1 Technical Note: The horizontal scale-dependence of the

2 cloud overlap parameter alpha

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8

9 Abstract

10 The cloud overlap parameter alpha relates the combined cloud fraction between two 11 altitude levels in a grid box to the cloud fraction as derived under the maximum and random overlap assumptions. In a number of published studies in this and other Journals it is found 12 13 that alpha tends to increase with increasing scale. In this Technical Note, we investigate this 14 analytically by considering what happens to alpha when two grid boxes are merged to give a 15 grid box with twice the area. Assuming that alpha depends only on scale then, between any 16 two fixed altitudes, there will be a linear relationship between the values of alpha at the two 17 scales. We illustrate this by finding the relationship when cloud cover fractions are assumed to be uniformly distributed, but with varying degrees of horizontal and vertical correlation. 18 19 Based on this, we conclude that alpha increases with scale if its value is less than the vertical 20 correlation coefficient in cloud fraction between the two altitude levels. This occurs when the clouds are deeper than would be expected at random (i.e. for exponentially distributed cloud 21 22 depths).

1 1 Introduction

2

Clouds tend to be represented in GCMs as plane-parallel and horizontally homogeneous, with the combined horizontal cloud fraction between clouds at different altitudes specified according to various overlap schemes (e.g. Smith, 1990; Tiedtke, 1993). These schemes are generally based on a combination of maximum and random overlap. In maximum overlap the clouds are maximally overlapped in height resulting in the minimum of interaction between clouds and downward radiation. Where clouds are randomly overlapped in height the interaction with radiation is greater.

10

11 Taking advantage of the fact that clouds close together in altitude are likely maximally 12 overlapped and those significantly different in altitude are likely randomly overlapped Hogan 13 and Illingworth (2000) introduced a cloud overlap scheme that has since been widely taken up 14 within GCMs. In this scheme, the mean combined cloud fraction between two altitude levels 15 is taken to be a weighted average (with weight α) of the mean values given by maximum and 16 random overlap assumption respectively.

17

18 The value of α is generally taken to be a function of the height separation (Δz) between 19 the two altitudes and is found to often have an inverse exponential dependence on Δz (e.g. 20 Hogan and Illingworth, 2000). The rate of fall is then determined by a cloud 'decorrelation 21 length' *L* (i.e. $\alpha = e^{-\frac{\Delta z}{L}}$). Since this initial study of Hogan and Illingworth (2000) many others 22 have investigated how α (and *L*) depend on horizontal scale (e.g. Mace and Benson-Troth 23 2002; Oreopoulos and Khairoutdinov 2003; Pincus *et al.* 2005; Willén et al. 2005; Barker 2008a & 2008b; Shonk and Hogan 2010; Oreopoulos and Norris 2011; Oreopoulos *et al.* 2012). Though a number of different definitions for α and methods for deriving *L* have been
 used in such studies, they generally find that α (and, hence, *L*) increases with horizontal scale.

4

5 **2.** The overlap parameter α

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From the observed horizontal cloud fractions c_a and c_b at altitudes *a* and *b* (at a fixed scale) the horizontal cloud fractions c_{max} and c_{rand} can be formed, under the maximum and random overlap schemes, as:

10

$$c_{max} = \max(c_a, c_b) \tag{1}$$

$$c_{rand} = c_a + c_b - c_a c_b \tag{2}$$

12 From the definition as given by Hogan and Illingworth (2000) for α these are related to 13 the combined horizontal cloud fraction, c_t (jointly covered by the clouds at both altitudes) by:

14
$$\overline{c_t} = \alpha \overline{c_{max}} + (1 - \alpha) \overline{c_{rand}}$$
(3)

15 Where $\overline{c_t}$, $\overline{c_{max}}$ and $\overline{c_{rand}}$ are the averages (over time) of c_t , c_{max} and c_{rand} respectively. 16 For the idealised case given here the averaging period is not important. However, we do need 17 the mean and variance in the cloud cover to be stable and similar at both heights and most 18 published work on cloud overlap is based on seasonal averages (e.g. Hogan and Illingworth, 19 2000; Oreopoulos and Norris, 2011).

