



# 1 **Detecting long-term changes in point source fossil CO<sub>2</sub>** 2 **emissions with tree ring archives**

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## 8 9 **Abstract**

10 We examine the utility of tree ring <sup>14</sup>C archives for detecting long term changes in fossil CO<sub>2</sub>  
11 emissions from a point source. Trees assimilate carbon from the atmosphere during  
12 photosynthesis, in the process faithfully recording the average atmospheric <sup>14</sup>C content in  
13 each new annual tree ring. Using <sup>14</sup>C as a proxy for fossil CO<sub>2</sub>, we examine interannual  
14 variability over six years of fossil CO<sub>2</sub> observations between 2004-05 and 2011-12 from two  
15 trees growing near the Kapuni Natural Gas Plant in rural Taranaki, New Zealand. We quantify  
16 the amount of variability that can be attributed to transport and meteorology by simulating  
17 constant point source fossil CO<sub>2</sub> emissions over the observation period with the atmospheric  
18 transport model WindTrax. We compare model simulation results to observations and  
19 calculate the amount of change in emissions that we can detect with new observations over  
20 annual or multi-year time periods given both measurement uncertainty of 1ppm and the  
21 modelled variation in transport. In particular, we ask, what is the minimum amount of change  
22 in emissions that we can detect using this method, given a reference period of six years? We  
23 find that changes of 42% or more could be detected in a new sample from one year at the  
24 same observation location, or 22% in the case of four years of new samples. This threshold  
25 lowers and the method becomes more practical the more the size of the signal increases. For  
26 point sources 10 times larger than the Kapuni plant (a more typical size for power plants  
27 worldwide), it would be possible to detect sustained emissions changes on the order of 10%  
28 given suitable meteorology and observations.



## 1    **1    Introduction**

2    Carbon dioxide (CO<sub>2</sub>) emitted by anthropogenic activity is the largest single contributor to the  
3    radiative forcing causing climate change (IPCC, 2014). It thus plays a crucial role in any  
4    attempt to prevent or mitigate further warming. Large point sources (mainly from electricity  
5    generation and industry) contribute around a third of the total fossil-fuel derived CO<sub>2</sub> (CO<sub>2</sub>ff)  
6    emissions (IPCC, 2014) and in many places are included in government regulatory schemes  
7    that aim to reduce emissions (e.g. European Union ETS, South Korea, Switzerland, and others  
8    at the city/state level; Serre et al., 2015). Emissions are typically reported on an annual basis,  
9    and reduction targets are commonly agreed as annual or multi-year caps, often requiring  
10    changes in emissions relative to a baseline year (e.g. the Kyoto Protocol and the new Intended  
11    Nationally Determined Contributions (INDC), UNFCCC, 2015).

12    Emissions are currently known from “bottom up” techniques such as self-reported data from  
13    fuel usage statistics (Boden et al., 2015) and/or continuous stack monitoring (U.S.  
14    Environmental Protection Agency, 2005; eGRID, 2014) and are subject to significant  
15    uncertainties (Ackerman and Sundquist, 2008; Gurney et al., 2009, 2012). This uncertainty  
16    might include not only methodological biases and possible deliberate underreporting but also  
17    simple error in compiling statistics. The integrity of regulation schemes and their  
18    effectiveness at limiting future climate change will require independent methods of evaluating  
19    reported emissions and improvement in the accuracy of emissions inventories (Tans and  
20    Wallace, 1999; Nisbet and Weiss, 2010; National Research Council, 2010; Gurney, 2013).

21    “Top-down” atmospheric observations can provide an independent method for evaluating  
22    emissions. This involves taking observations of atmospheric gas mole fractions in  
23    combination with atmospheric transport modelling to infer the magnitude of emissions from a  
24    source or region over a particular time period (e.g. McKain et al., 2012; Lindenmaier et al.,  
25    2014; Brioude et al., 2013). It can be quite challenging to quantify absolute values of  
26    emissions and CO<sub>2</sub> fluxes in general because of the large errors and biases typically  
27    encountered in transport models (e.g. Stephens et al., 2007; Lin and Gerbig, 2005; Gerbig et  
28    al., 2008; Prather et al., 2008; Geels et al., 2007; Liu et al., 2011; Kretschmer et al., 2012).  
29    However, *relative* changes in emissions are usually easier to determine, since any consistent  
30    biases in the model will cancel out. By establishing a baseline measurement over a reference  
31    period, we can compare future observations to this reference and calculate relative changes



1 that occur. In this manner, we can potentially verify relative emission reduction targets  
2 without requiring precise knowledge of the absolute levels of emissions.

3 One of the biggest challenges of atmospheric observations of CO<sub>2</sub>ff is distinguishing the  
4 fossil component from the considerable background level of CO<sub>2</sub> that occurs naturally in the  
5 atmosphere, currently about 400 parts per million (ppm; Mauna Loa observation record,  
6 <http://www.esrl.noaa.gov/gmd/ccgg/trends/index.html>, last access: 13 May 2015). In addition,  
7 there are large diurnally and seasonally varying CO<sub>2</sub> fluxes from the biosphere, which may  
8 result in changes in CO<sub>2</sub> mole fraction of tens of ppm within a single day at near-surface sites  
9 (e.g. Miles et al., 2012). This problem can be avoided by using the <sup>14</sup>C isotopic content as a  
10 tracer for CO<sub>2</sub>ff. CO<sub>2</sub>ff contains no <sup>14</sup>C: the half-life of <sup>14</sup>C is 5,730 years (Karlen et al.,  
11 1968), and all of the <sup>14</sup>C has decayed away from fossil fuels. Other sources of CO<sub>2</sub> have  
12 roughly the same <sup>14</sup>C content as the atmosphere. By measuring the <sup>14</sup>C content of CO<sub>2</sub> or a  
13 proxy for CO<sub>2</sub>, we can calculate the portion of observed CO<sub>2</sub> that comes from recently added  
14 fossil fuel emissions (Levin et al., 2003; Meijer et al., 1996; Turnbull et al., 2006).

15 Plant material can be used as a proxy for atmospheric CO<sub>2</sub>ff because plants assimilate carbon  
16 from the atmosphere during photosynthesis, in the process faithfully recording the <sup>14</sup>C content  
17 in new plant material. The radiocarbon content in tree rings has been well established as a  
18 tracer for fossil CO<sub>2</sub> emissions (Suess, 1955; Tans et al., 1979; Djuricin et al., 2012;  
19 Rakowski et al., 2013) and as a method to detect leaks from CO<sub>2</sub> geosequestration (Donders  
20 et al., 2013). Tree rings represent an integrated average of daytime CO<sub>2</sub> atmospheric mole  
21 fractions and <sup>14</sup>C content over the tree's annual growth period. This allows for a retroactive  
22 analysis of CO<sub>2</sub>ff mole fractions over many years, including any trends in emissions that  
23 occurred during the life of the tree.

24 In this study, we evaluate whether we can detect changes in CO<sub>2</sub>ff emission rates from a point  
25 source on an annual time scale using the CO<sub>2</sub>ff mole fraction derived from the <sup>14</sup>C content of  
26 tree ring archives. Variations in the observed CO<sub>2</sub>ff mole fraction at a given location are  
27 dependent on not only the emission rate but also on atmospheric transport, which in turn is  
28 subject to naturally varying meteorological conditions (e.g. wind speed and direction,  
29 temperature, pressure, etc.). Detecting a change in the emission rate requires disentangling  
30 this change from the natural variability in transport and meteorology as well as from  
31 measurement uncertainty in the observations. The question we ask in this paper is: can we use  
32 tree ring archives to detect changes in CO<sub>2</sub>ff emissions from a point source, and if so, what is



the minimum change in annual emissions that we can detect given the typical measurement uncertainty of 1ppm and natural variability in transport? A similar analysis was carried out by Levin and Rodenbeck (2007) at the regional scale, using a 20-year time series of  $^{14}\text{C}$  observations over Germany. McKain et al. (2012) also assessed the ability of an observation-model framework to detect changes in regional urban  $\text{CO}_2$  emissions on a monthly time scale. We re-examine this question on the scale of an individual point source with mean annual observations.

We calculate interannual variability in observations from tree ring archives of annual (growing season)  $\text{CO}_2\text{ff}$  between 2004-05 and 2011-12<sup>1</sup>, taken from two different trees growing south of the Kapuni Natural Gas Treatment Plant in rural New Zealand (Norris, 2015). We then use an atmospheric transport model, WindTrax, with local meteorological data to quantify the interannual variability that can be expected due to measurement uncertainty, transport and meteorology at different distances and orientations from the source, including the locations of the trees. Finally, we look at what this implies for detection limits in the context of emissions monitoring or verification and practical considerations in the presence of multiple sources of uncertainty.

## 2 Methods

### 2.1 Site

The site of our study is the Kapuni Natural Gas Treatment Plant in rural Taranaki, New Zealand (39.477° S, 174.1725° E, 170 m.a.s.l.) (Fig. 1). This site was chosen because it is located in flat terrain and is relatively isolated from other sources of  $\text{CO}_2\text{ff}$ , considerably simplifying measurement and analysis. The gas treatment plant, owned and operated by Vector, processes natural gas extracted from natural gas wells in the Taranaki Basin. The gas contains around 40%  $\text{CO}_2$ , which is removed during processing and vented to the atmosphere at a rate of  $\sim 0.1 \text{ TgC yr}^{-1}$  (NZMED, 2010). In addition, there is an ammonia urea manufacturing plant 500m to the west of the gas plant (Fig. 1), operated by Ballance Agri-Nutrients, which also releases  $\text{CO}_2\text{ff}$  to the atmosphere during the manufacturing process. This site emits roughly a third of the amount of the Vector gas plant ( $\sim 0.03 \text{ TgC yr}^{-1}$ )

<sup>1</sup> Henceforth in this paper, the growing season spanning 1 September to 30 April will be referred to by the year in which the season began, i.e. 2004-05 will be designated 2004.



