50 (Hingray et al., 2007; Mariotti et al., 2011). Therefore, investigations of the spatial and
51 temporal patterns of regional climate changes and regional climatic response
52 mechanisms to global change are crucial for increasing the accuracy of models designed
53 to detect and explain the causes of global climate change and predictions of future
54 regional climate change.

The spatial heterogeneity of climate warming can be attributed to local climate factors and anthropogenic factors (Karl et al., 1991). For the local climate factors, determining factors such as cloud cover and precipitation (P) can significantly influence the speed of regional warming (Hegerl and Zwiers, 2007; Lauritsen and Rogers, 2012). Spatial heterogeneities in climate-factor trends have an important influence on various changes in the land-surface energy balance. Studies have demonstrated that an increase in cloud cover can diminishes the surface solar radiation (R_s) and therefore reduces the daytime temperature (Dai et al., 1997; Zhou et al., 2010; Taylor et al., 2011), although it has the potential to increase night-time temperature by intercepting outgoing longwave radiation (Campbell and VonderHaar, 1997; Shen et al., 2014).

Precipitation (P) can alter the proportion of surface absorbed energy partitioned into sensible heat flux and latent heat flux; therefore it has an inevitable effect on both land-surface and near-surface air temperatures (Wang and Dickinson, 2012; Wang and Zhou, 2015). Additionally, P has a significant effect on soil thermal inertia and the response of surface vegetation, which results in an important feedback for regional and global warming (Seneviratne et al., 2010; Wang and Dickinson, 2012; Ait-Mesbah et al., 2015; Shen et al., 2015).

In addition to local climate factors, regional climate systems are significantly

affected by the anthropogenic emissions of aerosols. Studies have indicated that improvements in air quality in recent decades over North America and Europe have led to brightening effect (Vautard et al., 2009; Wild, 2012), whereas East Asia and India have led to declines in R_s (Xia, 2010; Menon et al., 2002; Wang et al., 2012; Wang et al., 2015). Consequently, variations in R_s may have an effect on both local and global climate change (Wild et al., 2007; Wang and Dickinson, 2013a).

Changes in land cover can also alter the energy exchange between the land surface and the atmosphere, and such changes have the potential to affect regional climates (Bounoua et al., 1999; Zhou et al., 2004; Falge et al., 2005). Previous studies have suggested that urbanization and other land-use changes contribute to promoting the warming effect caused by greenhouse gases (Kalnay and Cai, 2003; Lim et al., 2005; Chen et al., 2015). Overall, the effects of these factors on climate change may be very important at the regional scale and could lead to marked spatial differences in regional climate change; however, they are usually omitted from the detection and attribution of climate change at the global scale (Károly and Stott, 2006).

China is a vast territory that has an abundance of climatic zones stretching from tropical to cold temperate, and a special alpine climate is observed over the Tibet Plateau. Additionally, the dramatic economic development and explosive population

growth in China in recent decades have caused significant changes in land cover and severe air pollution, including frequent haze events (Yin et al., 2017; Cheng et al., 2014; Wang et al., 2016). The climatic diversity and intensive human activity in this region will likely lead to a unique response to global warming with obvious spatial differences in climate change.

Karl et al. (1991) analyzed the observational records for the period 1951–1989 and found that warming trends in China were faster than those of the United States but slower than those of the former Soviet Union. Several studies have revealed that the warming rate in Northwest China was approximately $0.33\text{--}0.39\text{ }^{\circ}\text{C }10\text{yr}^{-1}$ during the second half of the last century (Zhang et al., 2010; Li et al., 2012), which was significantly higher than the average warming rate over China of $0.25\text{ }^{\circ}\text{C }10\text{yr}^{-1}$ (Ren et al., 2005) or the average global rate of $0.13\text{ }^{\circ}\text{C }10\text{yr}^{-1}$ (Hegerl and Zwiers, 2007). The air temperature (T_a) over the Tibet Plateau has increased by $0.44\text{ }^{\circ}\text{C }10\text{yr}^{-1}$ over the last 30 years (Duan and Xiao, 2015), and this rate is considerably faster than the overall warming rate in the Northern Hemisphere ($0.23\text{ }^{\circ}\text{C }10\text{yr}^{-1}$) and worldwide ($0.16\text{ }^{\circ}\text{C }10\text{yr}^{-1}$) (Hartmann et al., 2013). To provide insights on global warming and improve the accuracy of future climate change predictions, understanding the characteristics and mechanisms of regional climate change is critical.

T_a is a common metric for determining climate change on the global or regional scales. The land surface temperature (T_s) is also important in climate change research because of its direct relationship with the land surface energy budget. Previously, T_s values used in regional climate research are primarily derived from satellite retrievals or reanalysis datasets (Weng et al., 2004; Peng et al., 2014), which both have satisfactory global coverage but questionable accuracy and integrity. Furthermore, satellite-derived T_s values are only available under clear sky conditions, thus limiting their applicability in climate change studies.

In China, both T_s and T_a are measured as conventional meteorological observation parameters by nearly all weather stations. An analysis of the spatiotemporal patterns of these parameters identified a close relationship between T_s and T_a , which indicates that T_s and T_a present equivalent accuracy when used to determine the characteristics of climate change. More importantly, T_s is more sensitive than T_a to the local land surface energy budget.

It is well known that the diurnal cycles in T_a and T_s are primarily determined by the surface energy budget. After sunrise, the surface absorbs solar radiation, and the surface net radiation becomes positive and heats the surface first. As a result, the air above the surface becomes unstable. Surface net radiation can be partitioned into three

parts: ground heat flux, sensible heat flux, and latent heat flux. Ground heat flux heats the surface and stores energy during the daytime, and this energy may be re-emitted at night. Sensible heat flux directly heats the air above the surface. Latent heat flux is the energy employed to vaporize water during the surface water evaporation and vegetation transpiration processes. How surface net radiation partitions into ground heat flux, sensible heat flux, and latent heat flux is determined by both surface and atmospheric conditions (Wang et al., 2010a; Wang et al., 2010b; Wang and Dickinson, 2012), i.e., surface wetness. Daytime surface net radiation is primarily determined by R_s (Wang and Liang, 2008) and precipitation or surface wetness control partition of surface net radiation into latent and sensible fluxes (Wang and Liang, 2008). Therefore, it is expected that changes in R_s and P play a key role in the variability of T_s and are key factors controlling the land surface energy budget; therefore, changes in these two factors most likely cause regional differences in the warming rate of T_a (Hartmann et al., 1986; Wild, 2012; Manara et al., 2015).

However, quantitative assessments of the impact of R_s on T_s and T_a are still lack due the shortness of high quality of long-term estimates of R_s . In this study, we used sunshine duration derived R_s (Wang, 2014; Wang et al., 2015) to quantitate the impact of R_s on the spatial pattern of T_a and T_s . To our knowledge, this study presents the first analysis of the relationship between R_s (and P) and T_a (and T_s) based on their

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spatiotemporal patterns and we further quantified the effect of variations of R_s (and P) on T_a (and T_s) in China for the period 1960–2003.

