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Evaluation and application of multi-decadal visibility data for trend analysis of atmospheric haze

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There are few multi-decadal observations of atmospheric aerosols worldwide. This study applies global hourly visibility (Vis) observations at more than 3000 stations to investigate historical trends in atmospheric haze over 1945–1996 for the US, and over 1973–2013 for Europe and Eastern Asia. A comprehensive data screening and processing framework is developed and applied to minimize uncertainties and construct monthly statistics of inverse visibility (1/Vis). This data processing includes removal of relatively clean cases with high uncertainty, and change point detection to identify and separate methodological discontinuities such as the introduction of instrumentation. Although the relation between 1/Vis and $b_{\rm ext}$ varies across different stations, spatially coherent trends of the screened 1/Vis exhibit consistency with the temporal evolution of collocated aerosol measurements, including the atmospheric extinction coefficient (b_{ext}) trend of -2.4 % yr⁻¹ (95 % CI: -3.7, -1.1 % yr⁻¹) vs. 1/Vis trend of $-1.6\% \text{ yr}^{-1}$ (95% CI: -2.4, $-0.8\% \text{ yr}^{-1}$) over the US for 1989–1996, and the fine aerosol mass (PM_{2.5}) trend of $-5.8 \% \text{ yr}^{-1}$ (95 % CI: -7.8, $-4.2 \% \text{ yr}^{-1}$) vs. 1/Vis trend of -3.4 % yr⁻¹ (95 % CI: -4.4, -2.4 % yr⁻¹) over Europe for 2006–2013. Regional 1/Vis and EDGAR sulfur dioxide (SO₂) emissions are significantly correlated over the eastern US for 1970–1995 (r = 0.73), over Europe for 1973–2008 ($r \sim 0.9$) and over China for 1973–2008 ($r \sim 0.9$). Consistent "reversal points" from increasing to decreasing in SO₂ emission data are also captured by the regional 1/Vis time series (e.g. late 1970s for the eastern US, early 1980s for Western Europe, late 1980s for Eastern Europe, and mid 2000s for China). The consistency of inferred 1/Vis trends with other in situ measurements and emission data demonstrates promise in applying these reconstructed 1/Vis data for historical air quality studies.

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Atmospheric aerosols have broad implications for air quality and climate change. The Global Burden of Disease (GBD) assessment attributed ambient exposure to aerosol particles with an aerodynamic diameter below 2.5 µm (PM_{2.5}) as the sixth largest overall risk factor for premature mortality with 3.2 million premature deaths per year (Lim et al., 2012). Aerosols are also considered as the most uncertain component for global radiative forcing (IPCC, 2013). Aerosols are formed from a variety of emission sources and chemical processes with a short tropospheric lifetime against different removal mechanisms, yielding a highly variable spatiotemporal distribution that is not well understood (Fuzzi et al., 2015). Information on long-term aerosol temporal evolution is crucially needed across a range of disciplines. Historical PM_{2.5} exposure and its trends are needed to understand changes in Global Burden of Disease (Brauer et al., 2012), and to guide mitigation actions (Apte et al., 2015; Wong et al., 2004). Observations are needed to evaluate historical emission inventories that are crucial to accurately represent the changes in aerosol sources and its consequent feedbacks on climate (Lu et al., 2011; S. J. Smith et al., 2011a; Xu et al., 2013). Aerosol trend analysis is also fundamental to assessing radiative forcing, evaluating model processes, and projecting future changes (Chin et al., 2014; Leibensperger et al., 2012; Li et al., 2014). Various studies have been carried out to investigate aerosol trends using in situ measurements (Collaud Coen et al., 2013; Hand et al., 2012a; Murphy et al., 2011), satellite/ground remote sensing (Hsu et al., 2012; Li et al., 2014; Zhang and Reid, 2010), and analysis of measurements with models (Boys et al., 2014; Chin et al., 2014; Pozzer et al., 2015; Turnock et al., 2015). However these studies are mostly limited to the recent 2 decades, since few satellite or in situ aerosol observations exist over land prior to the 1990s. Long-term observations of aerosols at the global scale are needed to place current knowledge of their spatial distribution and temporal evolution in a historical context for all these applications.

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Visibility observations offer an alternative information source to investigate historical aerosol trends. Horizontal visibility (Vis) from worldwide meteorological stations and airports is mainly determined by the optical extinction (b_{out}) of the atmospheric boundary layer, and has been recognized as a proxy of the atmospheric aerosol burden/loading (Husar et al., 2000). Historical Vis data from more than 3000 stations have been applied to characterize decadal trends in global aerosol optical depth (AOD) from 1973 to 2007 (Wang et al., 2009). Regional trend studies of Vis were also conducted for populated areas e.g. the US (Husar et al., 1981; Schichtel et al., 2001), Europe (Vautard et al., 2009) and China (Che et al., 2007; Chen and Wang, 2015; Lin et al., 2014; Wu et al., 2012, 2014), and the inferred trends were usually attributed to changes in anthropogenic emission. Another study employing Vis over desert regions (Mahowald et al., 2007) found an association of Vis with meteorology factors such as drought index (based on precipitation and temperature) and surface wind speeds. Trends in Vis data interpreted with other datasets also supported studies of several aerosol related climate trends such as the western Pacific subtropical high (Qu et al., 2013) and precipitation (Rosenfeld et al., 2007; Stjern et al., 2011).

Multi-decadal Vis data might contain possible variation or even reversal in haze trends as expected from historical emission and surface solar radiation (SSR) data (Lu et al., 2010; Stern, 2006; Streets et al., 2006; Wild et al., 2005). It is of particular interest how these changes would associate with the trends of air quality, and would be captured by the Vis data. Meanwhile detailed variation in global Vis trends are rarely reported in these previous studies. On the other hand, Vis data are inherently uncertain because most Vis are recorded through human observations with variable protocols. For example, an increase in inverse visibility (1/Vis) has been reported over the US during 1993-2010 (Wang et al., 2012) that is opposite in sign with the significant decline (> 10 % decade⁻¹) of observed PM_{2.5}, sulfate and b_{ext} (Attwood et al., 2014; Hand et al., 2012a, 2014; US EPA, 2012), and raises questions about the quality of Vis observations.

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Relationship between Vis and b_{ext}

Visibility is a measure of the transparency of the atmosphere, and is defined as the greatest distance at which a black object can be recognized against the horizon sky (WMO, 2008). The visibility of a particular object (i.e. visibility marker) is determined by the contrast C between the radiation intensity / of the background b and of the object o reaching an observer at distance x from the object:

$$C(x) = \frac{I_{b}(x) - I_{o}(x)}{I_{b}(x)}$$
 (1)

Under assumptions of a plane-parallel atmosphere and homogeneous background intensity (i.e. constant sky brightness), C exhibits an exponential decay based on Beer's law.

$$C(x) = C_0 \exp(-b_{\text{ext}}x)$$
 (2)

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$$C_{\text{crit}} = C_0 \exp(-b_{\text{ext}} \text{Vis}) \tag{3}$$

Rearranging to solve for b_{ext} yields

$$b_{\rm ext} = \frac{K}{\rm Vis} \tag{4}$$

where $K = -\ln \frac{C_{crit}}{C_0}$. This is the Koschmieder equation (Griffing, 1980), representing a linear relationship between 1/Vis and $b_{\rm ext}$. The slope K of this relationship is mainly determined by two factors: the inherent contrast at the object's position C_0 and the critical contrast of the observer's eye C_{crit} . This equation is only valid for a plane-parallel and homogeneous atmosphere. For situations with high gradients of $b_{\rm ext}$ (e.g. smoke plumes), this could readily break down. Even for ideal conditions, this relationship could vary due to the variation of C_0 (change of markers or observing conditions) and/or C_{crit} (change of observer or protocol). It is sometimes assumed that the object is perfectly black ($C_0 = 1$) so that K is only determined by C_{crit} . Nevertheless, K still varies from 1.5 to 3.9 (e.g. Husar and Wilson, 1993; Schichtel et al., 2001; Wang et al., 2009) because of different C_{crit} values or different observing conditions. Below we similarly find that even where 1/Vis is highly correlated with $b_{\rm ext}$ data, K still varies significantly for different stations.