(Taranaki Regional Council, 2013). Although the signal from the Vector plant is much stronger, especially to the east (downwind from the dominant westerly winds), emissions from the Ballance plant are potentially large enough to detect at some locations and are included in our simulations unless otherwise specified.

The surrounding terrain is flat and mostly free of obstructions, with elevation varying no more than 10m within 2km of the plant. The largest nearby topographic feature is a dip of ~5m into the Kapuni stream immediately east of the Vector emission source. The landscape is dominated by highly productive pasture grazed by dairy cows, with large and diurnally varying CO<sub>2</sub> fluxes. The prevailing wind direction is from the west, with a smaller proportion from the southeast and north (Figs. 2 and 3).

## 2.2 CO<sub>2</sub> emissions

Emissions data were supplied by Vector as monthly totals (Peter Stephenson, personal communication), which we have converted to average daily rates for the purpose of modelling. Mean annual daily emissions for each year between 2004 and 2011 from 1 September to 30 April are shown in Fig. 4; data are listed in Table S1. The long-term mean is 5341 gC s<sup>-1</sup>, with a standard deviation in annual means of 388 (7.3%). There are annual fluctuations but no long-term trend over the modelled period 2004-2011. The largest change during a single year occurred in 2008, when the emissions dropped by 14% relative to the mean. On a longer time scale, there are more significant changes, including the start of operations at the Vector Plant in 1971. However, we focus on the 2004-2011 period during which high resolution local meteorological data is available. There are no significant seasonal or diurnal variations in the emissions of which we are aware.

The Ballance Agri-Nutrients Plant emissions are reported on an annual basis (Taranaki Regional Council, 2013). Average daily rates in each growing season are depicted in Fig. 4. The mean daily rate of emissions over the period 2004-2011 is 1512 gC s<sup>-1</sup> with a standard deviation in annual means of 88 (18%), which is more variable than the emissions from the Vector plant, but smaller in absolute terms. Emissions vary somewhat from day to day according to production levels, but more detailed daily or monthly information is unavailable; for simplicity we assume a constant emissions rate in each year. We note that emissions are much lower in 2011, which is due to downtime after both a fire and scheduled maintenance (Taranaki Regional Council, 2013).



### 2.3 Tree ring observations

Tree rings faithfully record the  $^{14}\text{C}$  content of assimilated  $\text{CO}_2$ , so when the rings are independently dated by dendrochronology, we can determine an average  $^{14}\text{C}$  content and recently added  $\text{CO}_2\text{ff}$  in the local atmosphere for the period during which the tree ring was laid down. We use core samples from two trees located south of the plant, a pine tree (*Pinus radiata*) and a chestnut tree (*Castanea sativa*) (Fig. 1; Norris, 2015). The pine tree is located in a stand of trees within 5m of the Kapuni stream, with the crown reaching 10m above the associated terrain dip. The chestnut is isolated in a flat paddock.

Each tree ring is assumed to represent the Southern Hemisphere summer growth period from 1 September to 30 April, as this is when the majority of plant photosynthesis occurs and new plant material is produced. The sample preparation, measurement and determination of  $\text{CO}_2\text{ff}$  are described in detail by Norris (2015). In summary, alpha cellulose was extracted from individual rings, combusted, reduced to graphite and measured by accelerator mass spectrometry.  $\text{CO}_2\text{ff}$  was determined following Turnbull et al. (2014) from the isotopic difference between the measured tree ring and clean air background  $\text{CO}_2$  measured at Baring Head, Wellington (41.4167°S, 174.8667°E; Currie et al., 2011; extended dataset to 2015 will be presented in an upcoming publication). Baring Head, located at the southern end of New Zealand's North Island and approximately 220 km southeast of Kapuni, was chosen as the background for this study over more local sites because it provides a long-term record of background  $\text{CO}_2$  and  $^{14}\text{C}$ , dating back to the early 1970s. Background levels in tree rings measured at a site in Kapuni 2km upwind of the Vector plant are close to those measured at Baring Head in the same time period, justifying the use of the Baring Head dataset (Norris, 2015). Uncertainty in  $\text{CO}_2\text{ff}$  is dominated by  $\Delta^{14}\text{C}$  measurement uncertainty in both background and the observed sample and is typically ~1ppm for this dataset.

The process of  $\text{CO}_2$  adsorption in plants is extremely complex. For simplicity, we assume a constant assimilation rate over all daylight hours. In reality,  $\text{CO}_2$  adsorption varies with plant species and photosynthesis rates, being weighted towards sunny periods and midday (Bozhinova et al., 2013). There are also many different climatic and nutrient limitations that can only be properly accounted for with a full process-based biogeochemical model of plant growth, which is beyond the scope of this study. We do, however, take into consideration the fact that plant material will tend to underestimate mean  $\text{CO}_2\text{ff}$  when  $\text{CO}_2\text{ff}$  is variable, as in the case of a plume from a point source (see Sect. 2.7).



## 2.4 WindTrax model

WindTrax (WindTrax 2.0; Thunder Beach Scientific, Nanaimo, Canada, [www.thunderbeachscientific.com](http://www.thunderbeachscientific.com)) is a Lagrangian particle dispersion model used to estimate unknown trace gas concentrations or emission rates from a source over short distances (~1km). WindTrax has been applied to agricultural emissions from area sources, such as methane, ammonia, and other gasses from grazing dairy cows, cattle feedlots and farm waste (e.g. Flesch et al., 2005; Laubach and Kelliher, 2005; Bonifacio et al., 2013; Rhoades et al., 2010; Wilson et al., 2012; McBain and Desjardins, 2005). It has also been assessed in the context of CO<sub>2</sub> sequestration leakage detection (Leuning et al., 2008; Loh et al., 2009). Modelling integrated averages of CO<sub>2</sub>ff in plant material is a relatively new application. WindTrax was chosen for this study because it is easy to use and the distance scale is appropriate for our site. We previously used WindTrax to estimate CO<sub>2</sub>ff in grass samples at the Kapuni site (Turnbull et al., 2014), demonstrating that the model is capable of providing reasonable estimates of observed CO<sub>2</sub>ff. Here, we take the same approach to model CO<sub>2</sub>ff measured in tree rings. We note that WindTrax is not applicable to complex terrain or larger distance scales and caution is urged when applying our methodology to other sites.

WindTrax simulates the transport of trace gases by releasing a set number of particles at each time step and following each particle's trajectory downwind. Based on Monin–Obukhov similarity theory, the physics underlying the model is described in detail in Flesch et al. (2004) and Wilson and Sawford (1996). The model equations are valid in the atmospheric surface layer. It assumes wind and other meteorological observations are averaged over a suitable time interval representing a stable, mean atmospheric state (10-30 minute intervals are recommended). Intervals longer than one hour have been shown to be problematic (Flesch et al., 2004) because at these time intervals, large-scale fluctuations not built in to the model become important. In this study, we use one hour time steps to match the resolution of our meteorological dataset (see Sect. 2.5).

The model can be run in forward (fLS) or inverse/backward (bLS) mode, depending on whether the emissions or the trace gas mole fractions are unknown. In all simulations described here we start with known emission rates and use the fLS mode to estimate the CO<sub>2</sub>ff mole fraction at locations surrounding the plant. Model “concentration sensors” represent simulated measurements of mole fractions at designated locations and supply the main model output.



1 The model is stochastic, meaning that it introduces random turbulence into particle  
2 trajectories, and no two runs are identical, even with the same parameters and meteorological  
3 input. There is, therefore, inherent error in the model predictions due to the randomness  
4 introduced in the transport process. Only the average behaviour of a group of particles can be  
5 determined, and releasing more particles at each time step will tend to reduce the degree of  
6 uncertainty. Statistical error (or the standard deviation within each set of trajectories) is  
7 calculated and output by the model at each time step. However, any biases in the modelled  
8 transport or the meteorological input data used to drive the model are not accounted for.

## 9 **2.5 Meteorology**

10 Modelling with WindTrax requires at a minimum wind speed, wind direction, air temperature,  
11 and atmospheric pressure at each time step. We use hourly meteorological data from the  
12 Hawera Automatic Weather Station (AWS) (39.6117°S, 174.2917°E, 98 m.a.s.l), downloaded  
13 from the New Zealand National Climate Database (CliFlo, 2014). Hawera, approximately  
14 20km distance to the southwest of Kapuni, is the nearest location with a nearly complete long-  
15 term dataset of hourly wind direction and speed. Eight years of data (2004-2011) were  
16 available at the time of our study. We use only data from the growing season (1 September –  
17 30 April) and daylight hours (08:00 – 18:00 local daylight savings time) in the model  
18 simulations to correspond to the time period during which trees assimilate CO<sub>2</sub>.

19 The area to the northwest of Hawera and Kapuni is dominated by Mount Taranaki, a 2518m  
20 volcanic cone that rises steeply from relatively flat surrounding terrain. Wind direction and  
21 speed can be very different at sites only a few kilometres apart because of the local impact of  
22 the mountain on atmospheric flow. Thus we compared Hawera and Kapuni meteorological  
23 datasets to ensure that Hawera is representative of Kapuni over long (~1 year) time periods  
24 and the wind speed and direction distributions as a whole are similar at both locations. A wind  
25 rose for the eight years (2004-2011) of data at Hawera is shown in Fig. 2, together with a  
26 wind rose for one year (2013) of data at Kapuni. Wind speeds are on average higher at  
27 Hawera, but the distribution in direction is very similar, with a small overrepresentation of  
28 northerlies at Hawera. The wind speed and direction distributions at both locations are shown  
29 in more detail in Fig. S1.