This article is organized as follows. Section 2 introduces the data and methods used in the study. Section 3 describes the spatial and temporal patterns of climate warming over China, analyses the effect of the variation in R_s and P on T_a and T_s , and examines the spatial and temporal patterns of the warming trend of T_a and T_s after adjusting for the effects of R_s and P , which eliminated the effects of R_s and P on warming and highlighted the effects of large-scale warming caused by elevated concentrations of atmospheric greenhouse gases. Moreover, the spatial contrast in the warming trends of T_a and T_s in China was substantially reduced after adjusting for the effect of R_s and P , and this result is consistent with the expectations under global warming. Finally, Section 4 presents a summary and discussion.

2. Data and methods

2.1. Data

The meteorological observational data used in this study are included recently released daily meteorological datasets, such as the China National Stations' Fundamental Elements Datasets V3.0 (CNSFED V3.0), and they were downloaded

from China's National Meteorological Information Centre (<http://data.cma.gov.cn/data>) (Cao et al., 2016). These datasets included observations of T_s , T_a , barometric pressure, relative humidity, and sunshine duration. All of the observational records of the climate variables were subjected to quality control measures, and the data acquisition and compilation.

As shown in Fig. 1, the number of stations used in this study (1,977 stations selected from a total of 2,479 stations) was significantly higher than that of previous studies (i.e., 57–852 stations) (Kukla and Karl, 1993; Shen and Varis, 2001; Liu et al., 2004; Li et al., 2015). Therefore, the observational data provided better spatial coverage and higher confidence in the detection of regional climate change than in previous studies (Fig. 1).

Observations of T_s from weather stations are different from T_s data retrieved via other approaches, such as satellite images and reanalysis. The T_s observations were performed in 4×2 m square bare land plots proximal to the weather stations. The surface of the observational fields was loose, grassless and flat, and at the same level as the ground surface of the weather station. Three thermometers, including a surface thermometer, a surface maximum thermometer, and a surface minimum thermometer were placed horizontal to the surface of the observational field, with half of each

thermometer embedded in the soil and the other half exposed to the air. When the observational field was covered by snow, the thermometers were placed on the snow surface. Additionally, the exposed parts of the thermometers were cleaned to remove dust and dew.

We verified the reliability of the T_s observational records by analyzing the relationship between T_a and T_s during 1960–2003. As shown in Fig. S1, the mean Pearson ~~Correlation~~correlation ~~Coefficients~~coefficients between daily maximum land surface temperature (T_{s-max}) and daily maximum air temperature (T_{a-max}) calculated from the monthly anomalies were 0.775, 0.843, and 0.806 for the annual, warm, and cold seasonal scales, respectively, and these values were statistically significant (99% confidence level) for all stations. The mean correlation coefficients between the daily minimum land surface temperature (T_{s-min}) and daily minimum air temperature (T_{a-min}) were 0.861, 0.842, and 0.865 for the annual, warm, and cold seasonal scales, respectively, and these values were statistically significant (99% confidence level) for all stations. The high correlations indicated that observations of either T_s or T_a could be used for climate change detection.

The most fundamental energy resource for T_s and T_a is R_s . In most previous studies, the observed R_s have been used to analyze the relation between the variation in R_s and

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199 T_a over China. However, fewer sites were used for R_s observations than for other
200 climatic variables; for example, only 85 sites were used for R_s observations in Liu et al.
201 (2004) and only 90 sites were used in Li et al. (2015).

202 Importantly, sensitivity drift the instruments used for the R_s observations led to a
203 faster dimming rate before 1990, and instrument replacements from 1990 to 1993
204 resulted in a false sharp increase in R_s (Wang, 2014; Wang et al., 2015). The limited
205 distribution and low quality of R_s observations have impeded the wide scientific
206 application of this parameter.

207 Therefore, we used sunshine duration-derived R_s , which is based on an effective
208 hybrid model developed by Yang et al. (2006). This model has subsequently been
209 improved (Wang, 2014; Vose et al., 2005) and it has performed well in regional and
210 global applications (Tang et al., 2011; Wang et al., 2012). Sunshine duration-derived
211 R_s not only accurately reflects the effects of clouds and aerosols on the R_s but also more
212 exactly reveals long-term trends (Wang, 2014; Wang et al., 2015). Additionally,
213 sunshine duration-derived R_s values are better correlated with the satellite retrievals,
214 reanalysis, and climate model simulations than R_s values observed from observation
215 (Wang et al., 2015)(Wang et al., 2015a).

216 The data are collected by a total of 2,474 meteorological stations; however, the

lengths of the effective observation records for the stations are different. Additionally, only a small number of stations were installed before 1960, and the observational records of T_s at many stations were anomalous after 2003 because of automation. Therefore, in our analysis, we selected 1,977 meteorological stations (see Fig. 1) for which the observation records with valid data were longer than 30 years during the 43 years between 1960 and 2003.

The monthly anomalies relative to the 1961–1990 climatology were calculated based on a monthly mean value of the daily values, and when a month was missing more than 7 daily values, that month was classified as a missing value (Li et al., 2015; Sun et al., 2016). For the annual anomalies, the monthly anomalies were averaged for the entire year. The anomalies in the warm seasons were the averages of the monthly anomalies from May to October, and the anomalies in the cold seasons were the averages of the monthly anomalies from November to the next April.

2.2 Methods

As shown in Fig. 1, the spatial distribution of the weather stations throughout China is extraordinarily asymmetric and the density of weather stations in east China is far greater than that in west China. We used the area-weight average method to reduce these biases when calculating the national mean. First, we divided the study region into

235 $1^{\circ} \times 1^{\circ}$ grids (see Fig. S2) for a total 953 grids covering China. Second, we assigned all
236 selected stations to the grids, and this resulted in 627 grids containing stations, which
237 accounted for 65.79% of the total. Finally, the grid box value was the average of all
238 stations in the grid, and the national mean was the area-weight average of all effective
239 grids (Jones and Moberg, 2003).

240 The linear trends reported in this study were calculated via linear regression based
241 on the monthly anomalies of T , R_s , and P . Two national mean trends were calculated
242 from the anomalies of the grids. In the first method (Method I), the national mean
243 monthly anomalies were calculated using the area-weight of each grid first, and then
244 the national mean trend based on the time series of the national average anomalies was
245 calculated. In the second method (Method II), the trend at each grid was calculated first,
246 and then the national mean trend was calculated from the grid trends.

247 In our study, we calculated the national mean trends of the temperatures using
248 Method I and II because both methods have been used in previous studies (Gettelman
249 and Fu, 2008). The results for the two methods are expected to be the same when the
250 time series of all grids is integrated and data are not missing (Zhou et al., 2009);
251 however, when data are missing, small differences may occur (See Table 1). As shown
252 in Table 1, the absolute value of the difference between Method I and Method II ranged

253 from 0.011 to 0.033 °C 10yr⁻¹, which represented 3.4% to 14.3% of the trends (using
254 the results of Method I as the reference). For purposes of clarification, the trends
255 derived from Method I are discussed in the main text, whereas the results from both
256 methods are shown in Table 1.