20 Provided $\overline{c_{max}}$ and $\overline{c_{rand}}$ are not equal to each other, which is unlikely (as this could only 21 happen if the cloud cover fraction was always zero or one) Eq. 3 can be rearranged to give:

22
$$\alpha = \frac{\overline{c_t} - \overline{c_{rand}}}{\overline{c_{max}} - \overline{c_{rand}}}$$
(4)

1 As pointed out in Pincus *et al.* (2005), this is only one way to define α . Another method is 2 to determine a set of values for α using Eq. 3 based on the individual (unaveraged) values of c_t , c_{max} and c_{rand} and, from these, find an average value for α . However, this approach 3 leads to data being discarded, as (the values for) α are not uniquely defined when either 4 $c_a = 0$ or $c_b = 0$, potentially giving rise to truncated statistics. As the probability that $c_a = 0$ 5 or $c_b = 0$ decreases with increasing grid size (e.g. Astin and Girolamo; 1999) it seems 6 7 prudent, when considering the scale-dependence, to use Eq. 4 to define α (in which no data 8 are discarded).

9

10 3. The horizontal scale-dependence of α

11

12 To investigate the scale-dependence of α , we will consider what happens when two 13 horizontally adjacent grid boxes, which we label *i* and i+1 respectively, are combined to give 14 a single larger grid box with double the area. In the following there is no significance to j or i+1 except as labels to distinguish the original two grid boxes. However, zonal and meridional 15 16 anisotropies in real cloud regimes could make α directionally dependent. This wouldn't affect 17 the mathematics in this note, but could blur the signal when applied to real data, if arbitrary 18 pairs of adjacent grid boxes are combined. This could be handled by giving a direction to j 19 with, say, grid box j+1 being zonally (or meridionally) adjacent to grid box j. In either case, the cloud fractions C_a and C_b at the two altitudes (a and b) in the larger grid box are given by: 20

21
$$C_{a} = \left(\frac{c_{a}(j) + c_{a}(j+1)}{2}\right)$$

$$C_{b} = \left(\frac{c_{b}(j) + c_{b}(j+1)}{2}\right)$$
(5)

1 where $c_x(y)$ is the cloud fraction in grid box y at altitude x. Again, the cloud overlap C_{MAX} 2 and C_{RAND} (at the larger scale) are formed, under the maximum and random overlap 3 assumptions, by:

4
$$C_{MAX} = \max(C_a, C_b) \tag{6}$$

$$C_{RAND} = C_a + C_b - C_a C_b \tag{7}$$

6 The combined cloud fraction, C_T , at the large scale is given by:

7
$$C_T = \frac{c_t(j) + c_t(j+1)}{2}$$
 (8)

8 where $c_t(y)$ is the combined cloud fraction in grid box y.

9 To continue, let α₁ be the value of α at the original scale and α₂ be the value of α
10 when the two grid boxes are merged. As in Eq. 4, the value of α₂ is given by:

11
$$\alpha_2 = \frac{\overline{C_T} - \overline{C_{RAND}}}{\overline{C_{MAX}} - \overline{C_{RAND}}}$$
(9)

12 where $\overline{C_T}$, $\overline{C_{MAX}}$ and $\overline{C_{RAND}}$ are the time averages of C_T , C_{MAX} and C_{RAND} respectively.

Assuming that α depends only on scale (and the altitude between *a* and *b*) then (using
Eq. 3) Eq. 8 becomes:

15
$$\overline{C_T} = \frac{\alpha_1 \overline{c_{max}(j)} + (1 - \alpha_1)\overline{c_{rand}(j)} + \alpha_1 \overline{c_{max}(j+1)} + (1 - \alpha_1)\overline{c_{rand}(j+1)}}{2}$$
(10)

16 The averages in Eq. 10 are those for grid boxes *j* and *j*+1 respectively. If *a* and *b* are 17 fixed altitudes then Eqs. 9 and 10 together imply that $\alpha_2 = m \alpha_1 + c$, where *m* and *c* are 18 constants. This doesn't necessarily imply that a linear relationship between α_1 and α_2 will be 19 observed, since data from different altitudes (likely having differing values of *m* and *c*) may 20 be combined in published studies. For Eq. 10 we have implicitly assumed that α_1 is the same for both grid boxes *j* and *j*+1. To simplify the mathematics, in the following we will also assume that any average is the same whether it is for grid box *j* or *j*+1 (e.g. $\overline{c_{max}(j)} = \overline{c_{max}(j+1)} = \overline{c_{max}}$). In Eq. 10 this is equivalent to dropping the *j* and *j*+1 dependences, which together with Eq. 9 gives:

5
$$\alpha_2 = \frac{\overline{c_{max}} - \overline{c_{rand}}}{\overline{c_{MAX}} - \overline{c_{RAND}}} \alpha_1 + \frac{\overline{c_{rand}} - \overline{c_{RAND}}}{\overline{c_{MAX}} - \overline{c_{RAND}}}$$
(11)