30 We demonstrate correlation between the two sites using the only overlapping dataset that was  
31 available for direct comparison at the time of the study. We collected data at a temporary





1 meteorological station at Kapuni at 10-minute intervals during the period 14 August – 26  
2 October 2012, with some significant data gaps (Turnbull et al., 2014). These were averaged to  
3 hourly intervals and compared with the corresponding set of measurements at the Hawera  
4 AWS. Only daylight hours were included for consistency with the model simulations. Using  
5 these datasets, correlation in wind speed is good, with  $R^2 = 0.82$ , and correlation in wind  
6 direction is moderate ( $R^2 = 0.61$ ). Because wind direction is an angular measurement,  
7 correlation in wind direction was performed using the circular package v0.4-7 in R v3.0.2  
8 (Lund and Agostinelli, 2013; R Core Team, 2013) rather than the standard linear correlation  
9 function. Scatter plots comparing wind at Kapuni and Hawera directly at each time step are in  
10 Fig. S2. Wind speed is a good match, with Hawera on average having slightly higher speeds  
11 than Kapuni. With wind direction, most points are close to the 1:1 line or slightly below,  
12 indicating a small rotation in direction between the sites. Approximately 67% of data points  
13 (one sigma) are within  $30^\circ$  of each other, and 85% are within  $45^\circ$ . For the purpose of our  
14 simulation in which we focus on integrated averages rather than particular points in time, the  
15 Hawera dataset is sufficiently representative of typical conditions at Kapuni.

16 We expect variability in  $\text{CO}_2$ ff mole fraction to be strongly related to variability in wind  
17 speed and direction, and consequently sampling location. Annual mean wind speed does not  
18 vary by much; the mean wind speed over all eight years is  $6.3 \text{ m s}^{-1}$ , and the standard  
19 deviation in annual mean is  $0.11 \text{ m s}^{-1}$ , which is only 2% of the mean. Mean wind direction is  
20  $273^\circ$  (from the west), but there is also a significant amount of wind from the southeast and  
21 north-northeast (Figs. 2 and 3). This general pattern did not change from year to year over the  
22 eight years of the simulation, but relative proportions in each direction did sometimes vary  
23 considerably (Fig. 3). In particular, northerlies (the direction most relevant to our  
24 observations) range from 21-28% of the total, a 30% change in the northerly fraction. While  
25 always the largest category, the percentage of westerlies varies between 38-52%. It is notable  
26 that there are very few periods with calm winds; the region is in general very windy.

## 27 2.6 Model parameters

28 Several model parameters are held constant throughout all simulations. The modelled surface  
29 is short grass (surface roughness  $z_0 = 2.3\text{cm}$ ), since the majority of the surrounding area is  
30 grazed dairy pasture. The heights of the two emissions stacks are set to their known values:  
31 35m above ground level for Vector and 36m for Ballance. The model's atmospheric stability  
32 parameter is also held constant using the general class of 'moderately unstable'. While this is



not true for all modelled time periods, in the absence of measurements from a 3D sonic anemometer or other reliable indicators of atmospheric stability, a general stability class is a first approximation. We tested the model at a different constant stability class ('slightly unstable') and found no significant difference in the amount of variability (results not shown). We note, however, that atmospheric stability is a potential source of error; others have found that stability is an important parameter that can bias results, and model estimates are generally improved with input from a sonic anemometer or vertical profiles of wind speed and temperature (Flesch et al., 2004; Gao et al., 2009; Koehn et al., 2013).

Model concentration sensors at the locations of the pine and chestnut trees are placed at heights of 15.0m and 5.0m, respectively, reflecting the approximate height of the canopy. A single height at each tree was chosen to reduce model complexity and runtime; however, we recognize that in reality CO<sub>2</sub> is assimilated over a range of heights at each tree, corresponding to the vertical spread of the canopy. Some previous studies have indicated that concentrations modelled with WindTrax are sensitive to sampling height and/or the ratio of sampling height to distance from the source (e.g. McBain and Desjardins, 2005; Laubach and Kelliher, 2005; Laubach, 2010). To test for dependence on height, we simulated CO<sub>2</sub>ff along a 20m vertical profile at the location of the pine and chestnut trees (results not shown). Results vary somewhat according to height, and averaging over a 5m height range slightly reduces the mean and interannual standard deviation, but not enough to change our results significantly.

## 2.7 Simulations

We ran a "constant emissions, variable meteorology" simulation at an hourly time step with all eight years of available meteorological data from Hawera (excluding night time and winter months), concentration sensors placed at the locations of the trees, and both the Vector and Ballance plants as CO<sub>2</sub>ff point sources (Fig. 1). Because emissions are held constant, this simulation enables us to isolate contributions to variability from meteorology and transport. For each tree, four concentration sensors were placed on the vertices of a square, with sides of length 30m, centred on the location of the trees and averaged to reduce model transport error. The emission rate at each source was the reported mean rate over the entire modelled period.

In addition to the model sensors at the locations of the trees, we placed sensors at hypothetical locations in four directions and two horizontal distances from the emissions source to examine more general model sensitivity and variability due to meteorological conditions at



1 our site without being tied to the locations of specific observations. Eight additional sensors  
 2 were placed 1.5m above the ground in the four cardinal directions relative to the Vector plant,  
 3 one each at 300m and 600m horizontal distance from the source. Only one point source, the  
 4 Vector plant, was included in the results at these sensors to simplify analysis. Emissions are  
 5 constant at the Vector mean rate over the eight years.

6 We also ran a “constant meteorology, variable emissions” simulation in which we repeat the  
 7 meteorology from one year (2004) and allow emissions rates to vary according to the reported  
 8 values. This allows us to examine model annual variability due to emissions, independent of  
 9 transport.

10 We subsequently generated a “variable emissions, variable meteorology” simulation by  
 11 scaling modelled mole fractions at the tree rings from the constant emissions, variable  
 12 meteorology simulation according to reported emissions levels in each year (Fig. 4). This is  
 13 valid because the relationship between source strength and concentration flux passing through  
 14 a location downwind is linear (Leuning et al., 2008). In addition, under unstable atmospheric  
 15 conditions the emissions leave the model domain within one hour and do not return, so data in  
 16 a given year is not affected by the emissions from previous years. This simulation is used to  
 17 compare the model to observations.

18 Because plant material will underestimate mean CO<sub>2</sub>ff when CO<sub>2</sub>ff is variable, rather than  
 19 comparing the tree ring measurements to the raw model output of CO<sub>2</sub> mole fractions, we  
 20 calculate a modelled “CO<sub>2</sub>ff<sub>tree</sub>”. This is the CO<sub>2</sub>ff that the model would predict from the plant  
 21 material given measured background levels and the equations governing Δ<sup>14</sup>C. We use the  
 22 following equations:

$$23 \quad \Delta_i = \frac{\Delta_{bg} C_{bg} + \Delta_{ff} C_{ff_i}}{C_{bg} + C_{ff_i}} \quad (1)$$

$$24 \quad \Delta_{tree} = \frac{1}{N} \sum_{i=1}^N \Delta_i \quad (2)$$

$$25 \quad C_{ff_{tree}} = \frac{C_{bg} (\Delta_{tree} - \Delta_{bg})}{\Delta_{ff} - \Delta_{tree}} \quad (3)$$

26 where  $\Delta = \Delta^{14}\text{C}$ ,  $C_{ff_i}$  is the modelled CO<sub>2</sub>ff at the  $i^{\text{th}}$  time step,  $N$  is the total number of model  
 27 time steps,  $C_{bg}$  and  $\Delta_{bg}$  are measured (Norris, 2015), and  $\Delta_{ff} = -1000$ . The basic  
 28 derivation of this equation can be found in Turnbull et al. (2006). This accounts for the fact



1 that plant material will assimilate roughly the same amount of  $\text{CO}_2$  at each time step  
2 regardless of the variability in atmospheric  $\text{CO}_2$  mole fraction induced by the emission plume,  
3 and thus the  $\Delta^{14}\text{C}$  of the plant material represents a simple mean of the  $\Delta^{14}\text{C}$  in the assimilated  
4  $\text{CO}_2$  at each time step. In contrast, sampling of whole air across the same time period would  
5 collect more  $\text{CO}_2$  during times of high  $\text{CO}_2$  mole fraction, weighting the resultant  $\Delta^{14}\text{C}$   
6 towards these periods. This results in a  $\text{CO}_2\text{ff}_{\text{tree}}$  that is lower than would be obtained by  
7 determining the simple mean  $\text{CO}_2\text{ff}$  from the modelled mole fractions. Model results from the  
8 variable emissions simulation reported in Fig. 4 and Sect. 3 were derived using these  
9 equations.

10

## 11 **3 Results and Discussion**

### 12 **3.1 Observation and model comparison**

13 We first compare modelled  $\text{CO}_2\text{ff}_{\text{tree}}$  to the observed tree ring  $\text{CO}_2\text{ff}$  to evaluate the model's  
14 ability to estimate annual integrated averages in this context and to identify possible biases  
15 and error in the model. Our observations from tree rings consist of six annual measurements  
16 of  $\text{CO}_2\text{ff}$  from both the pine tree and the chestnut tree between 2004 and 2011 (2008 and 2010  
17 are missing) (Fig. 4). The means over this period are 5.4ppm (pine) and 2.1ppm (chestnut)  
18 (Table 1). Mean modelled  $\text{CO}_2\text{ff}_{\text{tree}}$  over the same six years (excluding the two years without  
19 observations, 2008 and 2010) is 6.1ppm and 2.2ppm for the pine and chestnut tree,  
20 respectively. The modelled mean is almost an exact match for the chestnut tree (difference of  
21 0.1ppm) and within error for the pine tree (difference of 0.7ppm). Figure 4 shows a direct  
22 comparison between measured and modelled  $\text{CO}_2\text{ff}$  for each year. At the pine tree, model  
23 performance is very good: four of the six (66%) annual observed values are within one sigma  
24 of the modelled values, and the remaining two are within two sigma. The agreement for  
25 individual years at the chestnut tree is poorer, but with large errors in the observations and the  
26 distance from the source close to the limit of model capabilities, this is expected.