257 The effect of R_s and P on T_{s-max} and T_{a-max} was determined via a multiple linear
258 regression (Roy and Haigh, 2011) of the monthly anomalies using the following
259 equation:

$$260 \quad T = S_{R_s} \cdot R_s + S_P \cdot P + c + \varepsilon \quad (1)$$

261 where T represents the monthly anomalies of T_{s-max} , T_{s-min} , T_{a-max} , and T_{a-min} ; R_s and P
262 represents the monthly anomalies of surface solar radiation and precipitation
263 respectively. S_{R_s} and S_P are the sensitivities of temperatures to R_s and P , respectively; c
264 is constant term; and ε indicates the residuals of the equation, which represents the
265 contribution from other factors such as longwave radiation flux and internal variability.
266 The coefficients of determination (R^2) for the multilinear regression equation (Eq (1))
267 are shown in Fig. S3, and they indicate the portion of the variance of T that could be
268 attributed to that of R_s and P . High coefficients of determination were obtained, which
269 showed that the linear regression performed well, particularly for South China and the
270 North China Plain. To separate the contributions of R_s and P , we further calculated the

271 partial correlation coefficients between R_s and T (or P and T), which are shown in Fig.
272 S4 and Fig. S5.

273 To determine the effect of R_s and P on the analyzed temperatures, we removed
274 their effects from their original time series of T_{s-max} and T_{a-max} based on the multilinear
275 relationship calculated in Eq (1). Then, we calculated the trends from both the original
276 and adjusted time series. By comparing the derived trends of the original and adjusted
277 time series, we quantitatively assessed the effect of R_s and P on T_{s-max} and T_{a-max} ,
278 particularly for the spatiotemporal pattern of their trends.

279 3. Results

280 3.1. Trends of surface temperature and air temperature

281 3.1.1 Temporal patterns in temperature variability

282 The long-term changes in T_{s-max} and T_{a-max} and T_{s-min} and T_{a-min} from 1960 to
283 2003 are shown in Fig. 2 and Fig. 3, respectively. In addition to the annual variability
284 (Fig. 2a and Fig. 3a), the temperature variability in both warm seasons (May–October;
285 Fig. 2b and Fig. 3b) and cold seasons (November to the following April; Fig. 2c and
286 Fig. 3c) were analyzed. In the annual records, all temperatures exhibited an obvious
287 warming trend throughout China (Fig. 2a and Fig. 3a).

As shown in Table 1, the national mean warming rate from 1960 to 2003 for T_{s-max} was 0.227 ± 0.091 °C 10yr^{-1} (95% confidence level) and T_{a-max} was 0.167 ± 0.068 °C 10yr^{-1} (95% confidence level). The warming rate of T_{a-max} based on the 1,977 stations examined in the current study was slightly higher than the global average (0.141 °C 10yr^{-1}) from 1950 to 2004 (Vose et al., 2005) and the rate obtained from a previous analysis (0.127 °C 10yr^{-1}) of temperatures from 1955 to 2000 based on 305 stations in China (Liu et al., 2004). Additionally, the increases in T_{s-max} and T_{a-max} in the cold seasons were much larger than those in the warm seasons, which is consistent with previous studies of China and other regions (Vose et al., 2005; Ren et al., 2005; Shen et al., 2014).

Similarly, the warming rates of T_{s-min} and T_{a-min} in the warm seasons were also clearly lower than those in the cold seasons. As shown in Fig. 3, T_{s-min} increased by 0.315 ± 0.058 °C 10yr^{-1} (95% confidence level) and T_{a-min} increased by 0.356 ± 0.057 °C 10yr^{-1} (95% confidence level) (see Fig. 3a) from 1960 to 2003. The warming trend of T_a is generally consistent with earlier studies (Liu et al., 2004; Shen et al., 2014; Li et al., 2015); however, these trends are considerably larger than the rates reported for the global average (0.204 °C 10yr^{-1}) (Vose et al., 2005). For the seasonal scales, the warming rate of T_{s-min} and T_{a-min} in the cold seasons was almost double that of the warm seasons from 1960 to 2003 (see Table 1).

The warming rate of T_{s-min} (T_{a-min}) was significantly faster than that of T_{s-max} (T_{a-max}) and the warming rates of all temperatures in the cold seasons were substantially greater than those in the warm seasons (Easterling et al., 1997; Liu et al., 2004; Li et al., 2015). Although previous studies have indicated that the microclimate (e.g. urban heat island) has a larger effect on minimum temperatures because of the lower and more stable boundary layer at night (Christy et al., 2009; Zhou and Ren, 2011), many investigators argue that variability in R_s is the primary reason for the daily contrast in warming rates (Makowski et al., 2009; Sanchez-Lorenzo and Wild, 2012).

3.1.2. Spatial patterns in temperature variability

As shown in Fig. 4, clear spatial heterogeneity was demonstrated in the warming rates for T_{s-max} and T_{a-max} in China from 1960 to 2003. The trends of T_{s-max} and T_{a-max} were statistically higher for the Tibet Plateau, and Northwest and Northeast China (see Fig. S6) compared with the North China Plain and South China. Cooling trends in T_{s-max} even detected for the Sichuan Basin, the Yangtze River Delta, and the Pearl River Delta. Lower rates of warming of T_{a-max} in South China and the North China Plain have also been previously reported (Liu et al., 2004; Li et al., 2015).

The warming rates of T_{s-max} and T_{a-max} in South China and the North China Plain in the warm seasons were considerably lower than those in the cold seasons, which

resulted in stronger spatial heterogeneity in the warm seasons (Fig. 4b and 4h). The spatial and seasonal patterns of T_{a-max} were similar, although they were not as similar as the patterns of T_{s-max} . The spatial contrast in the trends between T_{s-min} and T_{a-min} was much less than that between T_{s-max} and T_{a-max} , although a strong dependence on latitude was observed (Fig. 4d and 4j). Related studies suggested that this dependence was strongly associated with the mode variability in large-scale circulation, such as a negative trend in the North Atlantic Oscillation during this period (Wallace et al., 2012; Ding et al., 2014).

The correlation between T_s and T_a was highly significant. Based on the time series of the national mean yearly anomalies (see Fig. 2 and Fig. 3), the correlation coefficient between T_{s-max} and T_{a-max} was 0.877 and between T_{s-min} and T_{a-min} was 0.976 on the annual scale. In the spatial pattern of the trends (Fig. 4), the correlation coefficient between T_{s-max} and T_{a-max} was 0.488 and between T_{s-min} and T_{a-min} was 0.638 on the annual scale. All of these correlations between T_s and T_a were significant at the 95% significance level, which indicated a close relation between T_s and T_a for both interannual fluctuations and secular trends.

