



On the scaling of the solar incident flux

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This discussion paper is/has been under review for the journal Atmospheric Chemistry and Physics (ACP). Please refer to the corresponding final paper in ACP if available.

On the scaling of the solar incident flux

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Received: 20 February 2015 – Accepted: 27 March 2015 – Published: 15 April 2015

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Published by Copernicus Publications on behalf of the European Geosciences Union.

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Abstract

It was recently found that spectral solar incident flux (SIF) as a function of the ultra-violet wavelengths exhibits $1/f$ -type power-law correlations. In this study, an attempt was made to explore the SIF intrinsic dynamics vs. a wider range of wavelengths, from 115.5 to 629.5 nm. It seemed that the intermittency of SIF data set was very high and the revealed DFA- n exponents were close to unity thus again indicating $1/f$ power-law correlations. Moreover, the power spectral density was fitted algebraically with exponents close to unity. Eliminating the fitting of Planck formula at the Sun's effective temperature from SIF data set, scaling exponents very close to unity were derived, indicating that the $1/f$ scaling dynamics concern not the Planck's law but its fluctuations.

1 Introduction

As it is well known, electromagnetic radiation is continuously emitted by every physical body. This emitted radiation is adequately described by Planck's law near thermodynamic equilibrium at a definite temperature. There is a positive correlation between the temperature of an emitting body and the Planck radiation at every wavelength. As the temperature of an emitting surface increases, the maximum wavelength of the emitted radiation increases too. Smith and Gottlieb (1974) re-examined the subject of photon solar flux and its variations vs. wavelength and showed that variations in the extreme ultraviolet (UV) spectrum and in the X-ray of solar flux may reach high orders of magnitude causing significant changes in the Earth's ionosphere, especially during major solar flares (Kondratyev et al., 1995; Kondratyev and Varotsos, 1996; Varotsos et al., 2001; Melnikova, 2009; Tzanis et al., 2009; Xue et al., 2011).

Solanki and Unruh (1998) proposed simple models of the total solar irradiance variations vs. wavelength showing that variations on solar flux are mainly caused by magnetic fields at the solar surface. Solar observations may be reproduced by a model of

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three parameters: the quiet Sun, a facular component and the temperature stratification of sunspots.

Tobiska et al. (2000) developed a forecasting solar irradiance model, called SO-LAR2000, covering the spectral range of 1–1 000 000 nm. Using this tool, the authors attempted to describe solar variation vs. wavelength and through time from X-ray through infrared wavelengths, in order to predict the solar radiation component of the space environment.

Very recently, Varotsos et al. (2013a, b) suggested the existence of strong persistent long-range correlations in the spectral space of the solar flux fluctuations vs. the UV wavelengths in the range 278–400 nm. More precisely, by applying the detrending fluctuation analysis (DFA) to the initial SIF vs. UV wavelengths data set power-law correlations of the type $1/f$, which is omnipresent in nature, was found pointing to a scaling feature in the spectral domain.

However, Varotsos et al. (2013a) tried to formulate the above-shown finding, i.e., that the solar spectral irradiance obeys $1/f$ power-law as a function of UV wavelength, using the well-known Planck's law: $I(\tau, T) = \frac{2hc^2}{\tau^5(e^{\frac{hc}{\tau kT}} - 1)}$ which, in the limit of small wavelengths

tends to the Wien approximation: $I(\tau, T) = \frac{2hc^2}{\tau^5} e^{-\frac{hc}{\tau kT}}$, where $I(\tau, T)$ is the amount of energy emitted at a wavelength τ per unit surface area per unit time per unit solid angle per unit wavelength, T is the temperature of the black body, h is Planck's constant, c is the speed of light, and k is Boltzmann's constant. By applying the DFA method on the various values of $I(\tau, T)$ Varotsos et al. (2013a) showed that the calculated $I(\tau, T)$ values do not obey the $1/f$ -type scaling vs. τ . Thus, the latter may reflect a scaling in its fluctuations which might be related with the complex physical processes taking place at the solar atmosphere (e.g. see Avrett and Loeser (2008) and references therein).