27 The model is able to simulate both the long-term mean and the annual variation in  $\text{CO}_2\text{ff}_{\text{tree}}$   
28 with a reasonable degree of accuracy, and there are no significant biases apparent. Thus we  
29 can be confident that the model is representative of relative interannual variability in  
30 transport, which is the focus for the remainder of this paper.



## 3.2 Drivers of interannual variability in CO<sub>2</sub>ff

Detecting changes in emissions requires disentangling the changes in CO<sub>2</sub>ff due to emissions from other sources of interannual variability. We now examine the variability in our observations and turn to our model simulations to determine the relative contributions from emissions, transport, and measurement uncertainty.

The observed standard deviations of the six annual CO<sub>2</sub>ff values from the tree rings are 0.8ppm (14% of the six-year mean) and 1.1ppm (51%) for the pine and chestnut tree, respectively (Table 1). This includes not only variability in emissions but other sources of uncertainty such as meteorology and transport, variable <sup>14</sup>C assimilation rates in the trees, precision of measurements, and background corrections. Measurement uncertainty in particular is important at these relatively small concentrations. Given that the standard deviations are very close to the typical measurement uncertainty of ~1ppm, the scatter in annual means can be attributed in large part to this factor alone. For example, at the pine tree, we would expect at least four out of six measurements to be within 1ppm (one sigma) of the long-term mean, all other factors being constant. This is indeed true of four of the six observations. Measurement uncertainty is proportionally much higher in the case of the chestnut tree, which is ~1km from the Vector plant and where the average signal is only ~2ppm. At this distance measurement uncertainty would seemingly dominate other sources of variability. In contrast, the pine tree is much closer to the source (~400m), and the signal is two to three times larger. Variations in emissions will make up a larger proportion of the total variation and are more likely to be detectable at current measurement precision.

The standard deviations of modelled CO<sub>2</sub>ff<sub>tree</sub> in the variable emissions, variable meteorology simulation are 0.5ppm (7.8%) and 0.3ppm (15%) at the pine and chestnut tree, respectively (Table 1). Adding measurement uncertainty of 1ppm in quadrature, we would predict the standard deviations of the annual means in observed CO<sub>2</sub>ff to be 1.1ppm (18%) and 1.0ppm (47%) for the pine and chestnut, respectively, if variability in emissions, atmospheric transport and measurement uncertainty explain all of the interannual variability. In comparison, the observed standard deviations of the annual means are 14% of the long-term mean at the pine tree and 51% at the chestnut tree. Thus emissions, transport, and measurement uncertainty are able to explain the interannual variability in the observations within error.



1 We can estimate the relative proportion of interannual variability that is due to atmospheric  
2 transport using the constant emissions model simulation, in which the only source of  
3 variability is meteorology. The modelled mean CO<sub>2</sub>ff over the six years with observations is  
4 7.4ppm and 2.7ppm for the pine and chestnut, respectively, and modelled standard deviations  
5 are both 0.5ppm (6.6% and 19% of the respective means) (Table 1). Over the full eight years  
6 of the model simulation, the means and standard deviations are 7.7 / 0.9 ppm (12%) and 2.7 /  
7 0.5 ppm (19%), respectively.

8 Examining more general patterns of meteorological and transport variability at the Kapuni site  
9 apart from the locations of the trees reveals that the variation is highly dependent on the  
10 direction of the observation location relative to the source. The results at the eight  
11 hypothetical sensors averaged in each individual year and means for the entire eight years of  
12 simulation are compared in Fig. 5, and the long-term means and standard deviations are given  
13 in Table 2. The variation to the south of the plant (10-11% of the mean) is the lowest of any  
14 direction and consistent with the variation found at the pine tree in the constant emissions  
15 simulation over the full eight years (12%). Absolute CO<sub>2</sub>ff mole fractions are highest in the  
16 east (westerlies being dominant), but standard deviations are slightly higher at 14% of the  
17 mean. Concentrations in the west are low (~2ppm) and highly variable, the result of the low  
18 percentage of easterlies in any given year (Fig. 3). Variation is relatively insensitive to the  
19 distance from the source.

20 It is apparent that wind direction drives a large part of the variation in transport. Annual  
21 modelled CO<sub>2</sub>ff at the trees in the constant emissions simulation is correlated with the annual  
22 percentage of wind in the direction +/- 30° of the direct line between the source and the tree,  
23 corresponding to the plume trajectories that are most likely to pass through the tree locations  
24 (Fig. S3;  $R^2 = 0.56$  and  $0.72$  for the pine and chestnut tree, respectively). The same correlation  
25 between wind direction and modelled CO<sub>2</sub>ff at all eight hypothetical sensors combined gives  
26 an  $R^2$  of 0.58. Over half of the transport variability is thus explained solely by variation in the  
27 percentage of wind in each direction. However, other meteorological variables and model  
28 parameters (e.g. wind speed, temperature, pressure, etc.) still play a non-negligible role, as the  
29 annual variation in wind direction is not equivalent to the interannual variability in modelled  
30 CO<sub>2</sub>ff.

31 In the same manner, we can determine the contribution of changes in emission rates to the  
32 overall interannual variability with the constant meteorology simulation in which emissions



1 vary but transport is the same in each year. This results in interannual variability in CO<sub>2</sub>ff  
2 similar to the variability in the emissions themselves, with the magnitude roughly scaled to  
3 the distance from the emission source: the standard deviations are 0.5ppm (7.4%) and 0.2ppm  
4 (7.6%) for the pine and chestnut tree, respectively. In comparison, the standard deviation of  
5 the average daily emissions rate over the six years with observations is 7.9% of the mean for  
6 the Vector plant and 21% for the Ballance plant, with a standard deviation of 8.1% for the  
7 combined total (over the full eight years between 2004 and 2011, the standard deviations are  
8 7.3% and 18% of the 8-year mean for Vector and Ballance emissions, respectively, and the  
9 variation in the combined emissions is 7.7%). This is on the same order of magnitude of the  
10 variability due to transport at the pine tree but only about half the amount at the chestnut tree.

11 Looking at all of the factors together (Table 1), variations in emissions and transport  
12 contribute about equally to total variation at the pine tree. At the chestnut tree, transport  
13 makes up a larger proportion of the total, which likely reflects the greater variability in  
14 meteorology in that particular direction. The variability in emissions somewhat counter-  
15 balances the variability in transport, particularly at the chestnut tree, where the standard  
16 deviation with both variable emissions and meteorology (0.3ppm / 15%) is lower than that  
17 with constant emissions (0.5ppm / 19%). This is most likely coincidental to the particular  
18 years of observations, as there is no correlation between variations in emissions and variations  
19 in transport (not shown). Meteorological variation happens to be lowest in the south, where  
20 the trees are located, even though the largest signal occurs to the east (Table 2 and Fig. 5). In  
21 this respect, the trees are fortuitously located for our study. This underscores the benefit of  
22 analysing transport variability at a particular location before collecting observations, as the  
23 quality of results can be greatly influenced by meteorological patterns.

### 24 3.3 Detection limits

25 Given the amount of interannual variation in meteorology and transport that we can infer from  
26 the model and typical measurement uncertainty of 1ppm, what is the minimum change in  
27 emissions that it is possible to detect in a tree ring sample taken at Kapuni, representing an  
28 integrated average of CO<sub>2</sub>ff over a year or more? We use a student t-test to quantify the  
29 minimum amount of change in observations required (relative to the long-term average or  
30 reference period) that would allow us to conclude that there has been a change in emissions.  
31 The t-test calculates the minimum difference between the long-term mean and a new annual  
32 tree ring sample (or samples) that would be statistically significant above scatter or noise from



1 other factors. We make the assumption that our observations and simulated mole fractions are  
2 normally distributed. The results of the 2-sided test (representing change in either direction) at  
3 a 95% confidence level are given in Table 3 for “future” samples representing one, two and  
4 four years of integrated average CO<sub>2</sub>ff. All percentages are relative to the long-term mean  
5 over six years, our reference period for this study. We assume that the standard deviation in  
6 future samples due to interannual variability in meteorology is the same as the standard  
7 deviation over the reference period.

8 Using the modelled means and standard deviations from the constant emissions simulation of  
9 tree ring CO<sub>2</sub>ff and measurement uncertainty of 1.0ppm, the detection limits represent the  
10 minimum observed change that would indicate a driver of variability other than transport or  
11 measurement uncertainty, in this case CO<sub>2</sub>ff emissions. With a new observation representing  
12 one year (i.e. one tree ring), the difference between the long-term mean and the new sample  
13 would need to be more than 42% at the pine tree and 115% at the chestnut tree to have high  
14 confidence that the sample shows a change in emissions, rather than just natural variability or  
15 uncertainty. If we have four new annual observations at the new emission rate, the difference  
16 reduces by half to 22% and 62%, respectively. These detection thresholds are well above the  
17 reported annual changes in emission rates between 2004 and 2011 (Fig. 4). At the distance  
18 and location of the chestnut tree (~1km), it seems likely that the signal is too small and  
19 variable to be practical for detecting emission changes for a point source with emissions of  
20 this magnitude.

21 If we relax the condition to one sigma (or a 68% confidence level), would we be able to detect  
22 the largest change in emissions reported at the Vector Plant between 2004 and 2011? The  
23 student t-test at 68% confidence level gives corresponding detection limits listed in Table 3.  
24 For a one-year observation from the pine tree, this is 18%. The largest change in emissions in  
25 any single year at the Vector plant is in 2007, with a decline of 14% relative to the long-term  
26 mean, still below the detection limit. Indeed, looking at the results in Fig. 4, there is no  
27 significant decline at the chestnut tree in 2007; there is a small decline in CO<sub>2</sub>ff at the pine  
28 tree but it is too small to conclude that emissions have changed. If we were able to achieve a  
29 reduction in measurement uncertainty to 0.5ppm, however, the threshold for detection at the  
30 pine tree becomes an 11% change in emissions, and we would expect to be able to observe a  
31 14% decline in emissions. In this case, the small decline in CO<sub>2</sub>ff at the pine tree in 2007  
32 would be significant.