Data and processing

Visibility data

We begin with raw Vis data from synoptic observations in the Integrated Surface Database (ISD, https://catalog.data.gov/dataset/integrated-surface-global-33794

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hourly-data) archived at the National Climatic Data Center (NCDC). ISD data are generated through merging hundreds of data sources (A. Smith et al., 2011). The data from different networks have different report frequencies (e.g. hourly, 3 hourly, 6 hourly, etc.). We reject the daily averaged data called "global summary of the day" (GSOD) since an 5 arithmetic mean could bias the daily and monthly statistics because of threshold and discreteness issues, as discussed in Sect. 3.1.2. Figure 1 shows an overview of our processing flow designed to comprehensively identify and exclude questionable data. Each step is described below.

Conventional screening 3.1.1

We begin with "conventional screening" using algorithms adapted from prior studies. We eliminate effects on Vis of weather conditions such as fog, precipitation, low cloud and high relative humidity (RH > 90%) following the description in Husar et al. (2000). A sensitivity test that limited conditions to RH < 80 % reduced data density but yielded similar trend results without changing the main findings in this study. Potential human errors are reduced by statistical checks of daily spikes and non-repeating values following Lin et al. (2014). Duplicate stations with different names are combined, and stations lacking geolocation information are removed following Willett et al. (2013).

3.1.2 Threshold filtering

We develop a filter to address spatial and temporal variation in the threshold of reported Vis. The "threshold" is the maximum reported Vis at a station that often depends on the furthest employed Vis marker. Vis above this threshold is not resolved. Thus the threshold acts as an artificial detection limit. The ability of Vis data to capture the variation of $b_{\rm ext}$ is weak when the air is clean and/or the adopted threshold Vis at the station is low. We identify the 99th percentile of reported Vis in each year as the threshold for each station, and reject months with < 50 % of the data below the threshold. This approach differs from eliminating stations with low thresholds (e.g. Husar et al., 2000). **ACPD**

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Observations could still be meaningful at heavily polluted stations even if the threshold is low, while for clean stations with high thresholds most of the reported Vis could remain unresolved. To further ensure data representativeness and variability, data are removed for any month with less than 4 different days of data or with nearly identical percentile values (i.e. the ratio of 50th and 25th percentile Vis is less than 1.07 or the ratio of the 25th to 10th percentile Vis is less than 1.1) following Husar et al. (2000).

We describe the monthly Vis level with nonparametric statistics rather than arithmetic mean for a few reasons. First, an arithmetic mean would have biased monthly statistics due to the variable fraction (50–100 % after the threshold filtering) of Vis reported under the threshold in one month. Second, Vis is recorded as discrete values with coarse and uneven increments, and is not normally distributed (Schichtel et al., 2001). The protocol of reporting Vis varies across stations, depending on local regulations and available Vis markers. Both issues would affect the GSOD data or the monthly mean 1/Vis so we work with the raw data. We follow the convention to adopt the 75th percentile 1/Vis as the monthly representation of haziness (Husar et al., 2000; Qu et al., 2013). Other statistics, such as 50th and 90th percentile 1/Vis lead to similar trends and do not alter the conclusion of this study. However, the 50th percentile is closer to and more vulnerable to the detection limit, while the 90th percentile tends to be more susceptible to extreme events. Husar and Patterson (1987) assessed the effects of different choices of statistics. Below we commonly refer to the 75th percentile as "monthly 1/Vis" unless stated otherwise.

3.1.3 Completeness check

Completeness criteria are applied for further screening. A year of data is removed if less than 6 months in this year is available to guarantee annual representativeness. Short-term time series covering less than 7 years are also removed since they offer little information on trends.

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Sudden discontinuities in characteristics of the derived monthly time series of 1/Vis are frequently found even after the comprehensive filtering. Any change of the Vis marker (i.e. change of C_0) or observing standard (i.e. change of C_{crit}) could alter the relationship (K) between b_{ext} and 1/Vis, introducing inconsistency in the time series unrelated to actual $b_{\rm ext}$ change. For example, instrumentation (e.g. telephotometers, transmissometers and scatterometers) has replaced human observers at many sites in the US (Kessner et al., 2013) and to a lesser extent in Europe (Vautard et al., 2009). but there is a lack of documentation recording when and at which stations this switch occurred. Such artificial changes could seriously bias the inferred trends if not addressed. Various methods have been proposed to detect abrupt "change points" (Costa and Soares, 2009; Reeves et al., 2007). For example, the RHtest software package developed for multiple change point detection is based on penalized maximal t and F test (Wang, 2008a; Wang et al., 2007) embedded in a recursive testing algorithm (Wang, 2008b). We adopt the FindU function in the RHtest (version 4, available at http://etccdi.pacificclimate.org/) software to detect "type-1" change points (without reference time series). We manually examine all reported change points for possible false detections.

Figure 2 shows an example of change point detection based on the time series of 50th and 75th percentiles of monthly 1/Vis at one ISD station. The change points are reported in 3 different types (95% confidence): significant change, possibly significant (undetermined) change and insignificant change. In this example, although 4 significant changes for the 50th percentiles 1/Vis and 2 significant change points for the 75th percentiles 1/Vis are reported, only one candidate (February 1988) indicated by both time series is considered as an obvious discontinuity and chosen as the actual change point.

The candidate change points provided by RHtest allow greater efficiency than pure manual detection, which is prohibitive for thousands of stations. Any gap of more than

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4 years in a time series is also considered as a change point. Such a large gap could obscure protocol changes and introduce uncertainties in the derived trends without separation. We analyze separately the 1/Vis time series before and after the determined change points. Finally, we eliminate any year of data with annual 1/Vis (average of monthly 1/Vis) less than 40 M m⁻¹ to address the poor data variation and representativeness of clean environments, as will be discussed in Sect. 4.1.

We acknowledge that, although guided by RHtest results and a synthetic analysis based on the time series of 50th and 75th percentiles, this is still a subjective method. A small fraction of determined change points could be extreme events, while a few undetected change points missed by this subjective judgement might remain in the analysis. Several time series with irregular temporal variation are also removed during the visual examination. In summary, only 1/Vis time series considered as consistent and continuous are analyzed here. The threshold filtering (Sect. 3.1.2) and change point detection (Sect. 3.1.4) are designed to ensure basic representativeness and continuity of the derived monthly 1/Vis time series, and are the main differences of this processing from prior investigations.

3.1.5 Distribution of stations

Figure 3 (top) shows the ISD stations and the number of years with available data for 1929–2013 before and after data processing. A total of 3930 stations (5306 time series) remain after processing, most of which are located in the US (753), Europe (1625) and Eastern Asia (791). More than 6000 removed stations have less than 7 years of data as indicated in the left panel. Many other removed stations have small population density or harsh observing environment (e.g. islands and polar regions), which might correspond to poor observing conditions or maintenance.

Figure 3 (bottom) shows that most US stations are screened after the mid 1990s across the US. This is because more than 90 % of the ISD stations gradually switched to employ a low Vis threshold of 10 miles (\sim 16 km) after the mid 1990s (Fig. A1), probably due to the introduction of unified instrumentation (Kessner et al., 2013). A max-

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imum Vis of 16 km may be sufficient for airport navigation and weather reports, but this threshold Vis under clear sky conditions represents a moderate pollution level, and clean cases are not resolved. Thus most of the US stations with such low thresholds are rejected during the threshold screening. In contrast, screened stations remain densely distributed with long-term data over Europe and Eastern Asia after the mid 1990s because the adopted thresholds are generally higher and more consistent (Fig. A1).