6 We can use Eq.11 (or Eq. 10) to investigate the conditions in which $\alpha_2 > \alpha_1$ (i.e. 7 where α would increase with scale). As an example, consider the contrived case where the 8 cloud cover varies between grid boxes, but is always the same at both heights *a* and *b* (i.e. 9 $c_a(j) = c_b(j)$ and $c_a(j+1) = c_b(j+1)$, but $c_a(j)$ may not equal $c_a(j+1)$). This says 10 nothing about the horizontal distribution of clouds at each height. However, this would seem 11 most likely to be associated with particular cloud regimes, such as vertically deep convective 12 clouds. For this case:

13
$$\overline{c_{max}} = \overline{\max(c_a(j), c_b(j))} = \overline{\max(c_a(j), c_a(j))} = \overline{c_a(j)}$$
(12)

14 Leading to:

15

$$\overline{c_{max}} = \overline{c_a(j)} = \overline{c_a} \tag{13}$$

16 Similarly, from Eq. 5, $C_a = C_b$, and $C_{MAX} = \max(C_a, C_b) = C_a$ giving:

17
$$\overline{C_{MAX}} = \overline{C_a} = \overline{\left(\frac{c_a(j) + c_a(j+1)}{2}\right)}$$
(14)

As we are assuming that the averages are the same for both *j* and *j*+1 Eq. 14 implies that $\overline{C_{MAX}} = \overline{c_a} = \overline{c_{max}}$ and $\alpha_2 = m \alpha_1 + (1 - m)$. Hence, in this case, the value of *m* is uniquely defined by the value of α_2 when α_1 equals zero (e.g. if $\alpha_2 = 0.2$ when $\alpha_1 = 0$ then m = 0.8and $\alpha_2 = 0.8 \alpha_1 + 0.2$). 1 It is instructive to consider this case further by studying the value of *m* analytically. In this 2 case, we can uniquely define a mean, μ , and variance, σ^2 , in cloud cover that is the same at 3 both heights, i.e.,

4
$$\mu = \overline{c_a(j)} = \overline{c_b(j)}$$

$$\sigma^2 = \overline{c_a^2(j)} - \mu^2 = \overline{c_b^2(j)} - \mu^2$$
(15)

5 In this case $\overline{c_{rand}}$ is by definition (from Eq. 2):

6
$$\overline{c_{rand}} = \overline{c_a(j)} + \overline{c_b(j)} - \overline{c_a(j)c_b(j)} = \mu + \mu - \overline{c_a^2(j)}$$
(16)

7 With Eq. 15, this gives:

8
$$\overline{c_{rand}} = 2\mu - \sigma^2 - \mu^2 \tag{17}$$

9 From Eqs. 7 and 14, the average $\overline{C_{RAND}}$ is given by:

10
$$\overline{C_{RAND}} = \overline{C_a} + \overline{C_b} - \overline{C_a C_b} = \overline{C_a} + \overline{C_a} - \overline{C_a C_a} = 2\mu - \overline{C_a^2}$$
(18)

11 This leads (from Eq. 5) to:

12
$$\overline{C_{RAND}} = 2\mu - \overline{\left(\frac{c_a(j) + c_a(j+1)}{2}\right)^2}$$
(19)

13 Multiplying out gives:

14
$$\overline{C_{RAND}} = 2\mu - \frac{1}{4} \overline{\left(c_a(j)\right)^2} - \frac{1}{4} \overline{\left(c_a(j+1)\right)^2} - \frac{1}{2} \overline{c_a(j)c_a(j+1)}$$
(20)

Again, assuming that averages are the same in both grid boxes, the mean, μ , and variance, σ , in cloud cover are the same for both grid boxes *j* and *j*+1, and retain their definitions as given in Eq. 15. In this case, the labels *j* and *j*+1 are redundant in the second and third terms on the RHS of Eq. 20 and can be dropped to give:

19
$$\overline{C_{RAND}} = 2\mu - \frac{1}{4}\overline{c_a^2} - \frac{1}{4}\overline{c_a^2} - \frac{1}{2}\overline{c_a(j)c_a(j+1)}$$
(21)

1 From Eq. 15 this reduces to:

2

4

$$\overline{C_{RAND}} = 2\mu - \frac{1}{2}(\sigma^2 + \mu^2) - \frac{1}{2}\overline{c_a(j)c_a(j+1)}$$
(22)