1 Would we be able to detect this change at a different location (in direction and/or distance)  
2 around the Kapuni plant? Our hypothetical concentration sensors 300m and 600m from the  
3 source (Table 2) indicate that with a single one-year CO<sub>2</sub>ff observation, only a change in  
4 emissions of at least 36% would be detectable at 95% confidence, a much larger change than  
5 occurs in our observational dataset. The location of the pine tree (at 400m southeast of the  
6 plant) appears to provide as good a detection capability as any of our hypothetical sensors.  
7 However, if we have four years of observations (and the change in emissions was sustained  
8 over that time period) located either to the east or the south of the plant at a distance of 300m,  
9 we would be able to detect a change of 10% or more at the one-sigma confidence level.  
10 Changes of 20% or more would be detectable at these same locations with one year of  
11 observations, or alternately, four years of observations if we require high confidence.

12 This analysis uses the actual meteorology only to determine the interannual variability in  
13 CO<sub>2</sub>ff that we might expect due to meteorological variations. If we also know the  
14 meteorology in the year or years of the new observations, we can quantify the change in  
15 emissions by modelling transport at constant emissions. For example, attributing 15% of the  
16 one-year variation at the pine tree to the combined factors of transport and measurement  
17 uncertainty (Table 1) and assuming that the rest of the variation is due to emissions, this  
18 translates to a change in emissions of 27% over the one year. In this manner it is possible to  
19 get a more precise estimate of the long-term changes in emissions.

20 Additionally, if we have multiple measurements over the same period at different locations  
21 around the point source, measurement uncertainty reduces proportionally by  $1/\sqrt{n}$ , where  $n$  is  
22 the number of independent measurements. The resulting reduction in detection thresholds is  
23 more complex and depends on the long-term mean and variation at each of the observation  
24 locations. One could, for example, use a paired t-test to determine if the change detected in all  
25 of the measurements taken together is significant. This is beyond the scope of the current  
26 study, but the detection thresholds given in Tables 2 and 3, based on a single observation  
27 location, would overestimate the minimum change in emissions that it is possible to observe  
28 with multiple measurements designed to cover the area surrounding the point source.

### 29 **3.4 Applicability to other point sources**

30 The results presented here are specific to the meteorology and point sources at the Kapuni  
31 site, but the methodology can be extended to any point source with suitable trees growing



1 nearby. Ideally, observations would be made as close to the source as possible in the direction  
2 where the signal is strongest and/or most consistent. If measurement uncertainty of 1ppm is to  
3 be relatively unimportant compared to the combined transport and emissions variability of 8%  
4 at the pine tree (i.e. adding measurement uncertainty does not change the total variation in  
5 measured CO<sub>2</sub>ff by more than 1-2%), we require a signal around 20-30ppm, implying a  
6 required emission rate five times that of the Kapuni Vector plant. Alternatively, if we were  
7 able to reduce measurement uncertainty to 0.5ppm (for example, by increased measurement  
8 precision or taking measurements from multiple locations at the site), we would be able to  
9 detect changes with signals at around half the magnitude, and the method could be more  
10 feasible for emission sources the size of the Kapuni Vector Plant. Additionally, if we have  
11 multiple measurements from the same period at various locations surrounding the source,  
12 detection thresholds lower further and we can achieve the same sensitivity with a smaller  
13 point source.

14 Our case study involves point sources that are fairly small on an international scale; for  
15 comparison, the world's largest power plant, in Taiwan, emits about 300,000 gC s<sup>-1</sup> or 9.5  
16 TgC yr<sup>-1</sup> (Ummel, 2012), which is 95 times as much as the Vector plant at Kapuni. There are  
17 approximately 800 power plants worldwide that emit more than 10 times the annual total  
18 CO<sub>2</sub>ff at Kapuni (CARMAv3.0, 2009; Wheeler and Ummel, 2008; Ummel, 2012). The  
19 typical emission rates seen at these larger power plants would produce signals in which  
20 measurement uncertainty is only a small proportion of the total. With annual signals  
21 theoretically 10 times that observed at the Kapuni pine tree and the same amount of  
22 meteorological variation, all other things being equal, the detection threshold for a one-year  
23 measurement at the location of the pine tree would be 19%, or 10% with four years of  
24 measurements. This is a plausible reduction target, and the method would be useful for  
25 verifying emissions changes in such cases.

26

## 27 4 Conclusions

28 We have examined sources of interannual variability in CO<sub>2</sub>ff in samples from tree ring  
29 archives representing integrated averages over one year. We used the atmospheric transport  
30 model WindTrax to separate variability in meteorology and transport from other sources of  
31 variation in our observations. At the location of the pine tree, modelled variation in transport  
32 is 7% of the six-year reference mean. This is about the same magnitude as the variation in



1 emissions that were recorded over the same time period. At the chestnut tree, variation due to  
2 atmospheric transport is larger, at 19% of the mean, and is about twice the magnitude of the  
3 variation in emissions. Taking into account typical measurement uncertainty of 1ppm for  
4 radiocarbon samples, in order to conclude with high confidence that there has been a change  
5 in emissions and not just natural variation in meteorology, we would require an observed  
6 change of 42% from the mean in a new one-year sample from the pine tree. If we take a two-  
7 year or four-year sample average, this reduces to 30% and 22%, respectively. This is well  
8 above the largest single-year change in emissions at the Vector Plant, which is 14%.  
9 However, if we are able to reduce measurement uncertainty by half, to 0.5ppm, or if the point  
10 source doubles in strength, detection thresholds are closer to the actual level of variation in  
11 emissions. If we only require confidence at the one-sigma level, we would in this case be able  
12 to detect a 14% change in a single year.

13 Detection limits are highly dependent on the location of the observations and specific  
14 meteorology of the site. Wind patterns should be carefully considered before deciding where  
15 to take samples in any study, preferably in an area where the signal will be strongest and  
16 where wind patterns will be most consistent through time. A model analysis such as we have  
17 performed can give an idea of the baseline variability in transport and the size of the signal  
18 needed to observe changes in emissions. This makes it theoretically possible to separate the  
19 uncertainty in transport from other sources of uncertainty.

20 In general, this method will be most effective when observations are made in the dominant  
21 wind direction and/or in a direction with consistent winds, close enough to the point source so  
22 that natural variability in meteorological conditions and measurement uncertainty does not  
23 overwhelm the signal from the emissions. The larger the point source (the higher the emission  
24 rate) and the signal from CO<sub>2</sub>ff, the more able integrated averages from plant material will be  
25 to detect changes in emissions. For larger power plants or other point sources of a more  
26 typical size worldwide, detecting changes with this method could be feasible; with signals 10  
27 times or more the size of Kapuni, measurement uncertainty is relatively insignificant, and  
28 sustained changes in emissions on the order of 10% can be detected from a single sampling  
29 location given suitable meteorological conditions and observations.

30

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1   References

- 2   Ackerman, K. V., and Sundquist, E. T.: Comparison of two US power-plant carbon dioxide  
 3   emissions data sets, *Environ. Sci. Technol.*, 42, 5688-5693, 2008.
- 4   Boden, T. A., Marland, G., and Andres, R. J.: Global, Regional, and National Fossil-Fuel CO<sub>2</sub>  
 5   Emissions, Carbon Dioxide Information Analysis Center, Oak Ridge National Laboratory,  
 6   U.S. Department of Energy, Oak Ridge, Tenn., USA, doi 10.3334/CDIAC/00001\_V2015,  
 7   2015.
- 8   Bonifacio, H. F., Maghirang, R. G., Razote, E. B., Trabue, S. L., and Prueger, J. H.:  
 9   Comparison of AERMOD and WindTrax dispersion models in determining PM<sub>10</sub> emission  
 10   rates from a beef cattle feedlot, *J. Air Waste Manage.*, 63, 545-556, 2013.
- 11   Bozhinova, D., Combe, M., Palstra, S. W. L., Meijer, H. A. J., Krol, M. C., and Peters, W.:  
 12   The importance of crop growth modeling to interpret the  $\Delta^{14}\text{CO}_2$  signature of annual plants,  
 13   *Global Biogeochem. Cy.*, 27, 792-803, 2013.
- 14   Brioude, J., Angevine, W. M., Ahmadov, R., Kim, S.-W., Evan, S., McKeen, S. A., Hsie, E.-  
 15   Y., Frost, G. J., Neuman, J. A., Pollack, I. B., Peischl, J., Ryerson, T. B., Holloway, J.,  
 16   Brown, S. S., Nowak, J. B., Roberts, J. M., Wofsy, S. C., Santoni, G. W., Oda, T., and  
 17   Trainer, M.: Top-down estimate of surface flux in the Los Angeles Basin using a mesoscale  
 18   inverse modeling technique: assessing anthropogenic emissions of CO, NO<sub>x</sub> and CO<sub>2</sub> and  
 19   their impacts, *Atmos. Chem. Phys.*, 13, 3661-3677, doi:10.5194/acp-13-3661-2013, 2013.
- 20   CliFlo: NIWA's National Climate Database on the Web: <http://cliflo.niwa.co.nz/> , last access:  
 21   4 November 2014.
- 22   Currie, K. I., Brailsford, G., Nichol, S., Gomez, A., Sparks, R., Lassey, K. R., and Riedel, K.:  
 23   Tropospheric <sup>14</sup>CO<sub>2</sub> at Wellington, New Zealand: the world's longest record,  
 24   *Biogeochemistry*, 104, 5-22, 2011.
- 25   Djuricin, S., Xu, X., and Pataki, D.E.: The radiocarbon composition of tree rings as a tracer of  
 26   local fossil fuel emissions in the Los Angeles basin: 1980-2008, *J. Geophys. Res.-Atmos.*,  
 27   117, D12303, 2012.
- 28   Donders, T. H., Decuyper, M., Beaubien, S. E., van Hoof, T. B., Cherubini, P., and Sass-  
 29   Klaassen, U.: Tree rings as biosensor to detect leakage of subsurface fossil CO<sub>2</sub>. *Int. J.*  
 30   *Greenh. Gas Con.*, 19, 387-395, 2013.