Complimentary in situ data

We adopt complimentary data to evaluate and interpret the constructed monthly 1/Vis time series and trends. The measured and calculated aerosol optical data from the Interagency Monitoring of PROtected Visual Environments (IMPROVE) programme (http://vista.cira.colostate.edu/improve/Data/data.htm) are employed to evaluate the screened 1/Vis data and its trends after 1988. IMPROVE applies empirical mass extinction and RH growth factors to measured mass of aerosol components to calculate and report ambient b_{ext} in a 3-4 day frequency, and for several stations concurrent measurements of aerosol scattering coefficient (b_{sp}) are also made at hourly frequency using nephelometers. We generate monthly mean total $b_{\rm ext}$ (including aerosol extinction and Rayleigh scattering) and $b_{\rm sp}$ from data with RH < 90 % and status flags as "V0" (valid). Any month with less than 4 available days for averaging is abandoned. Pitchford et al. (2007) demonstrated that the estimated $b_{\rm ext}$ is consistent with measured $b_{\rm sn}$. We also find high correlation (r = 0.90, N = 3439) between monthly b_{ext} and b_{sp} across IMPROVE stations (Fig. A2).

The measurement of $b_{\rm ext}$ or $b_{\rm sp}$ is sparse outside the US. Therefore we obtain longterm measurements of fine particulate matter mass (PM25) from the European Monitoring and Evaluation Programme (EMEP, http://ebas.nilu.no) for comparison over Europe. Forty-five stations of data collected by filter-based ambient samplers are used. Similarly, these daily PM_{2.5} data are averaged monthly provided at least 4 valid measurements are available.

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We apply bottom-up SO_2 emission inventories to interpret historical 1/Vis trends. This approach exploits the close relation of sulfate aerosol concentration with SO_2 emission due to the short time scale of SO_2 oxidation (Chin et al., 1996, 2014; Daum et al., 1984; Hand et al., 2012a), the major $PM_{2.5}$ contribution from sulfate aerosols over land for most populated areas (Chin et al., 2014; Philip et al., 2014), and the dominance of sulfate for light extinction due to its hygroscopicity (Hand et al., 2014). We employ 3 different SO_2 emission datasets, including country-level data for 1850–2005 (S. J. Smith et al., 2011a, b), gridded data from EDGAR version 4.2 at 0.1° resolution for 1970–2008 (http://edgar.jrc.ec.europa.eu/), and data from Lu et al. (2011) at 0.5° resolution for 1996–2010 over China. The data from S. J. Smith et al. (2011a) are referred to as "Smith emissions" below. The data from Lu et al. (2011) are referred to as "Lu emissions".

3.4 Trend analysis

In this study, we separately calculate trends for several periods of 8–10 years to allow possible trend reversal, and to include stations with short-term data. The choice of study periods is mainly based on the historical SO_2 emission data. Figure A3 shows the Smith emission data for several representative countries. SO_2 emission trends in the US changed direction at \sim 1944, \sim 1954, and again at \sim 1973. Also, for most Eastern European countries, there is a sharp reduction of SO_2 emission starting from \sim 1989 after the breakdown of the communist system, while the 1997 Asian financial crisis affected the SO_2 emission trend in Korea. It is of particular interest to examine how Vis is affected by these emission changes. Data for most ISD stations outside US start from the year 1973, and representative coverage of Vis stations over the US starts from the year 1945, although the earliest records after screening start from 1929. Based on these transition points of SO_2 trends and Vis data availability, 8 periods (1945–1953, 1954–1963, 1964–1972, 1973–1980, 1981–1988, 1989–1996,

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1997–2005, 2006–2013) are chosen to be analyzed in detail over the US, while the latter 5 periods are studied for Europe and Eastern Asia. We also take a glimpse of two short periods before 1945 (1929-1934 and 1935-1944) over the US where stations are less spatially representative (not included in regional quantitative analysis) but still 5 show prominent trend information in 1/Vis.

We assess the linear trend and its significance (p value, two-tail test) in the deseasonalized monthly anomalies using Sen's slope (Sen, 1968) and the Mann-Kendall (MK) test (Kendall, 1975; Mann, 1945). All monthly data are deseasonalized by removing multi-year monthly means. Pre-whitening is introduced to reduce the effect of lag-1 autocorrelation (Yue et al., 2002), and 95 % confidence interval (CI) of the slope is calculated (Li et al., 2014). This nonparametric trend estimation method is insensitive to missing values and outliers in the time series, and does not require a normal distribution, thus it has been widely adopted to study aerosol trends in previous studies (Collaud Coen et al., 2013; Papadimas et al., 2008). Least square trends (Weatherhead et al., 1998) are also calculated, and are found to be consistent with the MK-Sen trends. For all the 8027 calculated slopes in 1/Vis, 88 % are unanimously diagnosed as significant (90 % confidence, p < 0.1) or insignificant by both methods. For the significant trends 76% of their differences are within 20%. Relative trends are calculated by normalizing the absolute MK-Sen trends to the multi-year mean to facilitate the comparison and interpretation with other in situ data.

Short-term trends of 8–10 years are expected to be less statistically robust and more sensitive to extremes. For each period, a time series is required to contain at least half of the total months and 2/3 of the total years (e.g. at least 60 monthly data in at least 7 years for a 10 year period) for the calculated trend to be representative. This step only reduces the number of stations at which trends are reported, but does not further screen the data.

We calculate composite trends based on monthly 1/Vis averaged from an ensemble of stations (e.g. for the time series of collocated stations in Sect. 4 or defined regions in Sect. 5). To ensure temporal representativeness, a station is considered in the average **ACPD**

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only if 2/3 of the total months of data are available for the study period. Qualified stations are gridded to 1° resolution before averaging to avoid biased averaging towards more densely distributed areas. To ensure spatial representativeness, only monthly data derived from at least 75% of the maximum number of grids for each study period 5 are used in the composite trend estimation. This strategy reduces sampling difference within each periods, however the composite 1/Vis for different periods might be averaged from a different distribution of stations. We expect the uncertainty from spatially variant K and data quality to be random, and to be reduced by spatial averaging and by normalizing the slopes into relative trends.

Evaluation against in situ data

Comparison with IMPROVE best and EMEP PM_{2.5}

We compare the monthly IMPROVE best data with the quality controlled 1/Vis from Sect. 3.1. Collocations are considered between IMPROVE and ISD time series within the distance of less than 1° and altitude difference of less than 500 m. One IMPROVE station could pair with more than one ISD station and vice versa. Fifty-nine collocations (each with at least 20 paired monthly values) are made. We expect a maximum correlation of 0.9 given the relation between measured $b_{\rm sp}$ and calculated $b_{\rm ext}$ (Fig. A2). Similarly, we create collocations between ISD 1/Vis and EMEP PM_{2.5} on a monthly basis, and expect a weaker correlation due to variation of aerosol water and mass extinction efficiency.

Figure 4 shows the comparison results between collocated 1/Vis and b_{ext} over the US. This evaluation highlights several major findings:

1. The mean $b_{\rm ext}$ level of collocated IMPROVE stations after 1990 is below 50 M m⁻¹ for the western US, and below 120 Mm⁻¹ for the eastern US (bottom left). As discussed in Sect. 3.1, the low threshold Vis of ~ 16 km (equivalent to $b_{\rm ext} \sim 100$ – **ACPD**

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 $240 \,\mathrm{M\,m^{-1}}$ depending on K) recently adopted by most US stations fails to resolve actual $b_{\rm ext}$ variation under this relatively clean environment. Thus many stations are rejected by the threshold filtering.