3 By definition, the co-variance of $c_a(j)$ and $c_a(j+1)$ is given by:

$$Cov(c_a(j), c_a(j+1)) = \overline{c_a(j)c_a(j+1)} - \mu^2$$
(23)

5 Similarly, by definition, the (horizontal) cross-correlation coefficient, *R*, in cloud cover between

6 the adjacent (smaller) grid boxes is given by:

7
$$R = \frac{Cov(c_a(j), c_a(j+1))}{\sqrt{Var(c_a(j))}\sqrt{Var(c_a(j+1))}} = \frac{Cov(c_a(j), c_a(j+1))}{\sigma^2}$$
(24)

8 Eqs, 22, 23 and 24 together give:

9
$$\overline{C_{RAND}} = 2\mu - \frac{1}{2}(\sigma^2 + \mu^2) - \frac{R}{2}\sigma^2 - \frac{1}{2}\mu^2 = 2\mu - \frac{1}{2}(1+R)\sigma^2 - \mu^2$$
(25)

10 Putting these into Eq. 11 gives:

11
$$m = \frac{\overline{c_{max}} - \overline{c_{rand}}}{\overline{c_{MAX}} - \overline{c_{RAND}}} = \frac{\mu - \sigma^2 - \mu^2}{\mu - \frac{1}{2}(1+R)\sigma^2 - \mu^2}$$
(26)

As an example, if the cloud fraction can be modelled as a Beta(*p*,*q*) distribution (e.g. Falls
13 1974; Tompkins 2002) then:

14
$$m = \frac{2(p+q)}{2(p+q)+(1-R)}$$
 (27)

15
$$\alpha_2 = \frac{2(p+q)}{2(p+q)+(1-R)} \alpha_1 + \frac{(1-R)}{2(p+q)+(1-R)}$$
(28)

In the simplest case, where the cloud fraction in each grid box is uniformly or Beta(1,1)
distributed (e.g. LeTreut and Li, 1991), Eq. 28 gives:

18
$$\alpha_2 = \frac{4}{5-R}\alpha_1 + \frac{1-R}{5-R}$$
(29)

19 (Thus, where R = 0 then $\alpha_2 = 0.8 \alpha_1 + 0.2$). Hence, in this contrived case (where the cloud 20 cover is the same at both heights) α will always increase with scale (i.e. $\alpha_2 > \alpha_1$) provided the horizontal correlation coefficient, *R*, in cloud fraction between adjacent grid boxes is
positive and less than 1.

3 Trivially, when R = 1 there is no scale-dependence to alpha (as m = 1). However, as 4 *R* decreases to zero the degree of the scale-dependence increases and maximises where R = 0. 5 This is displayed in Fig. 1, which shows the relationship between between α_1 and α_2 for a 6 range of values for *R* in the case where the cloud fraction in the adjacent grid boxes are 7 assumed to be uniformly distributed. The scale-dependence is strongest when R = 0, in which 8 $\alpha_2 = 0.8 \alpha_1 + 0.2$.

So far, we have looked at the scale-dependence where the cloud fraction varies from grid box to grid box, but doesn't vary with altitude. This implies that the vertical correlation between the cloud fractions at the two altitudes is $\rho = 1$. Let us now consider what happens when $\bar{c_a} = \bar{c_b}$, but $c_a(j)$ need not equal $c_b(j)$ (i.e. $\rho \neq 1$). For illustration, and to simplify the mathematics we will take the extreme case where R = 0 and assume that the cloud cover fractions at heights *a* and *b* are correlated uniform distributions, with (vertical) correlation coefficient ρ . This implies that mean cloud fraction at each height is $\mu = \frac{1}{2}$.

16 By Clarke (1961) or Nadarajah and Kotz (2008) for example, the mean ($\overline{c_{max}}$) of the 17 maximum of two correlated normally distributed random variables with mean $\mu = \frac{1}{2}$, standard 18 deviation σ and correlation coefficient ρ is given by:

19
$$\overline{c_{max}} = \frac{1}{2} + k(1-\rho)^{1/2}$$
(30)

20 where $k = \sigma^2 \pi^{-0.5}$.

21 We couldn't find a reference for the mean of the maximum of two correlated uniform 22 random variables so we will use Eq. 30, with *k* chosen to give the correct answer for $\overline{c_{max}}$ 1 when $\rho = 0$. (Eq. 30 will always give the correct answer when $\rho = 1$.) We will comment 2 later on the accuracy of this assumption.