In the present study, focusing on these fluctuations, we examine whether the $1/f$ scaling feature is apparent in a wider spectrum of SIF including both UV and visible spectrum, namely for wavelengths (WL) between 115.5 and 629.5 nm.

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2 Data and analysis

As mentioned just above the available solar incident flux data for WL ranging from 115.5 to 629.5 nm with a step of 1 nm were employed. The spectrophotometric data of spectral extraterrestrial solar flux have been taken from the book by Makarova et al. (1991) (see also Makarova et al., 1994; Melnikova and Vasilyev, 2005). Figure 1a depicts SIF values for the wavelength range of 115.5–629.5 nm. The principal feature shown in this figure is the existence of “apparent” non-stationarities vs. WL into the solar spectral distribution and the strong upward trend up to about 450 nm. The detrending of this data set was accomplished (Fig. 1a) by applying the Planck formula $B_1 \left(\frac{b_1}{\tau}\right)^5 / \left[\exp\left(\frac{b_1}{\tau}\right) - 1\right]$ with $b_1 = 2486.4$ nm based on the Sun’s effective temperature ($T_{\text{sun}} = 5778$ K) reported by NASA (<http://nssdc.gsfc.nasa.gov/planetary/factsheet/sunfact.html>) and the derived parameter was found to be $B_1 = 85.8 \pm 0.7$ (0.82 %) $\text{mWm}^{-2} \text{nm}^{-1}$. Hereafter, we focus on these “detrended” SIF data which are shown with the blue line in Fig. 1a; these are the deviations from a pure black body spectrum. To estimate the scaling exponents, we applied the well-known DFA method (Peng et al., 1994; Weber and Talkner, 2001; Varotsos, 2005; Varotsos et al., 2008, 2012; Skordas et al., 2010; Efstathiou et al., 2011; Chattopadhyay and Chattopadhyay, 2014).

Furthermore, we calculated the power spectrum for the detrended SIF data set using Fast Fourier Transform (FFT) algorithm as well as the maximum entropy method (MEM) (Hegger et al., 1999).

A brief description of DFA- n tool may be given as follows:

1. Consider the SIF data set $x(i)$ of length N which is integrated over WL. Therefore, the integrated data set, $y(i)$, is consisting of the following points:

$$y(1) = x(1), \quad y(2) = x(1) + x(2) \dots, \quad y(i) = \sum_{k=1}^i x(k) \quad (1)$$

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2. We split the integrated data set into non-overlapping boxes of equal length, τ . In each box, a best polynomial local trend (of order n) is fitted in order to detrend the integrated profile (by subtracting the locally fitted trend).
3. The root-mean-square fluctuations $F_d(\tau)$ of this integrated and detrended profile is calculated over all scales (box sizes). The detrended fluctuation function F_k within the k th box (Kantelhardt et al., 2002) is defined by:

$$F_k^2(\tau) = \frac{1}{\tau} \sum_{i=k\tau+1}^{(k+1)\tau} [y(i) - z(i)]^2, \quad k = 0, 1, 2, \dots, \left(\frac{N}{\tau} - 1\right) \quad (2)$$

where $z(i)$ is the corresponding polynomial least-square fit to the τ data contained and (Peng et al., 1994):

$$F_d^2 = \frac{1}{(N/\tau)} \sum_{k=0}^{(N/\tau)-1} F_k^2(\tau). \quad (3)$$

In case the signals involve scaling, a power-law behavior for the root-mean-square fluctuation function $F_d(\tau)$ is observed:

$$F_d(\tau) \sim \tau^\alpha \quad (4)$$

where α is the scaling exponent, a self-affinity parameter that represents the long-range power-law correlation (Ausloos and Ivanova, 2001).

The exponent α is equal to the generalized Hurst exponent of the integrated process. Due to this integration, one must subtract one to obtain the scaling exponent H of the process itself: $H = \alpha - 1$. In the special case where the process is quasi-Gaussian then – unlike the general multifractal case – a single exponent is sufficient to characterize the scaling of the process. If the process is exactly Gaussian, then when $0 < \alpha < 1$

($-1 < H < 0$), the process is a fractional Gaussian noise (fGn) and when $1 < \alpha < 2$ ($0 < H < 1$), it is a fractional Brownian motion (fBm).