- 2. As shown in the top left panel, correlation coefficients of monthly values vary from \sim 0 to 0.85. About half of the collocations (29 out of 59) have r < 0.5, while 10 collocated ISD stations have r > 0.7. The overall moderate correlation is not unexpected, as is similarly found in previous studies (Mahowald et al., 2007; Wang et al., 2012). Correlations are expected to differ from station to station, due to the inherent difference in observing conditions, protocols, and residual uncertainties. This preliminary evaluation suggests that Vis data at individual stations can be unreliable, and in the following discussion we focus on interpreting regionally coherent observations.
- 3. Correlations generally exceed 0.5 in the eastern US, where the mean $b_{\rm ext}$ is higher due to higher aerosol concentration (Hand et al., 2012b; van Donkelaar et al., 2015) and to a larger fraction of hygroscopic sulfate aerosols (Hand et al., 2012b). The correlation increases significantly with the mean $b_{\rm ext}$, indicating the tendency for better 1/Vis representativeness in more polluted regions. As previously discussed, at lower $b_{\rm ext}$ more reported Vis are close to the threshold Vis, thus the true 1/Vis tends to be less well resolved. Also, because the Vis data are reported in discrete values, clean stations with a narrow dynamic range of $b_{\rm ext}$ have few reportable Vis to capture the continuous $b_{\rm ext}$ variation. Moreover, the increment of adjacent reportable Vis is relatively coarse in cleaner conditions (WMO, 2008), and atmospheric homogeneity might break down for longer distances. All these factors weaken the ability of Vis to capture b_{ext} variation in clean environments. Wang et al. (2012) found low correlation of 1/Vis with PM₁₀ over the US and Canada, and similarly attributed this to low aerosol concentrations and higher Vis uncertainty over North America. Thus we apply the 40 Mm⁻¹ threshold of annual 1/Vis to further filter the data as introduced in Sect. 3.1.4. Without

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this screening, 7 of 8 stations with mean 1/Vis < $40\,\mathrm{M\,m^{-1}}$ were found to exhibit low correlations (r < 0.25) with collocated b_ext . Different thresholds from 10 to $70\,\mathrm{M\,m^{-1}}$ were tested, and thresholds above $40\,\mathrm{M\,m^{-1}}$ ceased to improve the consistency with the few sites reporting b_ext .

4. The slope of fitted linear relationship (top right) between 1/Vis and $b_{\rm ext}$ varies from \sim 0.8 to \sim 2 even over the eastern US where correlations are higher. This supports the expectation that this slope (K) would differ spatially with observing conditions (Griffing, 1980; Husar et al., 2000; Schichtel et al., 2001), as discussed in Sect. 2. Thus in the later analysis we focus on the relative trend of 1/Vis which is independent of K.

Figure 4 (bottom right) also shows the correlation between monthly 1/Vis and PM $_{2.5}$ over Europe. Although the relation of 1/Vis with PM $_{2.5}$ is expected to be more uncertain than with $b_{\rm ext}$, we find more stations with high correlation (r > 0.5) over Europe (93 out of 129, 72%) than over the US (51%). Wang et al. (2012) similarly found higher correlation of 1/Vis with PM $_{10}$ over Europe and China than over the US and Canada. The higher thresholds and higher concentration of fine aerosol over Europe (van Donkelaar et al., 2015) allow 1/Vis to better resolve PM $_{2.5}$ variation there. These findings suggest more reliability of Vis observations at areas with both higher aerosol loading and sufficiently high thresholds to resolve $b_{\rm ext}$ variation, e.g. the three populated regions investigated in this study.

4.2 Trend evaluation

Figure 5 shows the spatial distribution of relative trends in 1/Vis, from IMPROVE estimated $b_{\rm ext}$ and measured $b_{\rm sp}$ over the US for 1989–2013. The overall decrease across the US is consistent with recent trend studies employing IMPROVE $b_{\rm ext}$ (Hand et al., 2014) and $b_{\rm sp}$ (Collaud Coen et al., 2013) data, and is determined by the reduction of both aerosol mass and hygroscopicity (Attwood et al., 2014). The 1/Vis trends over

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1989–1996 generally reproduce the $b_{\rm ext}$ trends, with decreasing tendencies in the eastern and western US. For this period, 15 ISD stations and 9 IMPROVE stations with significant trends are collocated and labeled.

Figure 6 (top) shows the composite time series of the collocated 1/Vis and $b_{\rm ext}$ stations. The seasonal variation of the averaged $b_{\rm ext}$ is well reproduced by that of collocated 1/Vis, with a correlation of 0.77 between these two time series. Both composite 1/Vis and $b_{\rm ext}$ show a peak in summer months, due mostly to increased aerosol concentration in warm months because of increased photochemical activity and biogenic emission (Chen et al., 2012; Hand et al., 2012b). The trend of collocated 1/Vis (-1.6 % yr⁻¹; 95 % CI: -2.4, -0.8 % yr⁻¹) is within the confidence intervals of the decrease of $b_{\rm ext}$ (-2.4 % yr⁻¹; 95 % CI: -3.7, -1.1 % yr⁻¹). The slight underestimation may reflect the weak sensitivity of discrete 1/Vis data to the continuous decrease of $b_{\rm ext}$ in clean environments due to the threshold and discreteness issues.

Figure 5 (top) indicates that for the last 2 periods, the number of available ISD stations for trend analysis is dramatically reduced by their detection limit and improved air quality. Although the remaining sparse ISD stations still show overall consistency in trends with nearby $b_{\rm ext}$ and $b_{\rm sp}$, they cannot provide spatially coherent and aggregated trend information. We thus suggest that the ISD Vis data over the US are not appropriate for studying haze trends after the mid 1990s, and limit our analysis to data before 1996 for this region. Overall, the trend maps of 1/Vis, $b_{\rm ext}$ and $b_{\rm sp}$ show a dominant trend of decreasing haziness over the whole US after 1988, which reflects reduction of aerosol sources (Hand et al., 2014; Leibensperger et al., 2012).

Figure 7 shows the spatial distribution of relative trends in 1/Vis and $PM_{2.5}$ over Europe for 2006–2013. The dominant decreasing trends of $PM_{2.5}$ is well captured by the 1/Vis trends, especially at the 19 ISD and 10 EMEP collocated sites with significant trends.

Figure 6 (bottom) shows composite time series of $PM_{2.5}$ and 1/Vis of these collocated stations. High correlation (0.80) between these two time series indicates consistent seasonal variation. The winter maximum in the composite 1/Vis over Europe

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well represents the PM_{2.5} seasonality at most collocated EMEP sites, which could be attributable to near surface inversion and low surface winds (Yttri et al., 2012), to greater nitrate aerosol formation (Aas et al., 2012; Yttri et al., 2012), and to higher carboneceous aerosol emission from residential wood combustion (Denier van der Gon et al., 2015). The CI of the 1/Vis trend (-3.4 % yr⁻¹, 95 % CI: -4.4, -2.4 % yr⁻¹) overlaps with that of the PM_{2.5} trend (-5.8 % yr⁻¹, 95 % CI: -7.8, -4.2 % yr⁻¹), but underestimates the relative decrease of PM_{2.5}. In addition to the weak sensitivity of discrete 1/Vis to resolve aerosol variation under clean environment (the collocated EMEP stations are mostly in the cleaner Western Europe), the inclusion of Rayleigh scattering in 1/Vis and the non-linear association between ambient 1/Vis and dry PM_{2.5} (fixed at 50 % RH) also contribute to this bias.

In summary, 1/Vis exhibits spatially variant K (i.e. relationship with $b_{\rm ext}$) and data quality that suggests uncertainty in the information of one station especially at clean locations. However the aggregated 1/Vis time series successfully capture the seasonal variation and trends of collocated in situ data. The high correlation between composite time series and the overall consistency of composite trends suggest that the interpretation value of 1/Vis data benefits from averaging over multiple stations.