- 1 eGRID: Technical Support Document for the 9th Edition of eGRID with Year 2010 Data
- 2 (Emissions & Generation Resource Integrated Database), Washington, D.C., 2014.
- 3 Flesch, T. K., Wilson, J. D., Harper, L., Crenna, B., and Sharpe, R.: Deducing ground-to-air
- 4 emissions from observed trace gas concentrations: a field trial, *J. Appl. Meteorol.*, 43, 487–
- 5 502, 2004.
- 6 Flesch, T. K., Wilson, J. D., Harper, L., and Crenna, B.: Estimating gas emissions from a farm
- 7 with an inverse-dispersion technique, *Atmos. Envir.*, 39, 4863–4874, 2005.
- 8 Gao, Z., Mauder, M., Desjardins, R. L., Flesch, T. K., and van Haarlem, R. P.: Assessment of
- 9 the backward Lagrangian stochastic dispersion technique for continuous measurements of
- 10 CH<sub>4</sub> emissions, *Agr. Forest Meteorol.*, 149, 1516–1523, 2009.
- 11 Geels, C., Gloor, M., Ciais, P., Bousquet, P., Peylin, P., Vermeulen, A. T., Dargaville, R.,
- 12 Aalto, T., Brandt, J., Christensen, J. H., Frohn, L. M., Haszpra, L., Karstens, U., Rödenbeck,
- 13 C., Ramonet, M., Carboni, G., and Santaguida, R.: Comparing atmospheric transport models
- 14 for future regional inversions over Europe – Part 1: mapping the atmospheric CO<sub>2</sub> signals,
- 15 *Atmos. Chem. Phys.*, 7, 3461–3479, doi:10.5194/acp-7-3461-2007, 2007.
- 16 Gerbig, C., Körner, S., and Lin, J. C.: Vertical mixing in atmospheric tracer transport models:
- 17 error characterization and propagation, *Atmos. Chem. Phys.*, 8, 591–602, doi:10.5194/acp-8-
- 18 591-2008, 2008.
- 19 Gurney, K. R.: Beyond Hammers and Nails: Mitigating and Verifying Greenhouse Gas
- 20 Emissions, *Eos Trans. AGU*, 94, 199, 2013.
- 21 Gurney, K. R., Mendoza, D. L., Zhou, Y., Fischer, M. L., Miller, C. C., Geethakumar, S., and
- 22 de la Rue du Can, S.: High resolution fossil fuel combustion CO<sub>2</sub> emission fluxes for the
- 23 United States, *Environ. Sci. Technol.*, 43, 5535–5541, 2009.
- 24 Gurney, K. R., Razlivanov, I., Song, Y., Zhou, Y., Benes, B., and Abdul-Massih, M.:
- 25 Quantification of fossil fuel CO<sub>2</sub> emissions on the building/street scale for a large US City,
- 26 *Environ. Sci. Technol.*, 46, 12194–12202, 2012.
- 27 IPCC Core Writing Team, Pachauri, R. K. and Meyer, L. A. (Eds.): *Climate Change 2014:*
- 28 *Synthesis Report. Contribution of Working Groups I, II and III to the Fifth Assessment*
- 29 *Report of the Intergovernmental Panel on Climate Change*, IPCC, Geneva, Switzerland, 151
- 30 pp., 2014.



- 1 Karlen, I., Olsson, I. U., Kllburg, P., and Kilici, S.: Absolute determination of the activity of  
2 two  $^{14}\text{C}$  dating standards, *Arkiv Geofysik*, 4, 465–471, 1968.
- 3 Koehn, A. C., Leytem, A. B., and Bjorneberg, D. L.: Comparison of atmospheric stability  
4 methods for calculating ammonia and methane emission rates with WindTrax, *Transactions of*  
5 *the American Society of Agricultural and Biological Engineers*, 56, 763-768, 2013.
- 6 Kretschmer, R., Gerbig, C., Karstens, U., and Koch, F.-T.: Error characterization of  $\text{CO}_2$   
7 vertical mixing in the atmospheric transport model WRF-VPRM, *Atmos. Chem. Phys.*, 12,  
8 2441-2458, doi:10.5194/acp-12-2441-2012, 2012.
- 9 Laubach, J.: Testing of a Lagrangian model of dispersion in the surface layer with cattle  
10 methane emissions, *Agr. Forest Meteorol.*, 150, 1428-1442, 2010.
- 11 Laubach, J., and Kelliher, F. M.: Methane emissions from dairy cows: Comparing open-path  
12 laser measurements to profile-based techniques, *Agr. Forest Meteorol.*, 135, 340-345, 2005.
- 13 Leuning, R., Etheridge, D., Luhar, A., and Dunse, B.: Atmospheric monitoring and  
14 verification technologies for  $\text{CO}_2$  geosequestration, *Int. J. Greenh. Gas Con.*, 2, 401-414,  
15 2008.
- 16 Levin, I., and Rödenbeck, C.: Can the envisaged reductions of fossil fuel  $\text{CO}_2$  emissions be  
17 detected by atmospheric observations?, *Naturwissenschaften*, 95, 203-208,  
18 doi:10.1007/s00114-007-0313-4, 2007.
- 19 Levin, I., Kromer, B., Schmidt, M., and Sartorius, H.: A novel approach for independent  
20 budgeting of fossil fuel  $\text{CO}_2$  over Europe by  $^{14}\text{CO}_2$  observations, *Geophys. Res. Lett.*, 30,  
21 2194, doi:10.1029/2003GL018477, 2003.
- 22 Lin, J. C., and Gerbig, C.: Accounting for the effect of transport errors on tracer inversions,  
23 *Geophys. Res. Lett.*, 32, L01802, doi:10.1029/2004GL021127, 2005.
- 24 Lindenmaier, R., Dubey, M. K., Henderson, B. G., Butterfield, Z.T., Herman, J. R., Rahn, T.,  
25 and Lee, S. H.: Multiscale observations of  $\text{CO}_2$ ,  $^{13}\text{CO}_2$ , and pollutants at Four Corners for  
26 emission verification and attribution, *P. Natl. Acad. Sci. USA*, 111, 8386-8391, 2014.
- 27 Liu, J., Fung, I., Kalnay, E., and Kang, J.-S.:  $\text{CO}_2$  transport uncertainties from the  
28 uncertainties in meteorological fields, *Geophys. Res. Lett.*, 38, L12808,  
29 doi:10.1029/2011GL047213, 2011.



- 1 Loh, Z., Leuning, R., Zegelin, S., Etheridge, D., Bai, M., Naylor, T., and Griffith, D.: Testing  
2 Lagrangian atmospheric dispersion modelling to monitor CO<sub>2</sub> and CH<sub>4</sub> leakage from  
3 geosequestration, *Atmos. Environ.*, 43, 2602–2611, doi:10.1016/j.atmosenv.2009.01.053,  
4 2009.
- 5 Lund, U., and Agostinelli, C.: circular: Circular Statistics, R package version 0.4-7, available  
6 at: <http://cran.r-project.org/web/packages/circular/index.html> (last access: 29 Jan 2015), 2013.
- 7 McBain, M. C., and Desjardins, R. L.: The evaluation of a backward Lagrangian stochastic  
8 (bLS) model to estimate greenhouse gas emissions from agricultural sources using a synthetic  
9 tracer source, *Agr. Forest Meteorol.*, 135, 61-72, 2005.
- 10 McKain, K., Wofsy, S. C., Nehrkorn, T., Eluszkiewicz, J., Ehleringer, J. R., and Stephens, B.  
11 B.: Assessment of ground-based atmospheric observations for verification of greenhouse gas  
12 emissions from an urban region, *P. Natl. Acad. Sci. USA*, 109, 8423-8428, 2012.
- 13 Meijer, H. A. J., Smid, H. M., Perez, E., and Keizer, M. G.: Isotopic characterisation of  
14 anthropogenic CO<sub>2</sub> emissions using isotopic and radiocarbon analysis, *Phys. Chem. Earth*, 21,  
15 483-487, 1996.
- 16 Miles, N. L., Richardson, S. J., Davis, K. J., Lauvaux, T., Andrews, A. E., West, T., Bandaru,  
17 V., and Crosson, E. R.: Large amplitude spatial and temporal gradients in atmospheric  
18 boundary layer CO<sub>2</sub> mole fractions detected with a tower-based network in the US Upper  
19 Midwest, *J. Geophys. Res.*, 117, G01019, 2012.
- 20 National Research Council: Verifying Greenhouse Gas Emissions: Methods to Support  
21 International Climate Agreements, The National Academies Press, Washington, D.C., USA,  
22 124 pp., 2010.
- 23 Nisbet, E., and Weiss, R.: Top-down versus bottom-up, *Science*, 328, 1241-1243, 2010.
- 24 Norris, M. W.: Reconstruction of historic fossil CO<sub>2</sub> emissions using radiocarbon  
25 measurements from tree rings, M.S. thesis, Victoria University of Wellington, Wellington,  
26 New Zealand, 155 pp., 2015.
- 27 NZMED: New Zealand's Energy Outlook 2010, New Zealand Ministry of Economic  
28 Development, Wellington, New Zealand, 12 pp., 2010.
- 29 Prather, M. J., Zhu, X., Strahan, S. E., Steenrod, S. D., and Rodriguez, J. M.: Quantifying  
30 errors in trace species transport modeling, *P. Natl. Acad. Sci. USA*, 105, 19617-19621, 2008.