The correlation between T_{s-min} and T_{a-min} was significantly higher than that between T_{s-max} and T_{a-max} . T_{s-min} is closely related to the land-atmosphere longwave

wave radiation balance at night, which is closely associated with the atmospheric
 greenhouse effect (Dai et al., 1999). During the day, T_s is directly determined by the
 land surface energy balance, i.e., the incoming energy (including R_s) and atmospheric
 longwave radiation (Wang and Dickinson, 2013b), and it is partitioned into latent and
 sensible heat fluxes (Zhou and Wang, 2016). Although T_a is dependent on the land-
 atmosphere sensible heat flux, it is also affected by local and/or large-scale circulation.
 Thus, the changes in the land surface energy balance caused by R_s and P have different
 levels of effect on T_s and T_a during the day, which most likely caused the lower
 correlation between T_{s-max} and T_{a-max} than that between T_{s-min} and T_{a-min} .

3.2. Effect of R_s and P on temperatures

3.2.1 Effect of R_s

As shown in Fig. S4, R_s is closely linked with T_{s-max} and T_{a-max} but not with T_{s-min}
 and T_{a-min} , and the correlation between T_{s-max} and R_s was higher than that between $T_{a-
 max}$ and R_s . For the seasonal scales, the partial correlation between T_{s-max} and T_{a-max} and
 R_s in the warm seasons was ~~higher~~ **stronger** than that in the cold seasons, and this
 correlation was stronger in South China and the North China Plain. South China has
 high soil moisture; therefore, the relationship between the energy used for
 evapotranspiration and R_s is approximately linear (Zhou et al., 2007; Wang and

Dickinson, 2013a). However, northwest China presents dry soil over most of the year; thus the energy used for evapotranspiration is more dependent on P . As a result, the energy available for heating the surface and air temperatures is not as closely correlated with R_s . Therefore, the correlation coefficients between R_s and T_{s-max} and T_{a-max} were higher-lower in South-the northwest China.

To quantify the effect of R_s on temperature, the sensitivity of the studied temperatures to changes in R_s was calculated (Eq. (1)). As shown in Fig. S7 shows, T_{s-max} was the most sensitive to R_s , followed by T_{a-max} , and the national means for T_{s-max} was $0.092 \pm 0.018 \text{ } ^\circ\text{C (W m}^{-2}\text{)}^{-1}$ (95% confidence level) and T_{a-max} was $0.035 \pm 0.010 \text{ } ^\circ\text{C (W m}^{-2}\text{)}^{-1}$ (95% confidence level). T_{s-min} and T_{a-min} were not sensitive to R_s because these temperatures are primarily affected by atmospheric longwave radiation night.

Based on the above analysis, we calculated the effect of changes in R_s on the studied temperatures. From 1960 to 2003, the calculations of the monthly anomalies at 1,977 stations indicated that the national mean rate of decrease of R_s was $-1.502 \pm 0.42 \text{ W m}^{-2} 10\text{yr}^{-1}$ (95% confidence level), and the trend was significant in most regions of China (see Fig. S8). Our rate of decrease was considerably less than the global average diminishing rate (form approximately -2.3 to $-5.1 \text{ W m}^{-2} 10\text{yr}^{-1}$) between the 1960s and the 1990s (Gilgen et al., 1998; Stanhill and Cohen, 2001; Liepert, 2002; Ohmura,

2006) and the national mean dimming rate across China (from approximately -2.9 to $-5.2 \text{ W m}^{-2} 10\text{yr}^{-1}$) between the 1960s and the 2000s based on radiation station observations (Che et al., 2005; Liang and Xia, 2005; Shi et al., 2008; Wang et al., 2015).

As noted in the data section, the sensitivity drift and replacement of instruments used for the R_s observations resulted in a significant homogenization of the station observation records (Wang, 2014; Wang et al., 2015), which introduced considerable uncertainty to the trend estimations. Tang et al. (2011) used quality-controlled observational data from 72 stations and two radiation models based on 479 stations to determine that the rate in China decreased from approximately -2.1 to $-2.3 \text{ W m}^{-2} 10\text{yr}^{-1}$ during 1961–2000, and they also showed that R_s values have remained essentially unchanged since 2000. These findings are generally consistent with our results.

Because of the decreasing trend in R_s , the national mean warming trends of $T_{s-\max}$ and $T_{a-\max}$ decreased by $0.139 \text{ }^{\circ}\text{C } 10\text{yr}^{-1}$ and $0.053 \text{ }^{\circ}\text{C } 10\text{yr}^{-1}$, respectively. Spatially, the decreasing rate of R_s in South China and the North China Plain was significantly higher than that in other regions, particularly in the warm seasons (Fig. 5b). Therefore, the cooling effect of decreasing R_s on $T_{s-\max}$ and $T_{a-\max}$ was more significant in South China and the China North Plain, and it resulted in significantly lower warming rates

of T_{s-max} and T_{a-max} in those regions than in the other regions (see Fig. 4). The spatial consistency between the decreasing R_s trend and the slowdown of T_{s-max} and T_{a-max} warming implied that variations in R_s were the primary reason for the spatial heterogeneity of the warming rate in T_{s-max} and T_{a-max} .

3.2.2 Effect of P

As shown in Fig. S5, a significant negative correlation was detected between T_{s-max} and P , and the correlation was more significant in the warm seasons than in the cold seasons. P negatively correlated with temperature because P reduces temperatures by increasing the surface evaporative cooling (Dai et al., 1997; Wang et al., 2006). The national mean sensitivities of T_{s-max} and T_{a-max} to P were -0.321 ± 0.098 °C 10 mm⁻¹ and -0.064 ± 0.054 °C 10 mm⁻¹ (95% confidence level), respectively. As shown in Fig. S9, seasonal and spatial changes in the sensitivity of T_{s-max} and T_{a-max} to P were apparent (Fig. S9a–c and Fig. S9g–i). The sensitivities of T_{s-max} and T_{a-max} were significantly higher in arid regions (dry seasons) than humid regions (rainy seasons) (Wang and Dickinson, 2013a). As expected, T_{s-min} and T_{a-min} were both less sensitive to variations in the P .

The trend in P from 1960 to 2003 over the 1,977 stations showed obvious spatial heterogeneities. A slight increasing trend in P was observed in China during this period

415 at rate of $0.112 \pm 0.718 \text{ mm } 10\text{yr}^{-1}$ (95% confidence level). An increasing P trend was
416 observed in northwestern China and southeastern China, whereas a decreasing trend
417 was observed in the North China Plain, the Sichuan Basin, and parts of northeastern
418 China. However, the P trends were not significant in most regions (see Fig. S8).
419 Variations in P significantly differed by season (see Fig. 6b and Fig. 6c). The seasonal
420 and spatial variations in P are consistent with those of previous studies (Zhai et al.,
421 2005; Wang et al., 2015).