5 Historical trends of 1/Vis

5.1 United States

Figure 8 presents the calculated relative trend of 1/Vis of all qualified stations over the US for 1945–1988. Figure 9 shows the regionally averaged time series and trends of 1/Vis over the eastern US for 1945–1996, superimposed with the evolution of SO_2 emission data. Historically, 1/Vis in the eastern US experienced a pronounced decrease (-2.8 % yr⁻¹, p < 0.001) after World War II until the mid 1950s, a consistent upward trend afterwards (0.9–1.8 % yr⁻¹, p < 0.001) until the early 1970s, variable tendencies during 1973–1980, and a significant decreasing trend (-1.1 to -2.0 % yr⁻¹, p <

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0.005) from the early 1980s until 1996 (also in Fig. 5). This 1/Vis trend evolution resembles the SO₂ emission trend. Industrial activity gradually decreased after World War II until mid 1950s, followed by economic growth until the early 1970s with the emergence of both the oil crisis and the Clean Air Act (Greenstone, 2001). The emission of SO₂ 5 starts to consistently decrease after 1973 for the Smith inventory, and after 1977 for the EDGAR inventory. For the period 1973–1980 the regional 1/Vis is generally consistent with these two inventories except for an anomalous peak of annual 1/Vis in 1977-1979. The NOAA Climate Extremes Index (http://www.ncdc.noaa.gov/extremes/cei/) describes the winters of 1977–1979 as the coldest during 1945–1996 across the US. Increased emissions from domestic heating, as well as stagnant weather may contribute to the 1/Vis peak. After 1978, the three annual time series uniformly exhibit a downward tendency.

Table 1 contains the correlation of annual 1/Vis with SO₂ emissions. Annual 1/Vis over the eastern US exhibits a correlation of 0.66 with the Smith SO₂ emissions over the entire US (1946–1995), and of 0.73 with the EDGAR SO₂ emissions over the eastern US (1970-1995). The 1/Vis trends over the western US (where organic aerosols dominate in PM_{2.5} and forest fires are more prevalent) are less consistent than over the eastern US with the SO₂ emission data, given the influence of other sources. In summary, the 1/Vis time series successfully capture large-scale haze evolution over the eastern US from 1945 to 1996, which is consistent with changes in SO₂ emissions as well as previous investigations on 1/Vis for this region (Husar and Wilson, 1993; Schichtel et al., 2001).

Figure A4 shows the calculated 1/Vis trends over the US for two short periods prior to 1945. Although the stations are sparsely distributed, the nearly uniform trends in 1/Vis strongly suggest a prominent decrease over 1929–1934, and then a rapid increase over 1935-1944. This evolution reflects the significant drop in industrial activity following the 1929 Great Depression, and the economic recovery after ~ 1933 during the New Deal programs and World War II. The Smith SO₂ emissions of the US (Fig. A3) also reflect these socioeconomic events.

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Figure 10 presents the spatial distribution and temporal evolution of haze trends over Europe as derived from the 1/Vis data for 1973–2005. The historical trend pattern of 1/Vis is quite different between Western and Eastern Europe. The large-scale 1/Vis trend over Western Europe is consistently decreasing for the 4 periods after 1981 (also in Fig. 7). Some countries such as the UK and France begin decreasing prior to 1981, consistent with the SO₂ emission decrease over these countries (Fig. A3). Prior analysis also indicated Vis improvements after ~ 1973 for most sites over the UK (Doyle and Dorling, 2002). Meanwhile stations over Eastern Europe have significantly increased 1/Vis for 1973–1980, a mostly decreasing trend in its western part for 1981–1988, and then a decrease-dominant trend after 1989.

Figure 11 shows the regionally composite time series of 1/Vis as well as SO_2 emissions over Western and Eastern Europe for 1973–2013. Table 2 lists the specific country names included in the Smith emissions for the two regions. The evolution of 1/Vis over Western and Eastern Europe is broadly consistent with the SO_2 emissions, and reflects the lag of emission reduction in Eastern vs. Western Europe. Stjern et al. (2011) similarly reported later improvement in Vis over Eastern vs. Western Europe. The SO_2 emission reduction extends from the 1980s to the end of data record for Western Europe, and primarily over 1989–2000 for Eastern Europe. The composite 1/Vis time series successfully capture the significant reduction of haze over Western Europe (–1.1 to –1.7 % yr⁻¹, p < 0.08). For Eastern Europe the decrease of 1/Vis is stronger before 1997 (–2.0 % yr⁻¹, p < 0.001) than after 2006 (–1.1 % yr⁻¹, p = 0.03), and the calculated trend over 1997–2005 is insignificant, consistent with the SO_2 emission evolution. There is an obvious peak in 1/Vis from October 1995 to March 1996 especially over Eastern Europe, which is consistent with the peak sulfate concentration that Stjern et al. (2011) attributed to the anomalously cold winter of 1996 with stagnant air.

Table 1 shows that the annual 1/Vis time series exhibit a correlation of 0.91 (0.92) with the Smith Emissions for 1973–2005, and of 0.92 (0.92) with the EDGAR emissions

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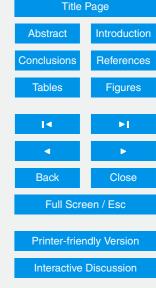
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5.3 Eastern Asia

Figure 12 shows the calculated relative trends of 1/Vis over Eastern Asia after 1973. A persistent increasing trend of 1/Vis dominates over eastern China for more than 30 years. A prominent feature in the trends over China is more heterogeneity in the spatial distribution compared to the trend maps over the US and Europe. This could be a result of asynchronous economic development, as several studies reported "lagging" of Vis impairment in rural sites (from ~ 1990s) compared to urban sites (from ~ 1960s) in China (Quan et al., 2011; Wu et al., 2012). The overall increasing trend in 1/Vis reverses in the last period of 2006–2013, when most stations in southern China and many in northern China show a statistically significant decreasing trend of 1/Vis. This is consistent with the implementation of fuel-gas desulfurization facilities in power plants after ~ 2007. This recent reduction was also supported by satellite observations of SO₂ and AOD (Li et al., 2010; Lu et al., 2010, 2011; S. Wang et al., 2015; Zhao et al., 2013).

Figure 12 also shows a consistent increase of 1/Vis over Korea from 1973 to 1996. After 1997 when the SO_2 emission transits to decrease (Fig. A3), the increase in 1/Vis levels off and reverses. The aerosols over China also affect areas downwind through long-range transport (Aikawa et al., 2010). For the 1997–2005 period, most eastern stations of Korea show a downward trend, in contrast with the increasing 1/Vis over the west, which is more strongly influenced by pollutant transport from China. Similarly, stations over the western and coastal areas of Japan consistently exhibit an upward 1/Vis trend before 2006, despite the continuous decrease of local SO_2 emission and concentration since 1970 (Wakamatsu et al., 2013). Aikawa et al. (2010) found a zonal gradient in terms of both the magnitude and trend of measured SO_2 and sulfate concentrations over Japan, and in the modeled contribution from China to the sulfate concentration in Japan. Lu et al. (2010) reported that most EANET stations over Japan

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and Korea have increasing trends in SO₂ and sulfate aerosols from 2001 to 2007. For the last period 2006–2013, 1/Vis shows a dominant decreasing trend over Japan and Korea that may reflect in part China's SO₂ emission controls. Itahashi et al. (2012) reported a trend reversal of MODIS fine AOD over the Sea of Japan from increasing to ₅ decreasing at ~ 2006 that is more consistent with China's SO₂ emission than the local emission. This analysis highlights the sensitivity of 1/Vis to long range transport.