- 1 R Core Team: R: A language and environment for statistical computing, R Foundation for  
2 Statistical Computing, Vienna, Austria, available at: <http://www.R-project.org/> (last access:  
3 21 August 2014), 2013.
- 4 Rakowski, A. Z., Nadeau, M. J., Nakamura, T., Pazdur, A., Pawełczyk, S., and Piotrowska,  
5 N.: Radiocarbon method in environmental monitoring of CO<sub>2</sub> emission, Nucl. Instrum. Meth.  
6 B, 294, 503-507, 2013.
- 7 Rhoades, M. B., Parker, D. B., Cole, N. A., Todd, R. W., Caraway, E. A., Auvermann, B. W.,  
8 Topliff, D.R. and Schuster, G. L.: Continuous ammonia emission measurements from a  
9 commercial beef feedyard in Texas, Transactions of the American Society of Agricultural and  
10 Biological Engineers, 53, 1823-1831, 2010.
- 11 Serre, C., Santikarn, M., Stelmakh, K., Eden, A., Frerk, M., Kachi, A., Unger, C., Wilkenig,  
12 K., and Haug, C. (Eds): Emissions Trading Worldwide: International Carbon Action  
13 Partnership (ICAP) Status Report 2015, International Carbon Action Partnership, Berlin,  
14 Germany, 71 pp., 2015.
- 15 Stephens, B. B., Gurney, K. R., Tans, P. P., Sweeney, C., Peters, W., Bruhwiler, L., Ciais, P.,  
16 Ramonet, M., Bousquet, P., Nakazawa, T., Aoki, S., Machida, T., Inoue, G., Vinnichenko, N.,  
17 Lloyd, J., Jordan, A., Heimann, M., Shibistova, O., Langenfelds, R. L., Steele, L. P., Francey,  
18 R. J., Denning, A. S.: Weak Northern and Strong Tropical Land Carbon Uptake from Vertical  
19 Profiles of Atmospheric CO<sub>2</sub>, Science, 316, 1732-1735, 2007.
- 20 Suess, H. E.: Radiocarbon concentration in modern wood, Science, 122, 415, 1955.
- 21 Tans, P. P., and Wallace, D. W.: Carbon cycle research after Kyoto, Tellus B, 51, 562-571,  
22 1999.
- 23 Tans, P. P., De Jong, A. F. M., and Mook, W. G.: Natural atmospheric <sup>14</sup>C variation and the  
24 Suess effect, Nature, 280, 826-828, 1979.
- 25 Taranaki Regional Council: Ballance Agri-Nutrients (Kapuni) Ltd Monitoring Programme  
26 Biennial Report 2010-2012: Technical Report 2012-91, Taranaki Regional Council, Stratford,  
27 New Zealand, 63 pp., June 2013.
- 28 Turnbull, J. C., Miller, J. B., Lehman, S. J., Tans, P. P., Sparks, R. J., and Southon, J. R.:  
29 Comparison of <sup>14</sup>CO<sub>2</sub>, CO and SF<sub>6</sub> as tracers for determination of recently added fossil fuel



- 1 CO<sub>2</sub> in the atmosphere and implications for biological CO<sub>2</sub> exchange, *Geophys. Res. Lett.*,  
2 33, L01817, doi:10.1029/2005GL024213, 2006.
- 3 Turnbull, J. C., Keller, E. D., Baisden, W. T., Brailsford, G., Bromley, T., Norris, M. W., and  
4 Zondervan, A.: Atmospheric measurement of point source fossil fuel CO<sub>2</sub> emissions, *Atmos.*  
5 *Chem. Phys.*, 14, 50001-5014, doi: 10.5194/acp-14-5001-2014, 2014.
- 6 Ummel, K.: CARMA revisited: an updated database of carbon dioxide emissions from power  
7 plants worldwide, Center for Global Development Working Paper 304, Center for Global  
8 Development, Washington, D.C., USA, 2012.
- 9 United Nations Framework Convention on Climate Change (UNFCCC): Kyoto Protocol,  
10 available at: [http://unfccc.int/kyoto\\_protocol/items/2830.php](http://unfccc.int/kyoto_protocol/items/2830.php), and Intended Nationally  
11 Determined Contributions, available at: [http://unfccc.int/focus/indc\\_portal/items/8766.php](http://unfccc.int/focus/indc_portal/items/8766.php)  
12 (last access: 29 May 2015), 2015.
- 13 U.S. Environmental Protection Agency: Plain English Guide to the Part 75 Rule, U.S.  
14 Environmental Protection Agency, Washington, D.C., USA, 118 pp., 2005.
- 15 Wheeler, D., and Ummel, K.: Calculating CARMA: global estimation of CO<sub>2</sub> emissions from  
16 the power sector, Center for Global Development Working Paper 145, Center for Global  
17 Development, Washington, D.C., USA, 2008.
- 18 Wilson, J. D., and Sawford, B. L.: Review of Lagrangian stochastic models for trajectories in  
19 the turbulent atmosphere, *Bound.-Lay. Meteorol.*, 78, 191–210, 1996.
- 20 Wilson, J. D., Flesch, T. K., and Crenna, B. P.: Estimating Surface-Air Gas Fluxes by Inverse  
21 Dispersion Using a Backward Lagrangian Stochastic Trajectory Model, in: *Lagrangian*  
22 *Modeling of the Atmosphere*, Lin, J., Brunner, D., Gerbig, C., Stohl, A., Luhar, A., and  
23 Webley, P. (Eds.), American Geophysical Union, Washington, D. C., USA, doi:  
24 10.1029/2012GM001269, 2012.
- 25



1 Table 1. Observed and modelled CO<sub>2</sub>ff means and standard deviations. All means and  
 2 standard deviations (SD) include six years (2008 and 2010 are omitted because there are no  
 3 observations available for these years). Measurement uncertainty (MU) of 1.0ppm is  
 4 explicitly added to the modelled results in the far right column. Observations implicitly  
 5 include this uncertainty.

Observation or simulation (2004-2011)	Mean (ppm)	SD (% of mean)	SD + 1.0ppm MU (% of mean)
<b>Pine</b>			
Observed	5.4		0.8 (14%)
Modelled CO <sub>2</sub> ff <sub>free</sub> : variable meteorology, variable emissions	6.1	0.5 (7.8%)	1.1 (18%)
Modelled CO <sub>2</sub> ff : variable meteorology, constant emissions	7.4	0.5 (6.6%)	1.1 (15%)
Modelled CO <sub>2</sub> ff : constant meteorology, variable emissions	7.3	0.5 (7.4%)	1.1 (15%)
<b>Chestnut</b>			
Observed	2.1		1.1 (51%)
Modelled CO <sub>2</sub> ff <sub>free</sub> : variable meteorology, variable emissions	2.2	0.3 (15%)	1.0 (47%)
Modelled CO <sub>2</sub> ff : variable meteorology, constant emissions	2.7	0.5 (19%)	1.1 (41%)
Modelled CO <sub>2</sub> ff : constant meteorology, variable emissions	2.3	0.2 (7.6%)	1.0 (43%)

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1 Table 2. Eight-year modelled mean CO<sub>2</sub>ff and standard deviation (SD) of eight hypothetical  
 2 sensors for eight years of constant emissions simulation and detection limits at the two-sigma  
 3 (95%) and one-sigma (68%) confidence level (CL) for samples representing an average of  
 4 one, two, or four years.

Model Sensor	Mean (ppm)	SD (% of mean)	SD + 1ppm MU (% of mean)	% change detectable (95% CL)			% change detectable (68% CL)		
				1 yr	2 yr	4 yr	1 yr	2 yr	4 yr
North 300m	12.2	2.4 (20%)	2.6 (21%)	53%	38%	29%	24%	18%	13%
North 600m	4.6	0.8 (18%)	1.3 (29%)	72%	52%	39%	33%	24%	18%
East 300m	22.8	3.2 (14%)	3.3 (15%)	37%	27%	20%	17%	12%	9.4%
East 600m	9.0	1.3 (14%)	1.6 (18%)	45%	33%	24%	20%	15%	12%
South 300m	11.7	1.3 (11%)	1.7 (14%)	36%	26%	20%	16%	12%	9.2%
South 600m	4.7	0.5 (10%)	1.1 (24%)	60%	43%	33%	27%	20%	15%
West 300m	1.6	0.8 (50%)	1.3 (81%)	204%	148%	111%	92%	68%	52%
West 600m	0.34	0.16 (50%)	1.0 (300%)	744%	540%	405%	337%	250%	190%

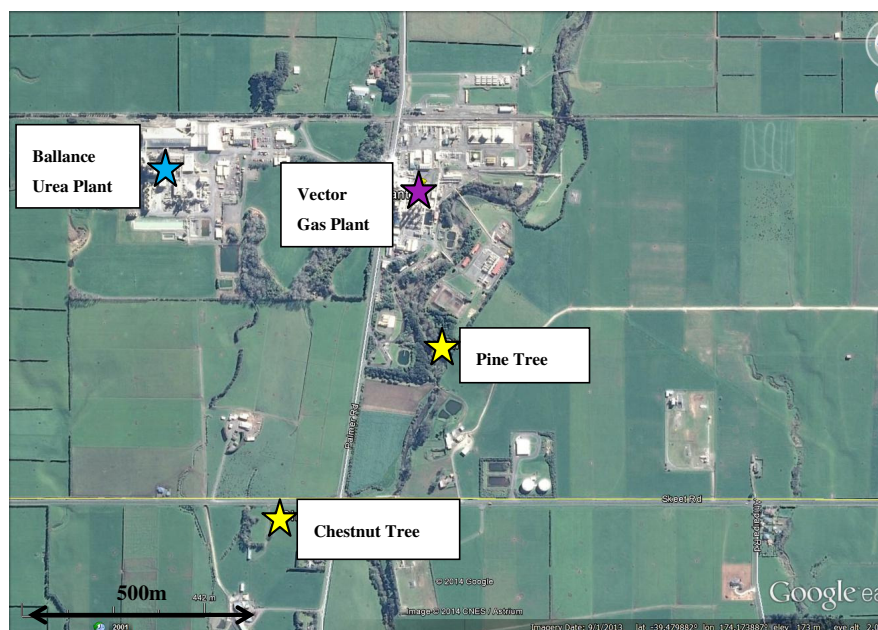
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1 Table 3. Detection limits for samples at trees, calculated with modelled CO<sub>2</sub>ff at constant  
 2 emissions and six years of observations in reference period. Limits are given at the two-sigma  
 3 (95%) and one-sigma (68%) confidence level (CL) for samples representing an average of  
 4 one, two, or four years. Measurement uncertainty (MU) of 1.0ppm or 0.5ppm is added in  
 5 quadrature to the standard deviation of modelled CO<sub>2</sub>ff before limits are calculated.