422 For $T_{a-\max}$ and $T_{s-\max}$, the warming trend in the North China Plain, the Sichuan
423 Basin, and parts of northeastern China was aggravated by the reduction in P , whereas
424 the warming trend in northwestern China and in the Mongolian Plateau were slowed by
425 increases in P (Fig. 6d). For the national average, the effect of increasing P resulted in
426 decreases in the warming trends of $T_{s-\max}$ and $T_{a-\max}$ by $-0.007 \text{ }^{\circ}\text{C } 10\text{yr}^{-1}$ and $-0.002 \text{ }^{\circ}\text{C}$
427 10yr^{-1} , respectively. However, the effect of P on $T_{s-\max}$ was approximately an order of
428 magnitude less than that of R_s .

429 3.3. Trends of surface and air temperature after adjusting for the effect of R_s and 430 P

431 Based on the above analysis of the effect of R_s and P on temperatures, we found
432 that variations in R_s and P had little effect on $T_{s-\min}$ and $T_{a-\min}$. However, R_s and P had

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important effect on the trends of T_{s-max} and T_{a-max} (see Fig. S3), particularly in central and South China, where T_{s-max} and T_{a-max} ~~was were~~ more closely related to R_s (see Fig. S4). Therefore, only the effects of R_s and P on T_{s-max} and T_{a-max} were analyzed. After adjusting for the effect of R_s and P (Fig. 7), the warming rates of T_{s-max} and T_{a-max} increased by $0.146\text{ }^{\circ}\text{C}\text{ }10\text{yr}^{-1}$ (64.3%) and $0.055\text{ }^{\circ}\text{C}\text{ }10\text{yr}^{-1}$ (33.0%), respectively. Additionally, the increasing amplitude of warming rates in the warm seasons was significantly higher than that in the cold seasons, which resulted in a seasonal contrast in warming rates, with T_{s-max} and T_{a-max} decreasing by 45.0% and 17.2% respectively (see Table 1).

More importantly, after adjusting for the effect of R_s and P , the spatial coherence of the warming rates of T_{s-max} and T_{a-max} in South China and the North China Plain clearly improved (Fig. 8). The regional differences among the North China Plain, South China, and other regions in China significantly decreased because of the increase in warming rates in South China and the North China Plain. Additionally, the warming trends of T_{s-max} and T_{a-max} became more statistically significant in the North China Plain and South China (see Fig. S10).

To clearly illustrate these changes, we selected two regions in China for further investigation: R1 primarily included the North China Plain and R2 primarily included

the Loess Plateau (see Fig. 9a). Although these regions share the same latitudes, the trend for R_s were substantially different (see Fig. 9b). After adjusting for the effect of R_s and P , the annual trends for T_{s-max} and T_{a-max} in R1 increased by 0.304 and 0.118 °C 10yr⁻¹, respectively, whereas those in R2 increased by only 0.025 and 0.016 °C 10yr⁻¹, respectively. Therefore, following the adjustment, the differences in the warming rates of T_{s-max} and T_{a-max} between R1 and R2 were significantly reduced (see Fig. 9d).

Following the adjustment in R1, the seasonal and diurnal differences in the warming rates of T_{s-max} and T_{a-max} significantly decreased. The differences in warming rates between the warm seasons and cold seasons decreased by 68.7% for T_{s-max} and by 50.8% for T_{a-max} after the adjustment. Additionally, the differences in the warming rates between T_{s-max} and T_{s-min} decreased by 93.4% and between T_{a-max} and T_{a-min} decreased by 59.6% in R1. In R2, the adjustment did not significantly change the seasonal and diurnal differences in temperatures. Overall, the trends for R1 and R2 became more consistent after adjusting for difference in R_s and P (see Fig. 9d).

4. Conclusions and Discussion

Although a general warming trends has been observed throughout China, the

469 regional warming trends show significant spatial and temporal heterogeneity. In this
470 study, we analyzed the spatial and temporal patterns of T_s and T_a from 1960 to 2003
471 and further analyzed and quantified the effects of R_s and P on these temperatures. The
472 primary results of the study are as follows.

473 The national mean warming rates from 1960 to 2003 of T_{s-max} , T_{s-min} , T_{a-max} , and
474 T_{a-min} were 0.227 ± 0.091 , 0.315 ± 0.058 , 0.167 ± 0.068 , and 0.356 ± 0.057 °C 10yr⁻¹,
475 respectively. The warming rates of T_{s-max} and T_{a-max} in South China and the North
476 China Plain were significantly lower than those in the other regions, and the spatial
477 heterogeneity in the warm seasons was greater than that in the cold seasons.

478 During the study period, the R_s value decreased by -1.502 ± 0.042 W m⁻² 10yr⁻¹
479 (95% confidence level), and higher ~~diminishing-dimming~~ rates were observed in South
480 China and the North China Plain. Using a partial regression analysis, we found that R_s
481 plays a distinctly important role in the spatial warming patterns of T_{s-max} and T_{a-max} .

482 After adjusting for the effect of R_s and P , the warming rates of T_{s-max} and T_{a-max} in
483 South China and the North China Plain significantly increased and the regional
484 differences in warming rates in China clearly decreased (see Fig. 8). After the
485 adjustments, the warming rates of T_{s-max} and T_{a-max} in the North China Plain increased
486 by 0.304 and 0.118 °C 10yr⁻¹, respectively, whereas those on Loess Plateau increased

only by 0.025 and 0.016 $^{\circ}\text{C } 10\text{yr}^{-1}$, respectively. Therefore, the differences in warming rates of $T_{s-\text{max}}$ and $T_{a-\text{max}}$ between the North China Plain and the Loess Plateau were almost eliminated (see Fig. 9d).

After adjusting for the effect of R_s and P , the warming trend of $T_{s-\text{max}}$ increased by 0.146 $^{\circ}\text{C } 10\text{yr}^{-1}$ and that of $T_{a-\text{max}}$ increased by 0.055 $^{\circ}\text{C } 10\text{yr}^{-1}$. In addition, the trends of $T_{s-\text{max}}$ and $T_{a-\text{max}}$ became 0.373 ± 0.068 and 0.222 ± 0.062 $^{\circ}\text{C } 10\text{yr}^{-1}$ respectively. Reduction in R_s resulted in decreases in the warming rates of $T_{s-\text{max}}$ and $T_{a-\text{max}}$ by 0.139 $^{\circ}\text{C } 10\text{yr}^{-1}$ and 0.053 $^{\circ}\text{C } 10\text{yr}^{-1}$, respectively, which accounted for 95.0% and 95.8% of the total effect of R_s and P , respectively. For the seasonal contrast, the warming rates of $T_{s-\text{max}}$ and $T_{a-\text{max}}$ decreased by 45.0% and 17.2%, respectively. For the daily contrast, the warming rates of T_s and T_a decreased by 33.0% and 29.1%, respectively.

In addition to R_s and P , temperature warming rates may be affected by many other factors, such as land cover and land use changes; however those factors have not been discussed in this study because of lack of data (Liu et al., 2005; Zhang et al., 2016). After adjusting for the effect of changes in R_s and P changes, the spatial differences in the warming trends clearly decreased; however, certain regional differences remained. The warming rate of $T_{s-\text{max}}$ in the Sichuan Basin remained significantly lower than that in other regions after adjusting for these effects. Additionally, the differences in the

505 warming rates of T_{s_min} and T_{a_min} between the northern and southern areas were not
506 explained by the effects of R_s and P ; further study is required.