3 If c_a and c_b are *independent* uniformly distributed random variables then $\rho = 0$ and 4 c_{max} follows a Beta(2,1) distribution, which has mean $\overline{c_{max}} = \frac{2}{2+1} = \frac{2}{3}$. Hence, Eq. 30 gives 5 the correct value for $\overline{c_{max}}$ if $k = \frac{1}{6}$. This leads to:

$$\overline{c_{max}} \cong \frac{1}{2} + \frac{1}{6} (1 - \rho)^{1/2}$$
(31)

Also, when c_a and c_b are independent uniformly distributed random variables their average C_a has the standard symmetric triangular distribution as does C_b . Hence $\overline{C_{MAX}}$ is the mean of the maximum of two independent triangularly distributed random variables. In this case $\overline{C_{MAX}} = \frac{37}{60}$ and Eq. 30 gives the correct value if $k = \frac{7}{60}$. This leads to:

11
$$\overline{C_{MAX}} \cong \frac{1}{2} + \frac{7}{60} (1 - \rho)^{1/2}$$
(32)

12 In a similar way to R, the vertical correlation coefficient ρ is defined as:

13
$$\rho = \frac{Cov(c_a(j),c_b(j))}{\sqrt{Var(c_a(j))}\sqrt{Var(c_b(j))}} = \frac{Cov(c_a(j),c_b(j))}{\sigma^2} = \frac{\overline{c_ac_b} - \mu^2}{\sigma^2}$$
(33)

14 Based on Eq. 2, Eq. 33 gives:

15
$$\overline{c_{rand}} = \overline{c_a} + \overline{c_b} - \overline{c_a c_b} = 2\mu - \mu^2 - \sigma^2 \rho$$
(34)

16 [This is identical to Eq. 17 when $\rho = 1$.] For a uniform distribution $\sigma^2 = \frac{1}{12}$, giving:

17
$$\overline{c_{rand}} = \frac{3}{4} - \frac{1}{12}\rho \tag{35}$$

18 Similarly:

1
$$\overline{C_{RAND}} = \overline{C_a} + \overline{C_b} - \overline{C_a C_b} = 2\mu - \frac{\overline{(c_a(j) + c_a(j+1))} (c_b(j) + c_b(j+1))}{2}$$
(36)

2 Multiplying out gives:

3
$$\overline{C_{RAND}} = 2\mu - \left(\frac{\overline{c_a(j)c_b(j)}}{4} + \frac{\overline{c_a(j)c_b(j+1)}}{4} + \frac{\overline{c_a(j+1)c_b(j)}}{4} + \frac{\overline{c_a(j+1)c_b(j+1)}}{4}\right)$$
(37)

4 As we are only considering the case where R = 0 (i.e. no horizontal correlation) this simplifies 5 (Eq. 23) to:

6
$$\overline{C_{RAND}} = 2\mu - \frac{\overline{c_a(j)c_b(j)}}{4} - \frac{\mu^2}{4} - \frac{\mu^2}{4} - \frac{\overline{c_a(j+1)c_b(j+1)}}{4}$$
(38)

7 As the averages are the same for both j and j+1:

8
$$\overline{C_{RAND}} = 2\mu - \frac{1}{2}\mu^2 - \frac{1}{2}\overline{c_a c_b} = 2\mu - \frac{1}{2}\mu^2 - \frac{1}{2}(\mu^2 + \sigma^2 \rho)$$
(39)

9
$$\overline{C_{RAND}} = 2\mu - \frac{1}{2}\mu^2 - \frac{1}{2}\overline{c_a c_b} = 2\mu - \mu^2 - \frac{1}{2}\sigma^2\rho$$
(40)

10 With
$$\sigma^2 = \frac{1}{12}$$
 this gives:

11
$$\overline{C_{RAND}} = \frac{3}{4} - \frac{1}{24}\rho \tag{41}$$

12 Putting the above values into Eq. 11 gives:

13
$$\alpha_2 \simeq \alpha_1 \left(\frac{30 - 10\rho - 20(1 - \rho)^{1/2}}{30 - 5\rho - 14(1 - \rho)^{1/2}} \right) + \left(\frac{5\rho}{30 - 5\rho - 14(1 - \rho)^{1/2}} \right)$$
(42)