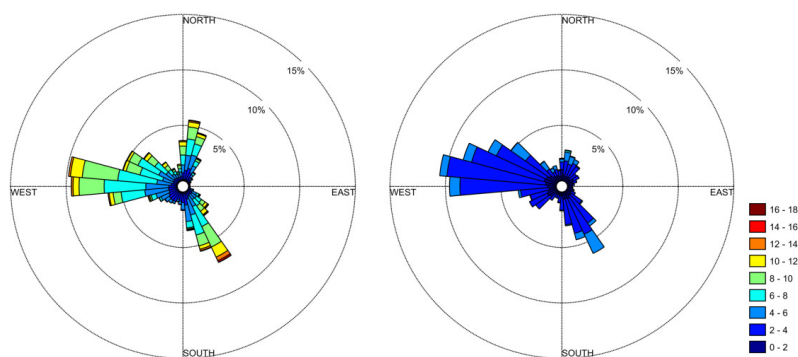
Modelled CO <sub>2</sub> ff: variable meteorology constant emissions	% change detectable (95% CL)			% change detectable (68% CL)		
	1 yr	2 yr	4yr	1 yr	2 yr	4yr
<b>Pine</b>						
Modelled CO <sub>2</sub> ff + 1.0 MU	42%	30%	22%	18%	13%	10%
Modelled CO <sub>2</sub> ff + 0.5 MU	27%	19%	14%	11%	8.5%	6.5%
<b>Chestnut</b>						
Modelled CO <sub>2</sub> ff + 1.0 MU	115%	83%	62%	92%	68%	52%
Modelled CO <sub>2</sub> ff + 0.5 MU	89%	64%	48%	38%	28%	22%

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Figure 1. Aerial view of Kapuni area, with the sampled pine and chestnut trees and Vector Gas Treatment Plant and Ballance Agri-Nutrient Urea Plant labelled.

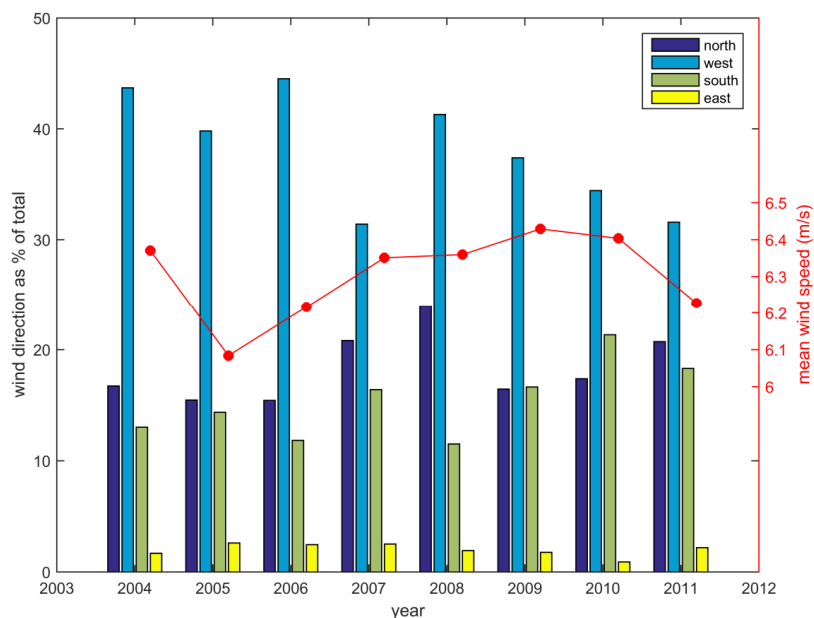


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3 Figure 2. Wind roses during the growing season at Hawera 2004-2011 (left) and Kapuni 2013  
 4 (right), daylight hours only (8:00am – 6:00pm).

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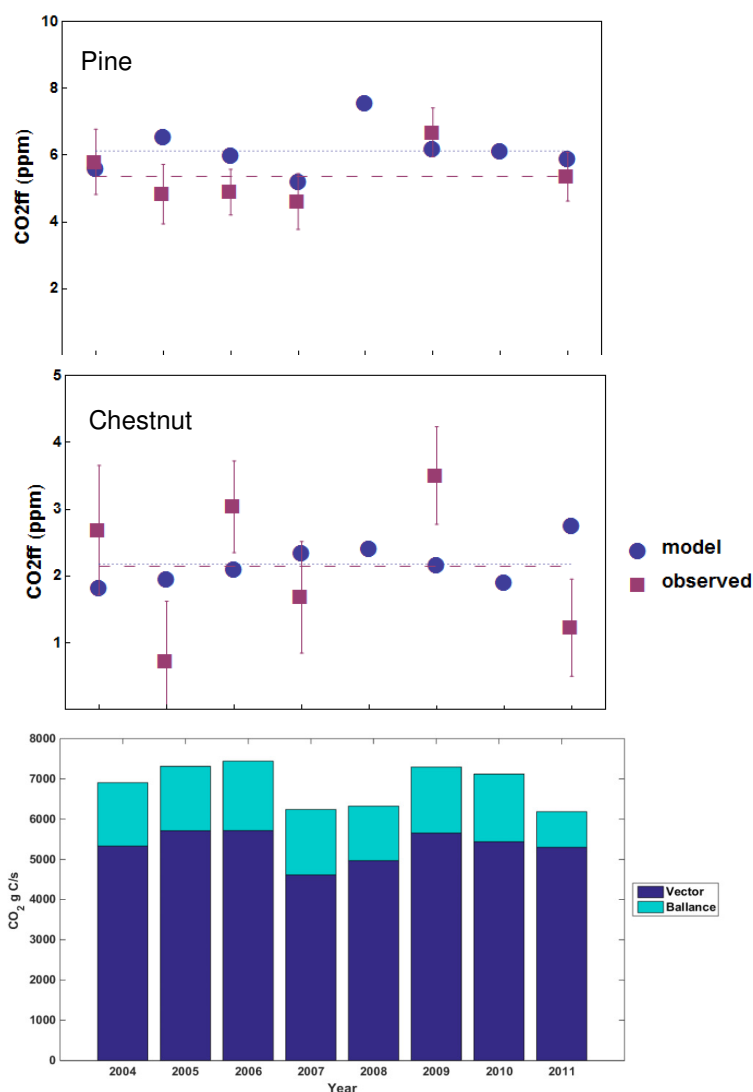
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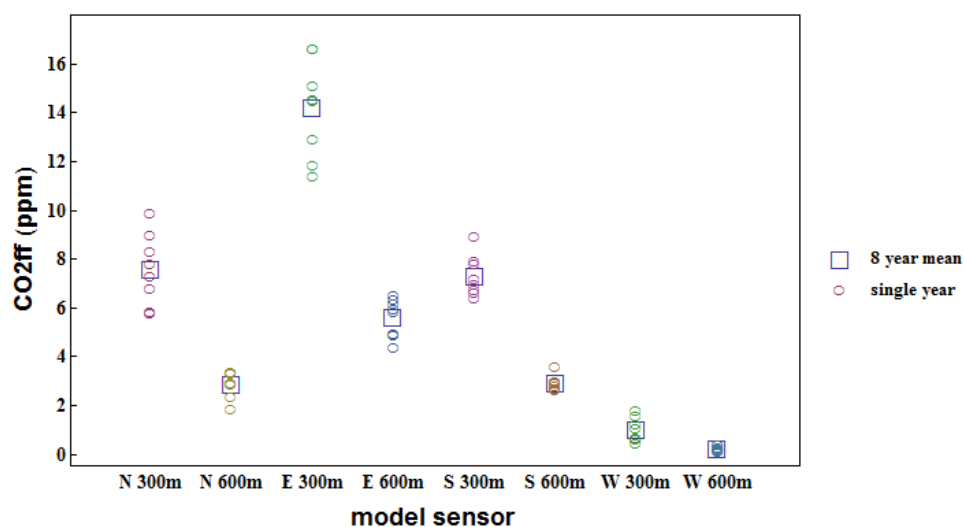
3 Figure 3. Percentage of wind in each of four directions (left axis) and mean wind speed (right  
 4 axis) by growing year between 2004 and 2011 (daylight hours only, 8:00am – 6:00pm).  
 5 Directions are defined by +/- 30 degrees due north, west, south, and east (i.e. west is defined  
 6 as wind from 240° to 300°). Note that this does not comprise the complete 360° circle so  
 7 percentages do not add to 100.

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 3 Figure 4. Pine tree (top) and chestnut tree (middle) modelled  $\text{CO}_2\text{ff}_{\text{tree}}$  vs. tree ring observed  
 4  $\text{CO}_2\text{ff}$ . Dashed lines show modelled and observed six-year means. Bottom panel shows the  
 5 average emissions rate for Vector and Ballance in each year.  
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3 Figure 5. Constant emissions, variable meteorology simulation results for hypothetical  
 4 sensors: CO<sub>2</sub>ff mole fraction averaged over all eight years of simulation (squares) and  
 5 individual annual averages (circles). Sensors are labelled by direction (N, E, S or W) and  
 6 distance (300m or 600m) from the source.