For uncorrelated quasi Gaussian data, the scaling exponent is $\alpha = 1/2$, ($H = -1/2$). Scaling generally implies the existence of long range statistical dependencies. However in the special Gaussian case, when $\alpha = 1/2$, then one has a Gaussian white noise and there are no long range dependencies. However, the fact that $\alpha = 1/2$, can only be used to infer an absence of long range dependencies if it can be shown that the process is indeed quasi-Gaussian (in practice it has to be shown that it is not significantly multifractal) in a certain range of τ values. If $0 < \alpha < 0.5$ (and if the data set is nonintermittent), power-law anticorrelations are present (antipersistence). When $0.5 < \alpha \leq 1.0$ (and if again the data set is nonintermittent), then persistent long-range power-law correlations prevail (the case $\alpha = 1$ corresponds to the so-called $1/f$ noise) (Weber and Talkner, 2001).

Finally, the scaling properties of SIF-WL data set were also studied using Haar fluctuation analysis (Lovejoy and Schertzer, 2012a, b). In the DFA method, fluctuations are defined by the SD of the residues of the polynomial regressions of the integrated process (the F_k function). In the more general framework of wavelets, they are defined by convolutions with respect to (almost) arbitrarily shaped “mother wavelets”. While both DFA and wavelet based fluctuations give essentially the same accuracy when used for estimating scaling exponents (such as H , see e.g., Lovejoy and Schertzer, 2012a), if the wavelet is appropriately chosen, then the interpretation of the fluctuations becomes very simple so that the analysis can be used much more generally (i.e. when there is more than a single scaling regime, or even there is no scaling at all). The difficulty in interpreting the DFA fluctuations is the reason why published plots units are typically not even provided for the fluctuation axes! In contrast, the simple (and indeed first wavelet) – Haar fluctuations $\Delta X(\Delta t)$ – are defined simply as the difference between the averages of a series over the first and second halves of an interval. They have the property that in regions where $-1 < H < 0$, they can be interpreted simply as anomalies. Whereas in regions where $0 < H < 1$, they can be interpreted as differences. Once

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the Haar fluctuations have been determined one can characterize them statistically by taking averages of various powers q ; the “generalized” q th order structure function $S_q(\Delta t) = \langle \Delta X(\Delta t)^q \rangle$, where the symbol $\langle \cdot \rangle$ stands for ensemble averaging. In a scaling regime, $S_q(\Delta t) \approx \Delta t^{\xi(q)}$, where the exponent $\xi(q) = qH - K(q)$ and $K(q)$ is in general non linear and convex and characterizes the intermittency (satisfying $K(1) = 0$). In the previous DFA discussion, the process was assumed to be quasi-Gaussian, this implies the assumption $K(q) = 0$ (the DFA fluctuations are of course valid without this restriction). For universal multifractals (e.g. Tessier et al., 1993),

$$K(q) = C_1 \frac{(q^\alpha - q)}{(\alpha - 1)}, \quad (5)$$

where $C_1 (\geq 0)$ indicates the scaling intermittency and $\alpha (0 \leq \alpha \leq 2)$ – not to be confused with the DFA exponent – is the multifractality index (the Lévy index of the generator).

3 Discussion and results

Varotsos et al. (2013a) studying the high-resolution observations of SIF reaching the ground and the top of the atmosphere, suggested that SIF vs. UV WL exhibits $1/f$ -type power-law correlations. This result was derived by applying the DFA method on the SIF dataset obtained from the Villard St. Pancrace station of the Lille University of Sciences and Technology and was based on the slope (i.e., $\alpha = 1.02 \pm 0.02$, hence $H = 0.02 \pm 0.02$) of the log–log plot of the root mean square fluctuation function of SIF vs. the WL segment size τ .

In the present study, the scaling dynamics of a wider spectrum of SIF vs. WL data set was studied, for wavelengths between 115.5 and 629.5 nm. Firstly, DFA- n seemed to can take care of the trends revealing a DFA-exponent close to unity (after DFA-1), as shown in Fig. 1b.