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512 (CMA, <http://data.cma.gov.cn/data>).

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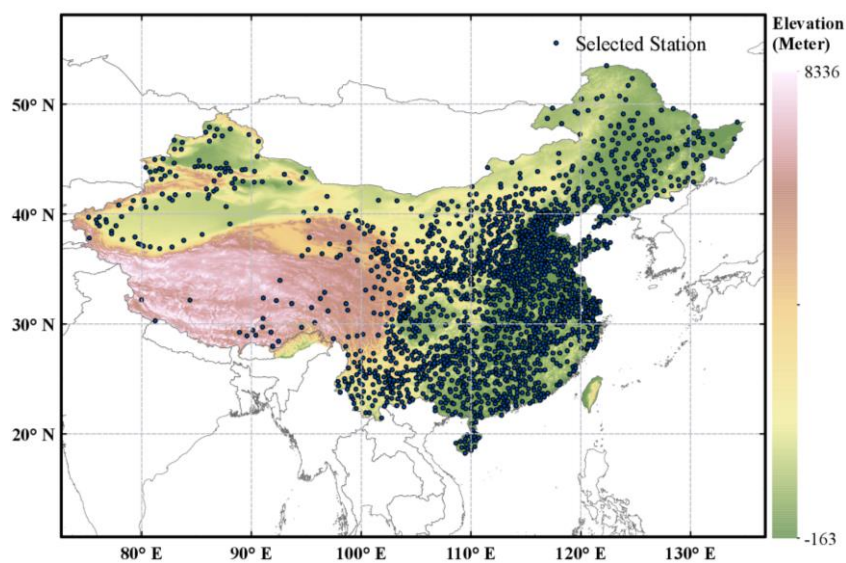
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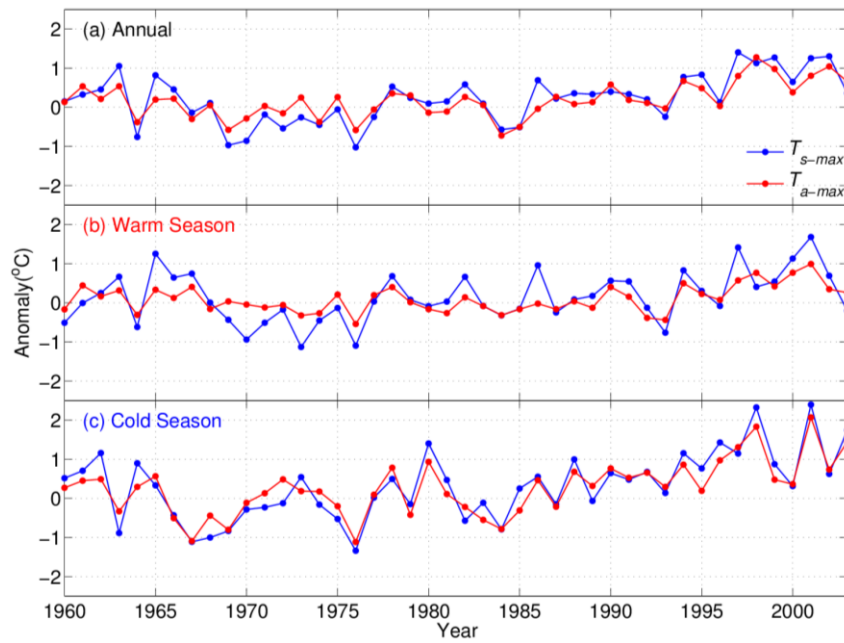
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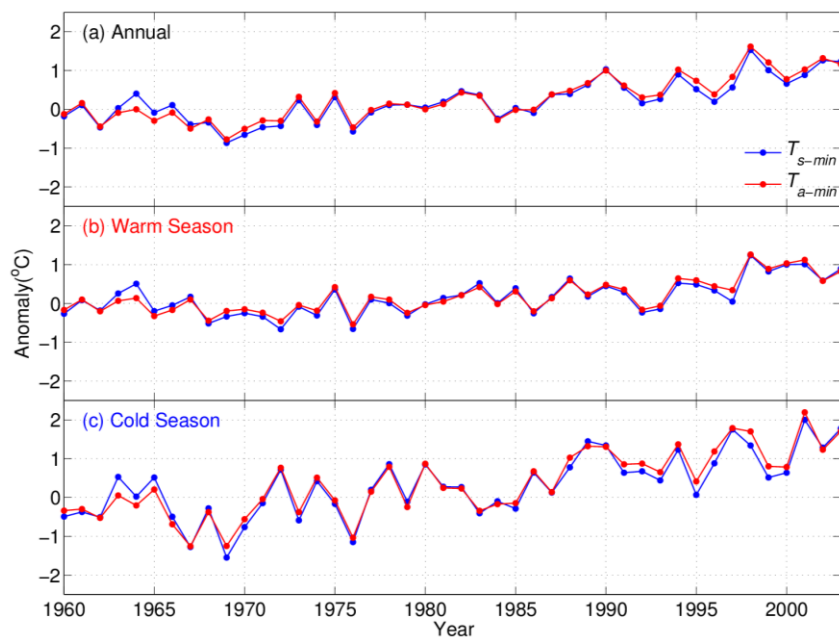
Figure 1. Elevation maps of mainland China and spatial distribution of the 1977 meteorological stations used in this study. The datasets were provided by China's National Meteorological Information Centre (You et al., 2016) (<http://data.cma.gov.cn/data>).

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775 Figure 2. National mean yearly anomalies of daily maximum land surface temperature
 776 (T_{s-max} , blue line) and daily maximum air temperature (T_{a-max} , red line) for the annual
 777 (a), warm (b), and cold (c) seasonal scales for the reference period from 1961 to 1990.



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779 Figure 3. National mean yearly anomalies of daily minimum land surface temperature
 780 (T_{s-min} , blue line) and daily minimum air temperature (T_{a-min} , red line) for the annual
 781 (a), warm (b), and cold (c) seasonal scales for the reference period 1961–1990.