Figure 13 presents a regional analysis of averaged 1/Vis time series over northern and southern China, and the evolution of SO₂ emissions from two inventories. The overall Vis impairment trend in China for 1973-2005 reflects the consistent SO₂ emission increase. Both the north and south show a steady and significant (p < 0.001) increase of haziness for 1973-1980 period, and southern China shows an even faster impairment (2.9 % yr⁻¹) than the north (1.2 % yr⁻¹). For the next 2 decades (1980-2000) the 1/Vis increase slows down in both the south and the north, in accordance with other investigations using Vis and SSR data (Chen and Wang, 2015; Luo et al., 2001; Wu et al., 2014). The south exhibits a slower (0.2 % yr⁻¹) and less significant (p > 0.3) increase than the north $(0.5-0.6 \% \text{ yr}^{-1})$. This difference is determined not only by the slower increase of SO₂ emissions in the south (Lu et al., 2010), but also by more precipitation and ventilation in the south that favors the removal of aerosols and their precursors (Xu, 2001; Ye et al., 2013). The decline of SO₂ emissions from 1996 to 2000 reflects both the 1997 Asian financial crisis, and a decline in coal use and sulfur content (Lu et al., 2011). Both regions show a leveling off or even reversal of 1/Vis increase during this short period, which is again more significant in the south. The period 2000–2006 exhibits significant growth (> 1 % yr⁻¹) of 1/Vis in both the north and south, resembling the steady growth in SO₂ emissions. The recent reduction of SO₂ emissions is reflected in the Lu emissions while not in the EDGAR emissions. After 2006 significant (p < 0.05) decreasing trends in 1/Vis are apparent $(-0.9 \text{ to } -1.6 \text{ \% yr}^{-1})$ for both northern and southern China, which is more consistent with the Lu emissions. As shown in Table 1, the annual 1/Vis time series exhibit a high

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5.4 Connections to global dimming and brightening

Historical records of surface solar radiation (SSR) serve as another data resource to study and interpret changes in air pollution during the last few decades. SSR is determined by the total columnar extinction of aerosols and clouds while 1/Vis represents the extinction level at the surface. Moreover, the direct scattering and absorption of solar radiation by aerosols could be amplified in less polluted regions or dampened over highly polluted stations, due to aerosol-cloud interaction (Fuzzi et al., 2015; Wild, 2009). Despite these uncertainties, the transitions of 1/Vis trends in this study agree well with the observed reversals of SSR from "dimming" to "brightening" in the 1980s for the US and Europe (Streets et al., 2006; Turnock et al., 2015; Wild, 2012). Over China, the recently reported decadal SSR variation shows dimming before the 1990s and no significant trend afterwards (Tang et al., 2011; K. Wang et al., 2015). The latter phenomenon may reflect compensation of more aerosol extinction by less cloud cover (Norris and Wild, 2009).

6 Conclusions

This study examines Vis observations as a trend indicator of haziness and air quality over the US (1945–1996), Europe (1973–2013), and Eastern Asia (1973–2013). We comprehensively process the raw data from over 20 000 stations considering effects from meteorological factors, protocol design, and human errors. We develop filters to exclude relatively clean cases (i.e. months with $\leq 50\,\%$ records below the threshold Vis, or years with annual $1/{\rm Vis} \leq 40\,{\rm M\,m^{-1}}$) with weaker sensitivity to $b_{\rm ext}$ variation, and apply change point detection and separation to largely reduce the intrinsic discontinuities. Nearly 4000 stations remain after the processing with 753 over the US, 1625

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over Europe, and 791 over Eastern Asia. The composite time series of 1/Vis over the US for 1989–1996 generally agrees with the collocated IMPROVE $b_{\rm ext}$ in terms of both seasonal variation (r = 0.77) and trends ($-1.6 \% \, yr^{-1}$, 95 % CI: -2.4, $-0.8 \% \, yr^{-1}$) in 1/Vis vs. b_{ext} (-2.4 % yr⁻¹, 95 % CI: -3.7, -1.1 % yr⁻¹). Similarly, for 2006–2013 over Europe, the seasonal variation (r = 0.80) and significant decrease ($-5.8 \% \text{ yr}^{-1}$, 95 % CI: -7.8, $-4.2 \% \text{ yr}^{-1}$) in PM_{2.5} are captured by collocated 1/Vis ($-3.4 \% \text{ yr}^{-1}$, 95 % CI: -4.4, -2.4 % yr⁻¹). This consistency highlights the benefits of thorough data screening to reduce uncertainties brought by the inherent issues in Vis observations such as threshold choices, discreteness and discontinuities. As discussed in Sect. 3.1, the inclusion of unresolved values in the mean 1/Vis and the contaminants of discontinuities could dampen the ability of 1/Vis to correctly resolve aerosol trends. Admittedly, the derived 1/Vis trends are still subject to several uncertainties, e.g. the spatially variant K and data quality, the less robust short-term trends, sampling differences and direct averaging in composite time series. Nevertheless, the interpretation value of 1/Vis data is shown to be enhanced by the comprehensive screening and spatial averaging. Therefore we focus on the trend results that are regionally coherent and aggregated, and avoid drawing strong conclusions based solely on the 1/Vis trends. Although at individual stations the 1/Vis changes might be affected by these above stated artificial factors, regionally coherent trend signals suggest these derived 1/Vis trends represent actual changes in b_{ext}. Upon final publication, our filtered monthly 1/Vis data will be freely available as a public good (http://fizz.phys.dal.ca/~atmos/martin/?page_id=2527).

Analysis of the 1/Vis trends for several short periods reveals haze trend evolution and reversals. These historical 1/Vis trends and their evolution also exhibit compelling consistency with SO₂ emissions and SSR studies. For example, 1/Vis shows statistically significant decreasing trends from the late 1970s to the mid 1990s over the eastern US $(-1.1 \text{ to } -2.0 \text{ \% yr}^{-1})$, from the early 1980s to 2013 over Western Europe $(-1.1 \text{ to } -2.0 \text{ \% yr}^{-1})$ to $-1.7 \% \text{ yr}^{-1}$), in the early 1990s ($-2.0 \% \text{ yr}^{-1}$) and after the mid 2000s ($-1.1 \% \text{ yr}^{-1}$) over Eastern Europe, and after the mid 2000s over China $(-0.9 \text{ to } -1.6 \% \text{ yr}^{-1})$. These recent decreases in 1/Vis are attributable to emission changes in these populated ar**ACPD**

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eas. Reversal points of 1/Vis trends also consistently reflect several historical socioe-conomic events e.g. the New Deal programs (from decrease to increase at \sim 1934), the end of World War II (from increase to decrease at \sim 1945) and the Clean Air Act (from increase to decrease at \sim 1979) in the US, the collapse of communism in Eastern Europe (from increase to decrease at \sim 1989), and the 1997 Asian financial crisis.

Therefore, the constructed 1/Vis data are applicable to resolve historical aerosol trends on a regional and annual basis, and provide complementary information about the historical changes in air quality. For instance, the annual 1/Vis time series exhibit high correlations (0.7–0.9) with SO_2 emissions for 5 large domains (Table 1). Apart from verifying the historical 1/Vis trends, this consistency also provides an evaluation of emission inventories. For example, after \sim 2006 1/Vis trends agree better with Lu et al. (2011) than the EDGAR emissions in capturing the SO_2 emission controls over China. Emission inventories differ significantly (S. J. Smith et al., 2011a), and 1/Vis data offer constraints on these inventories.

However, SO_2 emission inventories cannot fully explain the trends in ambient haze due to the influence of other emissions and meteorological factors. We have shown that occasional cold winters in the US and Europe, and the long-range transport of China's pollutants into Korea and Japan could affect the association between 1/Vis and local emission. Future work includes applying a chemical transport model to further interpret the observed 1/Vis $(b_{\rm ext})$ trends, as well as the contribution from meteorology and emissions.