In the following we plotted the power spectral density (using FFT) of the detrended SIF data set. The derived power spectral density showed that the power-law fitting

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gives an exponent $\beta = 0.99 (\pm 0.08)$ (see also below and Fig. 2). We can see that the slope inferred from the Haar analysis (below, $\beta = 1.46$) also fits quite well, this is discussed below. In terms of the structure function exponent, we have the relation $\beta = 1 + \xi(2)$ which follows from the Wiener–Khinchin theorem (the spectrum is the Fourier transform of the autocorrelation function, which is a second order statistic, hence the $q = 2$). On the other hand, the DFA-1 exponent was $1.09 (\pm 0.04)$, while by applying the DFA- n with $n > 1$ on the detrended SIF data, the derived exponents ranged from 0.98 to 1.01. Note that the DFA exponent for the fluctuation (a $q = 1$ statistic) cannot in general be used to estimate β which is a $q = 2$ statistic (unless one assumes $K(q) = 0$, the quasi-Gaussian assumption).

Next, to summarize our results we analysed the detrended SIF-WL data set by using Haar analysis (Lovejoy and Schertzer, 2012a, b) using the software available at <http://www.physics.mcgill.ca/~gang/software/doc/haarpack.zip>. According to Haar analysis, as also mentioned in the Sect. 2, the variation of SIF fluctuations vs. wavelength τ can be defined using the “generalized” q th order structure function $S_q(\tau) = \langle \Delta \text{SIF}(\tau)^q \rangle$, for which it holds that in a scaling regime $S_q(\tau) \approx \tau^{\xi(q)}$, where the exponent $\xi(q) = qH - K(q)$ and $K(q)$ illustrates the scaling intermittency (satisfying $K(1) = 0$ and $\xi(1) = H$). Figure 3b shows that the intermittency of SIF data set is very high ($C_1 = 0.16$), hence the RMS exponent $= \xi(2)/2 = 0.23$ is quite different from the $q = 1$ exponent (H) and the data are far from Gaussian. In the classical quasi-Gaussian case, $K(q) = 0$ so that $\xi(q)$ is linear. More generally, if the field is intermittent – for example if it is the result of a multifractal process – then the exponent $K(q)$ is generally non linear and convex and characterizes the intermittency, as already mentioned. The physical significance of H is thus that it determines the rate at which mean fluctuations grow ($H > 0$) or decrease ($H < 0$) with scale τ . According to Fig. 3a and b, the exponent $\xi(2)$ of the structure function equals to zero (at scales below 20 nm), a fact which means that the power spectrum exponent $\beta = 1 + \xi(2)$ equals to 1 ($1/f$ structure). On the other hand, at larger scales, the exponents $\xi(2)$ and β seem to equal to 0.46 and 1.46, respectively.

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high and the data were far from Gaussian. At scales below 20 nm, the power spectrum exponent β was almost 1 ($1/f$ structure), while at larger scales, the exponents $\xi(2)$ and β are equal to 0.46 and 1.46, respectively. This prompted us to revisit DFA-1 and search for such a crossover. Indeed a crossover at 20 nm can be observed leading to compatible β exponents.

- The results of the power spectral density for the detrended SIF-WL data set (using the MEM) vs. frequency are compatible with the aforementioned two β exponents.