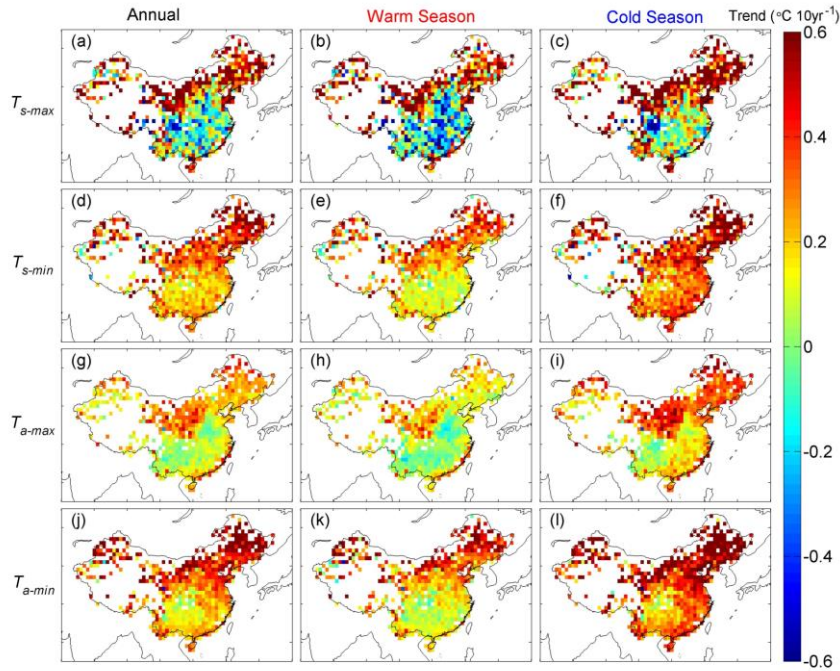


Figure 4. Maps of the trends of the monthly anomalies for daily maximum land surface temperature (T_{s-max} , a–c), daily minimum land surface temperature (T_{s-min} , d–f), daily maximum air temperature (T_{a-max} , g–i), and daily minimum air temperature (T_{a-min} , j–l) for the annual, warm (May–October), and cold (November–next April) seasonal scales. All trends reported in these figures were calculated using a linear regression based on the least square method.

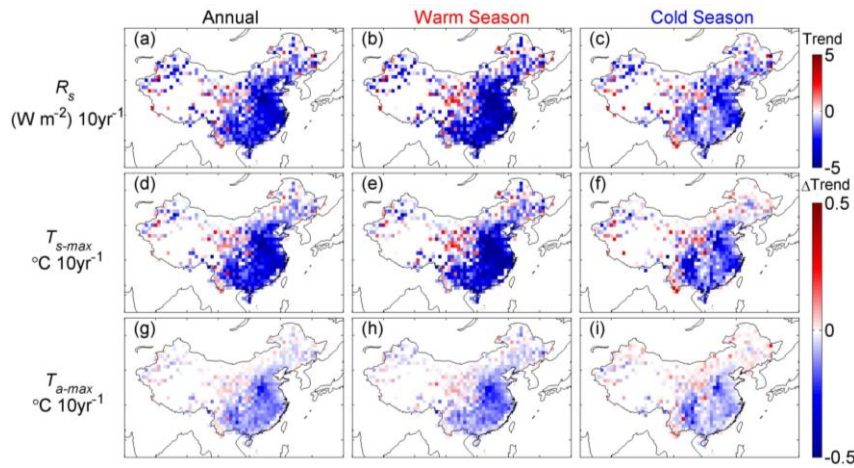


Figure 5. Maps of the trends in surface solar radiation (R_s , a–c) and its effect on the warming rates of daily maximum land surface temperature (T_{s-max} , d–f) and daily maximum air temperature (T_{a-max} , g–i). The first line (a–c) is the trend of R_s from 1960–2003; the second line (d–f) and the third line (g–i) are the trend changes caused by secular variations of R_s on T_{s-max} and T_{a-max} . Eq (1) was used to strip away the effect of R_s on temperatures, and we calculated the trend difference (Δ Trend, d–i) between the time series of temperatures before and after adjusting for the effect of R_s . Finally, the effect of R_s on the trends of T_{s-max} and T_{a-max} was quantified and analyzed (section 3.2.1).

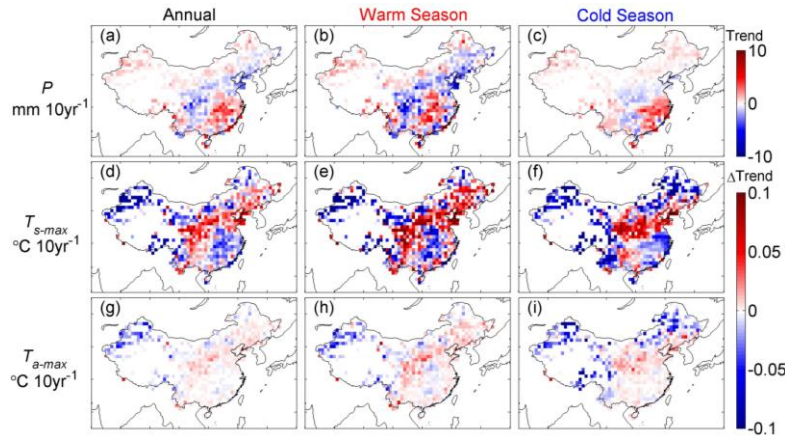


Figure 6. Maps of the trends in precipitation (P) (a–c) and their effect on the warming rates for daily maximum land surface temperature (T_{s-max} , d–f) and daily maximum air temperature (T_{a-max} , g–i). The first line (a–c) is the trend of P during 1960–2003; the second line (d–f) and the third line (g–i) are the trend changes caused by secular variations of P on T_{s-max} and T_{a-max} . We used Eq (1) to remove the effects of P on the temperatures, then calculated the trend difference ($\Delta Trend$, d–i) between the time series of temperatures before and after adjusting for the effect of P . Finally, the effect of P on the trends of T_{s-max} and T_{a-max} was quantified and analyzed (section 3.2.2).

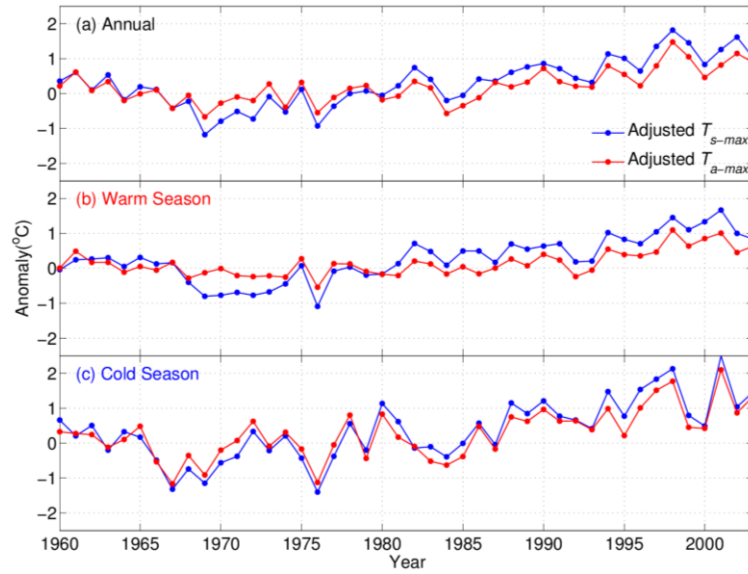


Figure 7. Regional average anomalies of daily maximum land surface temperature (T_{s-max} , blue line) and daily maximum air temperature (T_{a-max} , red line) for the annual (a), warm (b), and cold (c) seasonal scales for the reference period from 1961 to 1990. We used Eq (1) to simultaneously adjust for the effects of surface solar radiation (R_s) and precipitation (P) on T_{s-max} and T_{a-max} and then analyzed the changes in the interannual variation of T_{s-max} and T_{a-max} (section 3.3).