Appendix A

Four figures (Figs. A1–A4) are included for complementary interpretation.

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Table 1. Summary of Pearson correlation coefficients (r) between annual 1/Vis and SO₂ emissions for 5 regions.

Inventory	Period	Eastern US	
Smith	1946–1995	0.66	
EDGAR	1970–2008	0.73	
		Eastern Europe	Western Europe
Smith	1973–2005	0.92	0.91
EDGAR	1973–2008	0.92	0.92
		Northern China	Southern China
Lu	1996–2010	0.78	0.87
EDGAR	1973–2008	0.91	0.88

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Table 2. List of countries included to calculate regional SO₂ emission from the country-level emission data (countries with most parts inside the defined region) of S. J. Smith et al. (2011a).

Region	Countries	
Eastern US	United States	
Eastern Europe	Albania, Belarus, Bosnia and Herzegovina, Bulgaria, Czech, Croatia, Greece, Hungary, Latvia, Lithuania, Moldova, Poland, Romania,	
Western Europe	Serbia and Montenegro, Slovakia, Slovenia, Turkey, Ukraine Austria, Belgium, Denmark, France, Germany, Ireland, Italy, Netherland, Portugal, Spain, Switzerland, UK	

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Figure 1. Schematic flow of the data processing. Numbers in brackets indicate the remaining number of stations after this processing step.

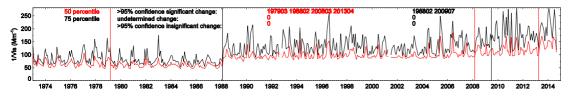


Figure 2. An example of change point detection and determination based on the time series of 50th (red) and 75th (black) percentiles of monthly 1/Vis. Automatically detected change points are represented by vertical lines. Text in the inset lists the dates of automatically detected points. February 1988 is determined as the separation point for further analysis, while other reported breaks are considered as false detections.

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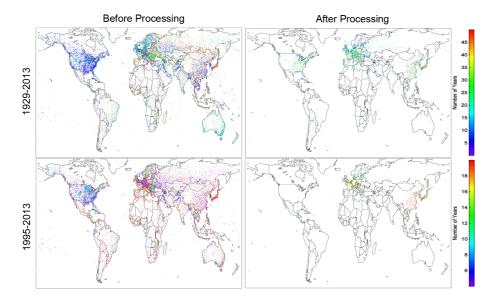


Figure 3. Distribution of Integrated Surface Database (ISD) stations before (left) and after (right) data screening. Colors indicate the number of years with available visibility data for (upper) 1929–2013 and (lower) 1995–2013.

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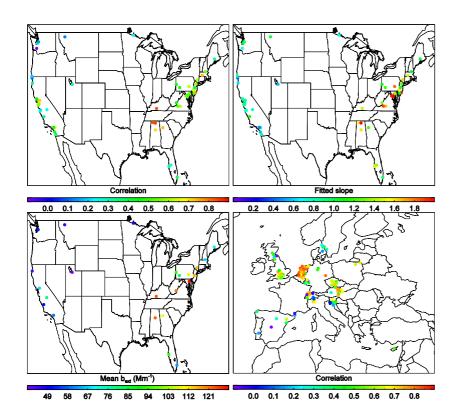


Figure 4. Spatial distribution of: (top left) Pearson correlation coefficients between collocated pairs of monthly ISD 1/Vis and IMPROVE b_{ext} , (top right) slope of monthly b_{ext} against monthly 1/Vis after linear fitting through the origin point using the reduced major-axis linear regression (Ayers, 2001), (bottom left) average of the collocated best of IMPROVE stations, and (bottom right) Pearson correlation coefficients between collocated pairs of monthly ISD 1/Vis and EMEP PM_{2.5}.

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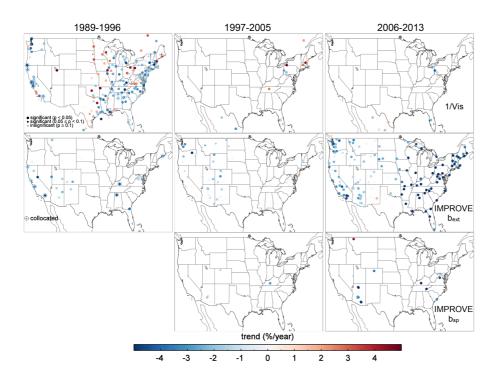


Figure 5. Spatial distribution of relative trends in 1/Vis (top row), IMPROVE b_{ext} (middle row), and IMPROVE $b_{\rm sp}$ (bottom row) over the US for 1989–2013. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90-95 % significance, and open circles indicate insignificant trends. Stations with cross and circle symbols are collocated between the ISD and IMPROVE networks over 1989-1996 for composite time series analysis in Fig. 6.

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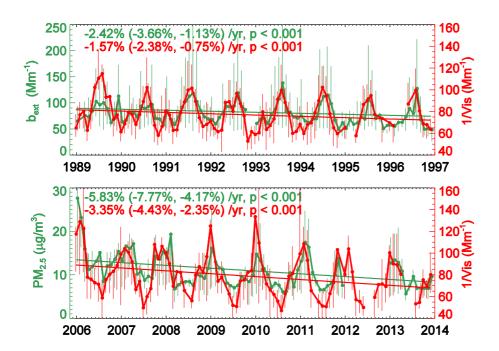


Figure 6. Composite time series and trends of (top) 1/Vis and $b_{\rm ext}$ for collocated ISD and IMPROVE stations (Fig. 5) over 1989–1996 and (bottom) 1/Vis and PM_{2.5} for collocated ISD and EMEP stations (Fig. 7) over 2006–2013. Only stations with significant trends of > 90 % confidence are collocated. The long ticks on the horizontal axis indicate the January of the year. Data gaps represent months with less than 75% of the total grids. Error bars show the 25th and 75th percentile of all monthly values of collocated stations.

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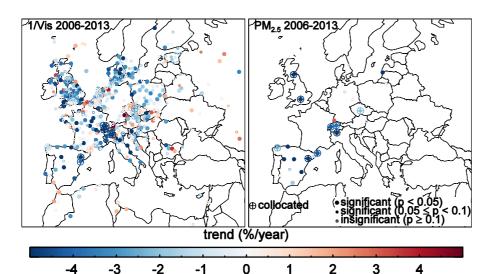


Figure 7. Spatial distribution of relative trends in 1/Vis and $PM_{2.5}$ over Europe for 2006–2013. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90–95 % significance, and open circles indicate insignificant trends. Stations with cross and circle symbols are collocated between the ISD and EMEP networks for composite time series analysis in Fig. 6.

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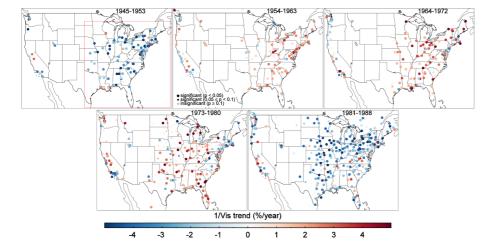


Figure 8. Spatial distribution of relative trends in 1/Vis over the US for 1945–1988. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90-95 % significance, and open circles indicate insignificant trends. The red rectangle defines the eastern US region for composite time series analysis in Fig. 9.