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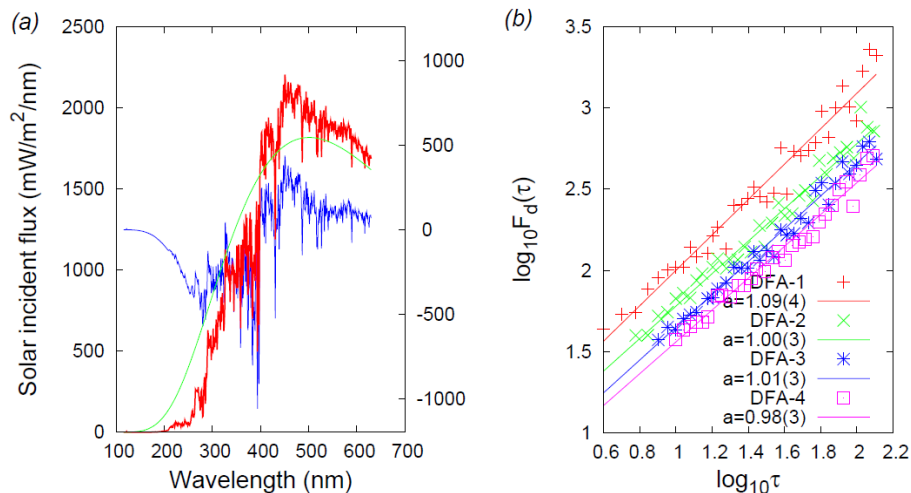


Figure 1. (a) SIF values (red, left scale) on the top of the atmosphere vs. WL from 115.5 to 629.5 nm together with the fitting employed (green, left scale). The detrended SIF data (blue, right scale) are also shown. (b) Log-log plot of the root mean square fluctuation function $F_d(\tau)$ of the detrended SIF vs. the WL segment size τ , for the wavelengths between 115.5 and 629.5 nm. The a values for DFA-1, DFA-2, DFA-3, DFA-4 are 1.09 (± 0.04), 1.00 (± 0.03), 1.01 (± 0.03), 0.98 (± 0.03), respectively.

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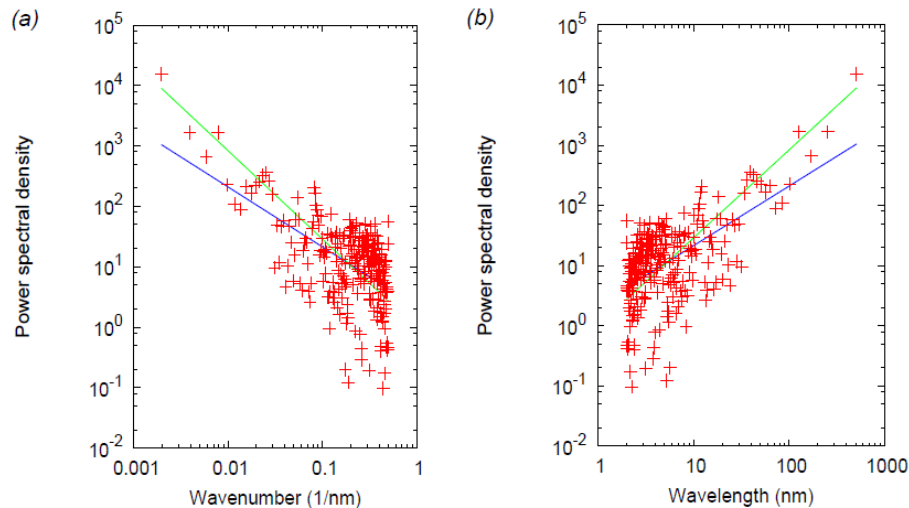


Figure 2. Power spectral density of the detrended SIF data set together with the least squares fit with power-law exponent $\beta = 0.99 (\pm 0.08)$ (blue line) and the $\beta = 1.46$ as determined with the Haar analysis (green line) vs. **(a)** the wavenumber, and **(b)** the wavelength.

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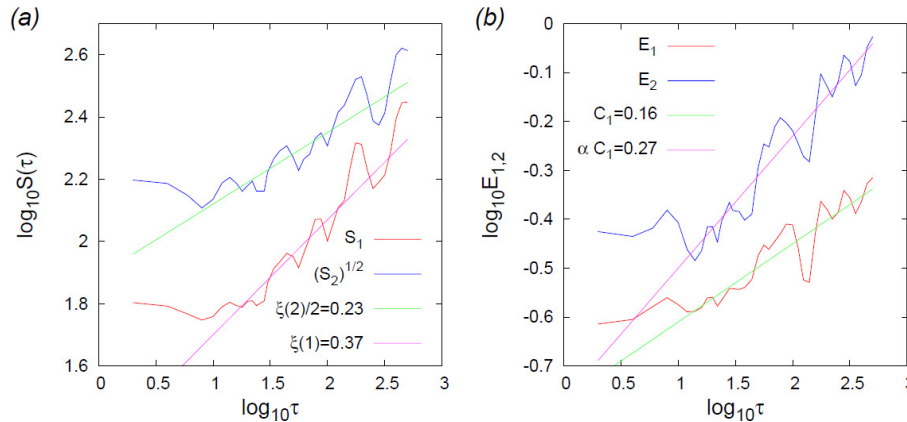