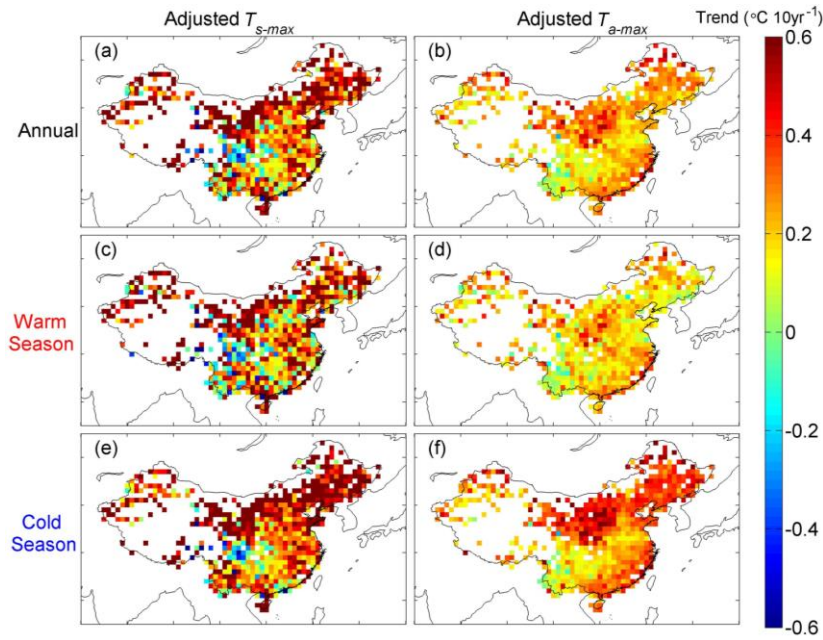


Figure 8. Maps of the trends of the monthly anomalies for the daily maximum land surface temperature (T_{s-max} , a, c, e) and daily maximum air temperature (T_{a-max} , b, d, f) for the annual, warm, and cold seasonal scales after adjusting for the effects of surface solar radiation (R_s) and precipitation (P). We used Eq (1) to simultaneously adjust the effects of R_s and P on T_{s-max} and T_{a-max} and then analyzed the changes in the secular trends of T_{s-max} and T_{a-max} (section 3.3).

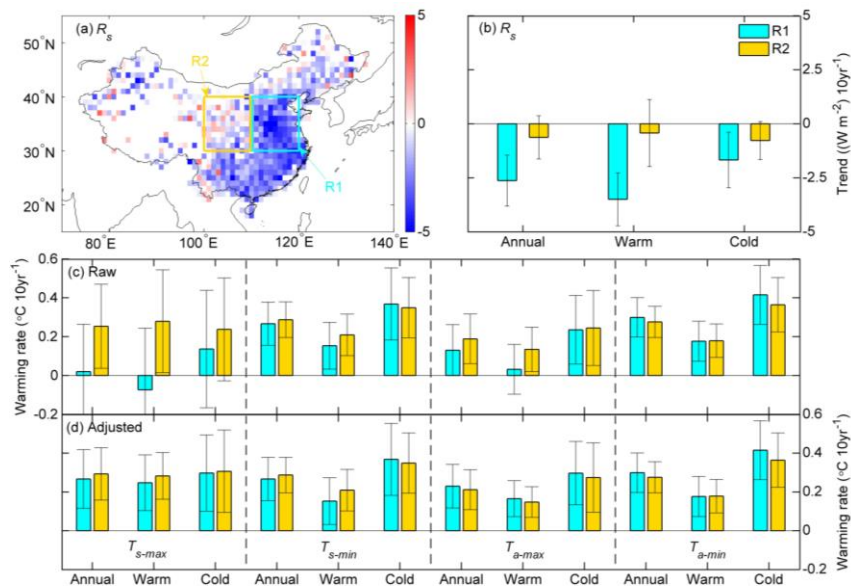


Figure 9. (a) Maps of the trends of surface solar radiation (R_s) and the location of the regions selected for further analysis: R1 (latitude: 30°–40° N; longitude: 110°–120° W) and R2 (latitude: 30°–40° N; longitude: 100°–110° W). (b) National mean trends for R1 and R2. (c) Annual, warm, and cold seasonal scale trends calculated based on the data before adjusting the effect of R_s and P . (d) Annual, warm, and cold seasonal scale trends calculated based on the data after adjusting the effect of R_s and P . All error bars indicate the 95% confidence interval.

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838 Table 1. Warming rates (unit: $^{\circ}\text{C } 10\text{yr}^{-1}$) of the temperatures (T_{s-max} , T_{s-min} , T_{a-max} , T_{a-min})
839 for the annual, warm and cold seasonal scales. Raw and Adjusted represent the warming
840 rates calculated for the data before and after adjusting for the effect of surface solar
841 radiation (R_s) and precipitation (P), respectively. In Method I, the national mean
842 anomalies were calculated first and then the national mean trend based on this time
843 series was calculated. In Method II, the trend of each grid was calculated first and then
844 the national mean value of the trends of all grids was calculated using the area-weight
845 average method. We calculated the national mean trends of the temperatures using both
846 methods.

			T_{s-max}	T_{s-min}	T_{a-max}	T_{a-min}
Method I	Raw	Annual	0.227 \pm 0.091	0.315 \pm 0.058	0.167 \pm 0.068	0.356 \pm 0.057
		Warm	0.172 \pm 0.103	0.221 \pm 0.054	0.091 \pm 0.056	0.245 \pm 0.049
		Cold	0.354 \pm 0.149	0.447 \pm 0.101	0.294 \pm 0.123	0.505 \pm 0.098
	Adjusted	Annual	0.373 \pm 0.068	--	0.222 \pm 0.062	--
		Warm	0.350 \pm 0.064	--	0.160 \pm 0.046	--
		Cold	0.450 \pm 0.119	--	0.329 \pm 0.114	--
Method II	Raw	Annual	0.254 \pm 0.197	0.328 \pm 0.094	0.183 \pm 0.103	0.368 \pm 0.082
		Warm	0.193 \pm 0.285	0.235 \pm 0.095	0.104 \pm 0.109	0.256 \pm 0.081
		Cold	0.321 \pm 0.267	0.415 \pm 0.159	0.264 \pm 0.167	0.476 \pm 0.139
	Adjusted	Annual	0.401 \pm 0.137	--	0.239 \pm 0.086	--
		Warm	0.374 \pm 0.173	--	0.174 \pm 0.082	--
		Cold	0.432 \pm 0.208	--	0.304 \pm 0.152	--

Units: °C 10yr⁻¹, ±95% Confidence interval.

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