14 Though this is an approximate result, the simulated values given in Fig. 2 show that 15 Eq. 42 can be taken as exact for all values of ρ . Thus, if $\rho = 0$ (i.e. the cloud cover at both 16 altitudes are uncorrelated) $\alpha_2 = \frac{5}{8}\alpha_1$ and so α will always decrease with scale (i.e. $\alpha_2 < \alpha_1$), 17 except where $\alpha_1 = 0$. It seems likely, given the linear relationship between the values of alpha at the two scales that for every value of ρ there will be a unique value for α that does not change with scale being the point-of-intersection with the $\alpha_1 = \alpha_2$ line. This is illustrated in Fig. 2, where the relationship between between α_1 and α_2 is displayed for a range of values for ρ (all with R = 0). From Fig. 2 this value seems to be where $\alpha_1 = \alpha_2 \approx \rho$. Also, where $\alpha_1 > \rho$ then α will decrease with scale and where $\alpha_1 < \rho$ then α will increase with scale.

8 4. Conclusions

9 Based on the definition of α and the scale invariance of the combined cloud fraction, if 10 α depends only on scale then the value of alpha, α_2 , at one scale is linearly related the value of alpha, α_1 , at the other scale (i.e. $\alpha_2 = m \alpha_1 + c$) provided the two altitudes are fixed. The 11 values of m and c depend on a number of parameters including the mean, μ , and variance, σ^2 , 12 13 in cloud fraction at each altitude. However, the most important parameters are the horizontal 14 correlation coefficient, R, between the cloud fractions in adjacent grid boxes (at a given 15 altitude) and the vertical correlation coefficient, ρ , between the cloud fractions at the two 16 altitudes.

17 If R, ρ , μ and σ^2 are found from real cloud data then this note allows the value of α_2 to 18 be calculated from α_1 directly. As horizontal cloud properties, R, μ and σ^2 can be found 19 directly from the passive or active remote sensing of clouds. However, ρ would require 20 knowledge of cloud vertical structure, which could come from active remote sensing (e.g. as 21 in Kato et al. (2010) from CloudSat and CALIPSO data).

22 Dependent on the relative values of α and ρ it is possible for α to increase, decrease or 23 stay the same with increasing scale. However, the strength of the dependence is controlled by 24 *R*. Published results tend to obscure the linear relationship between α_2 and α_1 by plotting them together on the same graph against height separation, rather than against one another (e.g. Oreopoulos and Norris, 2011). This also combines data from differing pairs of altitudes (*a* and *b*) together, where each pair could have a different linear relationship. However, our results indicate that an 'on average' increase of α with scale implies that on average α must generally be smaller than ρ .

6 In Astin and Di Girolamo (2006) we showed that on average $\alpha \approx \rho$ when cloud 7 depths follow an exponential distribution. Hence, we conclude that the published increase of 8 α with scale is a consequence of clouds being generally deeper than would be expected at 9 random (i.e. in a Random Markov Field).

10 Also, the scale-dependence disappears when R = 1 and is strongest when R = 0. Hence, 11 an increase in α with scale implies that R must be positive and less than 1. Based on published data on α , or directly from cloud data it is possible to determine R if there is enough data to 12 determine ρ , μ and σ^2 . As an illustration, Figure 1 of Oreopoulos and Norris (2011) gives 13 $\alpha_1 \approx 0$ (at 75 km scale) and $\alpha_2 \approx 0.04$ (at 150 km scale) for an altitude separation of 10 km 14 15 when averaged over June, July and August. Based on this note, this would indicate that if 16 $\rho = 0$ then R has a maximum value of 0.8 (our figure 1). However, R could equal zero, provided that $\rho \ge 0.2$ (our figure 2). As ρ is likely to be close in value to α_1 this would seem 17 to imply that R is closer to 0 than 0.8. This is a wide range for R, but could be made narrower 18 19 if ρ is known.

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Fig. 1. The dependence of α_2 on α_1 for cloud fractions (in adjacent grid boxes) that are uniformly distributed, where the vertical correlation coefficient in cloud cover $\rho = 1$ and the horizontal correlation coefficient in cloud cover is *R* (solid line). The dashed line is where there would be no scale dependence to α (i.e. $\alpha_2 = \alpha_1$). The circles are values given by simulation.



Fig. 2. The dependence of α₂ on α₁ for cloud fractions that are uniformly distributed (solid
line), where the horizontal correlation coefficient in cloud cover is *R* = 0, and the vertical
correlation coefficient in cloud cover is ρ. The dashed line is where there would be no scale
dependence to α (i.e. α₂ = α₁). The circles are values from simulation.