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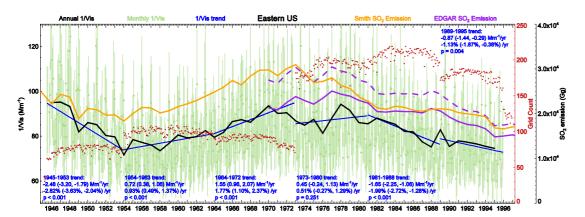


Figure 9. Composite time series of 1/Vis and SO₂ emission over the eastern US region. The long ticks on the horizontal axis indicate January of the year, where all annual values are plotted. Light green dots represent the average monthly 1/Vis of all qualified stations (error bars showing the 25th and 75th percentile) in the defined region. Red dots show the number of grid cells for averaging, and data gaps indicate months with less than 75% of the total grids for each period. Blue lines and text represent the 1/Vis trends calculated using the monthly anomalies for each period. Trends in parentheses are the 95 % confidence intervals. Black lines are the annual 1/Vis averaged from at least 8 monthly values. SO₂ emissions for the entire US from S. J. Smith et al. (2011a) are in orange. Purple indicates EDGAR SO₂ emissions for the entire US (dashed) and for the defined region (solid).

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Figure 10. Spatial distribution of relative trends in 1/Vis over Europe for 1973–2005. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90–95 % significance, and open circles indicate insignificant trends. Red rectangles define the Eastern and Western Europe regions for composite time series analysis in Fig. 11.

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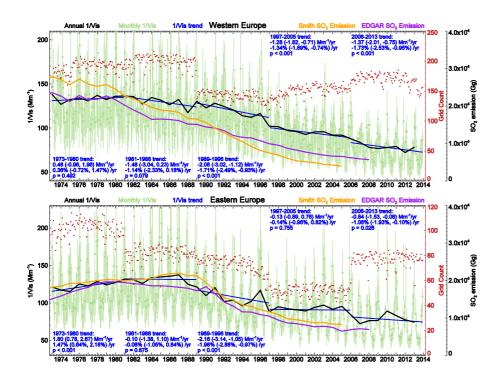


Figure 11. Regional time series analysis of 1/Vis and SO₂ emission over Western and Eastern Europe. The long ticks on the horizontal axis indicate January of the year, where all annual values are plotted. Light green dots represent the average monthly 1/Vis of all qualified stations (error bars showing the 25th and 75th percentile) in the defined region. Red dots show the number of grid cells for averaging, and data gaps indicate months with less than 75 % of the total grids for each period. Blue lines and text represent the 1/Vis trends calculated using the monthly anomalies for each period. Trends in parentheses are the 95 % confidence intervals. Black lines are the annual 1/Vis averaged from at least 8 monthly values. The Smith SO2 emissions in orange are the total emission of all countries listed in Table 2 for each region. The EDGAR SO₂ emissions in purple are summed from all pixels inside the defined region (Fig. 10).

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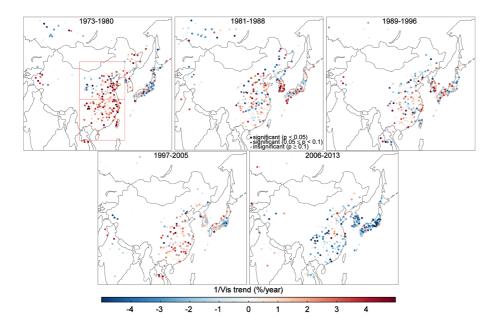


Figure 12. Spatial distribution of relative trends in 1/Vis over Eastern Asia for 1973–2013. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90–95 % significance, and open circles indicate insignificant trends. Red rectangles define the northern and southern China regions for composite time series analysis in Fig. 13.

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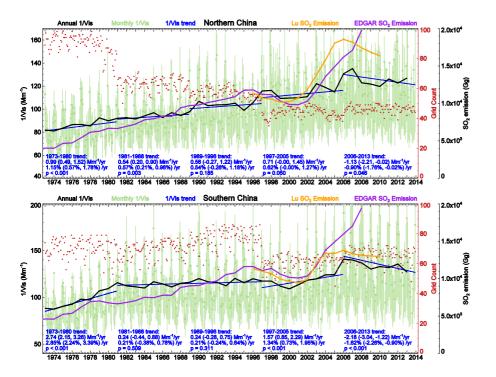


Figure 13. Regional time series analysis of 1/Vis and SO₂ emission over sorthern and nouthern China. The long ticks on the horizontal axis indicate January of the year, where all annual values are plotted. Light green dots represent the average monthly 1/Vis of all qualified stations (error bars showing the 25th and 75th percentile) in the defined region. Red dots show the number of grid cells for averaging, and data gaps indicate months with less than 75 % of the total grids for each period. Blue lines and text represent the 1/Vis trends calculated using the monthly anomalies for each period. Trends in parentheses are the 95% confidence intervals. Black lines are the annual 1/Vis averaged from at least 8 monthly values. The SO₂ emission in Lu et al. (2011) in orange and the EDGAR SO₂ emission in purple are summed from all pixels inside the defined region (Fig. 12).

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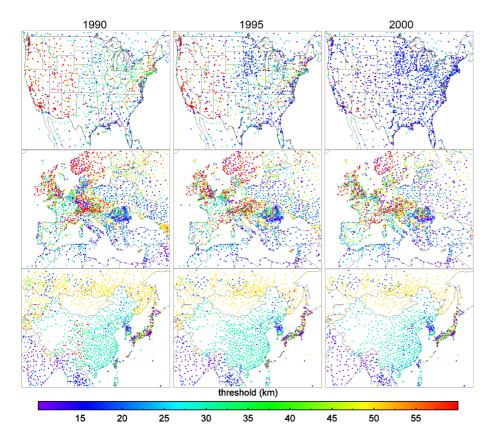


Figure A1. Threshold visbility of ISD stations over the US, Europe and Eastern Asia in 1990, 1995 and 2000.

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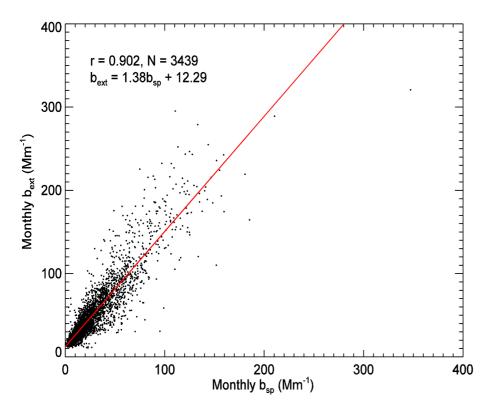


Figure A2. Scatter plot of monthly $b_{\rm sp}$ (measured by nephelometers) and $b_{\rm ext}$ (estimated from aerosol speciation data) from all IMPROVE stations with $b_{\rm sp}$ measurements for 1993–2013.

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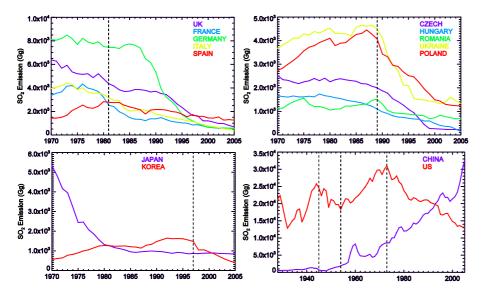


Figure A3. SO₂ emission for several major countries. Data are from S. J. Smith et al. (2011a). The top left and top right panels include major countries of Western and Eastern Europe, respectively. Vertical lines represent division years of the study periods that roughly indicate transition points of emission trend.

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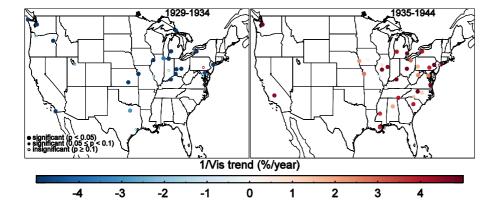


Figure A4. Spatial distribution of relative trends in 1/Vis over the US for 1929–1944. Larger filled points indicate trends with at least 95 % significance, smaller filled points represent trends with 90–95 % significance, and open circles indicate insignificant trends.

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