Figure 3. (a) Haar analysis on the detrended SIF data set for moments $q = 1$ (red) and RMS (blue). $\xi(1) = \xi(2) \approx 0$ at scales below 20 nm, whereas for larger scales up to 316 nm we have slopes $\xi(2)/2 = 0.23$ (green) and $\xi(1) = H = 0.37$ (magenta). **(b)** Red: log-log plot of $E_1 = S_1 / (S_{1+\delta q} / S_{1-\delta q})^{\frac{1}{2\delta q}}$ vs. τ whose slope is $K'(1) (= C_1)$, blue: log-log plot of $E_2 = (S_{1+2\delta q} S_{1-2\delta q} / S_1^2)^{-\frac{1}{2\delta q^2}}$ vs. τ whose slope is $K''(1) (= \alpha C_1)$ for $\delta q = 0.1$. The first yields an estimate $C_1 \approx 0.16$ indicating high intermittency, the ratio yields an estimate $\alpha = 1.7$. For smaller scales (up 40 nm), the corresponding slopes are close to 0 indicating Gaussianity. With these parameters and Eq. (5) (i.e., the universal multifractal equation for $K(q)$ in Sect. 2), we find $\xi(2)/2 = H - K(2)/2 \approx 0.37 - 0.14 = 0.23$ in agreement with the direct estimate in **(a)**.

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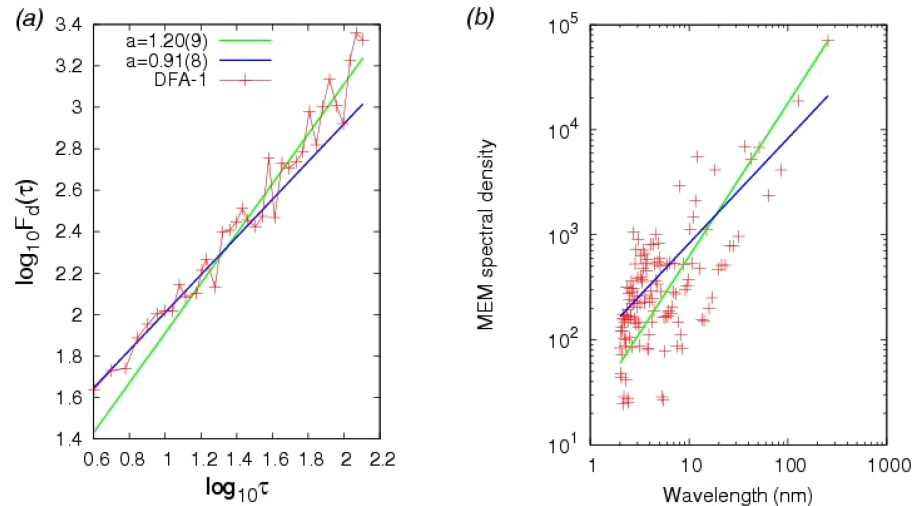


Figure 4. (a) Log-log plot of the root mean square fluctuation function $F_d(\tau)$ of the detrended SIF vs. the WL segment size τ for DFA-1 together with corresponding least-squares fits for $\tau \leq 15$ nm (blue) and $\tau > 20$ nm (green). (b) Power spectrum using the MEM for the detrended SIF data set with the two power-law behaviours fits $\beta = 1$ (blue) – from a least squares fit up to 20 nm – and $\beta = 1.46$ (green) for the $1/f$ and $1/f^{1.46}$ structure, respectively, vs. the wavelength.

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